

Probability and Statistics

for experimental physicists : part III

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Université Paris Diderot and IN2P3

Lecture 1 (Monday)

Basic concepts in Probability and Statistics

Lecture 2 (Tuesday)

**Maximum Likelihood theorem
Multivariate techniques**

Lecture 3 (today)

**An analysis example from *BaBar*
Hypothesis testing, limit settings**

Disclaimer

**Most, if not all of you, are already familiar with many of these topics...
for consistency, the scope spans from the very general concepts towards
more advanced developments...**

A COMPLETE ANALYSIS EXAMPLE FROM BABAR

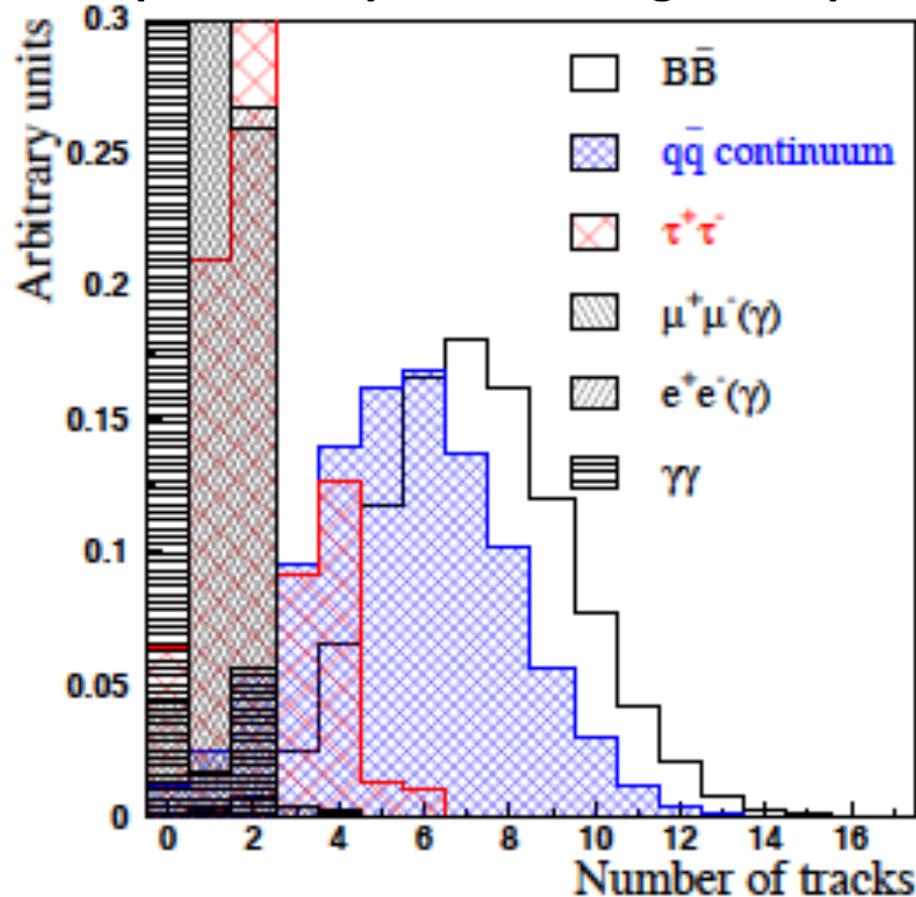
Amplitude analysis of neutral B mesons decaying into 3 particles : $B^0 / \overline{B}^0 \rightarrow K_s^0 \pi^+ \pi^-$

- **requires large levels of background suppression**
- **requires B-meson “flavor” tagging**
- **requires ML discrimination of signal against irreducible background**
- **requires composition of signal amplitude and interference patterns**

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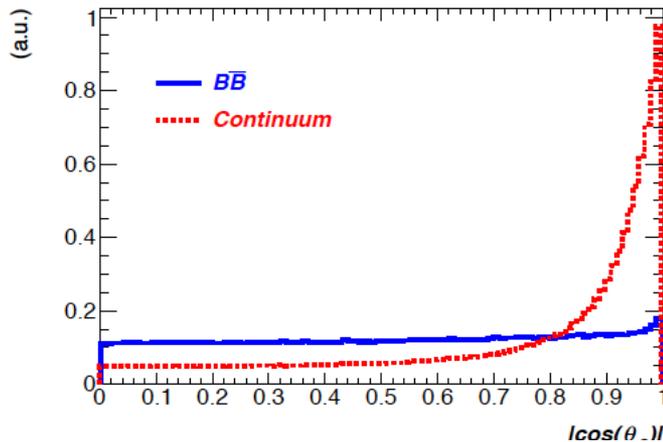
Cut-based selection, 1st stage :

filter on simple, robust variables

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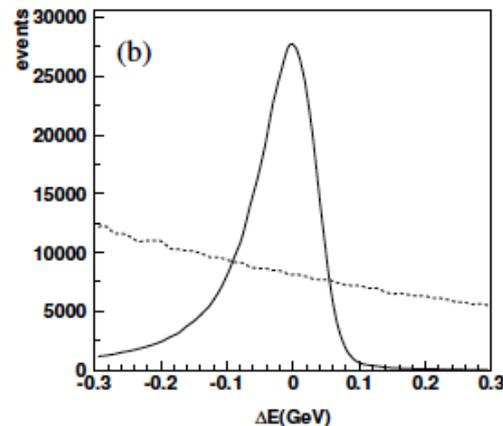
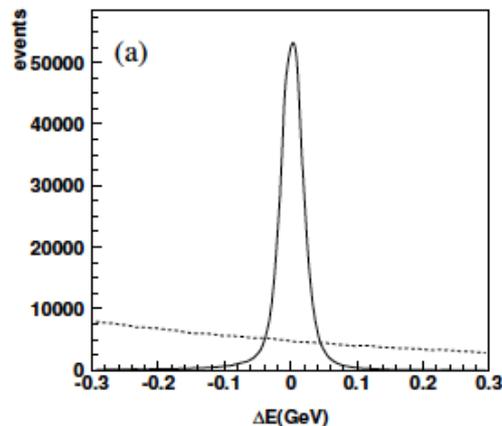


Cut-based selection, 2nd stage :

kinematical variables

“event-shape” variables

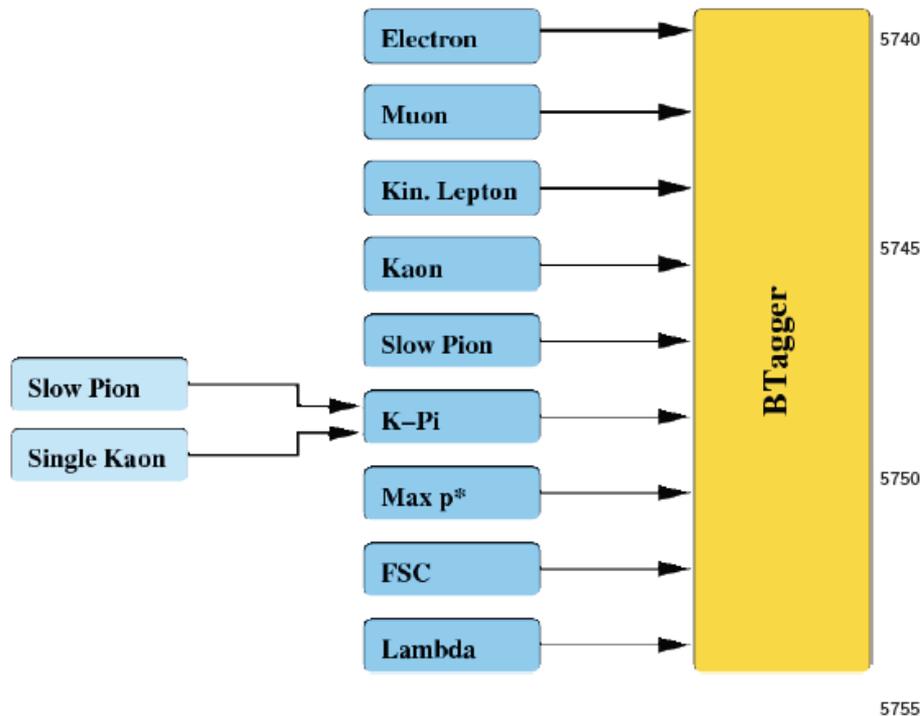
Define physics sample



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B-flavor tagging :

exploit available information

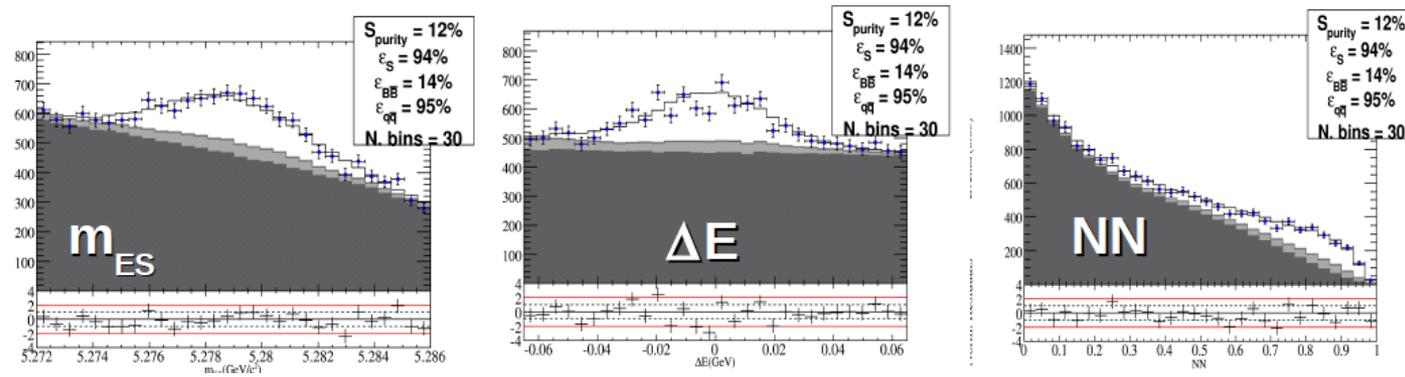
combine into a neural network

Fig. 8.5.1. Schematic overview of the *BABAR* tagging algorithm. Each box corresponds to a separate neural network.

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Information implemented in the Likelihood :

kinematical variables

various event-shape variables, combined into a neural network

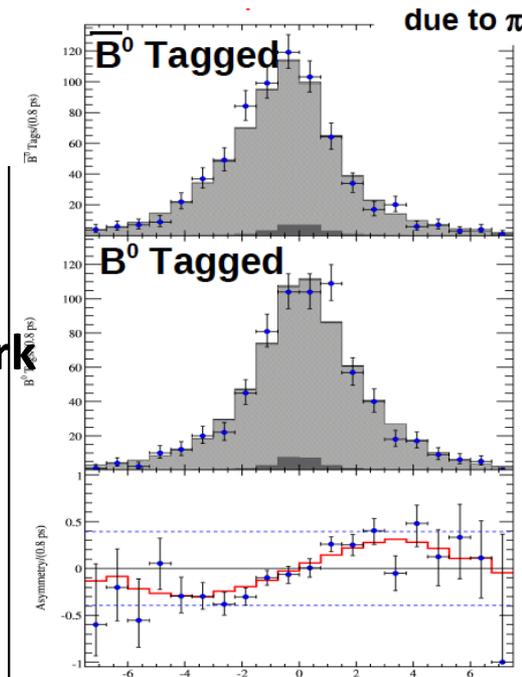
decay time information

...

signal composition in terms of amplitudes and phases

...

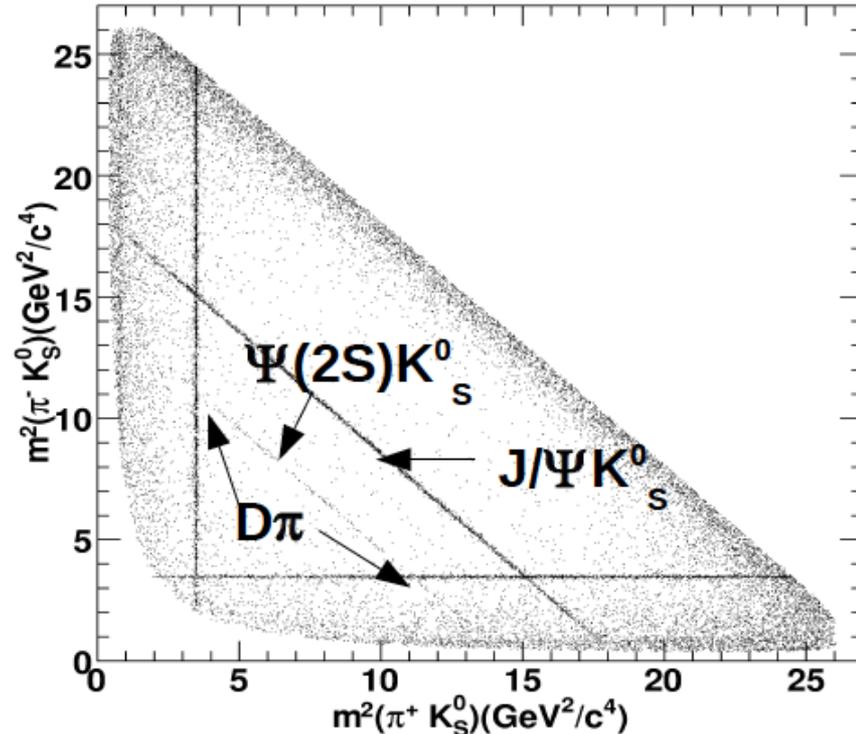
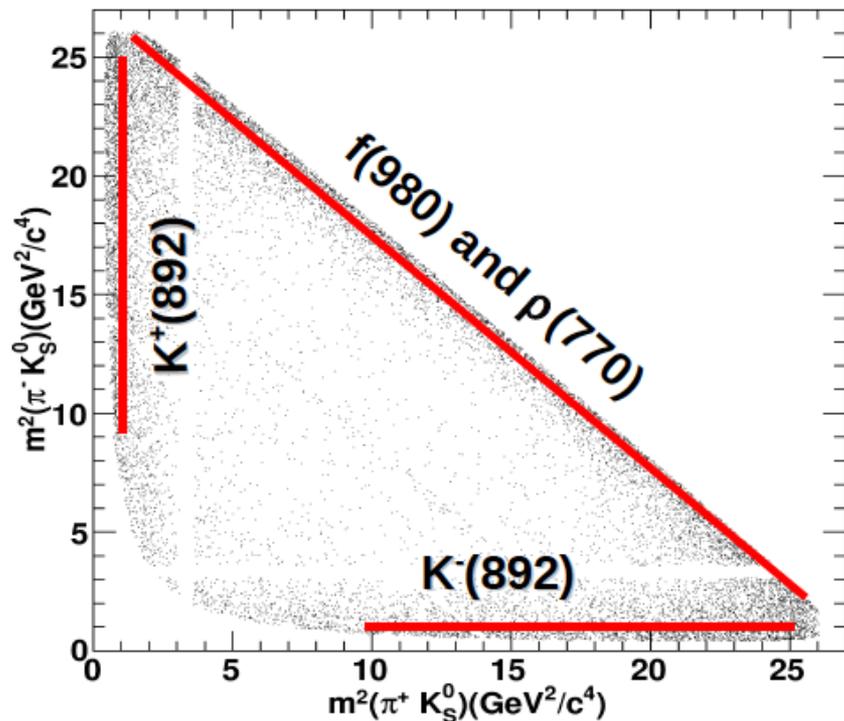
and perform a blind analysis !



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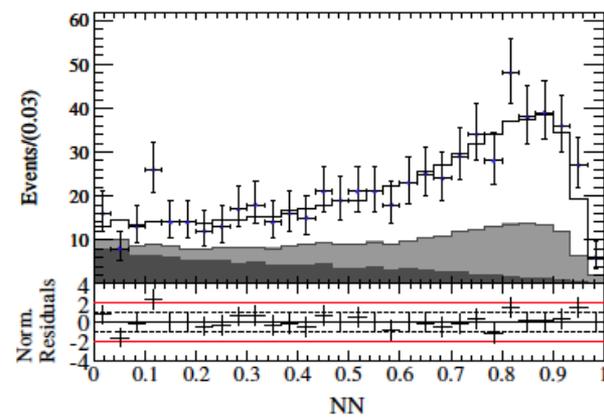
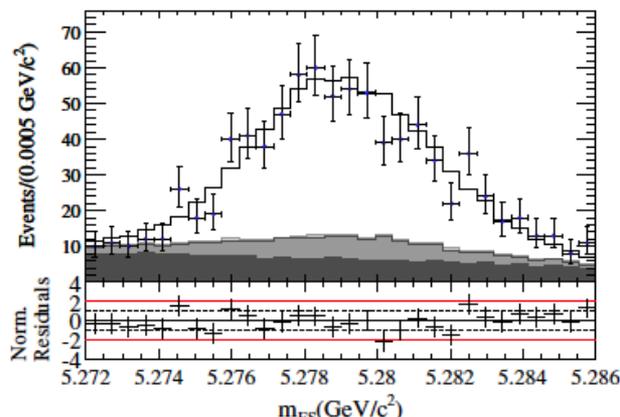
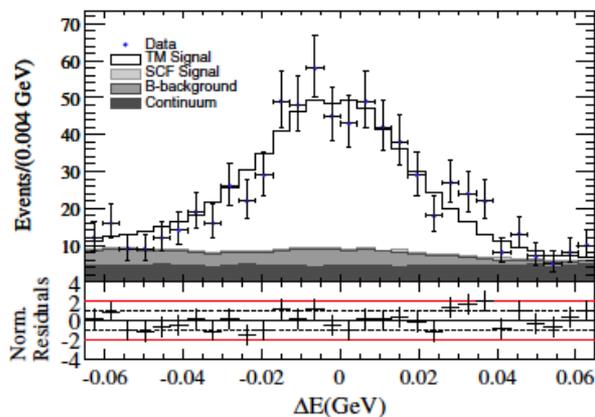
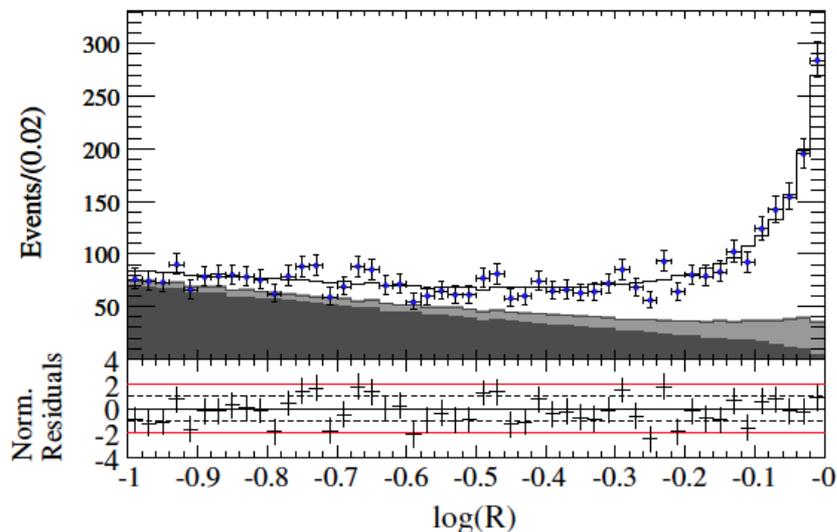
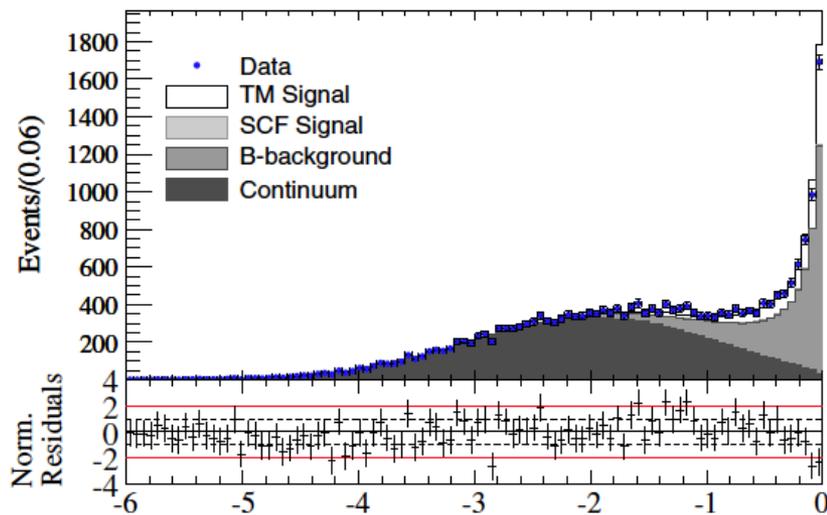
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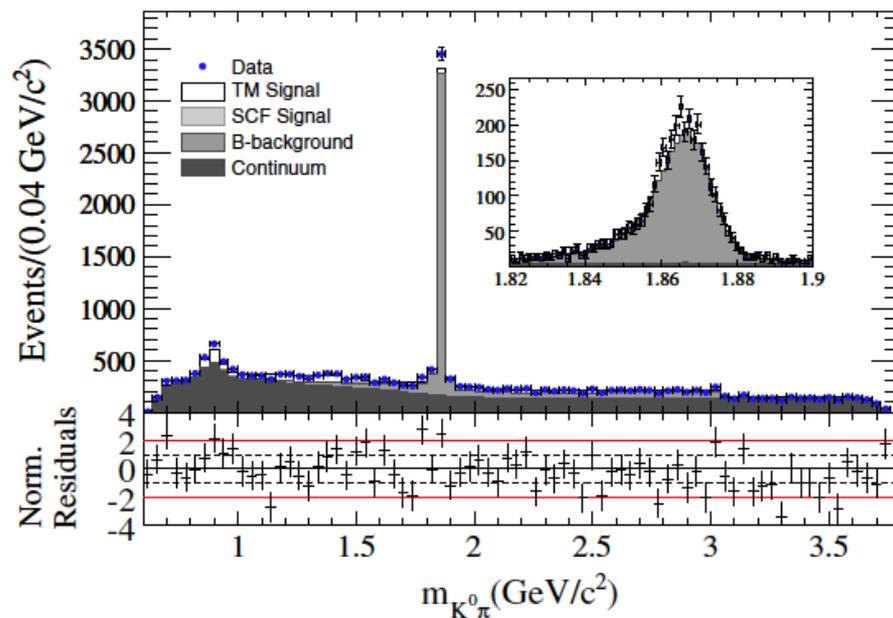
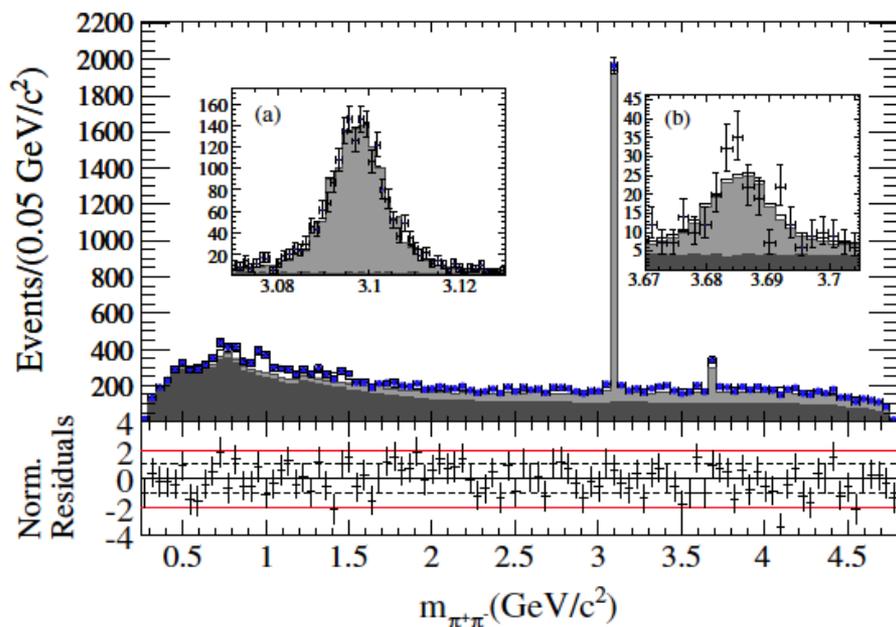
- Blind analysis : converge on all analysis criteria using control samples
- Goodness-of-fit studies : Likelihood ratio projections



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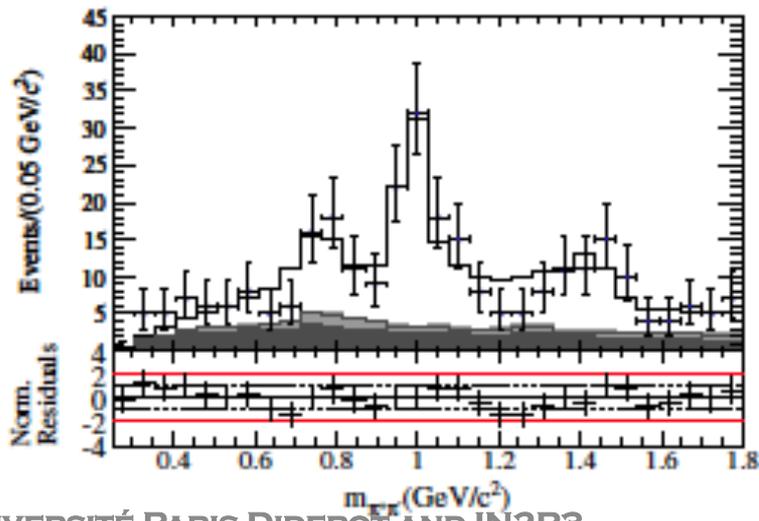
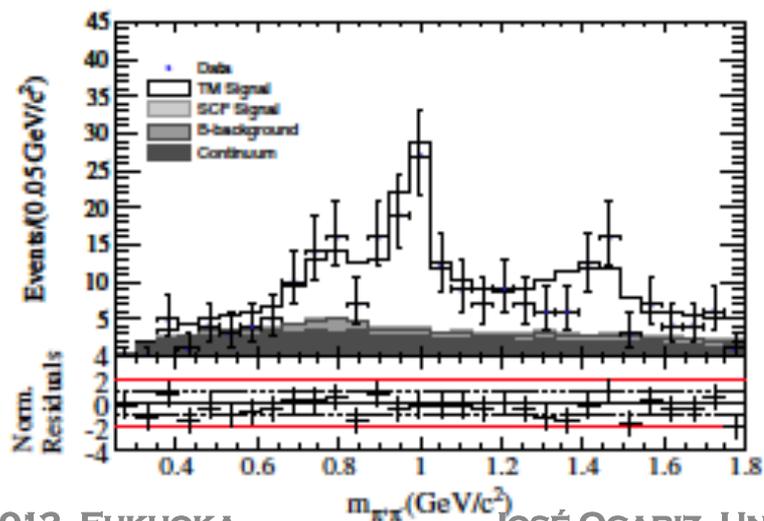
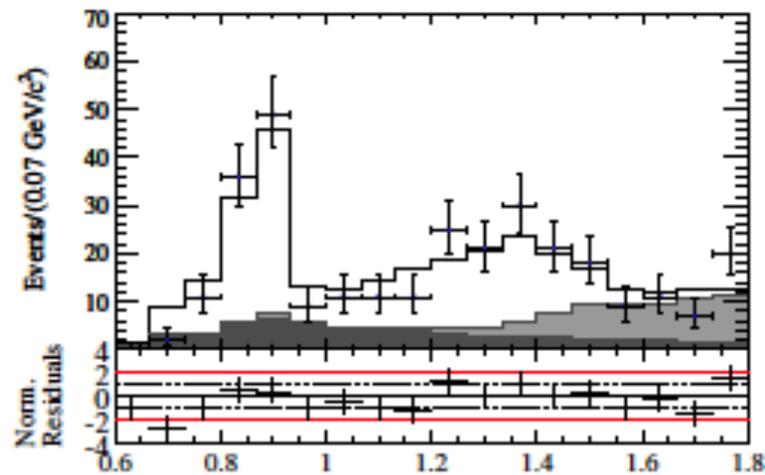
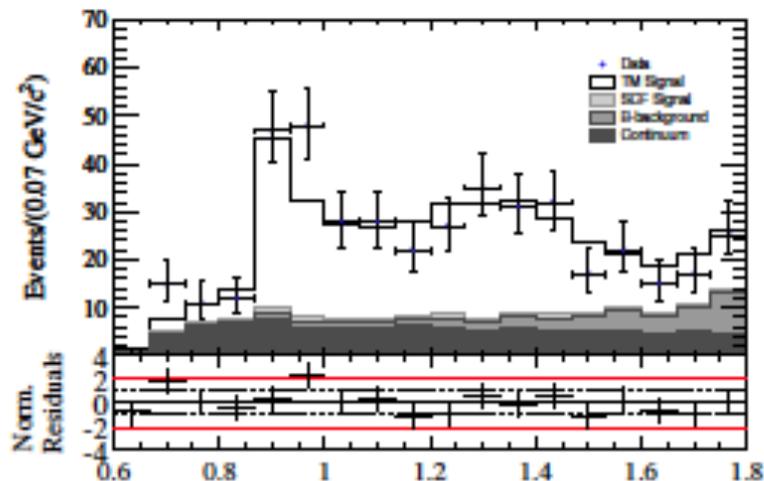
- Goodness-of-fit studies : Likelihood ratio projections : zoom on control samples



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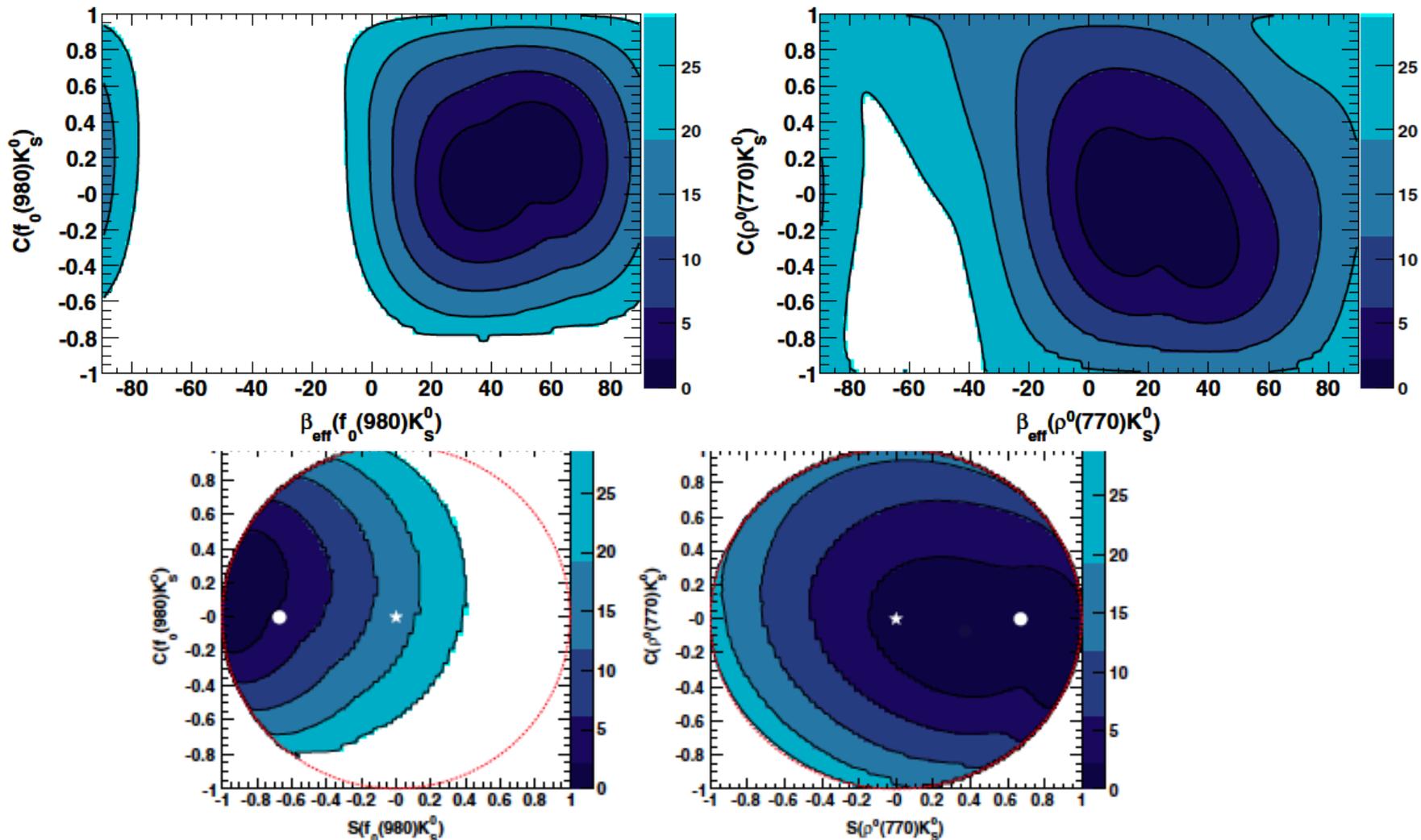
- Goodness-of-fit studies : Likelihood ratio projections on signal-enriched areas



A COMPLETE ANALYSIS EXAMPLE FROM BABAR

Amplitude analysis of neutral B mesons decaying into 3 particles : $B^0 / \bar{B}^0 \rightarrow K_s^0 \pi^+ \pi^-$

- Results : Likelihood contours on physical observables



HYPOTHESIS-TESTING (I)

Hypothesis (in physicist-oriented wording) : a model attempting at describing data

Some standard definitions :

- H_0 : the “null” hypothesis (e.g. background-only model)
- H_1 : the “alternative” hypothesis (e.g. signal-plus-background model)
- Hypothesis-testing :
 - build a statistics μ , define a confidence interval W , measure $\hat{\mu}$
 - if inside (outside) W , accept (reject) H_0
- Type-I error : false positive, reject H_0 despite being true (inefficiency)
 - α is the rate of Type-I error
 - α is also called “size”
- Type-II error : false negative, accept H_0 despite being false (fake-rate)
 - β is the rate of Type-II error
 - $1-\beta$ is also called “power”
- A good statistics : both α and β small (usually conflicting)
- Neyman-Pearson : at a fixed size, the *Likelihood Ratio* λ is optimal

$$\lambda(\mu) = \frac{\mathcal{L}(\mu | H_0)}{\mathcal{L}(\mu | H_1)} > \kappa(\mu)$$

beware of underlying assumptions ...

HYPOTHESIS-TESTING : P-VALUES

Statistical significance :

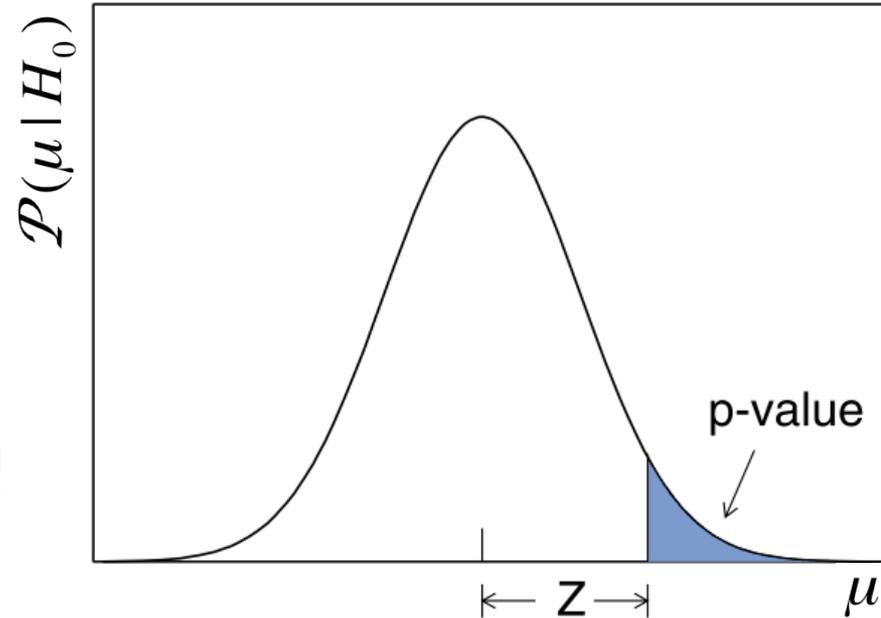
probability within H_0 of μ being larger than the observed $\hat{\mu}$ value

$$\text{p-value} = \int_{\hat{\mu}}^{\infty} d\mu \mathcal{P}(\mu | H_0)$$

Popular preference :

quote the p-value in terms of “sigmas”,

$$\text{p-value} = \int_{n\sigma}^{\infty} dx \frac{1}{\sqrt{2\pi}} e^{-x^2/2} = 1 - \frac{1}{2} \text{erf}\left(\frac{n}{\sqrt{2}}\right)$$



Usual benchmarks are asymmetric for exclusion and discovery :

- require $p < 0.05$ (i.e. 1.64σ , or 95% C.L.) applied to $H_0 = \text{sig} + \text{bkg}$
 - excluding a signal hypothesis
- require $p < 1.35 \times 10^{-3}$ (i.e. 3σ) applied to $H_0 = \text{bkg-only}$
 - “evidence” benchmark
- require $p < 2.87 \times 10^{-7}$ (i.e. 5σ) applied to $H_0 = \text{bkg-only}$
 - “observation” benchmark

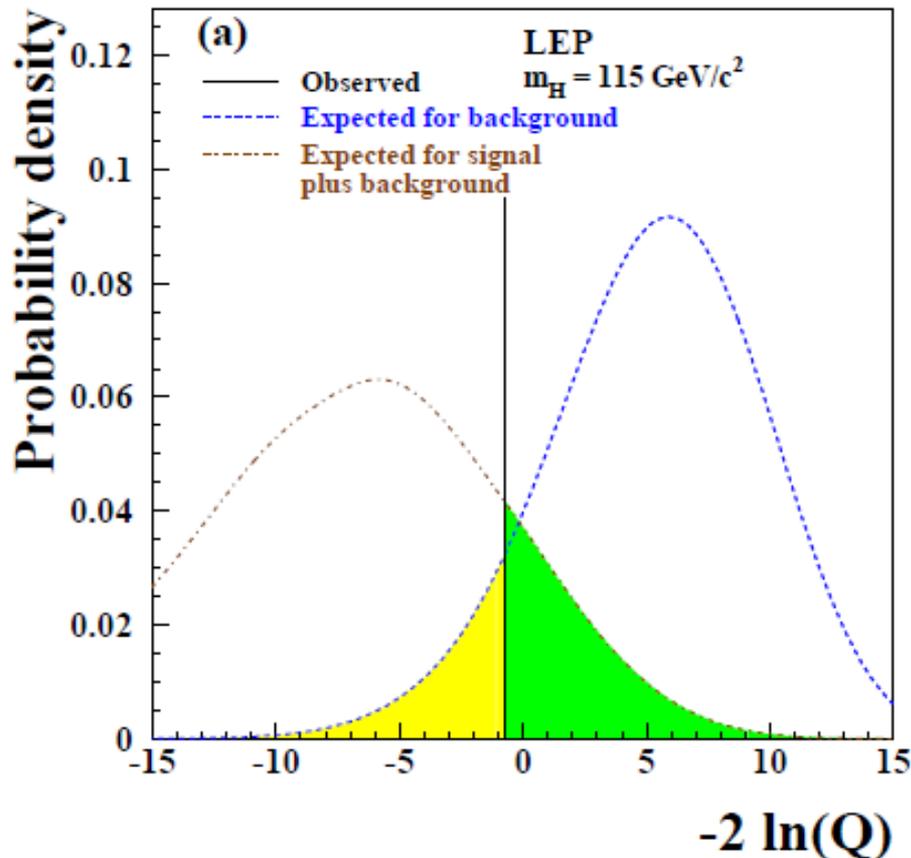
TEST STATISTICS : LEP AS AN EXAMPLE

$$Q = \frac{\mathcal{L}(H_1)}{\mathcal{L}(H_0)}, \text{ with}$$

$$\lambda = -2 \ln(Q)$$

$$\mathcal{L}(H_1) = \prod_{i=1}^{N_{ch}} \mathcal{P}_{\text{Poisson}}(n_i, s_i + b_i) \prod_{j=1}^{n_i} \frac{s_j \mathcal{S}(\vec{x}_j) + b_j \mathcal{B}(\vec{x}_j)}{s_j + b_j}$$

$$\mathcal{L}(H_0) = \prod_{i=1}^{N_{ch}} \mathcal{P}_{\text{Poisson}}(n_i, b_i) \prod_{j=1}^{n_i} \mathcal{B}(\vec{x}_j)$$



Standard definitions :

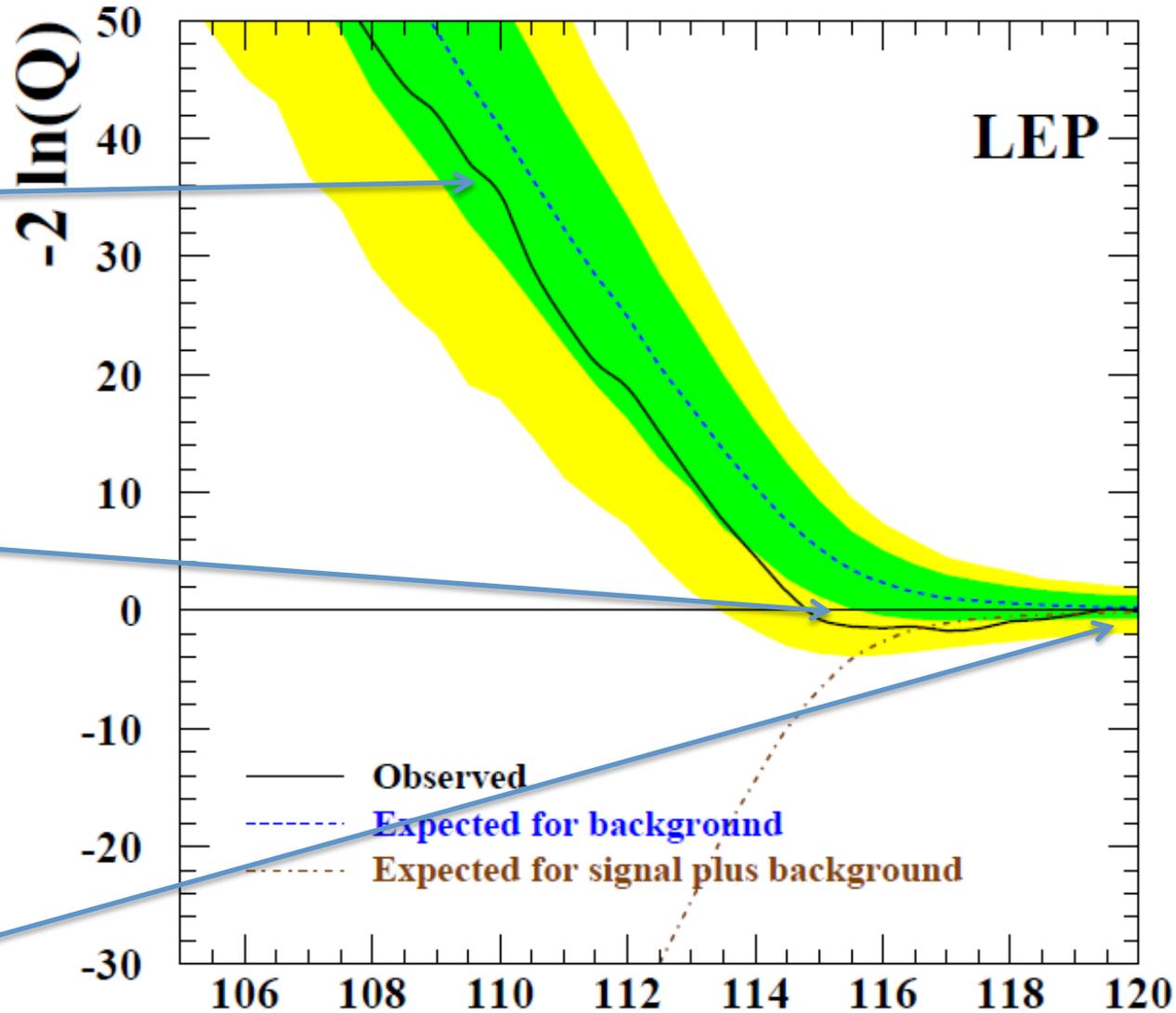
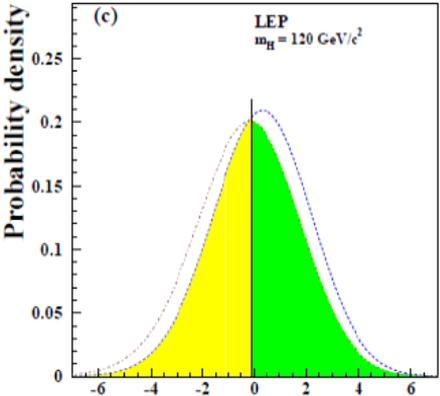
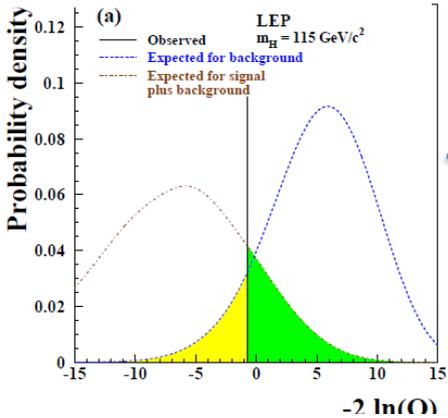
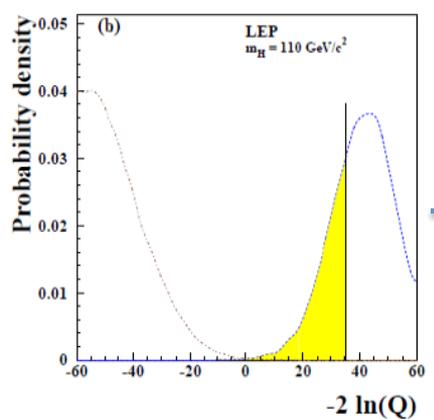
$CL(s+b)$: p-value under $H_0 = sig+bkg$
 probability of having sig+bkg yielding $-2\ln Q$
 larger than the observed one.

In the plot : green = $CL(s+b)$

$CL(b)$: p-value under $H_0 = bkg\text{-only}$
 probability of having bkg yielding $-2\ln Q$
 larger than the observed one.

In the plot : yellow = $1 - CL(b)$

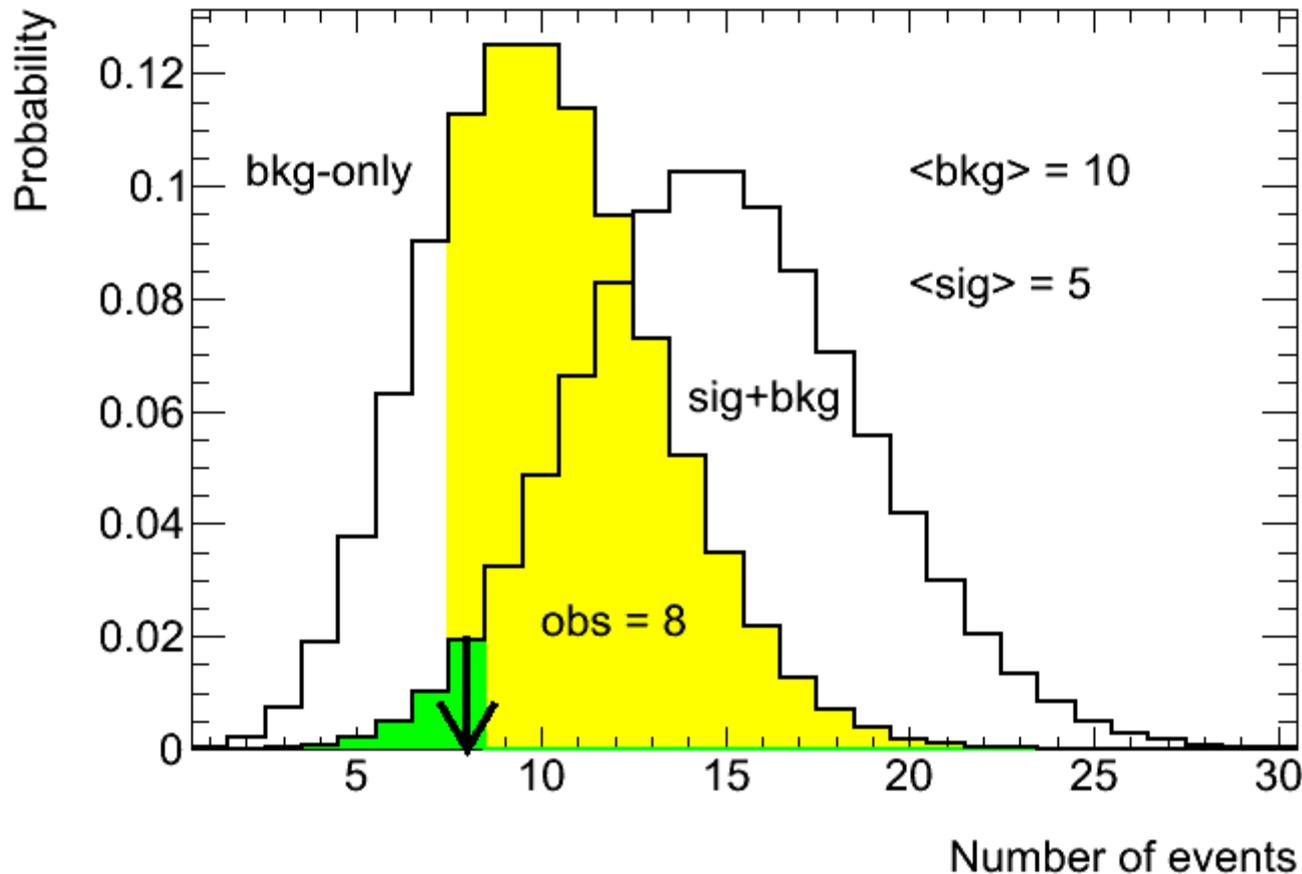
TEST STATISTICS : LEP AS AN EXAMPLE



Beware of color conventions ☺

$m_H (\text{GeV}/c^2)$

TEST STATISTICS : ANOTHER EXAMPLE



$$CL(s+b) = 3.7 \%$$

$$1-CL(b) = 33 \%$$

$$CL(s) = 11 \%$$

(check ☺)

Clearly, both signal and background had a (not-so-large) negative fluctuation ...
to avoid a 95% exclusion, define

(beware : NOT a p-value!)

$$CL(s) = \frac{CL(s+b)}{1-CL(b)}$$

PROFILE LIKELIHOOD REVISITED

Likelihood Ratio defined previously

- allows to extract POI's μ out of variables in sample $\{x\}$
- did not include nuisances

Profile Likelihood approach : Introduce a control sample $\{y\}$ to constrain nuisances :

$$\mathcal{L}(\vec{x}, \vec{y} | \mu, \theta) = \mathcal{L}(\vec{x} | \mu, \theta) \mathcal{L}(\vec{y} | \theta)$$

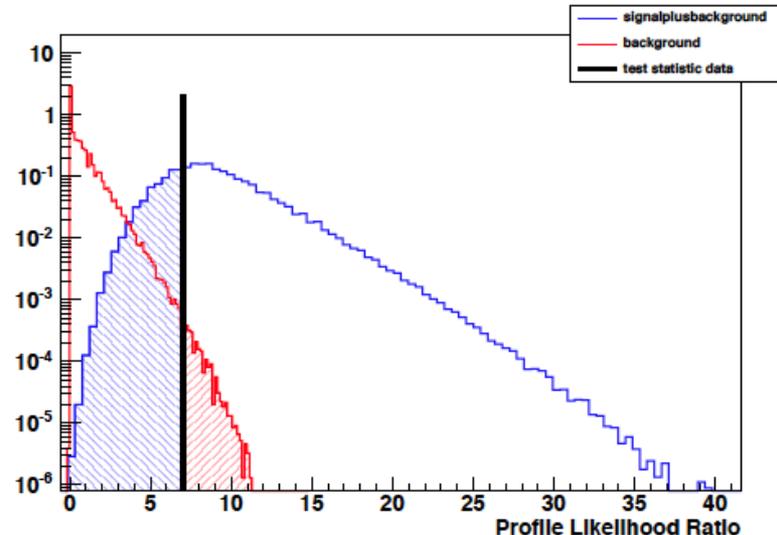
(n.b.: variables x and y , nor samples, need to be disjoint, i.e. sidebands...)

PL Test statistic: $q_{\mu}(\mu) = -2 \ln \lambda(\mu) = -2 \ln \frac{\mathcal{L}(\mu, \hat{\theta})}{\mathcal{L}(\hat{\mu}, \hat{\theta})}$ $\left(\begin{array}{l} \Rightarrow \text{fix } \mu, \text{ fit } \theta \\ \text{fit } \mu, \text{ fit } \theta \end{array} \right.$

tends to a χ^2 distribution with 1 ndof

p-value : $p(\mu^{OBS}) = \int_{q_{\mu}^{OBS}}^{\infty} dq_{\mu} \mathcal{P}(q_{\mu} | \mu)$

For complex likelihoods, distributions of q_{μ} via MonteCarlo integration: CPU-consuming...



A QUASI-ACADEMIC ILLUSTRATION : TESTING EXCLUSION

Explanatory figure (not actual data)

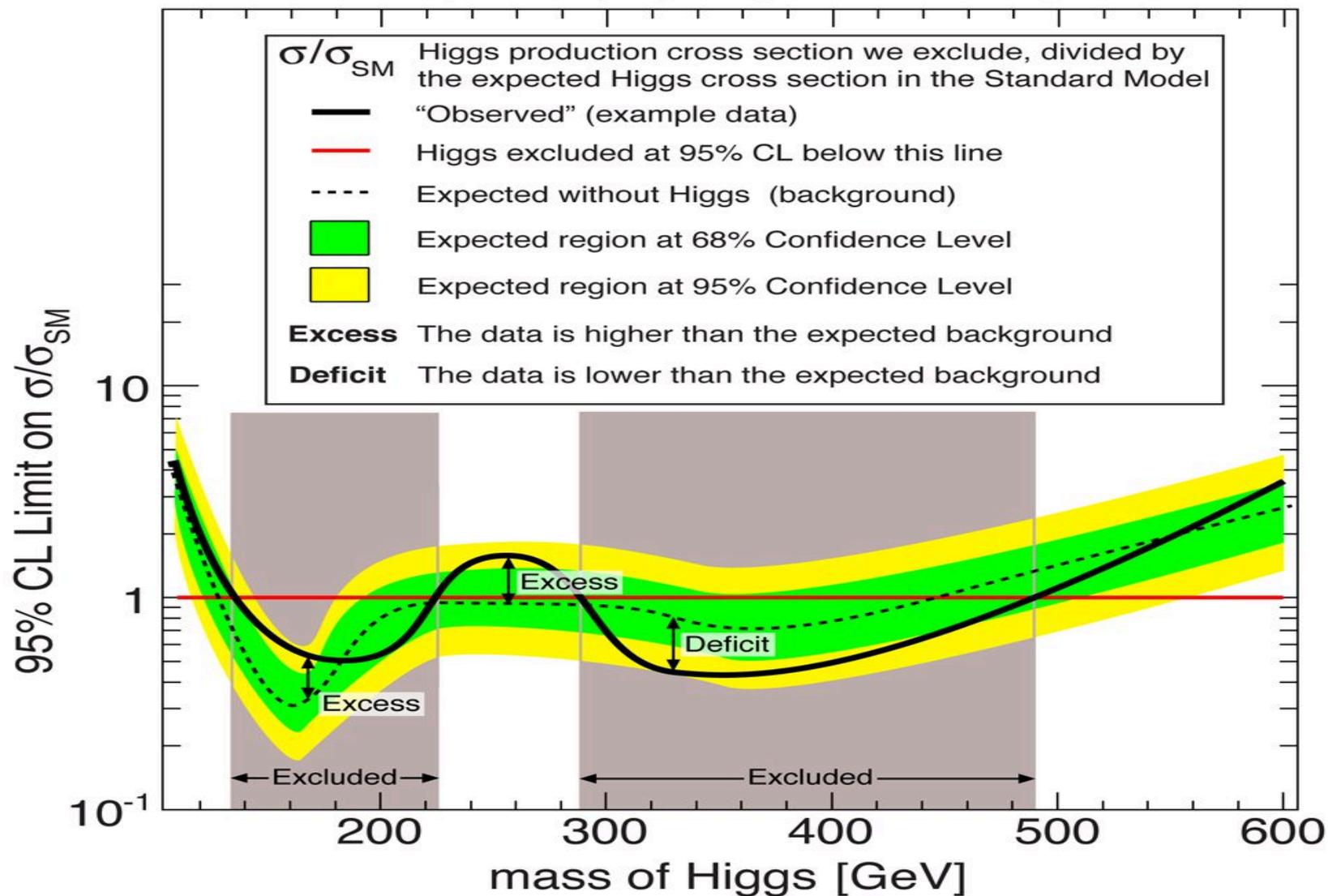
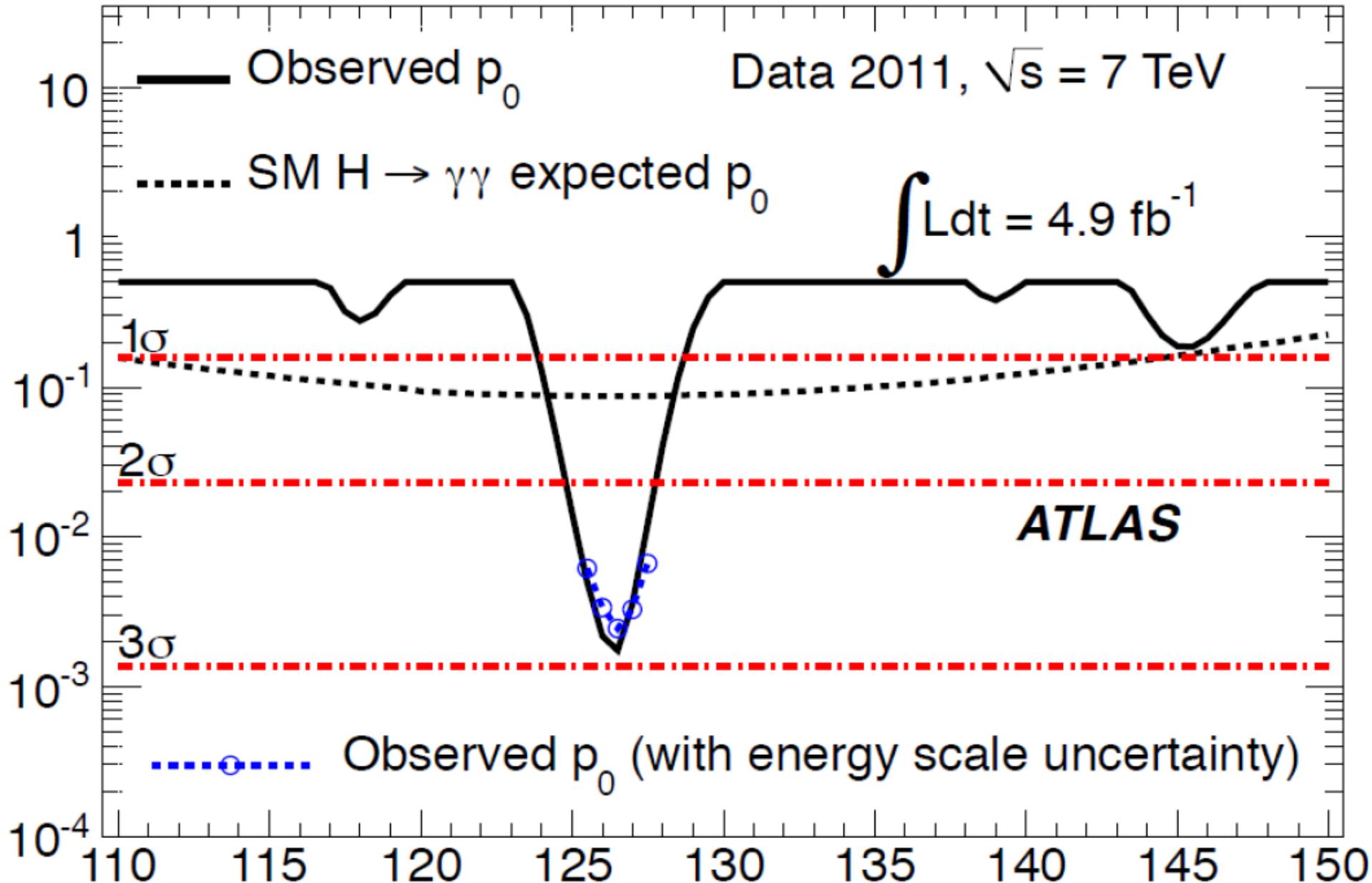


Figure A

LESS ACADEMIC ILLUSTRATION : TESTING OBSERVATION



ATLAS+CMS : PROFILE-LIKELIHOOD AND CL(S)

In practice, ATLAS and CMS agreed on using a modified version :

$$\tilde{q}_\mu(\mu) = -2 \ln \lambda(\mu) = -2 \ln \frac{\mathcal{L}(\mu, \hat{\theta})}{\mathcal{L}(\hat{\mu}, \hat{\theta})}, \quad 0 \leq \hat{\mu} \leq \mu$$

In fact it's slightly different, but this one is sufficient for our discussion here...

Constraint on $0 \leq \hat{\mu} \leq \mu$: upwards fluctuations (i.e. $\hat{\mu} > \mu$) should not be interpreted as a disfavor of signal hypothesis

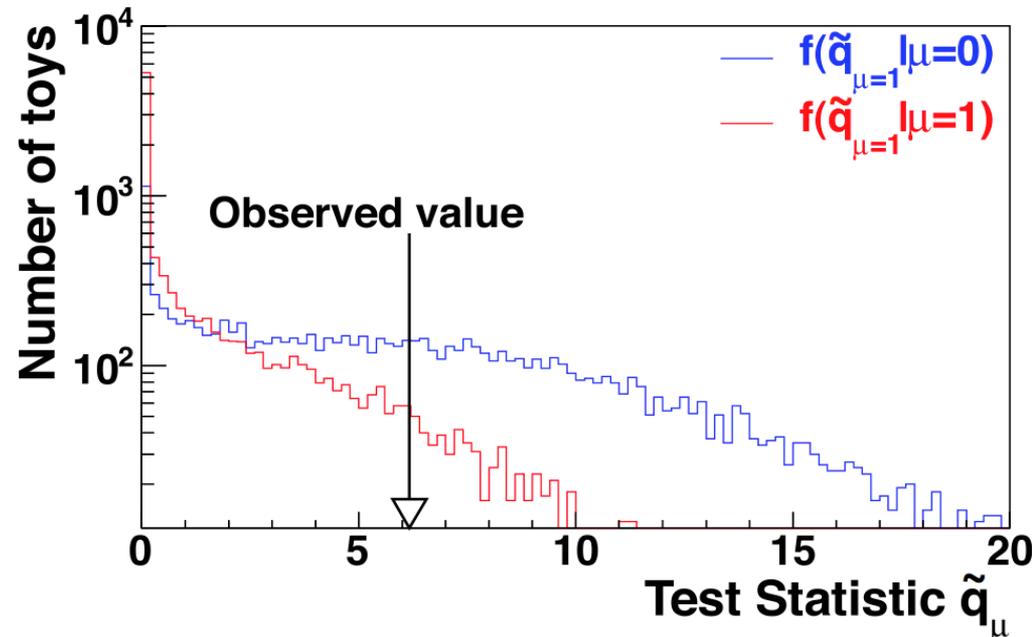
Also, ATLAS and CMS agreed on using

$$CL_s(\mu) = \frac{p_\mu}{1 - p_b}$$

Note : the modified test statistic

follows roughly a distribution $\frac{1}{2} \delta(\tilde{q}_\mu) + \frac{1}{2} \chi^2$

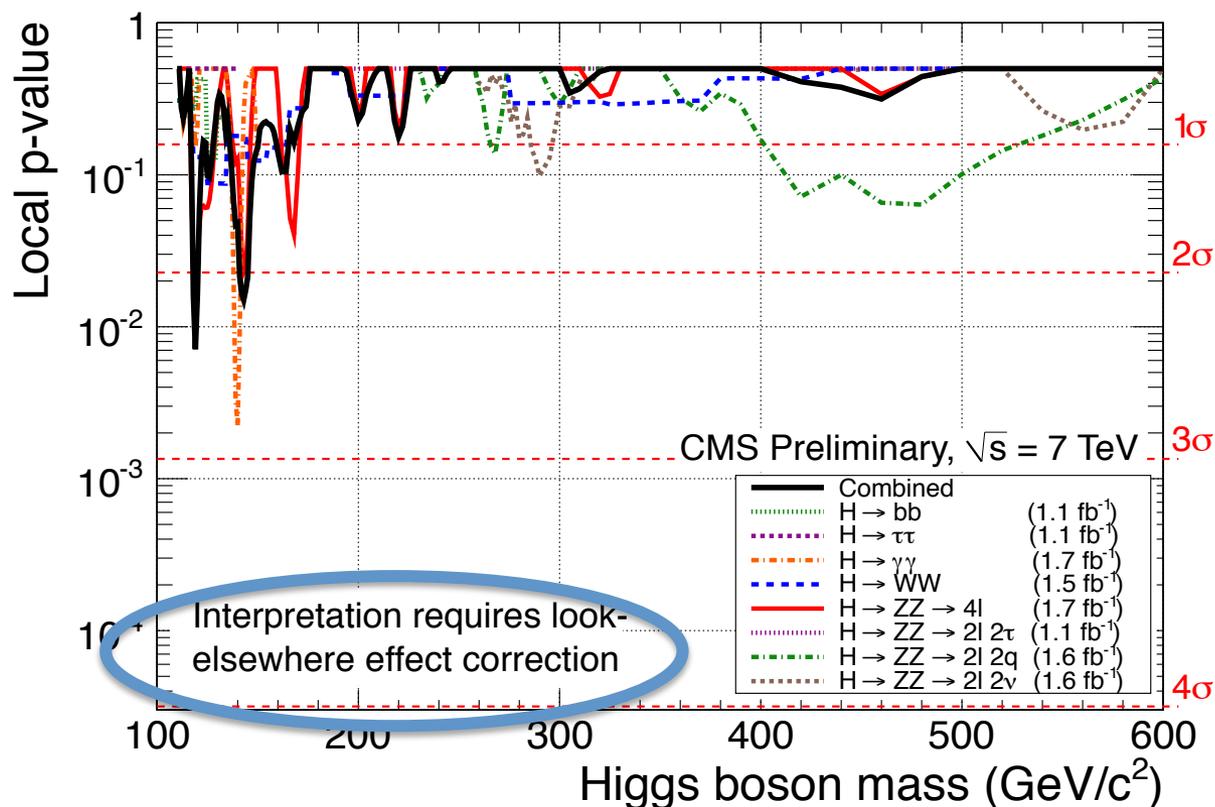
Useful asymptotic property, allows for CPU-saving approximations (i.e. the "Asimov" dataset)



THE “LOOK-ELSEWHERE-EFFECT”

Trials Factor : an apparently statistically significant hypothesis-testing conclusion

- may actually arise due to fluctuations
- should likely arise in cases of large populations or large parameter spaces
 - i.e. in a $\sim 1\text{K}$ sampling, an event showing a $\sim 3\sigma$ deviation is no surprise!
- To be specific : take one HEP-motivated scenario : the Higgs search



THE “LOOK-ELSEWHERE-EFFECT”

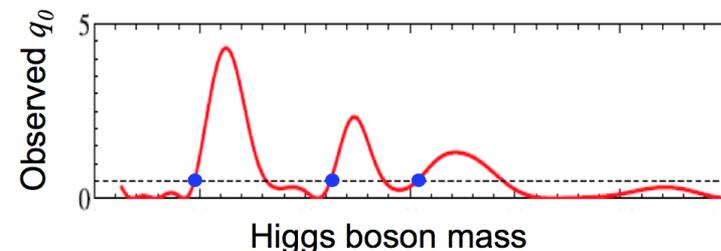
Scenario : apparently statistically significant hypothesis-testing conclusion

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The effect can be evaluated via MonteCarlo

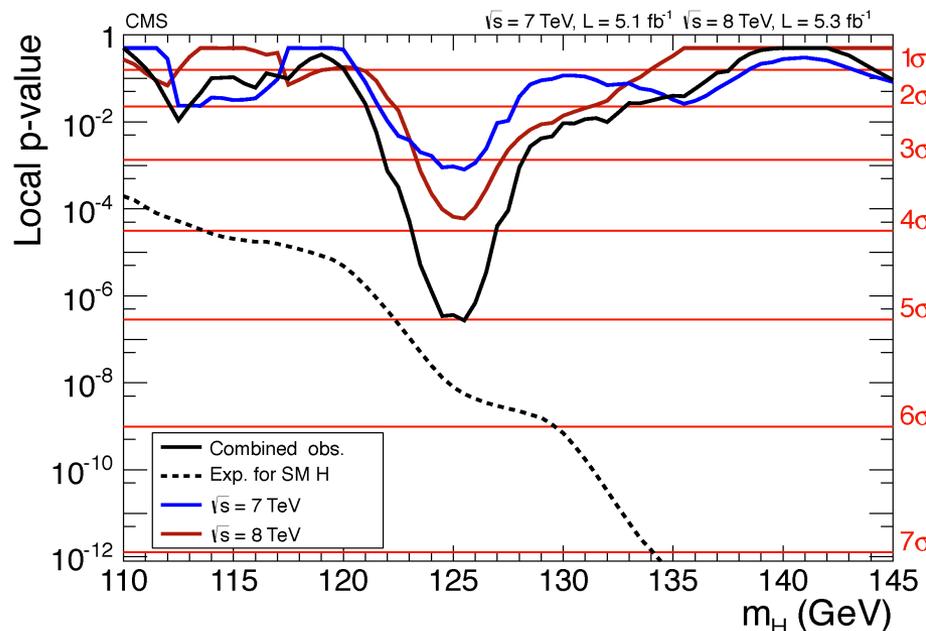
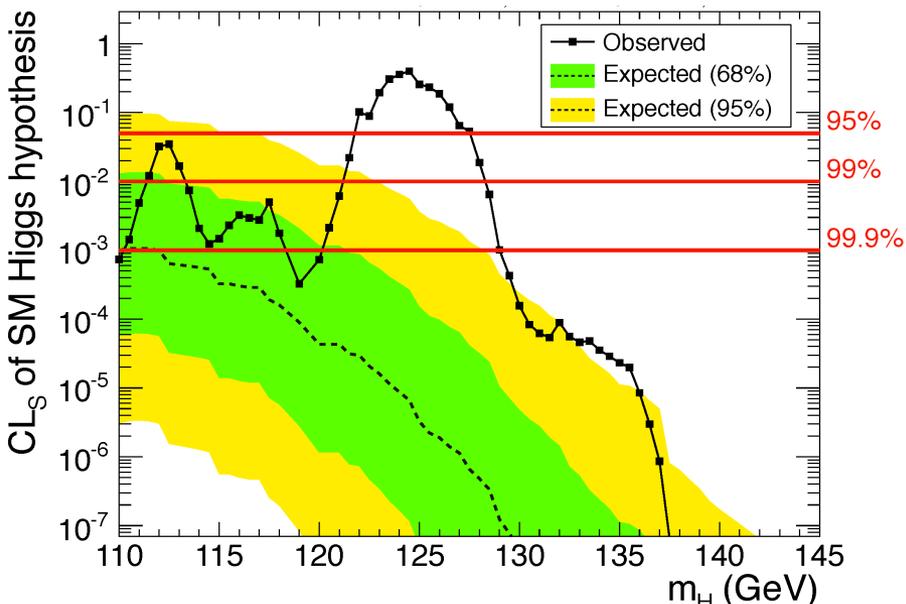
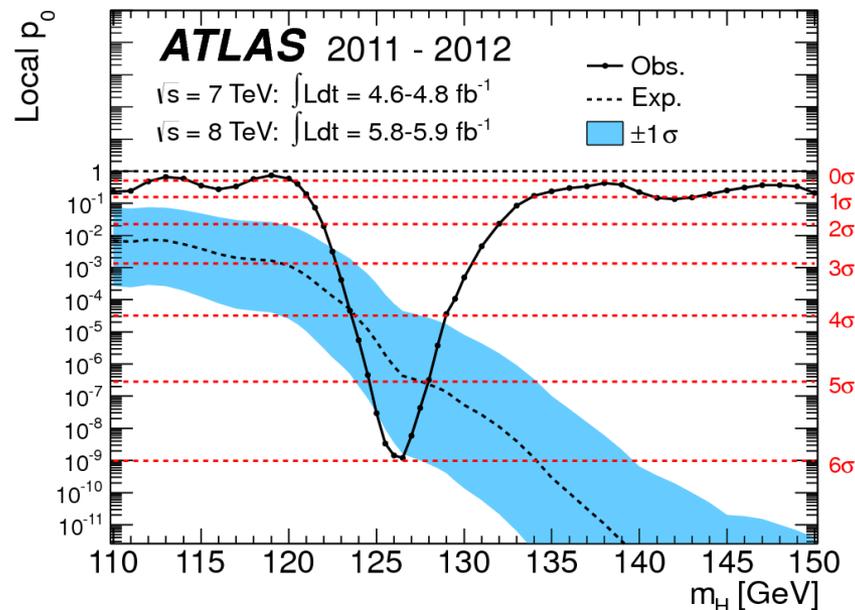
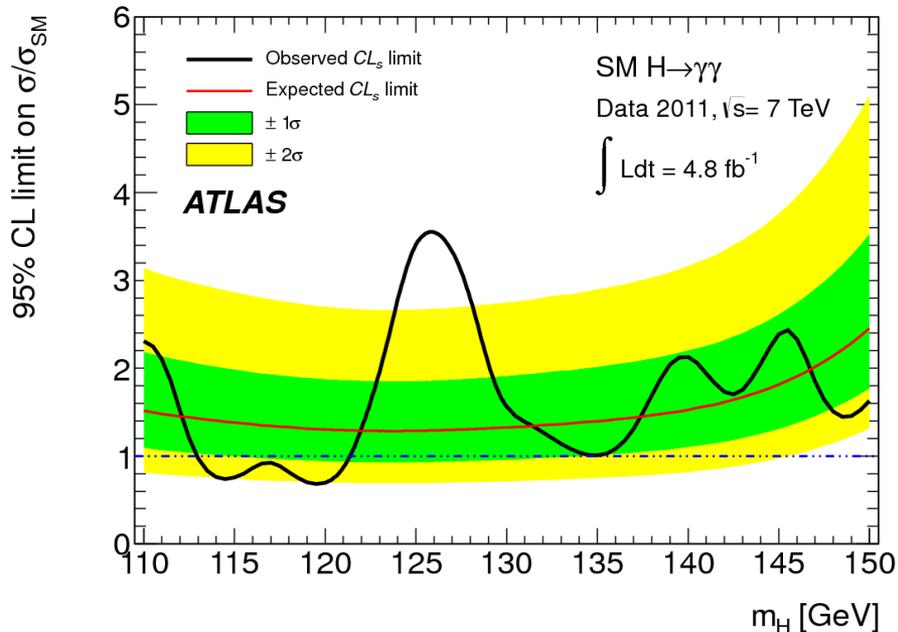
- generate simulated experiment outcomes out of a background-only PDF
- find the largest ‘local’ significance over the entire search range,
- extract the distribution of “local” significances

Can be approximated via asymptotic formula, and extrapolated from lower-statistics toy MonteCarlo

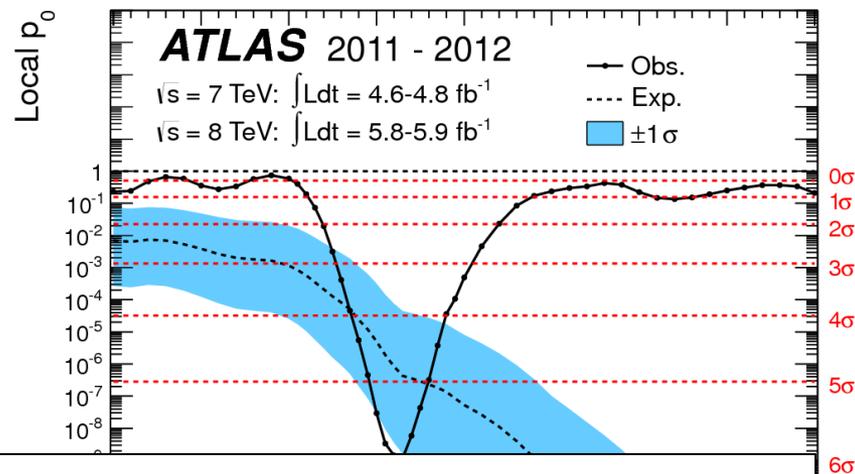
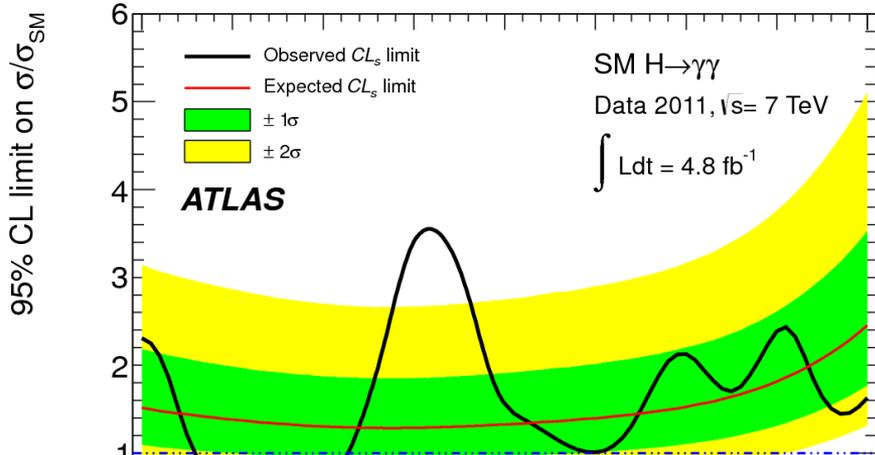


$$p_b^{global} = P(q_0(\hat{m}_H) > u) \leq \langle N_u \rangle + \frac{1}{2} P_{\chi_1^2}(u)$$

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THE END 😊

