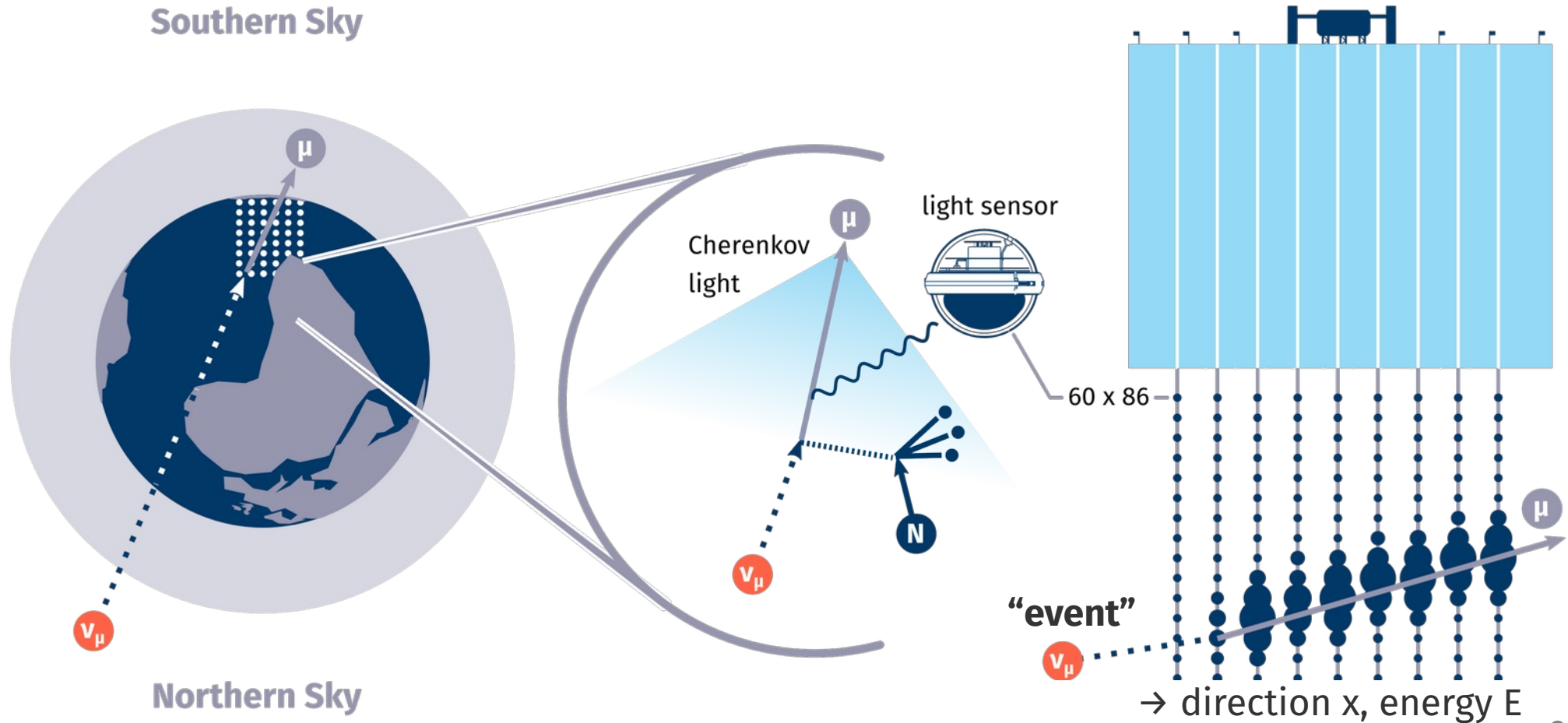




HOW TO SEARCH FOR
NEUTRINO SOURCES
(SIMPLIFIED VERSION)

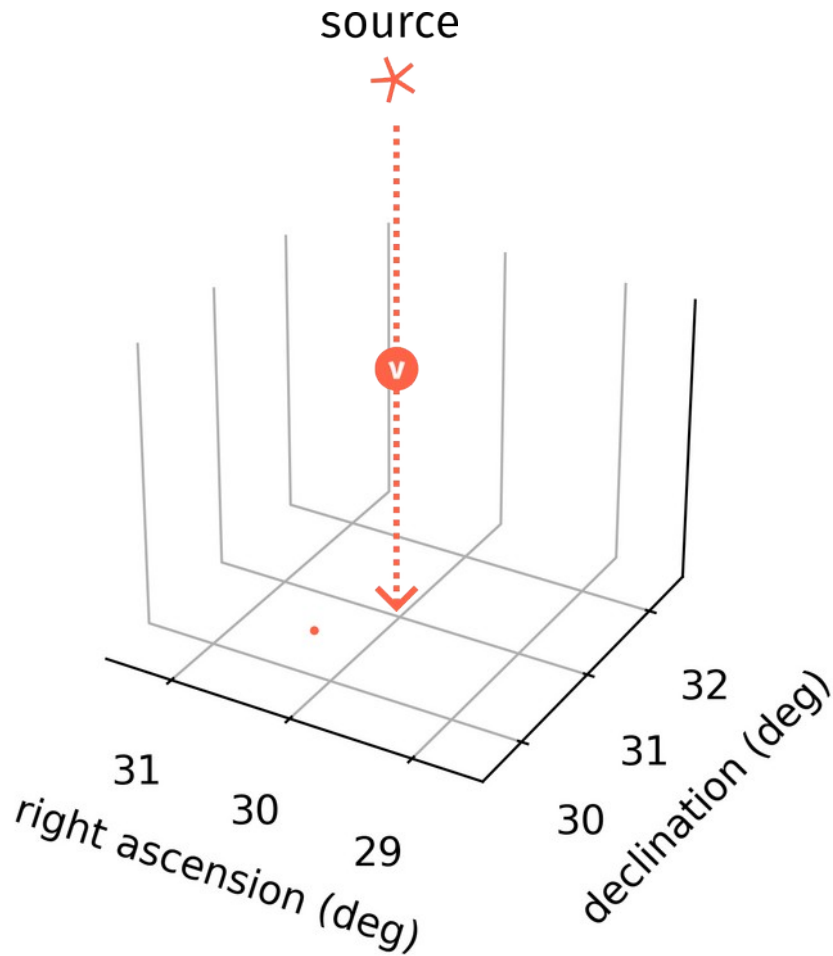
Detecting a neutrino signal with IceCube



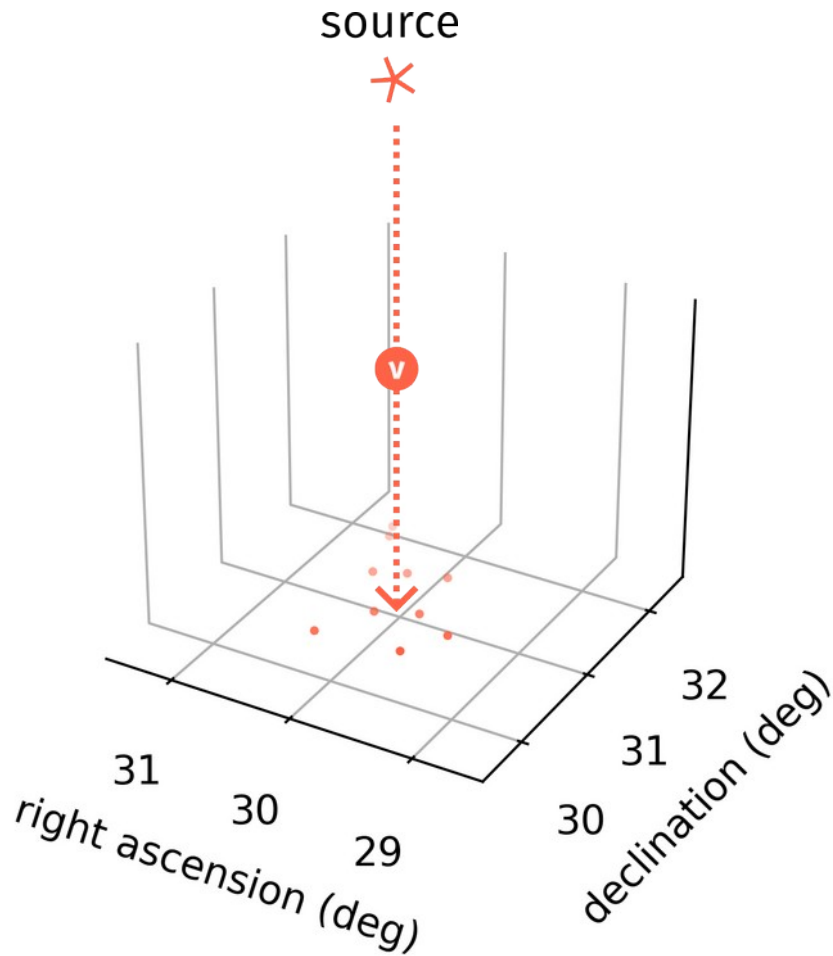
source



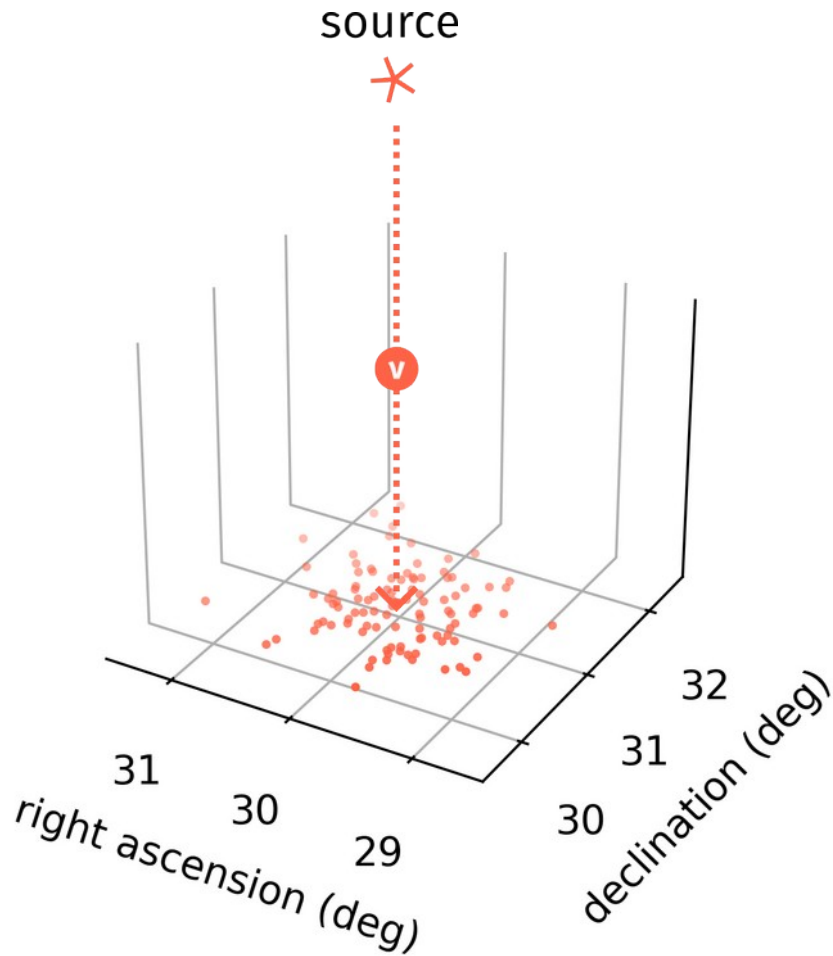
- Simulate (neutrino) events from a source
- Reconstructed directions are random
- More neutrinos
 - denser around the source
 - density as function of direction
 - reveals the **point spread function** (PSF)
- Can be approximated with
 - Kent distribution
(Gaussian on a sphere)
 - and a width parameter σ
(actually per event not per source)



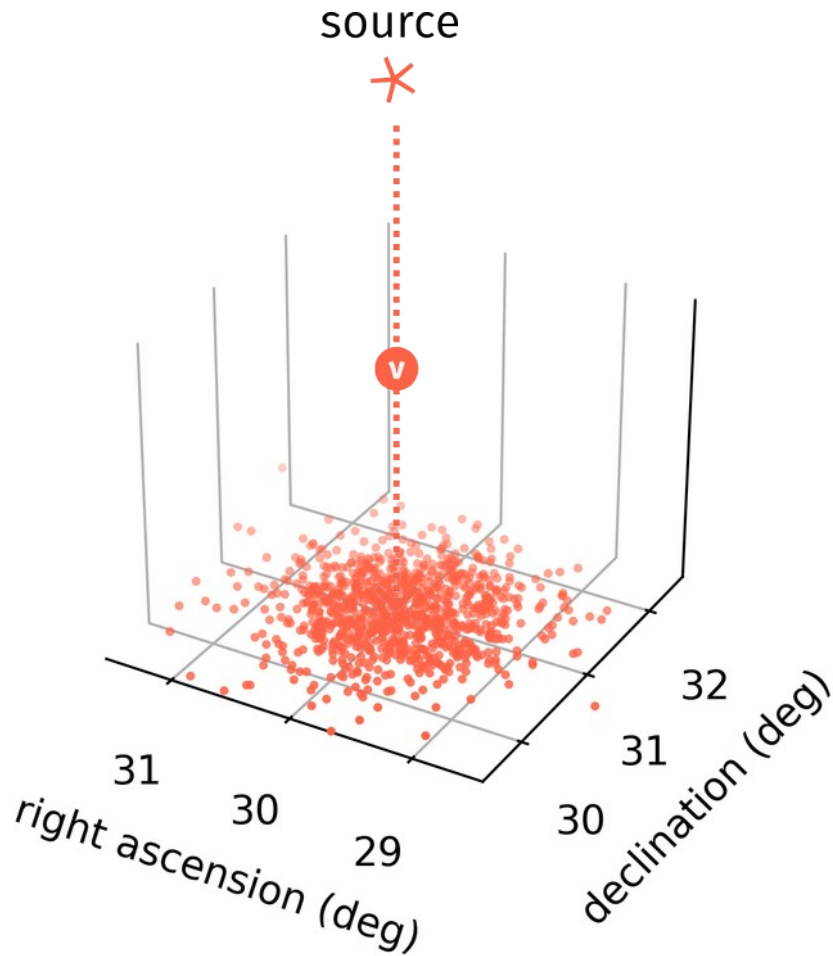
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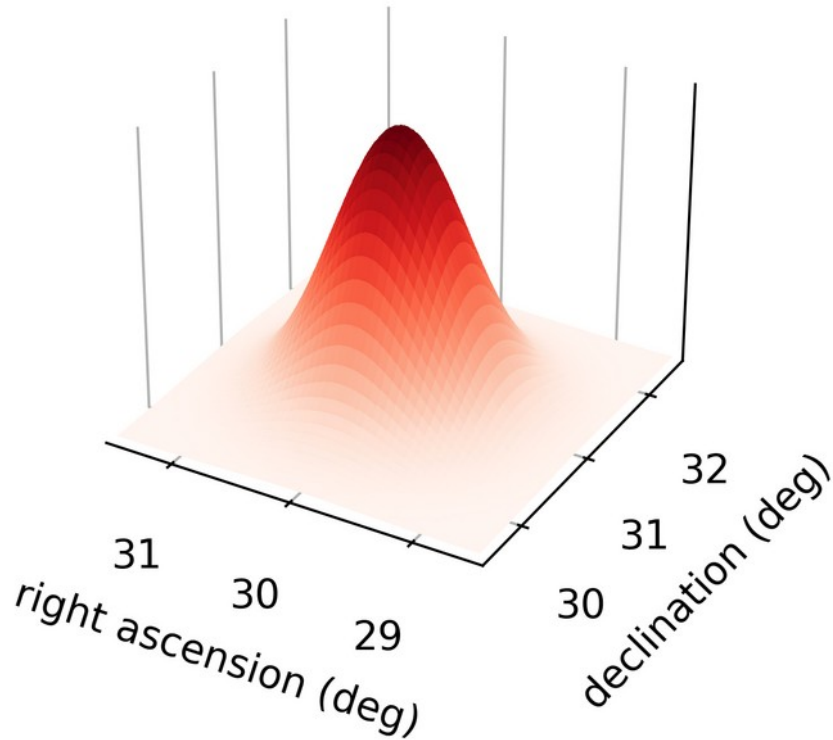
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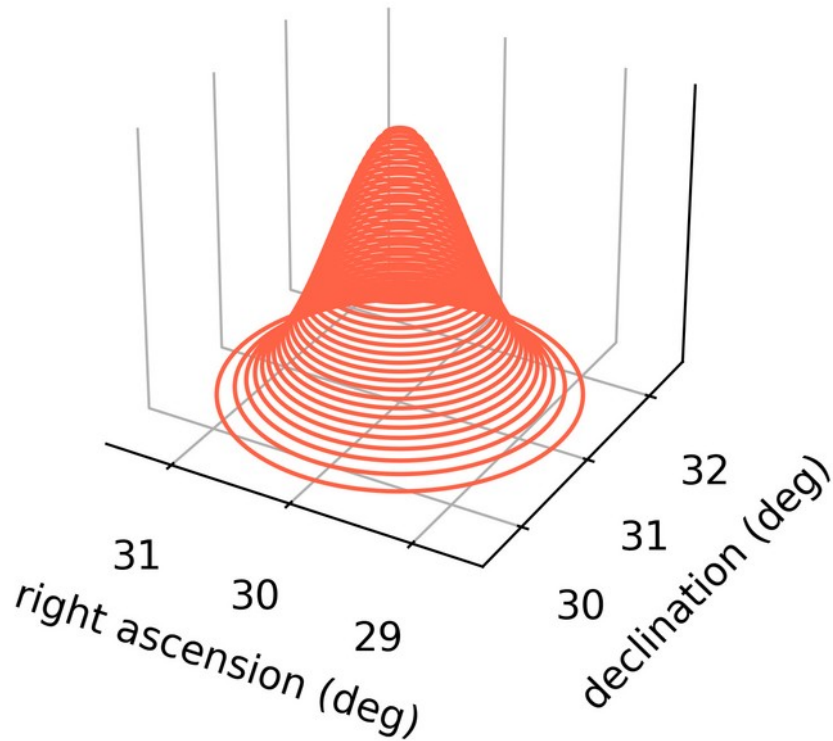
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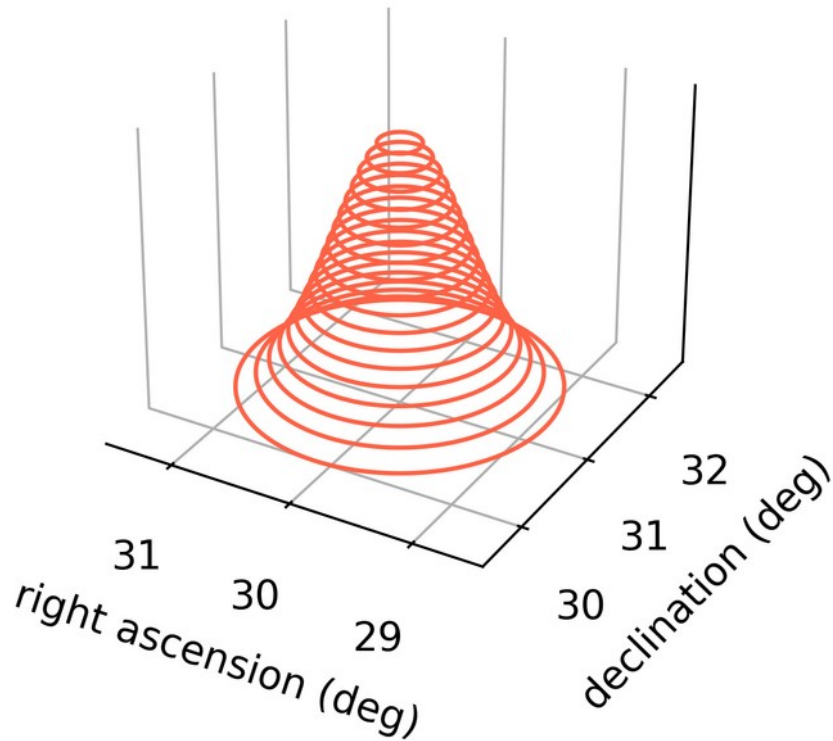
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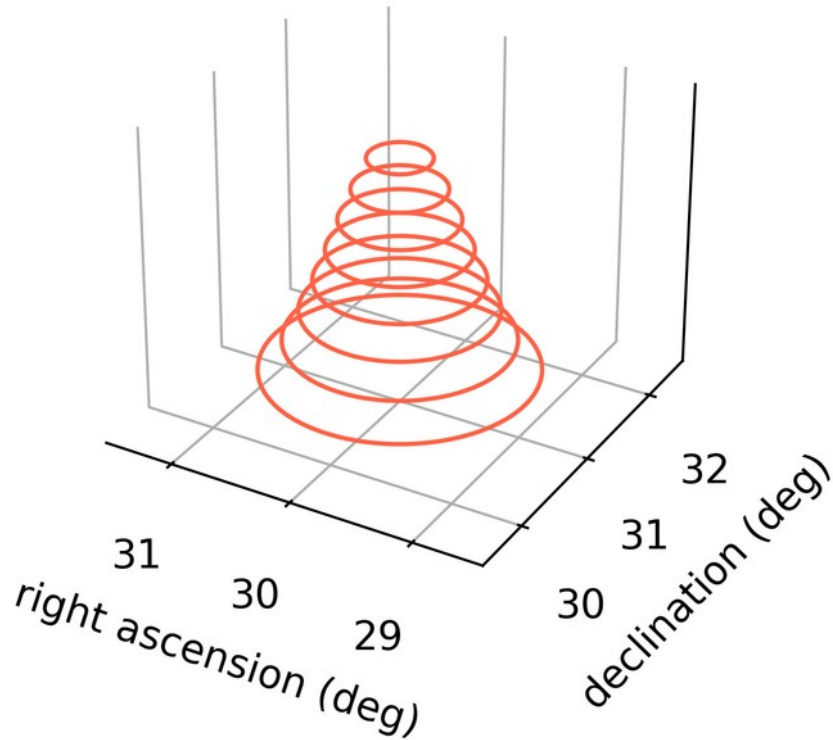
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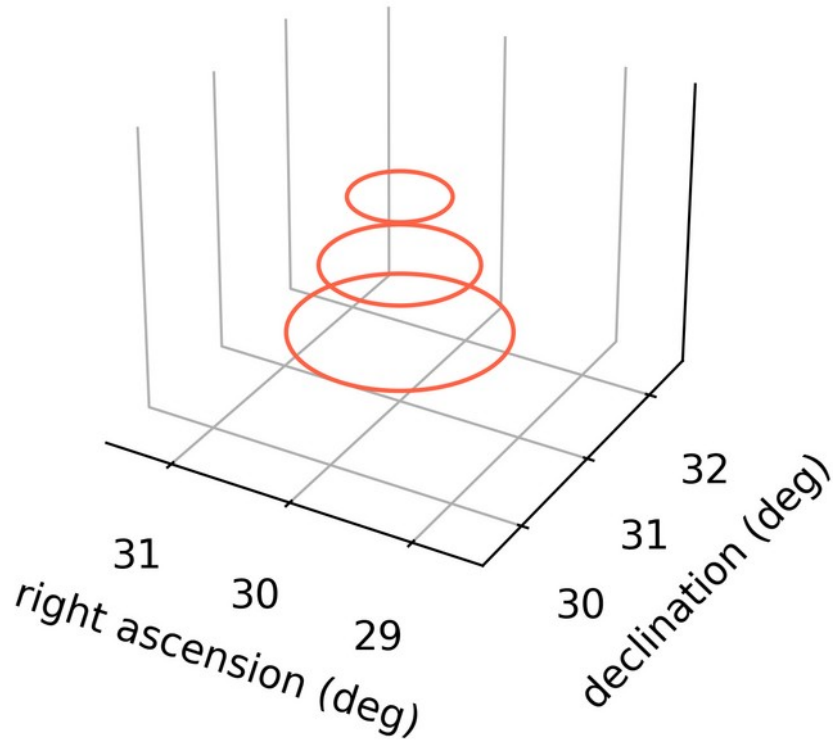
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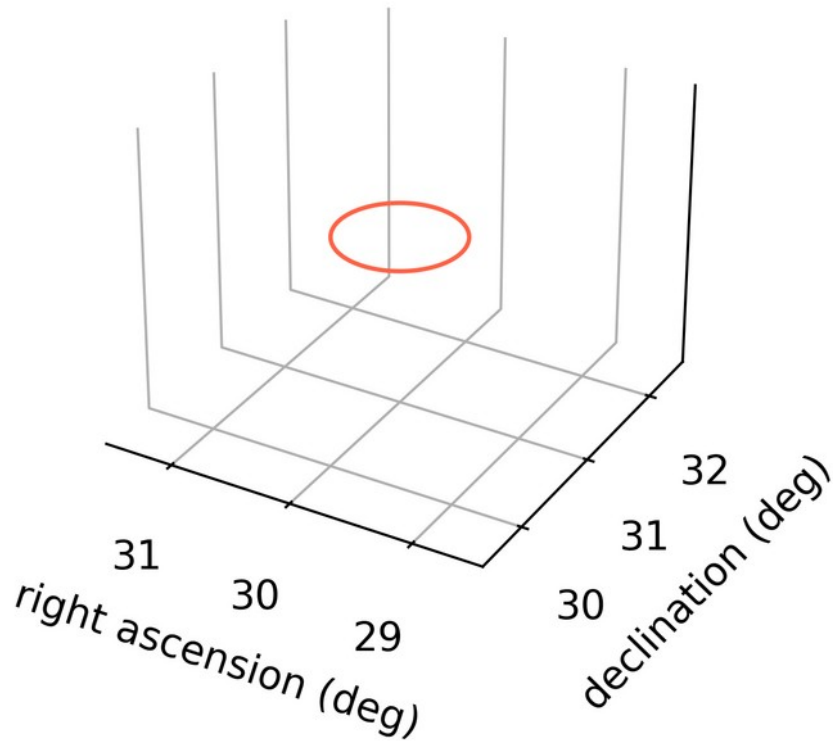
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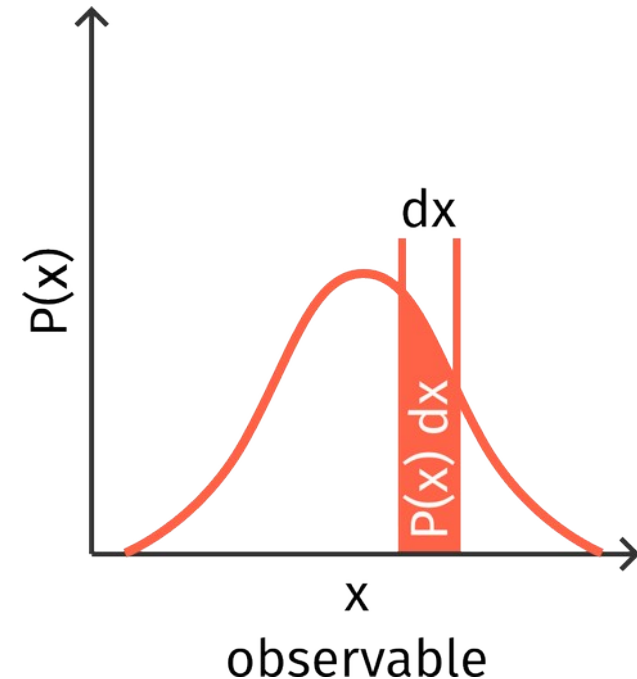


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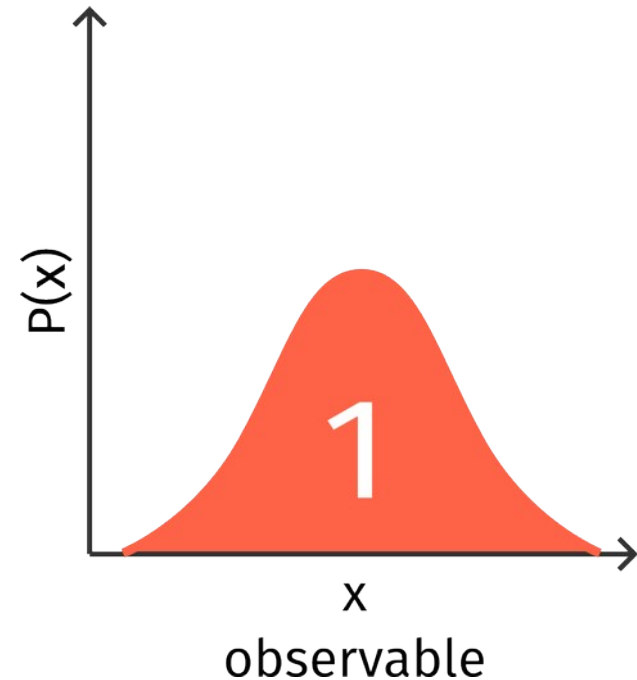


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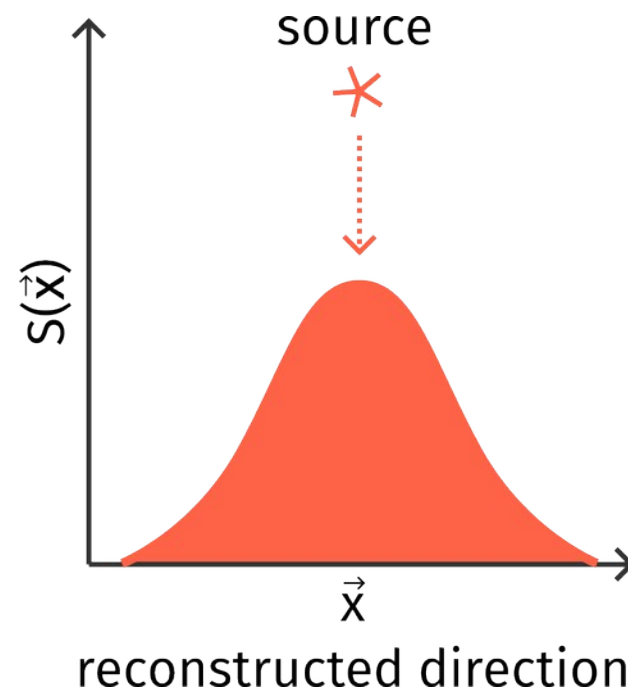
- A probability density function (PDF) P
 - describes the outcome of observable x
 - (during one event)
 - $\text{Prob}(\text{“event arrives in } dx\text{”}) = P(x) dx$
 - $\int P(x) dx = 1$
- Normalize PSF over direction $\mathbf{x} = (\text{RA}, \text{dec})$
 - centre on a source
 - signal PDF $\mathbf{S}(\mathbf{x})$ for that source
(probability density to be reconstructed at x)
- $S(x; \sigma) = \text{Kent}(\Psi; \sigma)$
with $\Psi = \text{angle}(\text{event } x, \text{source})$



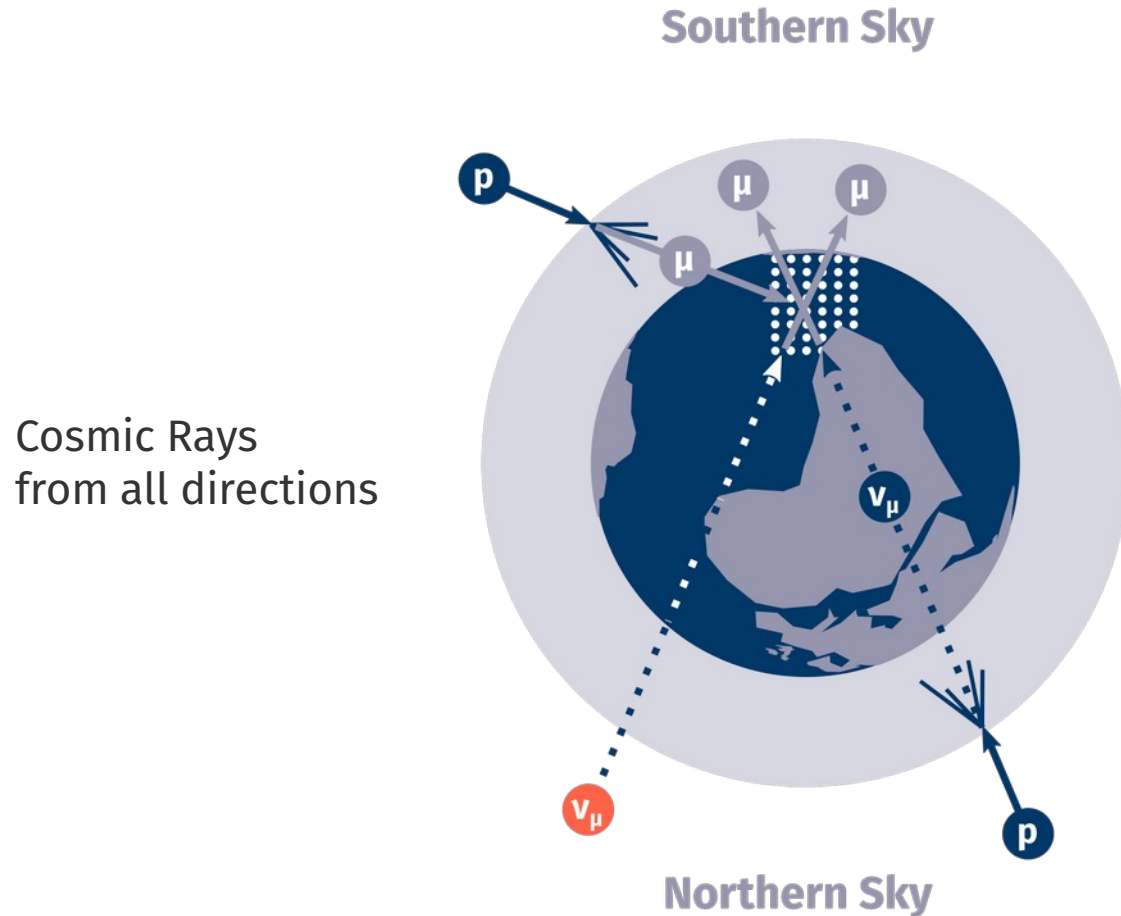
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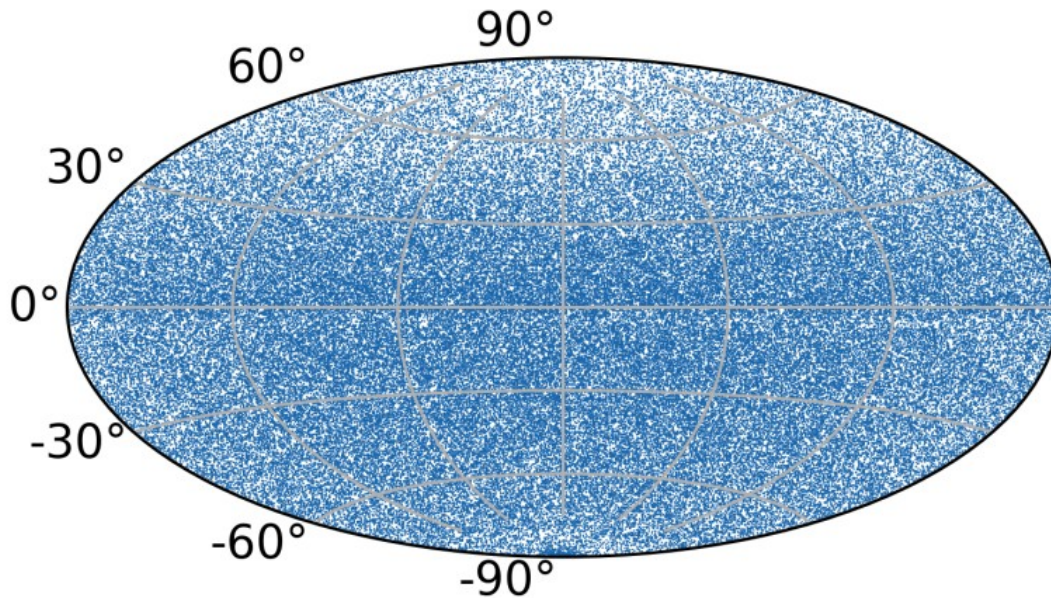
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- atmospheric muons (can reduce with event selection)
- atmospheric neutrinos (look like signal events)

→ individual events look like signal
this is our **background**

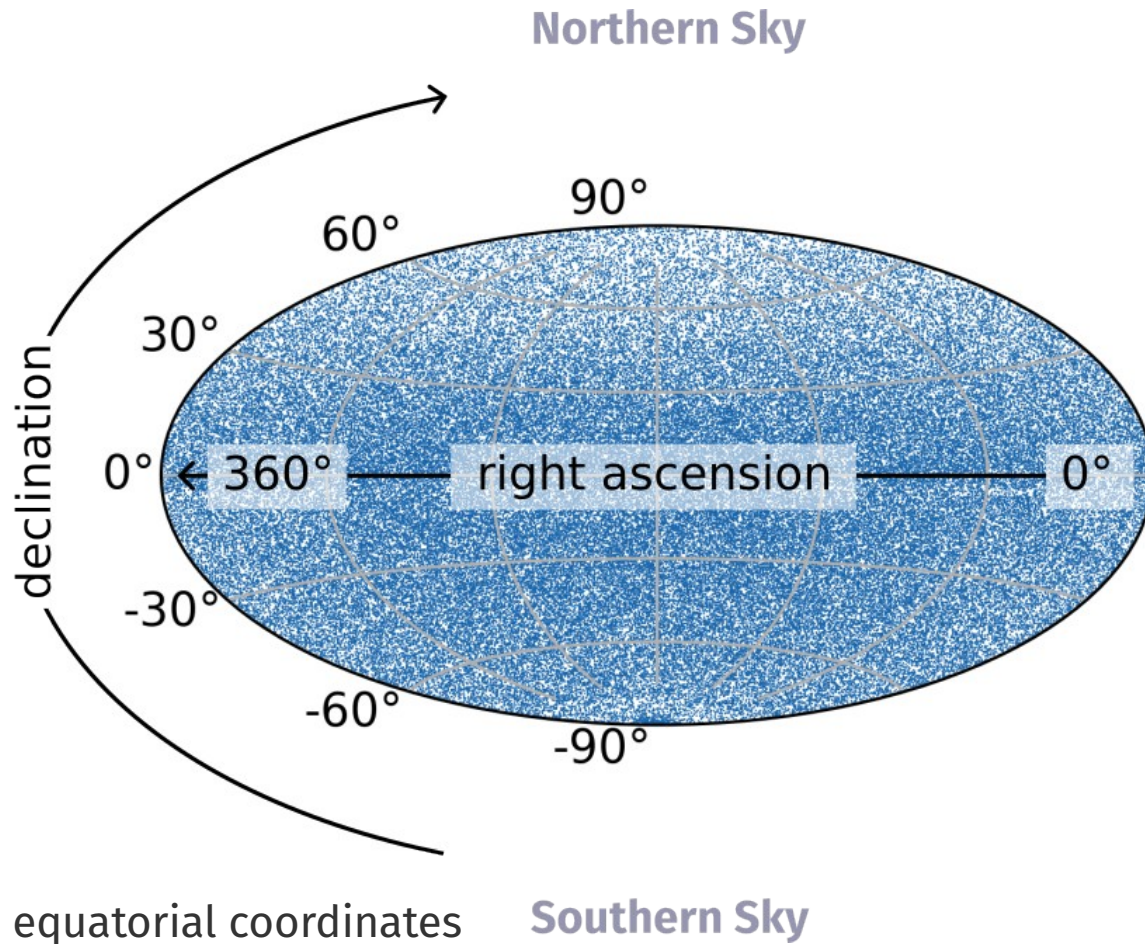
- atmospheric neutrinos
- (muons blocked by the Earth)



equatorial coordinates


- 1 dot = 1 event
- **reconstructed** direction
- equatorial coordinates

- One year of data
- Selected to reduce muons
- Still mostly background!
- Can use data to estimate background PDFs

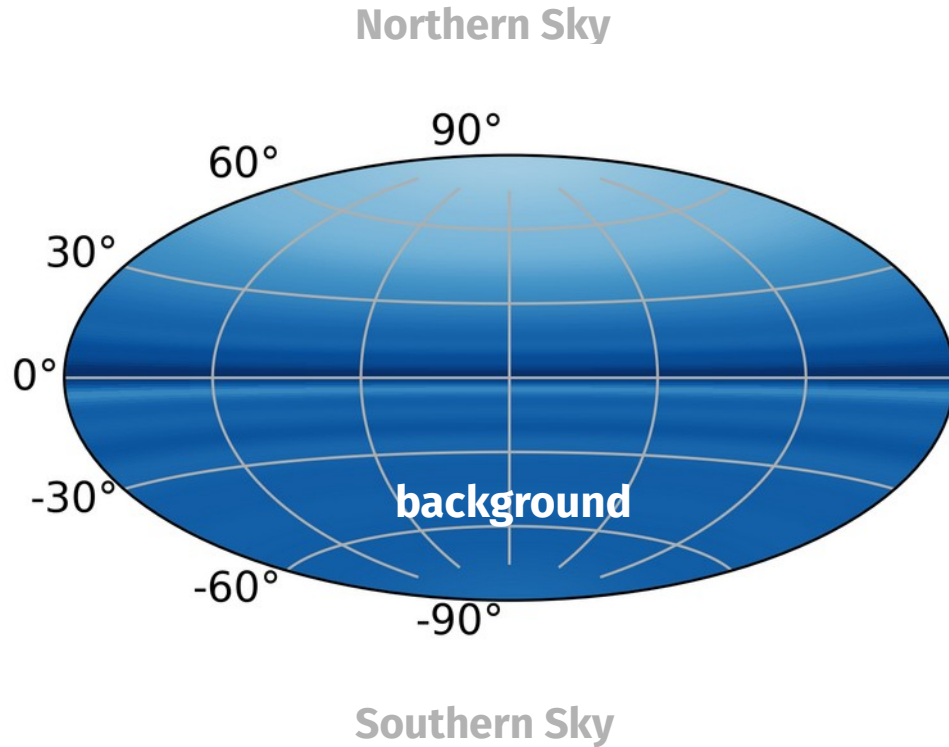


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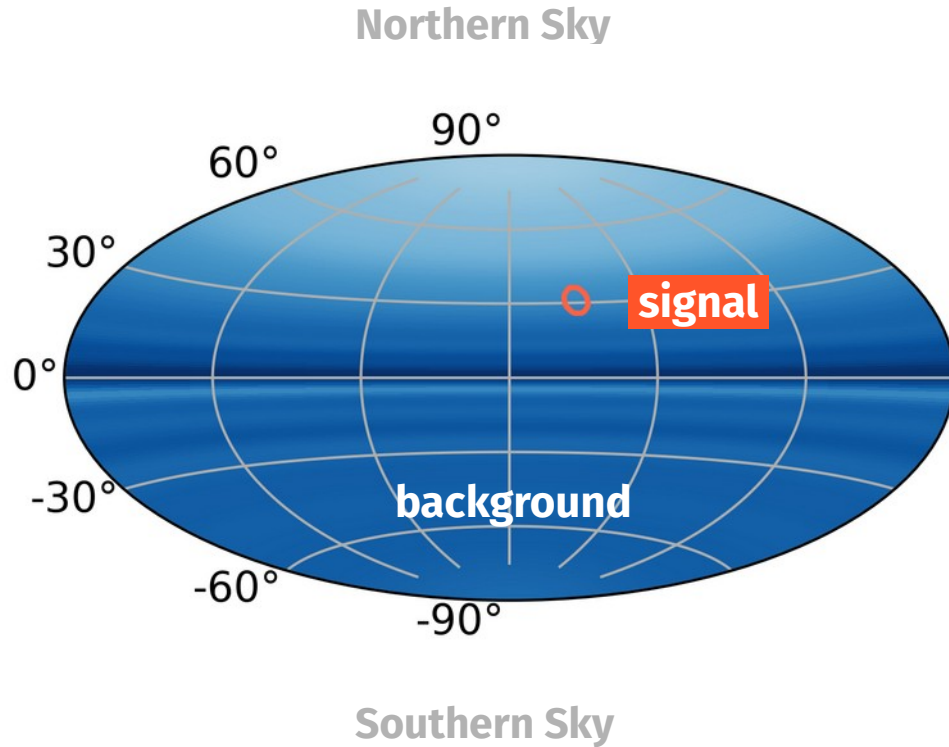
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Geographic South Pole

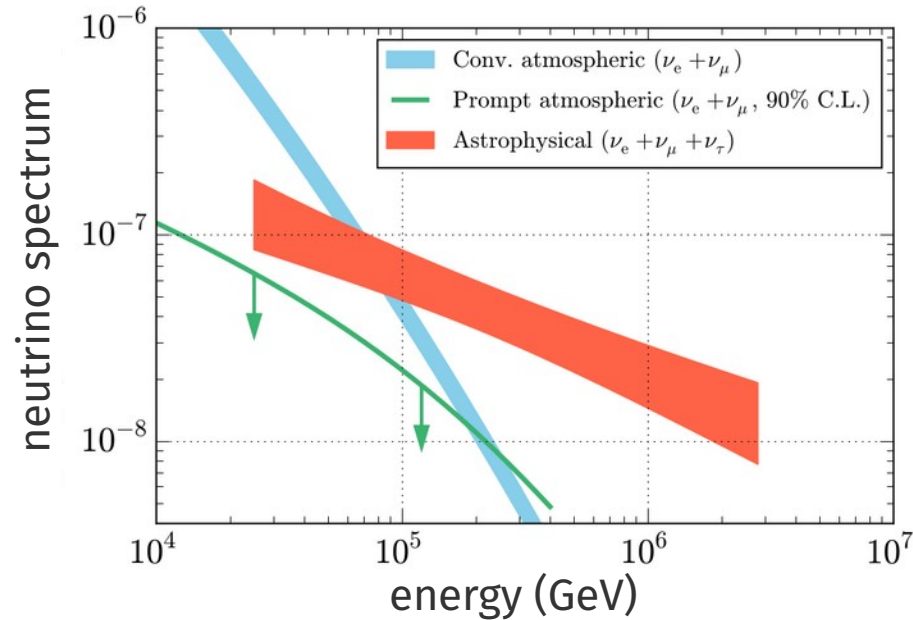


- distribution of data events in declination
- assume no dependence on right ascension (which is true for IceCube, **not** others)
 - background PDF **$B(x)$**
- shown here as sky map, darker = higher



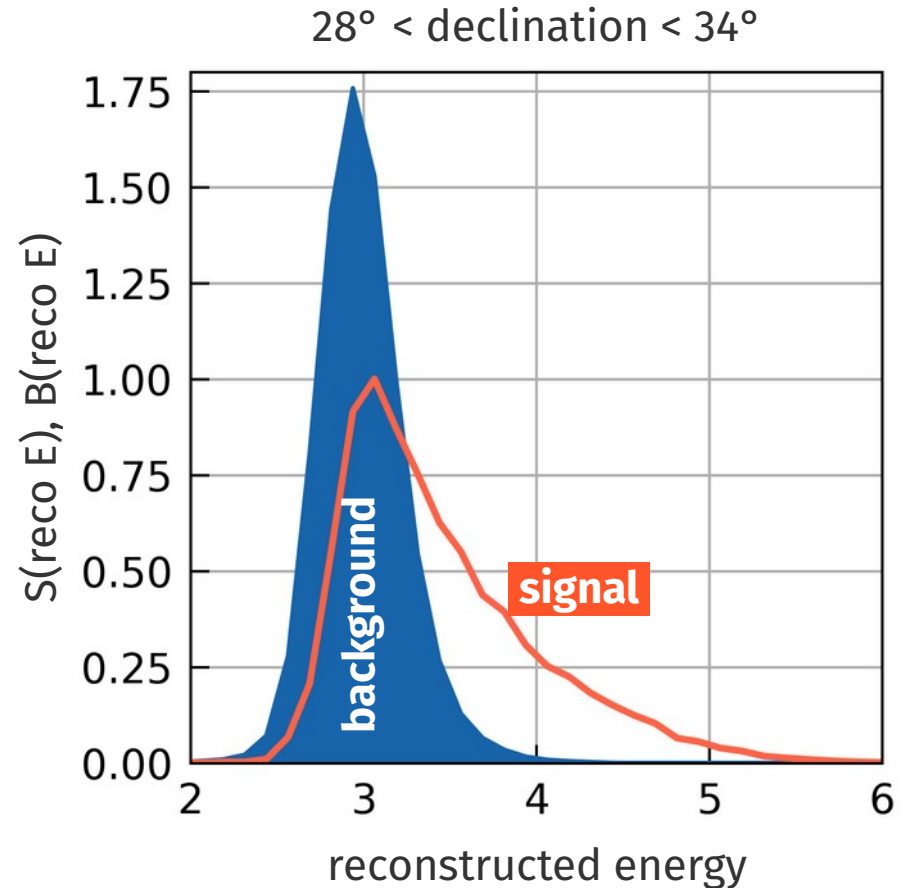
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Low energies more **atmospheric**



High energies more **astrophysical**

- Reconstructed energy \neq true energy
- Signal and background are different
- Example shown here
 - additional PDF parameter:
- $S(x) \rightarrow \mathbf{S(x, E)} = S(x) S(E; \text{declination})$
- $B(x) \rightarrow \mathbf{B(x, E)} = B(x) B(E; \text{declination})$

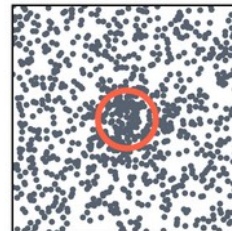
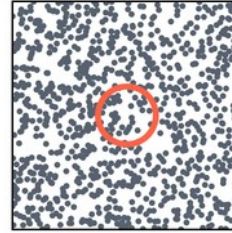


- Data sample of N events
- N is large $\rightarrow N \approx N_{\text{expected}}$
- Our hypothesis:
 - n_s signal events
 - $(N - n_s)$ background events
- PDF to describe this data:

$$\left(\frac{n_s}{N}\right)S(\vec{X}, E) + \left(1 - \frac{n_s}{N}\right)B(\vec{X}, E)$$

- Given the data, how can we tell if $n_s > 0$?

$N = 1000$

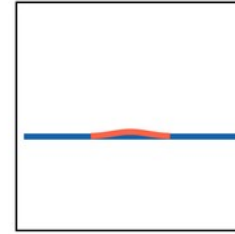
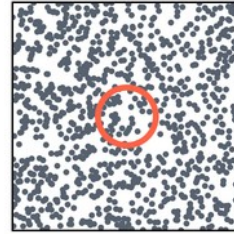


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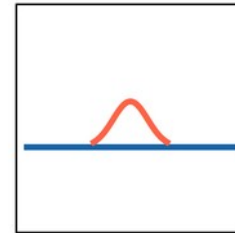
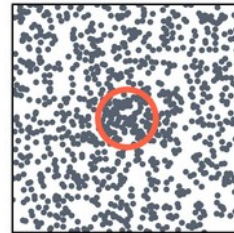
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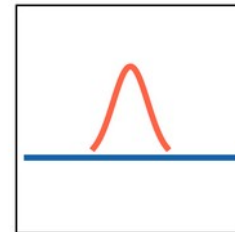
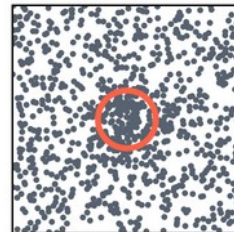
$N = 1000$



$n_s = 10$



$n_s = 100$



$n_s = 200$

Ideas?

- Mathematicians came up with this recipe:
 1. include the unknown parameters of the hypothesis in a PDF
 2. use PDF & data to define a “likelihood” (“probability to obtain this data”)
 3. find parameter values where L is largest
→ that is the “maximum likelihood estimate”

- In our case:
 1. Use the “signal + background” $P(x, E)$
 2. Multiply one term per IceCube event
 3. Find $n_S \geq 0$ where L is largest → \hat{n}_S
 4. Interpret the result?

$$L(\vec{a}) = \prod_{i=1}^N P(\vec{x}_i; \vec{a})$$

$$L(n_S) = \prod_{i=1}^N \left[\left(\frac{n_S}{N} \right) S(\vec{x}_i, E_i) + \left(1 - \frac{n_S}{N} \right) B_i(\vec{x}_i, E_i) \right]$$

$$L(\hat{n}_S) \geq L(n_S) \forall n_S > 0$$

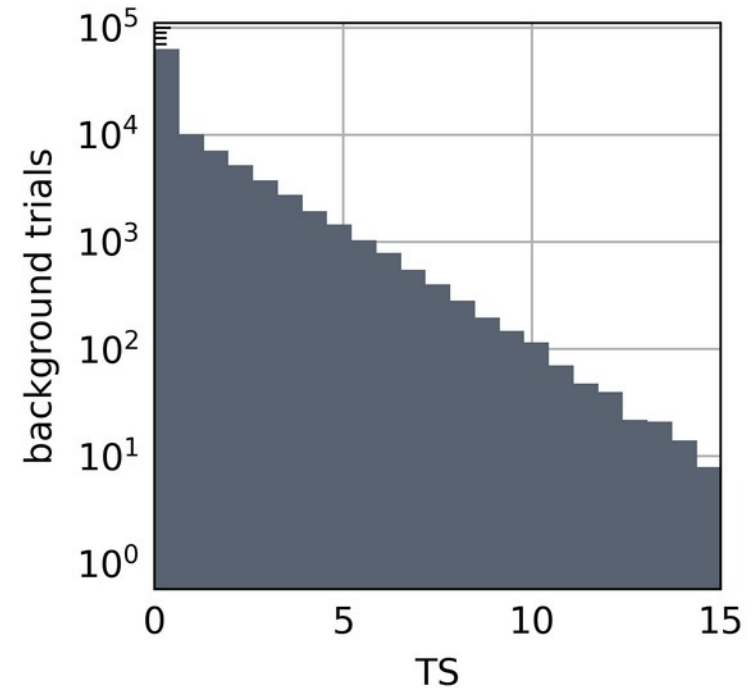
- Easier to calculate sums than products
→ logarithm of the likelihood
→ complete to Wilks' test statistic
- Properties:
 - $TS \geq 0$
 - the higher the TS, the more signal
 - If $\hat{n}_S = 0 \rightarrow TS = 0$
- But can get $TS > 0$ from pure background!

$$TS = 2 \log \frac{L(\hat{n}_S)}{L(n_S = 0)}$$

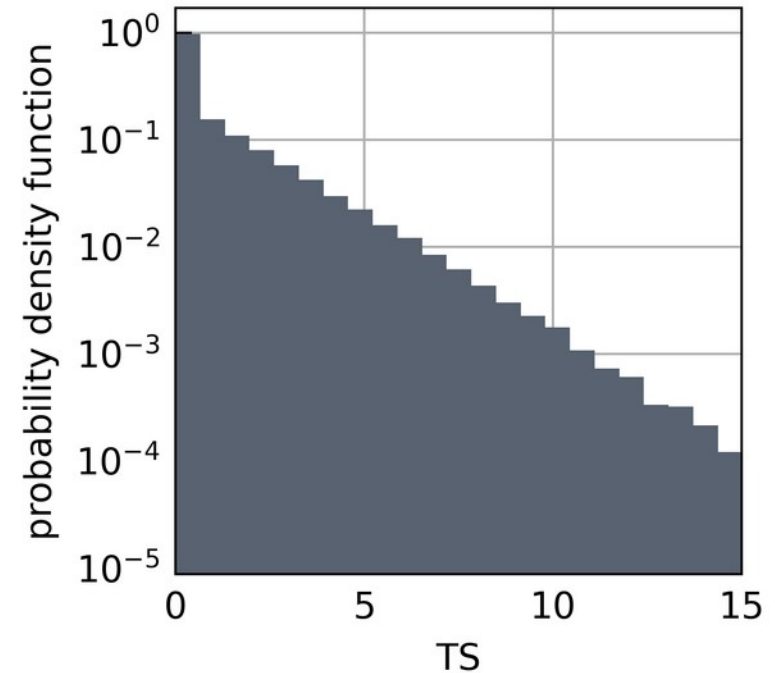
$$L(n_S = 0) = \prod_{i=1}^N B(\vec{x}_i, E_i)$$

$$\Rightarrow TS = 2 \sum_{i=1}^N \log \left[\left(\frac{n_S}{N} \right) \frac{S}{B}(\vec{x}_i, E_i) + \left(1 - \frac{n_S}{N} \right) \right]$$

- Background can fluctuate to $TS > 0$
- How can we quantify it?
- Need to simulate that we had multiple trials
- IceCube: randomize right ascension (“scrambling”)
 - different events end up near the source
 - these can not come from the source
 - equivalent to a pure-BG event sample (same detector and event selection)
- For each trial:
 - maximize TS for each “trial”
 - repeat to reveal TS distribution

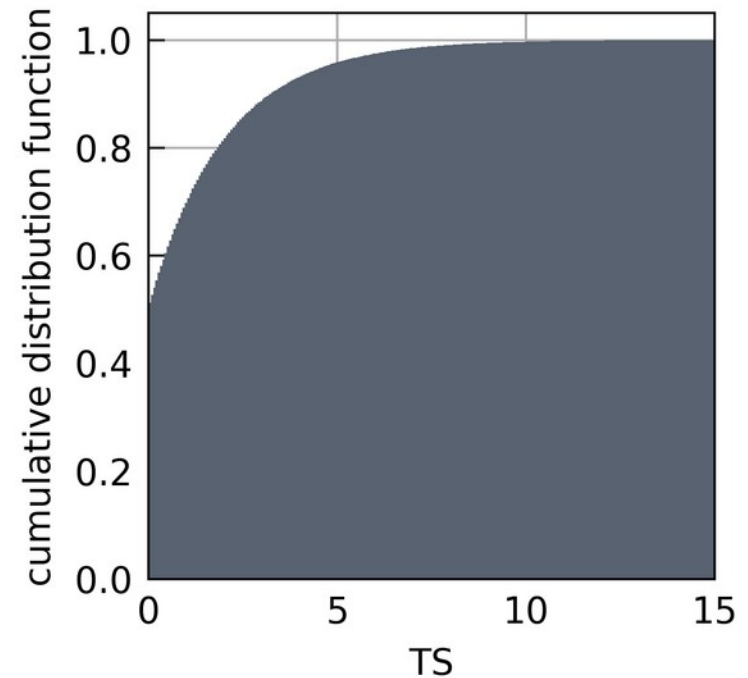


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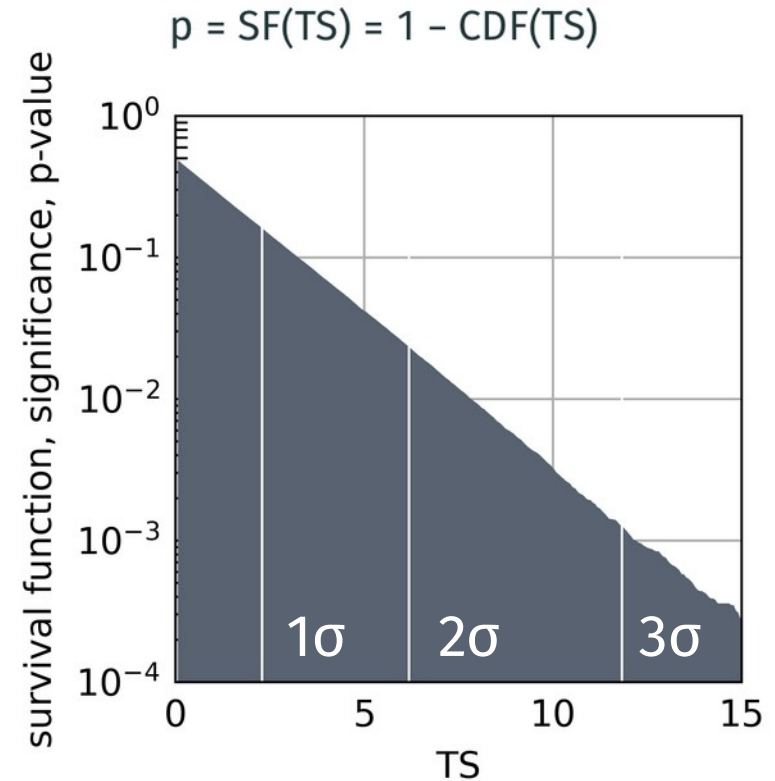


- Reject “pure BG” hypothesis for $TS > \text{threshold}$
- Probability to **falsely** reject it is the significance α
- A common (but arbitrary) choice:
 - $\alpha = 2.87 \times 10^{-7}$
 - “5 sigma”
- Set threshold = real data TS
 - combine with background trials
 - compute survival function
 - **p-value**
- Most of the time you don’t have such a clear result
- So instead we try to interpret

$$\text{CDF}(TS) = \int_0^{TS} \text{PDF}(TS') dTS'$$



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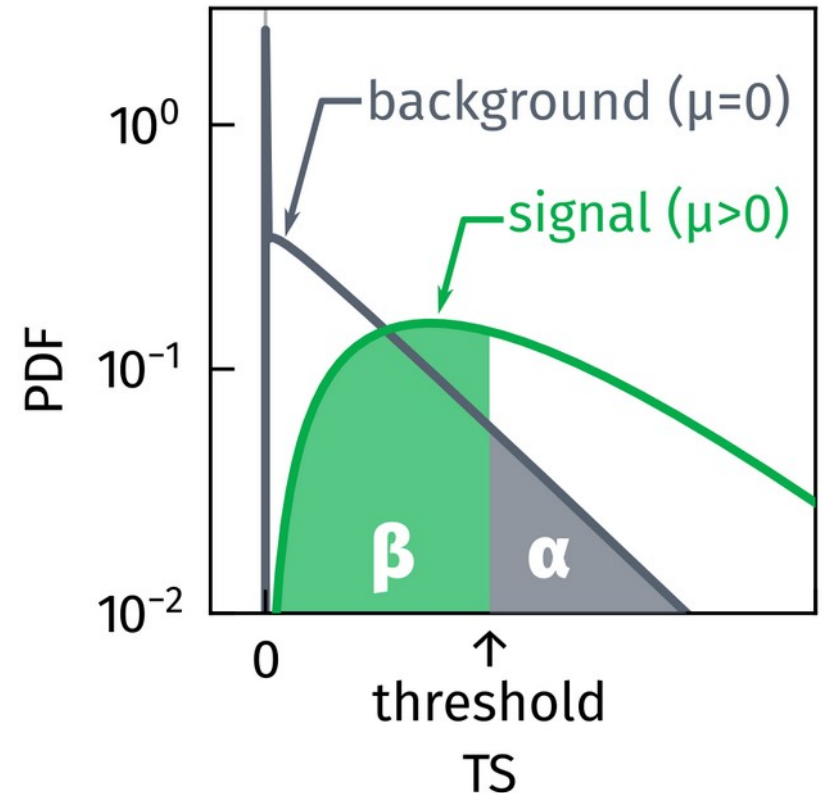


Do not be overimpressed by the attention and mathematical detail that goes into significance. This is largely due to the fact that it is under control and mathematically tractable. The power is usually far more important, but you can often say very little about it as the alternative hypothesis tends to be vague and unquantifiable. When overawed by a discussion of the significance of some test, it is salutary to remember that you can make a valid but meaningless significance test by choosing a random number r between 0 and 1 and rejecting the hypothesis if r is less than α .

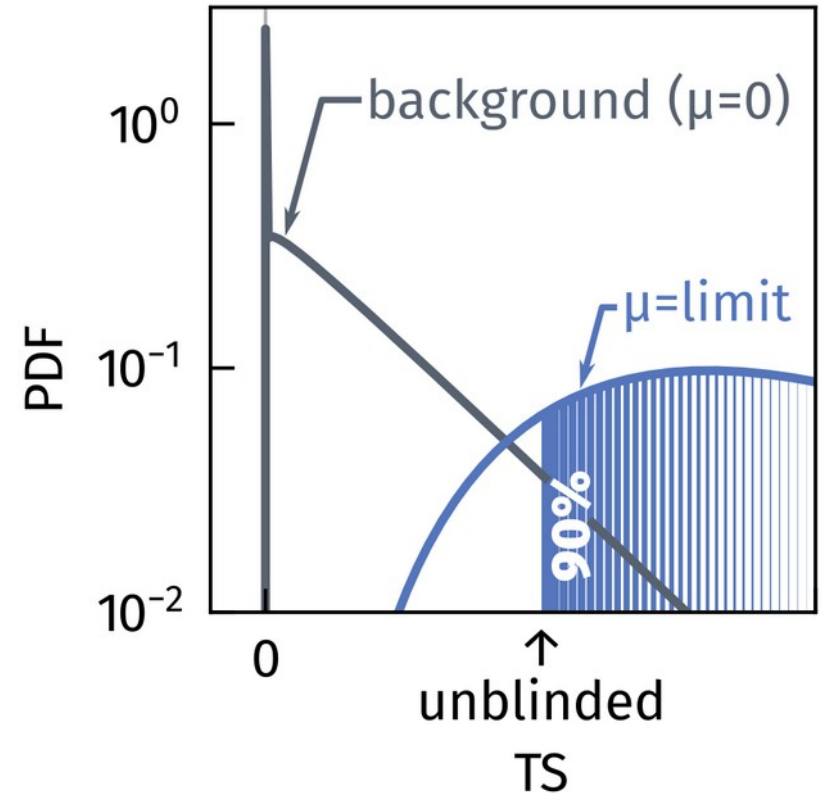
R. J. Barlow, "Statistics – A Guide to the Use of Statistical Methods in the Physical Sciences"

Less verbose: the work does not end at the p-value.

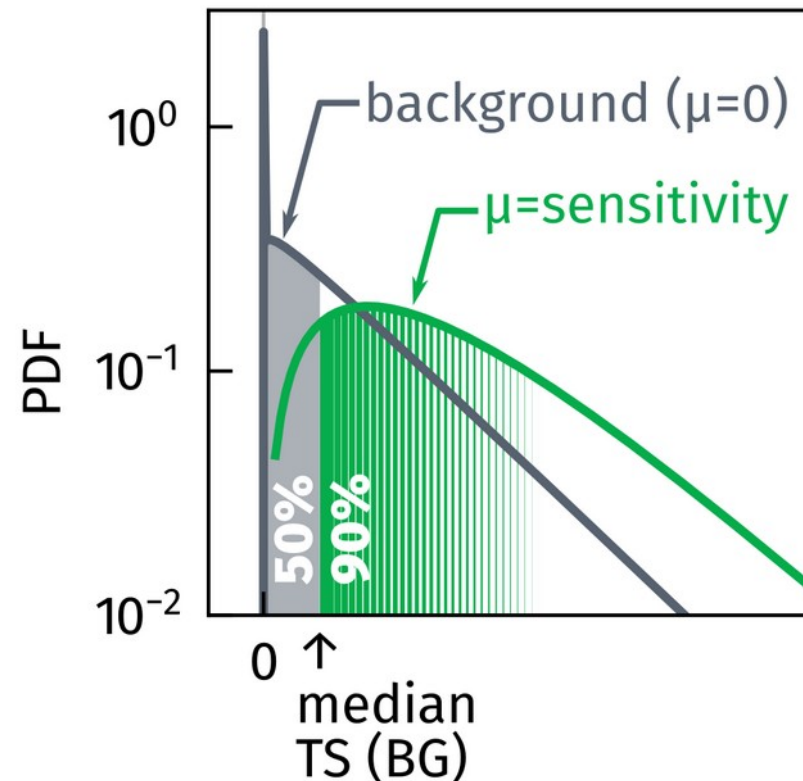
- Generate fake event samples:
background + signal $\times \mu$
i.e. scrambled data + simulation of signal
- Maximize test statistic for each event sample
→ distributions of TS given μ
→ measure power β of a hypothesis test



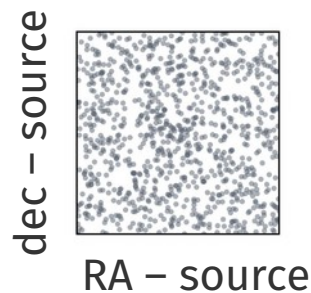
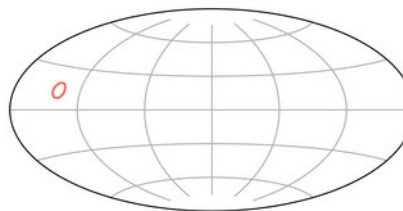
- Frequentist upper limit (for our case):
“If the true μ exceeds the upper limit, the probability to get a smaller TS than observed is 10% or less”
- Produce signal trials for a “ $\mu^{(90)}$ ”
- where 10th percentile of TS = TS(data)
- “ $\mu^{(90)}$ ” = the upper limit
 - on that hypothesis
 - at 90% “confidence level” (C.L.)



- If we could repeat the experiment,
- and there really was no signal,
- → sensitivity = median limit obtained
- Easy to construct in Neyman:
- Find “ $\mu^{(90)}$ ” where 90% of TS > median TS($\mu=0$)
- Characterizes the analysis
→ can develop it blindly



- Signal still indiscernible?
 - Collect more events
 - Wait for more data
 - or nearly equivalently:
 - Combine multiple sources
 - In mathematical terms:
 - add more PDFs
 - single signal PDF replaced by weighted sum
- can become clearer
(for the right choice of sources and weights)



$$S(\vec{x}, E)$$

- Signal still indiscernible?

→ Collect more events

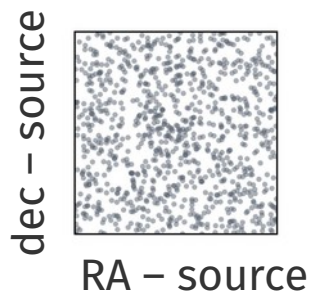
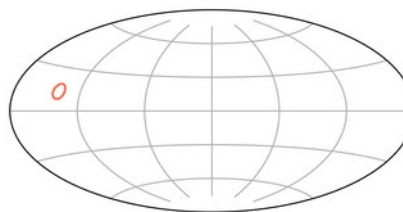
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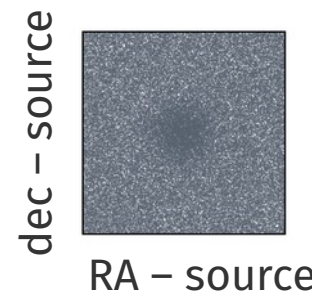
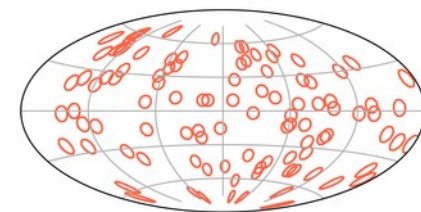
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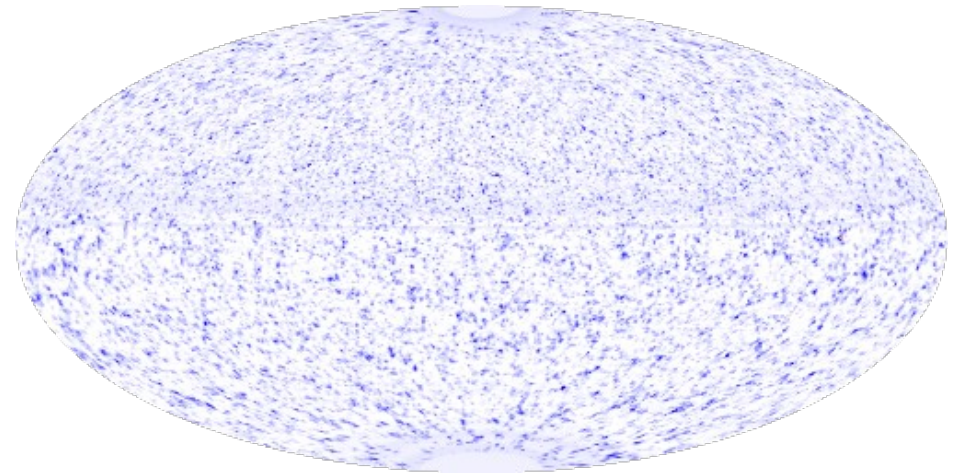
$$S(\vec{x}, E)$$



$$\rightarrow \sum_{k \in \text{sources}} w_k S_k(\vec{x}, E)$$

$$\sum_{k \in \text{sources}} w_k = 1$$

- Could also make fewer assumptions
 - no choice of source
 - no choice of source class
- Try all directions, calculate p-value each
- Getting back to a single result:
 - choose smallest p_{\min}
- Repeat background trials with this extra step
- Probability to obtain a lower p_{\min}
 - p_{post} larger
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