Goodness-of-Fit and Two-Sample Testing

Larry Wasserman

Department of Statistics and Data Science
and

Machine Learning Department
Carnegie Mellon University

We observe

$$Y_1,\ldots,Y_n\sim p$$

where

$$p = (1 - \lambda) \underbrace{b(y)}_{\text{background}} + \lambda \underbrace{s(y)}_{\text{signal}}.$$

We observe

$$Y_1,\ldots,Y_n\sim p$$

where

$$p = (1 - \lambda) \underbrace{b(y)}_{\text{background}} + \lambda \underbrace{s(y)}_{\text{signal}}.$$

We want to test

$$H_0: \lambda = 0$$

or, equivalently,

$$H_0: p=b.$$

We observe

$$Y_1,\ldots,Y_n\sim p$$

where

$$p = (1 - \lambda) \underbrace{b(y)}_{\text{background}} + \lambda \underbrace{s(y)}_{\text{signal}}.$$

We want to test

$$H_0: \lambda = 0$$

or, equivalently,

$$H_0: p = b.$$

And we have many flavors such as ...

Background:

Background:

• given: goodness-of-fit

Background:

• given: goodness-of-fit

• model: goodness-of-fit with nuisance parameters

Background:

• given: goodness-of-fit

• model: goodness-of-fit with nuisance parameters

• sampled: two-sample

Background:

• given: goodness-of-fit

• model: goodness-of-fit with nuisance parameters

• sampled: two-sample

When a signal model is given, this becomes model dependent search.

Background:

• given: goodness-of-fit

• model: goodness-of-fit with nuisance parameters

• sampled: two-sample

When a signal model is given, this becomes model dependent search.

My goal: pointers to the statistics literature that might be useful.

Assumed Background b(y)

If the background density b(y) is assumed, this is a goodness of fit test:

$$Y_1,\ldots,Y_n\sim p$$

$$H_0: p = b$$
 versus $H_1: p \neq b$.

Assumed Background b(y)

If the background density b(y) is assumed, this is a goodness of fit test:

$$Y_1,\ldots,Y_n\sim p$$

$$H_0: p = b$$
 versus $H_1: p \neq b$.

This is the classic goodness-of-fit problem but it is multivariate.

Is there an optimal test?

Is there an optimal test? Yes and No.

Is there an optimal test?

Yes and No.

Yes. (Ingster and Suslina 2003, Arias-Castro and Pelletier 2018, Balakrishnan and Wasserman 2019).

Is there an optimal test?

Yes and No.

Yes. (Ingster and Suslina 2003, Arias-Castro and Pelletier 2018, Balakrishnan and Wasserman 2019).

 $H_0: p = b$ versus $H_1: d(p, b) \ge \epsilon$ $p \in \mathcal{P}$ (nonparametric class: Sobolev space or Besov space) and some distance d.

Is there an optimal test?

Yes and No.

Yes. (Ingster and Suslina 2003, Arias-Castro and Pelletier 2018, Balakrishnan and Wasserman 2019).

 $H_0: p = b$ versus $H_1: d(p, b) \ge \epsilon$ $p \in \mathcal{P}$ (nonparametric class: Sobolev space or Besov space) and some distance d.

There exists a minimax test ϕ^* maximizes minimum power. That is, it achieves

$$\sup_{\phi} \inf_{d(p,b) \ge \epsilon} P(\phi = \text{reject})$$

Is there an optimal test?

Yes and No.

Yes. (Ingster and Suslina 2003, Arias-Castro and Pelletier 2018, Balakrishnan and Wasserman 2019).

 $H_0: p = b$ versus $H_1: d(p, b) \ge \epsilon$ $p \in \mathcal{P}$ (nonparametric class: Sobolev space or Besov space) and some distance d.

There exists a minimax test ϕ^* maximizes minimum power. That is, it achieves

$$\sup_{\phi} \inf_{d(p,b) \ge \epsilon} P(\phi = \text{reject})$$

It's optimal but the power is not high. Generally, the likelihood ratio test is not special!

Is there an optimal test?

Is there an optimal test?

No. Janssen (2000) showed that any omnibus test only has substantial power in finitely many directions.

Is there an optimal test?

No. Janssen (2000) showed that any omnibus test only has substantial power in finitely many directions.

Cannot distinguish close alternatives at a distance of $n^{-1/2}$.

Is there an optimal test?

No. Janssen (2000) showed that any omnibus test only has substantial power in finitely many directions.

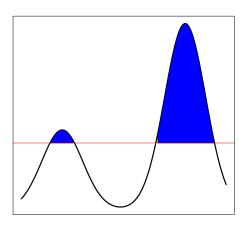
Cannot distinguish close alternatives at a distance of $n^{-1/2}$.

Nevertheless, there are some multivariate tests that you might not know which might be useful which we now review.

Let

$$\Gamma_t = \{y: \ b(y) \ge t\}$$

be the upper level set. This is a one-dimensional family of subsets. (VC dimension 1).



Then: P = B iff $P(\Gamma_t) = B(\Gamma_t)$ for all t.

Then: P = B iff $P(\Gamma_t) = B(\Gamma_t)$ for all t.

Let

$$T_n = \sup_t |P_n(\Gamma_t) - B(\Gamma_t)|$$

where

$$P_n(\Gamma_t) = \frac{1}{n} \sum_i I(Y_i \in \Gamma_t)$$
 $B(\Gamma_t) = \int_{\Gamma_t} b(y) dy$

Then: P = B iff $P(\Gamma_t) = B(\Gamma_t)$ for all t.

Let

$$T_n = \sup_t |P_n(\Gamma_t) - B(\Gamma_t)|$$

where

$$P_n(\Gamma_t) = \frac{1}{n} \sum_i I(Y_i \in \Gamma_t)$$
 $B(\Gamma_t) = \int_{\Gamma_t} b(y) dy$

Then

$$\sqrt{n}T_n \rightsquigarrow \sup_t |\mathbb{G}(t)|$$

where $\ensuremath{\mathbb{G}}$ is a Gaussian process. This is distribution free. Like a KS test.

Bickel-Breiman Nearest Neighbor Test

Bickel-Breiman Nearest Neighbor Test

Let

$$W_i = e^{-nb(Y_i)V_i}$$

where V_i is the volume of the ball containing the nearest neighbor.

Bickel-Breiman Nearest Neighbor Test

Let

$$W_i = e^{-nb(Y_i)V_i}$$

where V_i is the volume of the ball containing the nearest neighbor.

Let

$$F_n(t) = \frac{1}{n} \sum_i I(W_i \le t)$$

and

$$S = \int F_n^2(t) dt$$

which has a known limiting distribution. See Schilling (1983).

Neyman Smooth Test

This test targets certain directions by specifiying basis functions ϕ_1, ϕ_2, \ldots

Neyman Smooth Test

This test targets certain directions by specifiying basis functions ϕ_1, ϕ_2, \ldots

Model p(y) as

$$p(y) = b(y)e^{\sum_j \theta_j \phi_j(y) - Z}.$$

Neyman Smooth Test

This test targets certain directions by specifiying basis functions ϕ_1, ϕ_2, \ldots

Model p(y) as

$$p(y) = b(y)e^{\sum_j \theta_j \phi_j(y) - Z}.$$

Easy to estimate the θ_j 's and then test $\theta = 0$.

Neyman Smooth Test

This test targets certain directions by specifiying basis functions ϕ_1, ϕ_2, \ldots

Model p(y) as

$$p(y) = b(y)e^{\sum_j \theta_j \phi_j(y) - Z}.$$

Easy to estimate the θ_i 's and then test $\theta = 0$.

See Algeri (2020, 2021).

With a Given Signal (Model Dependent)

For a given signal s, the LRT is

$$T = \sup_{\lambda} \prod_{i} \left(1 - \lambda + \lambda \frac{s(Y_i)}{b(Y_i)} \right)$$

With a Given Signal (Model Dependent)

For a given signal s, the LRT is

$$T = \sup_{\lambda} \prod_{i} \left(1 - \lambda + \lambda \frac{s(Y_i)}{b(Y_i)} \right)$$

A possibly better test is the score test:

With a Given Signal (Model Dependent)

For a given signal s, the LRT is

$$T = \sup_{\lambda} \prod_{i} \left(1 - \lambda + \lambda \frac{s(Y_i)}{b(Y_i)} \right)$$

A possibly better test is the score test:

$$T = \frac{1}{n} \sum_{i} \frac{s(Y_i)}{b(Y_i)} - 1$$

which does not require estimating λ

This is a parametric family:

$$p(y) = (1 - \lambda)b(y) + \lambda s_{\theta}(y)$$

(or perhaps for a one-dimensional marginal such as mass).

This is a parametric family:

$$p(y) = (1 - \lambda)b(y) + \lambda s_{\theta}(y)$$

(or perhaps for a one-dimensional marginal such as mass).

 λ and θ can be estimated by maximum likelihood using the EM algorithm.

This is a parametric family:

$$p(y) = (1 - \lambda)b(y) + \lambda s_{\theta}(y)$$

(or perhaps for a one-dimensional marginal such as mass).

 λ and θ can be estimated by maximum likelihood using the EM algorithm.

Testing λ is tricky because of the boundary and because θ is not identified under H_0 . LRT has nonstandard limiting behavior.

This is a parametric family:

$$p(y) = (1 - \lambda)b(y) + \lambda s_{\theta}(y)$$

(or perhaps for a one-dimensional marginal such as mass).

 λ and θ can be estimated by maximum likelihood using the EM algorithm.

Testing λ is tricky because of the boundary and because θ is not identified under H_0 . LRT has nonstandard limiting behavior.

Max score:

$$T_n = \sup_{\theta} \frac{1}{n} \sum_{i} \frac{s_{\theta}(Y_i)}{b(Y_i)} - 1$$

and the null distribution can be obtained by simulation.

Target the bumps in a one dimensional marginal M = f(Y).

Target the bumps in a one dimensional marginal M = f(Y).

Test:

$$H_0: p(m) = b(m)$$
 for all m versus $H_1: p(m) > b(m)$ for some m .

Target the bumps in a one dimensional marginal M = f(Y).

Test:

 $H_0: p(m) = b(m)$ for all m versus $H_1: p(m) > b(m)$ for some m.

Don't use histograms! Use the local polynomial density estimator \hat{p} (Cattaneo, Jansson and Ma 2020).

Target the bumps in a one dimensional marginal M = f(Y).

Test:

$$H_0: p(m) = b(m)$$
 for all m versus $H_1: p(m) > b(m)$ for some m .

Don't use histograms! Use the local polynomial density estimator \widehat{p} (Cattaneo, Jansson and Ma 2020).

$$F(m) = P(M \le m).$$

Target the bumps in a one dimensional marginal M = f(Y).

Test:

$$H_0: p(m) = b(m)$$
 for all m versus $H_1: p(m) > b(m)$ for some m .

Don't use histograms! Use the local polynomial density estimator \hat{p} (Cattaneo, Jansson and Ma 2020).

$$F(m) = P(M \leq m).$$

For u near m:

$$F(u) \approx F(m) + (u - m)p(m) + \frac{(u - m)^2}{2}p'(m)$$

= $\beta_0(m) + (u - m)\beta_1(m) + (u - m)^2\beta_2(m)$

Let

$$F_n(m) = \frac{1}{n} \sum_i I(M_i \le m)$$

Let

$$F_n(m) = \frac{1}{n} \sum_i I(M_i \le m)$$

Let $\widehat{\beta}(m)$ minimize:

$$\min_{b} \sum_{i} (\widehat{F}_{n}(m) - r^{T}b)^{2} K\left(\frac{M_{i} - m}{h(m)}\right)$$

where $r = (1, M_i - m, (M_i - m)^2)$, K is a kernel, $h(m) = (C(m)/n)^{1/5}$ and C(m) is known.

Let

$$F_n(m) = \frac{1}{n} \sum_i I(M_i \le m)$$

Let $\widehat{\beta}(m)$ minimize:

$$\min_{b} \sum_{i} (\widehat{F}_{n}(m) - r^{T}b)^{2} K\left(\frac{M_{i} - m}{h(m)}\right)$$

where $r = (1, M_i - m, (M_i - m)^2)$, K is a kernel, $h(m) = (C(m)/n)^{1/5}$ and C(m) is known.

Let

$$\widehat{p}(m) = \widehat{\beta}_1(m).$$

This is optimal (under mild conditions) and boundary adaptive.

Let

$$F_n(m) = \frac{1}{n} \sum_i I(M_i \le m)$$

Let $\widehat{\beta}(m)$ minimize:

$$\min_{b} \sum_{i} (\widehat{F}_{n}(m) - r^{T}b)^{2} K\left(\frac{M_{i} - m}{h(m)}\right)$$

where $r = (1, M_i - m, (M_i - m)^2)$, K is a kernel, $h(m) = (C(m)/n)^{1/5}$ and C(m) is known.

Let

$$\widehat{p}(m) = \widehat{\beta}_1(m).$$

This is optimal (under mild conditions) and boundary adaptive.

Use

$$T = \sup[\widehat{p}(m) - b(m)].$$

There is a growing literature on robust tests:

$$H_0: Y_1, \ldots, Y_n \sim q, \ q \in N_{\epsilon}(b)$$

where $N_{\epsilon}(b)$ is a neighborhood of b.

There is a growing literature on robust tests:

$$H_0: Y_1, \ldots, Y_n \sim q, \ q \in N_{\epsilon}(b)$$

where $N_{\epsilon}(b)$ is a neighborhood of b.

Examples: Wasserstein neighborhood (Xie, Gao and Xie 2021) RKHS neighborhood (Sun and Zou 2022), Huber neighborhood (Huber 1965).

There is a growing literature on robust tests:

$$H_0: Y_1, \ldots, Y_n \sim q, \ q \in N_{\epsilon}(b)$$

where $N_{\epsilon}(b)$ is a neighborhood of b.

Examples: Wasserstein neighborhood (Xie, Gao and Xie 2021) RKHS neighborhood (Sun and Zou 2022), Huber neighborhood (Huber 1965).

Tradeoff between robustness and power.

With Simulated Background: Two Sample Test

$$X_1, \ldots, X_m \sim b$$

 $Y_1, \ldots, Y_n \sim p = (1 - \lambda)b + \lambda s$

With Simulated Background: Two Sample Test

$$X_1, \ldots, X_m \sim b$$

 $Y_1, \ldots, Y_n \sim p = (1 - \lambda)b + \lambda s$

Two sample test:

$$H_0: p = b$$
 versus $p \neq b$

With Simulated Background: Two Sample Test

$$X_1, \ldots, X_m \sim b$$

 $Y_1, \ldots, Y_n \sim p = (1 - \lambda)b + \lambda s$

Two sample test:

$$H_0: p = b$$
 versus $p \neq b$

Again there are many tests. There is no optimal test.

Let

$$\psi = \sup_{f \in \mathcal{F}} \left| \mathbb{E}_B[f(X)] - \mathbb{E}_P[f(Y)] \right|$$

where \mathcal{F} is a reproducing kernel Hilbert space (RKHS).

Let

$$\psi = \sup_{f \in \mathcal{F}} \left| \mathbb{E}_B[f(X)] - \mathbb{E}_P[f(Y)] \right|$$

where \mathcal{F} is a reproducing kernel Hilbert space (RKHS).

$$T_n = \frac{1}{n(n-1)} \sum_{i \neq j} H_{ij}$$

where

$$H_{ij} = K(X_i, X_j) + K(Y_i, Y_j) - K(X_i, Y_j) - K(X_j, Y_i).$$

Let

$$\psi = \sup_{f \in \mathcal{F}} \left| \mathbb{E}_B[f(X)] - \mathbb{E}_P[f(Y)] \right|$$

where \mathcal{F} is a reproducing kernel Hilbert space (RKHS).

$$T_n = \frac{1}{n(n-1)} \sum_{i \neq j} H_{ij}$$

where

$$H_{ij} = K(X_i, X_j) + K(Y_i, Y_j) - K(X_i, Y_j) - K(X_j, Y_i).$$

Null distribution is complicated.

$$h(z) = P(S = 1|Z = z)$$

$$= \frac{p(z|S = 1)P(S = 1)}{p(z|S = 1)P(S = 1) + p(z|S = 0)P(S = 0)}$$

$$= \frac{p(z|S = 1)\pi}{p(z|S = 1)\pi + p(z|S = 0)(1 - \pi)}$$

where $\pi = m/(m+n)$. Hence

$$\frac{p(y)}{b(y)} = \frac{h(z)}{1 - h(z)}$$

so we have estimated the density ratio. Chakravarti, Kuusela, Lei and Wasserman 2022

Which classifier?

Which classifier?

Current fashion: neural nets (deep learning)

Which classifier?

Current fashion: neural nets (deep learning)

Others:

random forests, logistic regression, ...

Which classifier?

Current fashion: neural nets (deep learning)

Others:

random forests, logistic regression, ...

Aside: why did everyone start calling classification and regression Machine Learning? It's statistics! We've been doing it for 100 years!

Which test?

Which test?

$$\prod_i \frac{\widehat{h}(Z_i)}{1-\widehat{h}(Z_i)}$$

is an estimate of the Neyman-Pearson test.

Which test?

$$\prod_{i} \frac{\widehat{h}(Z_{i})}{1 - \widehat{h}(Z_{i})}$$

is an estimate of the Neyman-Pearson test.

Really, the classifier is just a dimension reduction method. We have

$$\widehat{h}(X_1),\ldots,\widehat{h}(X_N)$$

and

$$\widehat{h}(Y_1),\ldots,\widehat{h}(Y_n)$$

Which test?

$$\prod_{i} \frac{\widehat{h}(Z_{i})}{1 - \widehat{h}(Z_{i})}$$

is an estimate of the Neyman-Pearson test.

Really, the classifier is just a dimension reduction method. We have

$$\widehat{h}(X_1),\ldots,\widehat{h}(X_N)$$

and

$$\widehat{h}(Y_1),\ldots,\widehat{h}(Y_n)$$

The data are now one-dimensional. We can use any one-dimensional two-sample test we want.

Which test?

$$\prod_{i} \frac{\widehat{h}(Z_{i})}{1 - \widehat{h}(Z_{i})}$$

is an estimate of the Neyman-Pearson test.

Really, the classifier is just a dimension reduction method. We have

$$\widehat{h}(X_1),\ldots,\widehat{h}(X_N)$$

and

$$\widehat{h}(Y_1),\ldots,\widehat{h}(Y_n)$$

The data are now one-dimensional. We can use any one-dimensional two-sample test we want.

For example:

classifier accuracy, density ratio (Neyman-Pearson), KS test, etc.

Constructing the classifier and doing the test on the same data can lead to invalid p-value.

Constructing the classifier and doing the test on the same data can lead to invalid p-value.

Two fixes: permuations and data splitting

Constructing the classifier and doing the test on the same data can lead to invalid p-value.

Two fixes: permuations and data splitting

Permutation: permute the labels, repeat the classifier K times, and the p-value is

$$\frac{1}{K}\sum_{j}I(T_{j}>t)$$

is a valid p-value. But this is expensive.

Constructing the classifier and doing the test on the same data can lead to invalid p-value.

Two fixes: permuations and data splitting

Permutation: permute the labels, repeat the classifier K times, and the p-value is

$$\frac{1}{K}\sum_{j}I(T_{j}>t)$$

is a valid p-value. But this is expensive.

Or: split the sample. Construct the classifier on first half. Conduct the test on the second half.

See Chakravarti, Kuusela, Lei and Wasserman (2022).

Are classifier tests better than other tests?

Are classifier tests better than other tests? No one knows.

Are classifier tests better than other tests?

No one knows.

The theoretical properties of black box classifiers (random forests, neural nets) are not understood.

Are classifier tests better than other tests?

No one knows.

The theoretical properties of black box classifiers (random forests, neural nets) are not understood.

Don't assume that neural nets are optimal.

Because there is no optimal test, we need to choose a test carefully.

Because there is no optimal test, we need to choose a test carefully. Classifier tests seem very promising.

Because there is no optimal test, we need to choose a test carefully.

Classifier tests seem very promising.

We have virtually no theory for these tests. (Some limited results in Kim, Ramdas, Singh and Wasserman 2021).

Because there is no optimal test, we need to choose a test carefully.

Classifier tests seem very promising.

We have virtually no theory for these tests. (Some limited results in Kim, Ramdas, Singh and Wasserman 2021).

THE END

Exact inference, no regularity conditions. (Wasserman, Ramdas, Balakrishnan 2020).

Exact inference, no regularity conditions. (Wasserman, Ramdas, Balakrishnan 2020).

Split data: \mathcal{D}_0 and \mathcal{D}_1 .

Exact inference, no regularity conditions. (Wasserman, Ramdas, Balakrishnan 2020).

Split data: \mathcal{D}_0 and \mathcal{D}_1 .

Compute $\widehat{\lambda}$ from \mathcal{D}_1 and likelihood \mathcal{L}_0 from \mathcal{D}_0 .

Exact inference, no regularity conditions. (Wasserman, Ramdas, Balakrishnan 2020).

Split data: \mathcal{D}_0 and \mathcal{D}_1 .

Compute $\widehat{\lambda}$ from \mathcal{D}_1 and likelihood \mathcal{L}_0 from \mathcal{D}_0 .

Let

$$U=\frac{\mathcal{L}_0(\widehat{\lambda})}{\mathcal{L}(0)}.$$

Exact inference, no regularity conditions. (Wasserman, Ramdas, Balakrishnan 2020).

Split data: \mathcal{D}_0 and \mathcal{D}_1 .

Compute $\widehat{\lambda}$ from \mathcal{D}_1 and likelihood \mathcal{L}_0 from \mathcal{D}_0 .

Let

$$U=\frac{\mathcal{L}_0(\widehat{\lambda})}{\mathcal{L}(0)}.$$

Repeat B times and let $U = B^{-1} \sum_{j} U_{j}$.

Exact inference, no regularity conditions. (Wasserman, Ramdas, Balakrishnan 2020).

Split data: \mathcal{D}_0 and \mathcal{D}_1 .

Compute $\widehat{\lambda}$ from \mathcal{D}_1 and likelihood \mathcal{L}_0 from \mathcal{D}_0 .

Let

$$U=\frac{\mathcal{L}_0(\widehat{\lambda})}{\mathcal{L}(0)}.$$

Repeat B times and let $U = B^{-1} \sum_{j} U_{j}$.

Reject of $U > 1/\alpha$.