Gaussian Processbased calculation of look-elsewhere trials factors

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Introduction

Recurring questions about how to efficiently and precisely (enough) calculate the global significance Z^{global} of a potential new resonance when a scan over a mass range yields a local significance Z^{local} at a particular mass value

- 1. Typical resonance search (bump hunt)
- 2. Look-elsewhere effect (trials factor)
- 3. Gaussian Processes
- 4. Leveraging Gaussian Processes for LEE/TF
 - a) Given a covariance, estimate trials factor
 - b) Estimating covariance in presence of background

Typical bump hunt



- Fit signal and background models to invariant mass spectrum
- Vary the mass parameter m of a hypothetical signal across the spectrum
- Calculate p₀, the probability to observe an excess of the background at least as big as the one we see in the data
 - Hope to see a narrow region with $p_0 \ll 1$
 - Report «local» significance ($Z \equiv N\sigma$)
 - Take into account trials factor (LEE)

Look-elsewhere effect (LEE) and Trials factor (TF)

- We focussed on the peak with maximum significance
- A peak from a background fluctation could arise anywhere in the spectrum: we have to «look elsewhere»
- The p_0 must be increased by a «trials factor»: This will give us a «global» p_0 and corresponding global Z with $Z^{global} < Z^{local}$

Trials factor
$$f_T \equiv p_o^{global}/p_0^{local}$$



3 LEE options

- **1.** Outdated rule-of-thumb $f_T \simeq \Delta m / \sigma_m$
- 2. Estimate by **brute force** fit a number of background-only pseudo-datasets or «toys»
- 3. Use a **cutting-edge** asymptotic approximation («G&V»)



Gross & Vitells Trials Factors (LEE)

- An elegant approximation based on average «up-crossings» $\langle N_{up} \rangle$
- Count «up-crossings» at low significance σ_t and extrapolate to high significance Z_{local} , i.e.

$$p_0^{global} \le p_0^{local} + \langle N_{up}(\sigma_t) \rangle e^{-\frac{Z_{local}^2 - \sigma_t^2}{2}}$$

- Can use relatively small number of MC experiments to estimate $\langle N_{up}(\sigma_t) \rangle$, but this can still be challenging
- In worst case can use the data for *rough* estimate of $\langle N_{up}(\sigma_t) \rangle$



Gross & Vitells Trials Factors (LEE)

- What to do if not comfortable with large statistical uncertainty on TF (Z^{global}), besides throwing a lot of carbon at the problem?
- What to do about obvious over-conservatism of G&V UL below $\sim 3\sigma$?
- Can we do better?!
- G&V article builds on extensive previous work on random χ^2 and Gaussian fields.
- I noticed the connection Rasmussen&Williams make in «Gaussian Processes for Machine Learning» (2006) between correlation length of Gaussian process and up-crossings.



Gaussian process





Rasmussen and Williams

- A GP is defined by a mean function $\mu(x)$ and a covariance function $\Sigma(x, x')$.
- In a typical (Bayesian) analysis one starts with priors and "trains" the process with data to obtain a posterior prediction $\mu^*(x), \Sigma^*(x, x')$.

Gaussian process

• Similar example

$$\mu(x) = 0$$

$$\Sigma(x, x') = A^2 e^{-(x-x')^2/(2l^2)},$$

with $A = l = 1$ and train on 1
noiseless and 1 noisy observation.

• Random samples of the posterior GP look a lot like signed local significance (Z) scans! Same is true for prior.



Gausssian process for LEE!

- Instead of defining priors and training the GP, we do *something completely different:* Construct the covariance directly and as efficiently as possible from information encoded in the fitting procedure used in the resonance search.
 - $\mu(x) \equiv \langle Z(m) \rangle = 0$ asymptotically under the background hypothesis
- Once we have $\Sigma(m, m')$, it is computationally cheap to sample the GP ("generate GP-toys") and measure with high precision:
 - up-crossings
 - directly the global $p_0,$ especially at low and moderate significance, but also at moderately high Z_{local}
- It was already worked out for us how to calculate the average up-crossings at any significance level directly from the covariance in *"Random Vibrations. Analysis of Structural and Mechanical Systems", Lutes, L.D. and Sarkani, S., Butterworth Heinemann, Boston (2004).*

A new (!) set of Asimov (background) datasets

- Sum the covariances $\Sigma_i(m, m')$ resulting from *individual background (bin i) fluctuations*.
- The standard background-only toy is a series of independent fluctuations!
 - Covariances for independent sources add linearly
 - There are cases with correlated data not covered here
- Fitted signal amplitude $\hat{\mu}(m)$ must be proportional to the amount of «signal». A single fluctuation is a placeholder for all possible fluctuations of the bin in question, just like the fit to a standard Asimov dataset is a placeholder for an ensemble of fits to random toy datasets.
- Last step is a normalization (inspired by Gibbs for his GP kernel function) to render $\Sigma(m,m) = 1$.



Covariances $\Sigma(m, m')$



Various paths to the trials factor (and Z_{global})



Various studies and cross-checks

- 3 different statistical models:
 - 1) G&V model based on background template,
 - 2) parametric background inspired by $H \rightarrow \gamma \gamma$, and
 - 3) inspired by $H \rightarrow \gamma \gamma$ with additional scan over width of mass peak
- We (more precisely ⁽ⁱ⁾) reproduce the published G&V example results (1)
- 2 different software implementations: python (published results) and MATLAB (cross-check)
- Effect of poor choice of bin width (affects sensitivity, but no bias)
- 10x larger data sample for G&V model (smaller diff. BF-GP toys)

SigCorr: Python software on gitlab.cern.ch

- A framework to study the trials factor
- Several ways to estimate the trials factor (see figure)
- Utilities that consistently operate on defined data structures that allow the user to build their own pipeline
- See <u>the docs</u> (sigcorr.docs.cern.ch/dev/) for installation details, usage examples and tutorials.



Conclusions

- We propose to model the covariance of the significance for bump hunts with fit-scans of special datasets consisting of Asimov background plus 1-bin-at-a-time fluctuations.
- With 3 quite different toy models of searches we have excellent agreement between large samples of brute-force toys and GP-toys (and reproduce the G&V published result on their toy model).
- The G&V approximation for high local significances based on $\langle N_{up} \rangle$ at low significance is still important, but now we have a GP-based method that gives precise and accurate results for moderate to low significances as well as precise and accurate estimates of the upcrossings at low significance.
- We will soon submit a publication showing how to obtain an estimate of the covariance with only a single Asmimov fit (and some derivative calculations) with the same accuracy as our previous, already carbon-friendly, procedure.



Backup

Significance scans of 2-D toys



Covariance for 2-D scan of mass and width



2 x 2-D (mass and width) → 4-D covariance matrix Unwrapped into 2-D



Gross & Vitells Trials Factors (LEE)

- Using high-Z approximation of χ_1^2 , can show that: $f_T \simeq 1 + \sqrt{\frac{\pi}{2}} \mathcal{N}Z_{local}$, where \mathcal{N} is «effective # of independent search regions»
 - Outdated rule-of-thumb is ~OK for $Z_{local} \simeq 2 3$, but wrong for large Z!
 - Bob Cousins called this «an important discovery»!

Prediction of up-crossings?

- G&V: «The function C(θ) [that $\langle N_{up}(\sigma_t) \rangle$ depends on] can in general be difficult to calculate.»
- The G&V results are based on the properties of random Gaussian fields, i.e. Gaussian Processes



Analytic prediction of up-crossings?

 Adler (1981) Theorem 4.1.1 states that expectation value of number of up-crossings per unit interval at level u is

$$\mathbb{E}(N_u) = \frac{1}{2\pi} \sqrt{\frac{-k''^{(0)}}{k(0)}} e^{-u^2/(2k(0))},$$

where k(x, x') = k(x - x'), i.e., a «stationary kernel» (i.e. covariance)

• For exp-squared kernel with correlation length l, $k(x) = \sigma^2 e^{-x^2/2l^2}$, we have

$$\mathbb{E}(N_u) = \frac{1}{2\pi l} e^{-u^2/(2\sigma^2)}$$

- For significance field $u \equiv Z$ and $\sigma \equiv 1$
- $\mathbb{E}(N_u) = \frac{1}{2\pi l} e^{-Z^2/2}$ (basis of G&V extrapolation from low to high Z)

Analytic prediction of up-crossings?

• For mass range Δm and constant Gaussian mass resolution $\sigma_m = l$ (*), the predicted average number of up-crossings at any threshold is

$$\mathbb{E}(n_u) = \frac{\Delta m}{2\pi\sigma_m} e^{-Z^2/2}$$

• (*) A comment in <u>Frate, Cramner et al., Modeling Smooth Backgrounds and Generic</u> <u>Localized Signals with Gaussian Processes, arXiv:1709.05681v1</u> about an additional $\sqrt{2}$ is misleading for the significance field.

Analytic prediction of up-crossings for $\sigma_m(m)$?

• Propose
$$\mathbb{E}(N_u) = \frac{\int_{m_0}^{m_1} \mathbb{E}(N_u(m))dm}{\int_{m_0}^{m_1} dm} = \frac{1}{2\pi} e^{-\frac{Z^2}{2}} \mathbb{E}(\frac{1}{l(m)})$$

• For linear $\sigma(m)$, e.g. «Gibbs kernel» (next page), we have

$$l(m) = l_0 + (m - m_0)(l_1 - l_0)/(m_1 - m_2)$$

and find

$$\mathbb{E}(n_u) = \frac{1}{l_1 - l_0} \ln(\frac{l_1}{l_0}) \frac{1}{2\pi} e^{-\frac{Z^2}{2}}$$

Two basic Gaussian Process kernels

- Gaussian signal with constant mass resolution
 - The unit «exponential-squared» kernel:

 $\Sigma(m,m',\sigma_m) = e^{-(m-m')^2/2\sigma_m^2}$

- Example with $\sigma_{\rm m}~=~5~{\rm GeV}~(8.3~{\rm bins})$
- Gaussian signal with mass-dependent resolution $\sigma_m(m)$:
 - The unit «Gibbs» kernel:

$$\Sigma(\mathbf{m},\mathbf{m}') = \sqrt{\frac{2\sigma_m(m)\sigma_m(m')}{\sigma_m^2(m) + \sigma_m^2(m')}} e^{-(m-m')^2/(\sigma_m^2(m) + \sigma_m^2(m'))}.$$

• Example with linear $\sigma_m(m)$ between 2-10 GeV (2.7-13.3 bins)



Gaussian Process toys prediction of up-crossings?

- A Gaussian Process for a set of points x is defined by a mean function $\mu(x)$ and a kernel k(x, x'), i.e. a covariance matrix $\Sigma(x, x')$.
- Fitting a signal model to an ensemble of background spectra at a particular mass hypothesis x_i should result in mean significance of 0 and standard deviation of 1.
- The shape of the signal model must strongly influence the covariance between 2 points $\Sigma(x, x')$, however, is it a good approximation to neglect the influence of background?
- With $\mu(\mathbf{x}) = 0$ and $\Sigma(\mathbf{x}, \mathbf{x}')$ in hand we can easily generate huge numbers of MC experiments for Z(m) from the multivariate Gaussian distribution or «GP-toys»!
 - No need to perform zillions of time-consuming fits!
 - Reliable results for small Z!



Gibbs kernel (Σ)

Exponential-squared kernel (Σ)

Random samples of

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Typical bump hunt

- At the LHC we often use $q_0 \equiv -2 \ln \frac{\mathcal{L}(0,m,\widehat{\theta})}{\mathcal{L}(\widehat{\mu},m,\widehat{\theta})}$
 - θ is a vector of nuisance parameters (background, systematics)
 - *m* is the mass parameter we are scanning over
 - μ is the amplitude or «strength» of the signal
- q_0 is distributed as χ_1^2 for background and large enough data samples (asymptotic regime)
- p_0 is the upper-tail probability $p_0 \equiv P(q_0 \ge q_0^{obs})$
- $\sqrt{q_0}$ corresponds to significance of excess, e.g. $q_0 = 16 \Longrightarrow Z = 4, i. e. 4\sigma$
- Use MC «toys» beyond asymptotic regime for $P(q_0 \ge q_0^{obs})$



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