PY410 / 505 Computational Physics 1

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Casino Monte Carlo



- A large number of processes in nature are random
- To some extent, we're all basically familiar with it
 - -Radioactive decay
 - -Flipping a coin
- In fact, quantum mechanics inherently is random

- What computational issues can come up with random numbers?
 - -Computers are, by very nature, NOT random
 - -So we need to make them LOOK random
 - -The question is, how random is "random enough"?
- If you want something truly random, you'll need to hook your computer up to a Geiger counter or something, and count decays (say, from atmospheric muons)

• Sounds silly, but it isn't



Chapter 7 in Numerical Recipes deals with generating random numbers

-"Deviates"

- How to formalize random number generation (RNG)?
 –Given a set S of N uniformly distributed random numbers x1...xN, then they must satisfy :
 - Given n generated numbers, the next number x_n+1 must be independent and uncorrelated
 - x_n+1 should be equally likely to be a member of S
 - -For the example of an unbiased coin toss, S has 1 or 0 (heads or tails)
 - -Each toss is independent of the previous and so a priori equally likely to be 1 or 0

- Why do computers have trouble here?
- They are inherently deterministic at the present time (that's why they're so great to use!)
- We generate "pseudo" random numbers
 - –The statistical properties (i.e. equally likely in some region of interest) is "good enough", but "good enough" depends on the situation
- For instance, if we use <u>linear congruential algorithms</u> of the typ x_{n+1} = (ax_n + c) mod m, "multiplier" "increment" "modulus"
 -This generates a sequence of random integers in the set {0,1,...,m-1}

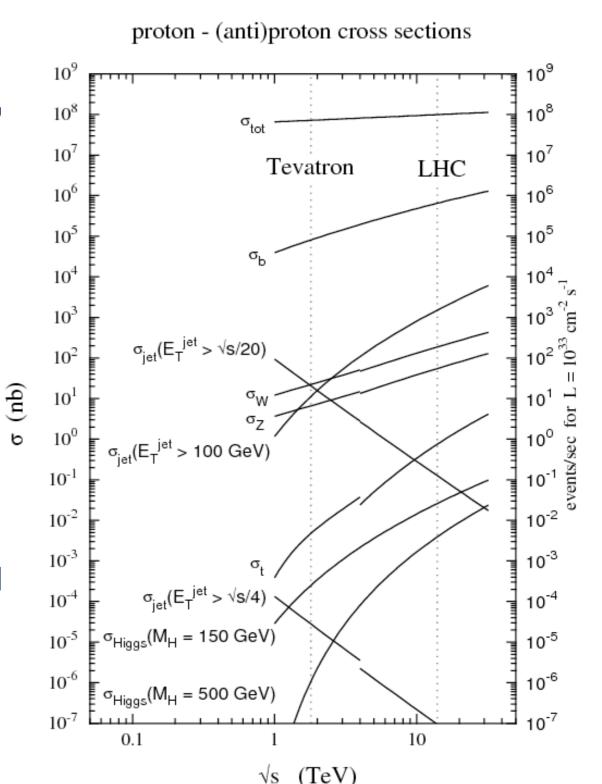
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- Why does this help?
 - -We typically choose a, c, and m to be large(ish) relatively prime numbers
 - -If we inappropriately choose a,c, or m, we can have repetition
 - -If any x_n = x_0, then the process will repeat
 - -We initialize this with a "seed" x0, and away goes the sequence

Source	m	(multiplier) a	(increment) c	output bits of seed in rand() / Random(L)
Numerical Recipes	2 ³²	1664525	1013904223	
Borland C/C++	2 ³²	22695477	1	bits 3016 in rand(), 300 in Irand()
glibc (used by GCC) ^[5]	2 ³¹	1103515245	12345	bits 300
ANSI C: Watcom, Digital Mars, CodeWarrior, IBM VisualAge C/C++ [6]	2 ³¹	1103515245	12345	bits 3016
Borland Delphi, Virtual Pascal	2 ³²	134775813	1	bits 6332 of (seed * L)
Microsoft Visual/Quick C/C++	2 ³²	214013 (343FD ₁₆)	2531011 (269EC3 ₁₆)	bits 3016
Microsoft Visual Basic (6 and earlier) ^[7]	2 ²⁴	1140671485 (43FD43FD ₁₆)	12820163 (C39EC3 ₁₆)	

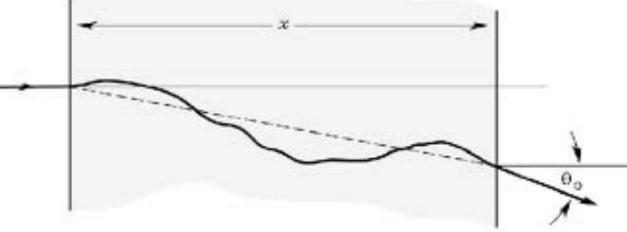
- There are several industrial-strength generators on the market
- But, if you need "really really" random numbers, use with extreme care
 - -C++11, R, Python, Ruby, IDL, Maple, Matlab, GNU MPAL, BOOST, Glib, and NAG :
 - <u>http://en.wikipedia.org/wiki/Mersenne_twister</u>
 - Long period of 2^19937 1
 - Passes lots of randomness tests
 - NOT suitable for cryptography : observing a certain number of iterations will allow you to predict the rest of the sequence
 - -Numerical recipes recommends
 - <u>http://en.wikipedia.org/wiki/Xorshift</u>
 - Period of 2^128 1

- Example: the "sparky" trigger
- In particle physics, we throw away almost all of our data
- Only 1 in 10⁵ is even "remotely interesting
- The really interesting stuff is only 1 in ~10¹⁶!
- Given this, we have to often have "REALLY" random numbers to test
- So, we set up a spark chamber to actually generate random numbers so our triggers could perform adequate tests

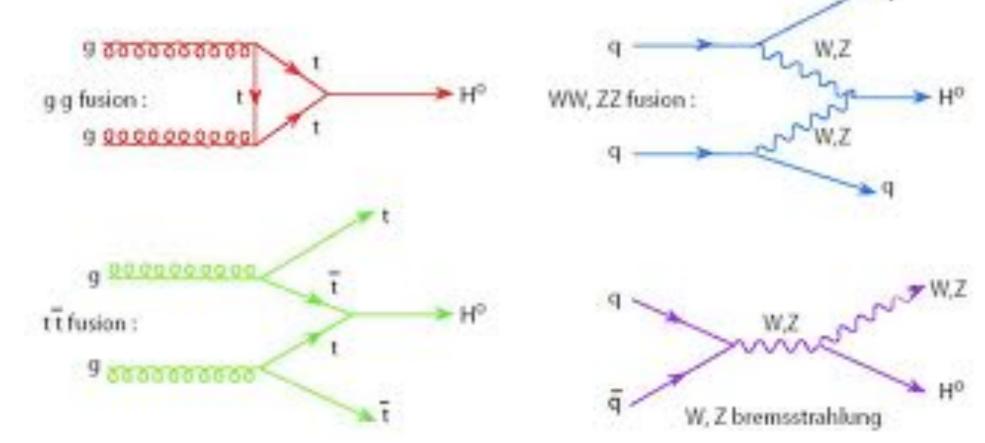


- A few tests will be performed here :
 - -The period of the generator should be much larger than the length of the generated sequence
 - A simple "eyeball test" (plotting (x_n,x_{n+1}) as (x,y) pairs) should reveal no structure
 - -The chi2 statistic should satisfy $\frac{\chi^2}{d.o.f.} \simeq 1$.

- Randomness is often used in simulations
- Lots of things are random in nature
- We often know their distribution, but cannot predict individual events :
 - -Standing a pencil on its side, what is the angle (wrt the desk) when it falls?
 - -What is the final angle between multiple scatterings of particles off of a lattice?



- So, to simulate these events, we need random (or pseudorandom) generators
- Example : Quantum mechanically, can produce Higgs bosons in several ways .



- Can't even ask (quantum mechanically) which will occur before it happens
- They all have a likelihood!
- So, to generate them, you have to randomly sample between these four (with appropriate weights)

 These methods are often called "Monte Carlo" after the Casino Monte Carlo in Monaco

 Much better name than "pseudo-random-number probabilistic event simulator"



- Why does this help?
- Suppose you take a system of 100 electrons in either "spin up" or "spin down" configurations
- Total number of states is $=2^{100} = 1.27 \times 10^{30}$
- Already intractable to list them (at one billion / second, would take 4.02×10^{13} years!)
- If you select one at random for any purpose, need to simply find a sufficiently good randomization algorithm!

- Markov-Chain MC (MCMC) :
 - -Sequence of elements chosen from a fixed set using a probabilistic rule
 - -Chain is constructed by adding the elements sequentially
 - -Given the most recently added element, next element only depends on most recent addition
 - –Formally, suppose x and y are members of a set S
 - -The transition probability function is

$$T(x \to y)$$
, $\sum T(x \to y) = 1$.

y

-Example : Random walks

- Random walks are simple examples of MCMC's
- Suppose a walker can occupy any site on an infinitelylong 1-d lattice
- The walker tosses a coin and decides to go left or right (1 or 0)
- The transition probability is therefore

$$T(x \to y) = \begin{cases} \frac{1}{2} & \text{if } y = x - 1, \\ \frac{1}{2} & \text{if } y = x + 1, \\ 0 & \text{otherwise.} \end{cases}$$

 In equilibrium, suppose the lattice is periodic now, with L lattice points on a circle

-Then we can take L->infinity to examine the properties

• Probability i: $P(x) = \frac{1}{L}$: (walker visits each site the same number of times)

• The walker's position after n steps depends on the sequence of tosses in the past, and cannot be predicted

$$x_n = \sum_{i=1}^n s_i , \qquad \qquad s_i = \pm 1$$

- Over a large number of n-step walks the average is zero –(From symmetry : left and right are equally likely!)
- However, each xn is not zero, and DOES increase:

$$\langle x_n^2 \rangle = \left\langle \sum_{i=1}^n \sum_{j=1}^n s_i s_j \right\rangle$$
$$= \left\langle \sum_{i=1}^n s_i^2 \right\rangle + \left\langle \sum_{i=1}^n \sum_{\substack{j=1\\j \neq i}}^n s_i s_j \right\rangle$$

Diffusion equation!

= n.

RMS displacement is

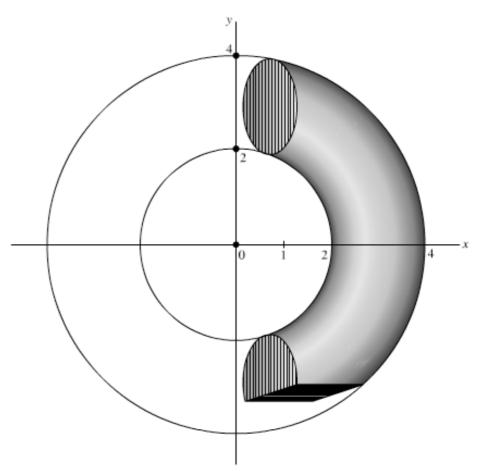
$$\sqrt{\langle x_n^2 \rangle} = \sqrt{n}$$
.

Diffusion constant is

$$\langle x_n^2 \rangle = 2Dn$$
.

• So, for a 1-d walk, D = 1/2

- Another example of MC methods : integration of functions
- Sprinkle the area you're interested with a "dust" of points uniformly distributed
- The fraction "below the curve" (or in N-dim, within the surface) is the integral!



• Formally, the integral is : I =

$$I = \int_{V} d^{d} \mathbf{x} f(\mathbf{x})$$

- Now we choose N uniformly distributed points in V, and estimate $V_{I} \simeq \frac{V}{N} \sum_{n=0}^{N-1} f(\mathbf{x}_{i})$
- The error of repeating this M times is :

$$I_m = \frac{b-a}{N} \sum_{n=0}^{N-1} f(\mathbf{x}_{m,n}) , \qquad m = 0 \dots M - 1$$

Can just compute the mean and standard deviation:

$$\bar{f} = \frac{1}{MN} \sum_{m=0}^{M-1} \sum_{n=1}^{N} f(\mathbf{x}_{m,n}) \qquad \sigma_M = \sqrt{\frac{1}{M} \sum_{m=0}^{M-1} I_m^2 - \left(\frac{1}{M} \sum_{m=0}^{M-1} I_m\right)^2} \\ = \sqrt{\frac{1}{M} \sum_{m=0}^{M-1} \left(I_m - \frac{1}{M} \sum_{m'=0}^{M-1} I_{m'}\right)^2}$$

• If the measurements are independent and randomly distributed about the mean then:

$$\sigma_M^2 = \frac{(b-a)^2}{N} \sigma_f^2$$

• where

$$\sigma_f^2 = \overline{f^2} - (\overline{f})^2 = \frac{1}{MN} \sum_{m=0}^{M-1} \sum_{n=1}^N f(\mathbf{x}_{m,n})^2 - \left(\frac{1}{MN} \sum_{m=0}^{M-1} \sum_{n=1}^N f(\mathbf{x}_{m,n})\right)^2$$

• Thus for MC integration:

$$I = \int_{V} d^{d} \mathbf{x} f(\mathbf{x}) \simeq V \left[\frac{1}{N} \sum_{n=1}^{N} f(\mathbf{x}_{i}) \pm \frac{\sigma_{f}}{\sqrt{N}} \right]$$

- Compare this with our previous integration methods
- For instance, Midpoint method

- We have
$$\label{eq:second} \frac{\sigma_f}{\sqrt{N}} < \frac{\overline{f''}}{N^{2/d}}$$

• MC integration "wins" when d > 4 (roughly)

- What if you want nonuniform samples?
 - -Example : Gaussian distribution, exponential distribution, line segment
 - -Still use the uniformly-distributed "deviates"
 - –Then change variables (x ->y(x)) to produce another distribution :

$$P(y) = P(x) \left| \frac{dx}{dy} \right|$$
, $\int dy P(y) = \int dx P(x) = 1$.

- Uniform line segment :
 - -Basically trivial, you just redefine (0,1) to (a,b) and you're done:

$$y(x) = a + (b - a)x$$
, $\left|\frac{dx}{dy}\right| = \frac{1}{|b - a|}$.

 $P(x) = \begin{cases} 1 & \text{for } 0 < x < 1 \\ 0 & \text{otherwise} \end{cases} \quad \Rightarrow \quad P(y) = \begin{cases} \frac{1}{|b-a|} & \text{for } a < y < b \\ 0 & \text{otherwise} \end{cases}.$

• Exponential distribution :

$$P(y) = \frac{1}{\lambda} e^{-y/\lambda} , \qquad 0 < y < \infty$$

Then we have

$$y(x) = -\lambda \log(x) , \qquad \left| \frac{dx}{dy} \right| = \frac{x}{\lambda} = \frac{e^{-y/\lambda}}{\lambda} = P(y) .$$
$$P(x) = \begin{cases} 1 & \text{for } 0 < x < 1\\ 0 & \text{otherwise} \end{cases}$$

so the distribution is

$$P(y) = \begin{cases} \frac{e^{-y/\lambda}}{\lambda} & \text{for } 0 < y < \infty \\ 0 & \text{otherwise} \end{cases}$$

.

Gaussian distribution ("normal" distribution)

$$P(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2/(2\sigma^2)} ,$$

- Cannot convert with elementary functions
- However, consider the product of TWO Gaussian distributions in x and y :

$$\begin{split} \int dx \int dy P(x) P(y) &= \frac{1}{2\pi\sigma^2} \int dx \int dy e^{-[(x-\mu)^2 + (y-\mu)^2]/(2\sigma^2)} \\ &= \int_0^{2\pi} \frac{d\theta}{2\pi} \int_0^\infty \frac{dr^2}{2\sigma^2} e^{-r^2/(2\sigma^2)} \,, \end{split}$$

• We have switched to polar coordinates:

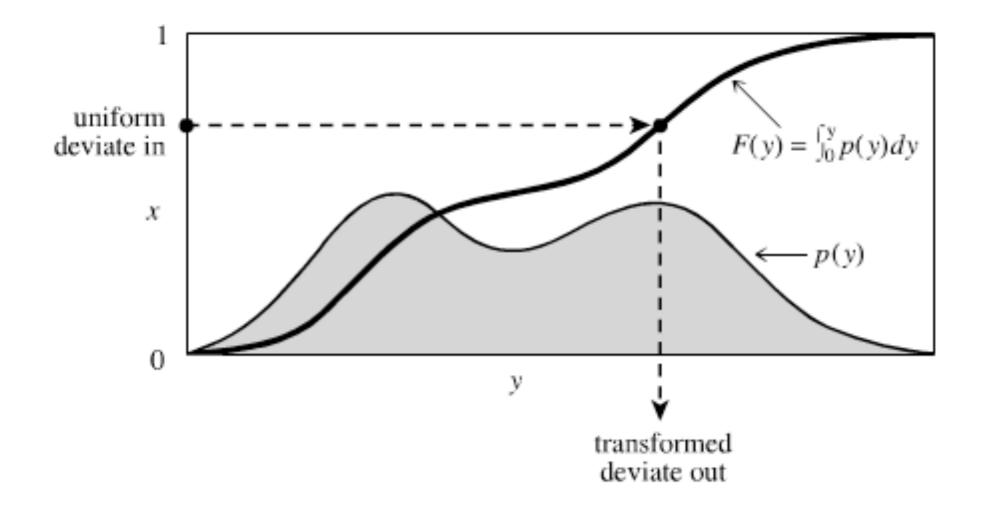
$$x - \mu = r \cos \theta$$
, $y - \mu = r \sin \theta$.

- This is the product of a uniform distribution on [0,2pi] with an exponential distribution in r²!
- So we can generate theta and r, convert to x,y!

$$\theta = 2\pi r_1 , \qquad r = \sqrt{-2\sigma^2 \log r_2} ,$$

-"Box-Muller" algorithm

• General version of transformation method:



-Can be used even if you only know F(y) numerically

- To generate this kind of sequence, though, the Metropolis-Hastings algorithm works well
 - -<u>http://en.wikipedia.org/wiki/Metropolis</u>-Hastings_algorithm
 - -Type of Markov Chain Monte Carlo (MCMC)
 - –One major advantage : it does not require the overall normalization of a distribution to be known to draw a random sequence from it!
 - Advantageous when using Bayesian statistics
- Original papers are from Metropolis et al and Hastings :
 - -J. Chem. Phys. 21, 1087 (1953)
 - -Biometrika 57, 97 (1970)

- This is a Markov Chain MC so the "next" step depends only on the "current" step
 - -This is a type of random walk, similar to other MCMC methods
- Suppose we want to generate a sequence from a probability distributic $P(\vec{x})$
- If we visualize this in d-dimensions, then P(x) is a mountainous terrain and the sequence is the steps of a hiker



- Heuristically :
 - -Explore the "terrain" a bit
 - Find peaks, valleys
 - -Then generate the steps for the MC application
 - -Record steps every so often
- As more and more sample values are produced, the distribution more closely approximates the desired distribution P(x)

• More specifically, we have steps :

"thermalization" steps to find the peaks

- Why sample?
 - -Individual steps may be correlated
 - If walker's step size is much smaller than typical distances between "peaks" in the terrain
 - To emulate a "truly" random sequence, we don't want to get "stuck" between peaks

- Walking algorithm :
- Suppose we have the walker at a point $ec{x}_{ ext{current}}$
- Then the next step $ec{x}_{ ext{next}}$ is determined by
 - –Choose a "trial" point randomly in the neighborhood of current x ($\vec{x}_{\rm trial}$)
 - 1-d : maximum step size (delta) may be used, and choose trial x in a uniform interval $[x_{\rm current}-\delta,x_{\rm current}+\delta]$
 - n-d : could try a fixed sphere of radius delta
 - -Requirement is that the probability $P_{try}(\vec{x}_{current} \rightarrow \vec{x}_{trial})$ should be SYMMETRIC: $P_{try}(\vec{x} \rightarrow \vec{y}) = P_{try}(\vec{y} \rightarrow \vec{x})$
 - In other language, the conditional probabilities must be equal :

$$P(x|y) = P(y|x)$$

-Calculate the next step by the ratio of probabilities (or "likelihoods") $r = P(\vec{x}_{\text{trial}})/P(\vec{x}_{\text{current}})$

• Given
$$r = P(\vec{x}_{\text{trial}})/P(\vec{x}_{\text{current}})$$

- If r >=1, walker is moving "uphill" (toward higher probability)
 - -Trial step is accepted, iterate
 - –The "acceptance probability" is 1 by definition (it's accepted, $P_{\rm accept}(\vec{x}_{\rm current} \to \vec{x}_{\rm trial}) = 1$)
- If r < 1, walker is moving "downhill" (toward lower probability)
 - -Accept only if displacement is not too large :
 - Generate uniform deviate $n_{
 m random}$
 - Accept if $r > n_{
 m random}$, iterate
 - Reject otherwise, remain at current "step"
 - "Acceptance probability" in this case is r

- What about those thermalization steps? Do they really thermalize?
- Imagine a large ensemble of walkers distributed with density $\rho(\vec{x})$
- Want to show that the ensemble density becomes proportional to P(x) and then remain invariant
- According to the ergodic hypothesis, the average properties of an ensemble of systems should be the same as the average over time in a single system in thermal equilibrium <u>-http://en.wikipedia.org/wiki/Ergodic_hypothesis</u>
- So, we'll use the former to demonstrate the latter

 Recall the transition probability for a walker to move from x to y is determined by two consecutive and independent decisions:

$$T(\vec{x} \to \vec{y}) = P_{\text{try}}(\vec{x} \to \vec{y}) \times P_{\text{accept}}(\vec{x} \to \vec{y})$$
.

• If these T(x->y) are symmetric, then:

$$\frac{T(\vec{x} \to \vec{y})}{T(\vec{y} \to \vec{x})} = \frac{P_{\text{accept}}(\vec{x} \to \vec{y})}{P_{\text{accept}}(\vec{y} \to \vec{x})} = \frac{P(\vec{y})}{P(\vec{x})}$$

• Second equality can just be verified directly: $-if P(y) > P(x), P_{accept}(\vec{x} \rightarrow \vec{y}) = 1$, $P_{accept}(\vec{y} \rightarrow \vec{x}) = P(\vec{x})/P(\vec{y})$ $-if P(y) < P(x), P_{accept}(\vec{x} \rightarrow \vec{y}) = P(\vec{y})/P(\vec{x})$, $P_{accept}(\vec{y} \rightarrow \vec{x}) = 1_{e}$

• The change in number density of walkers at x when all the walkers in the ensemble take a step:

$$\begin{split} \Delta\rho(\vec{x}) &= \int d^d y \left[\rho(\vec{y}) T(\vec{y} \to \vec{x}) - \rho(\vec{x}) T(\vec{x} \to \vec{y})\right] \\ &= \rho(\vec{x}) \int d^d y \ T(\vec{y} \to \vec{x}) \left[\frac{\rho(\vec{y})}{\rho(\vec{x})} - \frac{P(\vec{y})}{P(\vec{x})}\right] \,. \end{split}$$

- If rho(y)/rho(x) exceeds "equilibrium value" P(y)/P(x):
 –will get a + correction
- If rho(y)/rho(x) is too small :
- -will get a correction
- If rho is proportional to P, we'll have "detailed balance": $\rho(\vec{y})T(\vec{y} \rightarrow \vec{x}) = \rho(\vec{x})T(\vec{x} \rightarrow \vec{y})$

- Choices to make :
 - -Number of walkers
 - -Step size
 - -Thermalization steps
 - -Sampling frequency
- All of these depend on the problem at hand
- You have to generally use trial and error to make sure there aren't large dependencies on the parameters

- An excellent example of using probabilistic methods is in statistical mechanics
- Makes sense : it's all about probability and statistics!
- A few definitions :
 - -Microstates : configuration of a system

-Probability of ith microstate is

$$p_i = \lim_{N \to \infty} \frac{n_i}{N},$$
Number of systems in the ith microstate

-Average value (e.g. for energy) $i\langle E \rangle = \sum_{i} p_i E_i$.

-Variance is
$$\langle (E - \langle E \rangle)^2 \rangle = \sum_i p_i E_i^2 - \left(\sum_i p_i E_i\right)^2$$
.

- Examine the canonical and microcanonical ensembles:
 - -Canonical:
 - fixed number of constituents (N), fixed volume (V)
 - exchange energy with thermal reservoir at temperature T
 - Probability of a microstate with energy E is Bolzmann distribution:

$$p_i = \frac{e^{-E_i/(k_{\rm B}T)}}{Z} ,$$

Partition function is:

$$Z(N, V, T) = \sum_{i} e^{-E_i/(k_{\rm B}T)}$$

- -Microcanonical:
 - fixed number of constituents (N). fixed total energy (E). fixed
 - volume (V) Probability of a microstate $p_i = \begin{cases} \frac{1}{N} & \text{if } E = E_i \\ 0 & \text{otherwise} \end{cases}$ with energy E is :

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- Example of a system we can investigate is the hard-disk gas in two dimensions
 - Investigated by Metropolis et al with the MC methods above
 - -J. Chem. Phys. 21, 1087 (1953)
 - –"Equation of State Calculations by Fast Computing Machines"
 - "Fast" here meaning it could multiply two 40-bit ints in 1 millisecond (1 kHz)
 - For contrast, one core on your phone is 2.4 million times faster

 Look at the system as a Maxwell-Boltzmann gas at fixed volume and temperature

• Energy is
$$E = \text{K.E.} + \text{P.E.} = \frac{m}{2} \sum_{i} \mathbf{v}_{i}^{2} + \sum_{\text{pairs } ij} U(r_{ij})$$
,

• Pairwise potential energy function :

Num

$$U(r) = \begin{cases} 0 & \text{if } r > \sigma \\ \infty & \text{if } r \le \sigma \end{cases},$$

• In an ensemble of systems, at temperature T, probability that the system has energy E is a Boltzmann distribution

$$\sim W \exp\left[-\frac{E}{k_{\rm B}T}\right] = \exp\left[-\frac{E - TS}{k_{\rm B}T}\right] = \exp\left[-\frac{F}{k_{\rm B}T}\right],$$
Number of
microstates Entropy

Partition function is

$$Z = \sum_{E} W(E) \exp\left[-\frac{E}{k_{\rm B}T}\right]$$

.

.

 The equation of state relates pressure p, volume V and temperature T:

$$pV = Nk_{\rm B}T \left. \frac{\partial \log Z}{\partial \log V} \right|_{T,N}$$

 MC simulation of hard disks from Metropolis et al: -Close-pack N=224 disks of diameter d0 as follows

1090 METROPOLIS, ROSENBLUTH, ROSENBLUTH, TELLER, AND TELLER

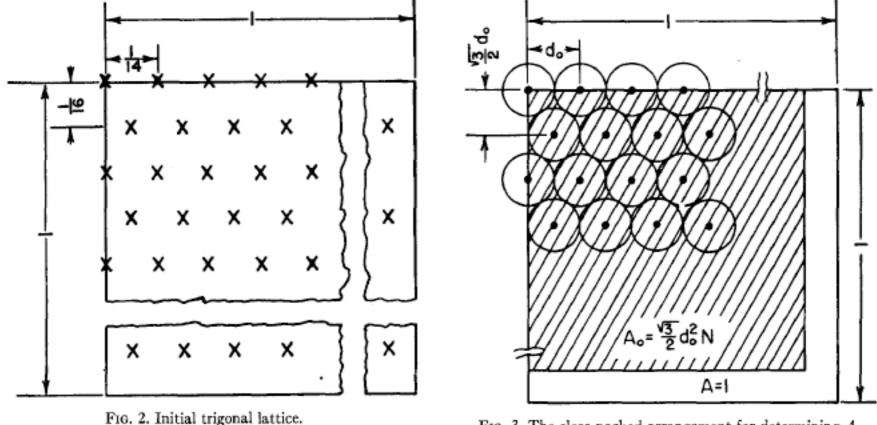


FIG. 3. The close-packed arrangement for determining A_0 .

- Fix $A_0 = \sqrt{3}d_0^2 N/2 = 1$ and vary d0, hence number is N
- Apply periodic boundary conditions

- Radial distribution function measures correlations between particles separated at distance r
- Can be used to distinguish solid, liquid, gases
 - -See <u>Gould-Tobochnik Chapter 8</u> Section 8.5
- Equation of state is deduced from the radial distribution function:

CALCULATION OF STATE BY FAST MACHINES

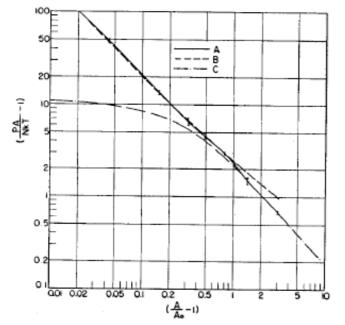


FIG. 4. A plot of (PA/NkT) = 1 versus $(A/A_0) = 1$. Curve A (solid line) gives the results of this paper. Curves B and C (dashed and dot-dashed lines) give the results of the free volume theory and of the first four virial coefficients, respectively.

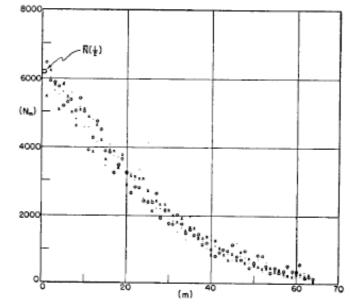
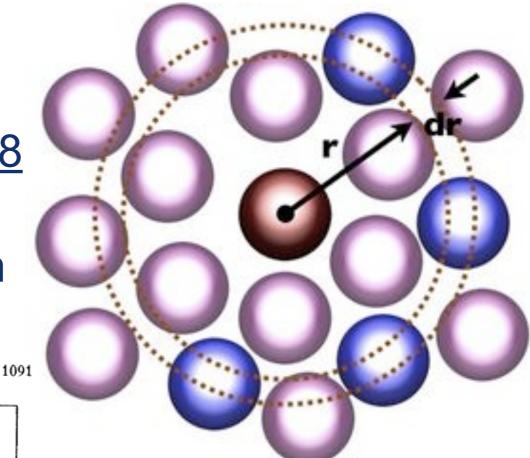


FIG. 5. The radial distribution function N_m for $\nu = 5$, $(A/A_0) = 1.31966$, K = 1.5. The average of the extrapolated values of N_{i} in $N_i = 6301$. The resultant value of (PA/NkT) - 1 is $64N_i/N^3(K^2-1)$ or 6.43. Values after 16 cycles, \bullet ; after 32, \times ; and after 48, \bigcirc .



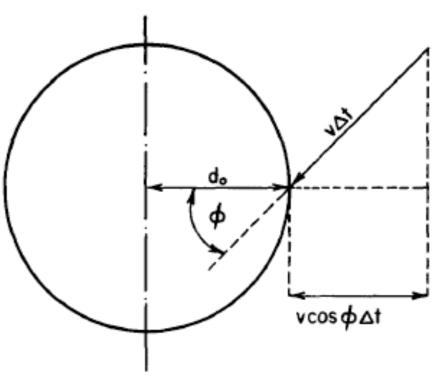
• They showed

$$PA = Nk_{\rm B}T\left(1 + \frac{\pi d_0^2 \overline{n}}{2}\right)$$
, where $\overline{n} = n(d)$,

• Using the Virial Theorem :

-States that
$$_{\text{K.E.}} = N \times \frac{1}{2}m\overline{v}^2 = pA + \frac{1}{2}\left\langle \sum_i \mathbf{r}_i \cdot \mathbf{X}_i^{\text{int}} \right\rangle$$

 Derivation is based on this figure, defines variables involved in a collision betweer two disks:



They showed

$$\left\langle \sum_{i} \mathbf{r}_{i} \cdot \mathbf{X}_{i}^{\text{int}} \right\rangle = -\frac{1}{2} \sum_{i} \sum_{j \neq i} r_{ij} F_{ij} = N \times \frac{1}{2} m \overline{v}^{2} \times \pi d_{0}^{2} \overline{n} \,.$$

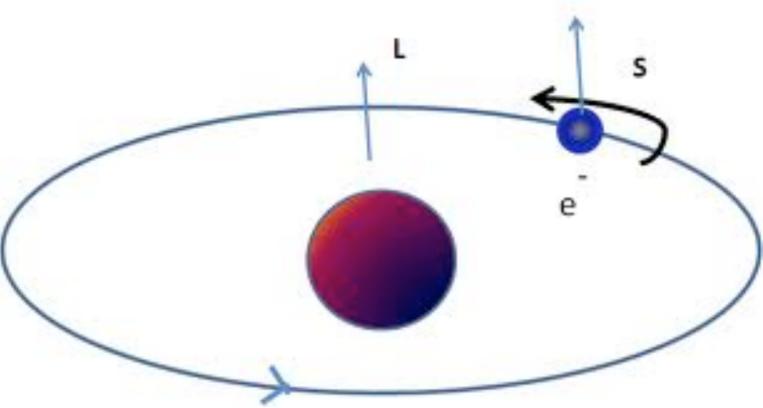
- MC simulation measured the radial distribution function as a histogram:
 - -For each MC configuration
 - For each disk:
 - Divide region from r=d to r=rmax into 64 annular zones of equal area
 - Count disks in each zone, store in histogram
 - -Average over configurations

-Fit histogram to model function and extrapolate to r=d

- Another very popular application of probabilistic methods is the simulation of idealized magnetic materials
- We use the Ising Model
 - -http://en.wikipedia.org/wiki/Ising_model
- William Lenz assigned this to his student Ernst Ising as a PhD thesis problem
- This was solved by Ising exactly for 1-d
 - -Didn't exhibit ferromagnetism
 - Details of 1-d Ising model : <u>Gould-Tobochnik Chapter 5</u>.
- Kramers, Wannier, Onsanger, et al worked on the 2-d model
 - Richer phenomenology :
 ferromagnetic phase at low temperatures, ↓
 paramagnetic phase at high temperatures ↑
 - –Second order phase transition at the Curie temperature $T_{\rm C}$

\downarrow	\downarrow	↑	↑	↑	\downarrow	\uparrow	\downarrow	\uparrow	↑
	\downarrow						\uparrow		
\downarrow	\uparrow	\downarrow	\uparrow	\downarrow	\uparrow	\downarrow	\downarrow	\uparrow	\uparrow
							\uparrow		
\uparrow	\downarrow	\uparrow	\downarrow	\downarrow	\downarrow	\uparrow	\uparrow	\downarrow	↑
\uparrow	↑	\downarrow	\downarrow	\downarrow	\downarrow	\uparrow	\uparrow	\uparrow	\downarrow
,↓	\downarrow	\uparrow	\downarrow	\uparrow	\downarrow	\uparrow	\downarrow	\uparrow	\downarrow
							\uparrow		
							\downarrow		
\downarrow	\downarrow	\uparrow	\downarrow	\downarrow	\uparrow	\downarrow	\uparrow	\uparrow	\uparrow

- Recall : magnetism caused by charged particles "spinning" in closed orbits or about their axes
- For elementary particles, of course, we mean "spinning" in the quantum mechanical sense, not the rotational Newtonian sense
- So atoms will have both an "L" and and "S" contribution to the magnetic properties



• We'll take a simple classical approximation called the "Ising spin" which has two values :

$$s_i = \begin{cases} +1, & \text{represents "spin up"} \uparrow \\ -1, & \text{represents "spin down"} \downarrow \end{cases}$$

2-d magnet is a set of Ns spins on a lattice

- The force between magnets falls of like r^{-3}
- So, approximate that the spins only interact with their nearest neighbors :

• The interaction energy is proportional to the alignment of the spins :

$$E = -J\sum_{\langle ij\rangle} s_i s_j - H\sum_i s_i$$

- If J > 0 : ferromagnetic
 - Energy is minimized if the spins point in the same direction
- If J < 0 : antiferromagnetic
 - Energy is minimized if spins locally point in the opposite direction
- "H" is an external field which couples to the total magnetization: $M = \sum s_i$,
- Spins will align to this external magnetization

- For H=0, the system will be in one of two states :
 - -Low temp (below Curie temperature) : magnetized
 - High temp (above Curie temperature) : sum of magnetization is zero
- Critical value is the second-order phase transition between ferromagnetic to paramagnetic phases

- We want to compute the observables (averages over the samples)
- Define a state as a specific set of spin values (up or down)

$$s_1 = +1, s_2 = -1, s_3 = +1, \ldots, s_{N_s} = +1$$

- Avg. of an observable calculated by weighting configurations by the Boltzmann factor
- Example : average magnetization :

$$\langle M \rangle = \frac{\sum_{\text{configs}} M e^{-E/k_{\text{B}}T}}{\sum_{\text{configs}} e^{-E/k_{\text{B}}T}}$$

- Total number of configurations is very large
- If we have N spins, we have 2^N configurations
- If L=20, N=20*20 = 400, so 2400 configurations (2.6e120 !!!)
- That's completely intractable to even list them, much less compute the total magnetization
- So instead, we use Monte Carlo methods
 - -We generate a reasonable number of configurations at random
 - Use the Boltzmann factors to define the probability of each
 - -Perfect for MC methods!

• The probability for each state is :

$$p(s_1, s_2, \dots, s_{N_s}) = \frac{e^{-E(s_1, s_2, \dots, s_{N_s})/k_{\rm B}T}}{\sum_{\rm configs} e^{-E/k_{\rm B}T}}$$

- Now, we need to generate N statistically independent configurations according to this probability
- The average magnetization and energy are :

$$\langle M \rangle = \frac{1}{N} \sum_{i=1}^{N} M(s_1^{(i)}, s_2^{(i)}, \dots, s_{N_s}^{(i)}) ,$$

$$\langle E \rangle = \frac{1}{N} \sum_{i=1}^{N} E(s_1^{(i)}, s_2^{(i)}, \dots, s_{N_s}^{(i)})$$

- Now use the Metropolis algorithm from last class to "map out" the configuration space
 - -Choose initial configuration
 - -For all spins :
 - Trial flip (+1 --> -1 or vice versa)
 - Compute change in energy
 - If $w = e^{-\Delta E/k_{\rm B}T} > r$, flip the spin (where r is a uniform deviate)
- Make sure to pick out the thermalization and skip steps

- IT is expensive to compute all of the Boltzmann factors at each step
- Instead, we can realize that for a simulation at fixed T and H, there are only 10 distinct values
- Can precompute them and then just refer to that
- First, consider the sum of the 4 neighboring spins:

$$\sum_{i=1}^{n} s_{i} = s_{(i_x+1,i_y)} + s_{(i_x-1,i_y)} + s_{(i_x,i_y-1)} + s_{(i_x,i_y+1)} \cdot s_{(i_x,$$

neighbors j

• The value is

- $\sum_{\text{neighbors } j} s_j = +4$ if all four neighbors point up, $\sum_{\text{neighbors } j} s_j = +2$ if three neighbors point up and one down, $\sum_{\text{neighbors } j} s_j = 0$ if two neighbors point up and two down,
- $\sum_{\text{neighbors } j} s_j = -2$ if one neighbor points up and three down, $\sum_{\text{neighbors } j} s_j = -4$ if all four neighbors point down.
- Product can only have one of these five values!

$$s_i \sum_{\text{neighbors } j} s_j = +4, +2, 0, -2, -4$$
.

- Now, consider H != 0
- Have terms like H*s_i
- For those we get the other 5 terms
- Can store this as a 2-d array of 5 rows and 2 columns

-First index :
$$2 + \frac{1}{2}s_i \sum_{\text{neighbors } j} s_j = 0, 1, 2, 3, 4$$
,
-Second index : $\frac{1+s_i}{2} = 0, 1$.

- Metropolis step :
 - -Choose a spin at random
 - -Take the metropolis step
 - -If the spin is at the boundary, use periodic boundary conditions
- Because the MC step of a single spin-flip will be highlycorrelated with "this" step, we need to make sure to choose sufficiently high sampling factors
 - -Typically need AT LEAST N steps
 - -Each spin has a chance to flip

 Below the Curie temperature, Kramers and Wannier showed that there is a duality argument to compute the exact value : <u>Phys. Rev. 60, 252 (1941)</u>,

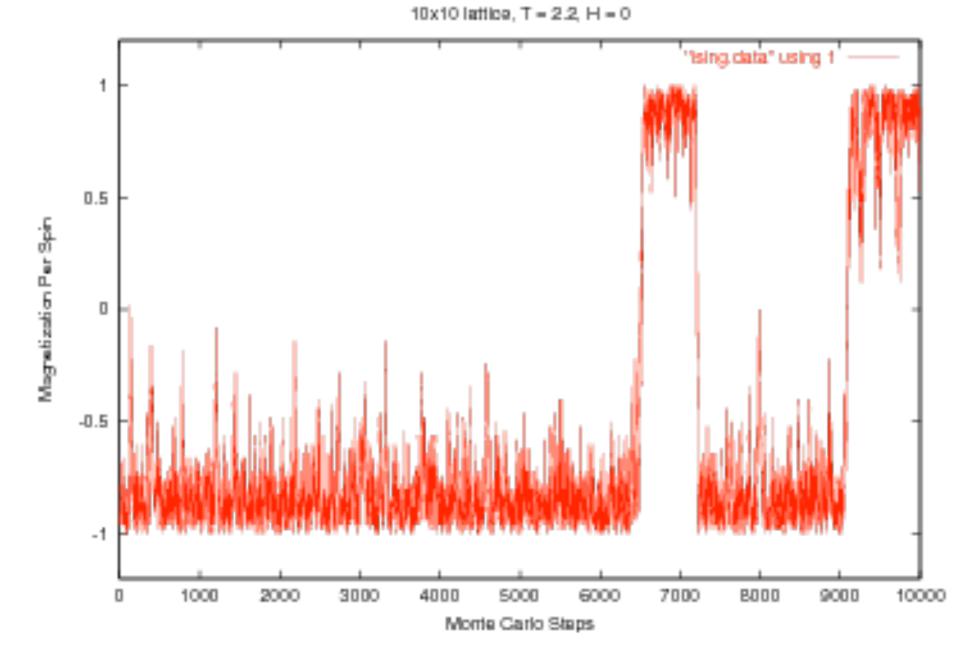
$$\frac{k_{\rm B}T_{\rm c}}{J} = \frac{2}{\log(1+\sqrt{2})} = 2.269\dots$$

 Osnager showed the same model in the thermodynamic limit N-> infinity with H=0 (Phys. Rev. 65, 117 (1944))

$$m = \lim_{N \to \infty} \frac{\langle \sum_i s_i \rangle}{N} = \begin{cases} \left[1 - \left\{ \sinh\left(\frac{2J}{k_{\rm B}T}\right) \right\}^{-4} \right]^{1/8}, & \text{for } T \le T_{\rm c} \\ 0, & \text{for } T > T_{\rm c} \end{cases}$$

- Near the Curie temperature $m \sim (T_c T)^\beta \;, \label{eq:mass}$
- Beta = 1/8 for the 2-d Ising model

• Here are some numerical results for T = 2.2, below Tc



• Note : for finite systems, spins can flip from time to time!

• And now for T=3.0, above Tc:

