

Drift vs Shift: Decoupling Trends & Changepoint Analysis

David S. Matteson, with Haoxuan Peter Wu & Sean Ryan

Cornell University (TRIPODS w/URochester) & the National Institute of Statistical Sciences (NISS)

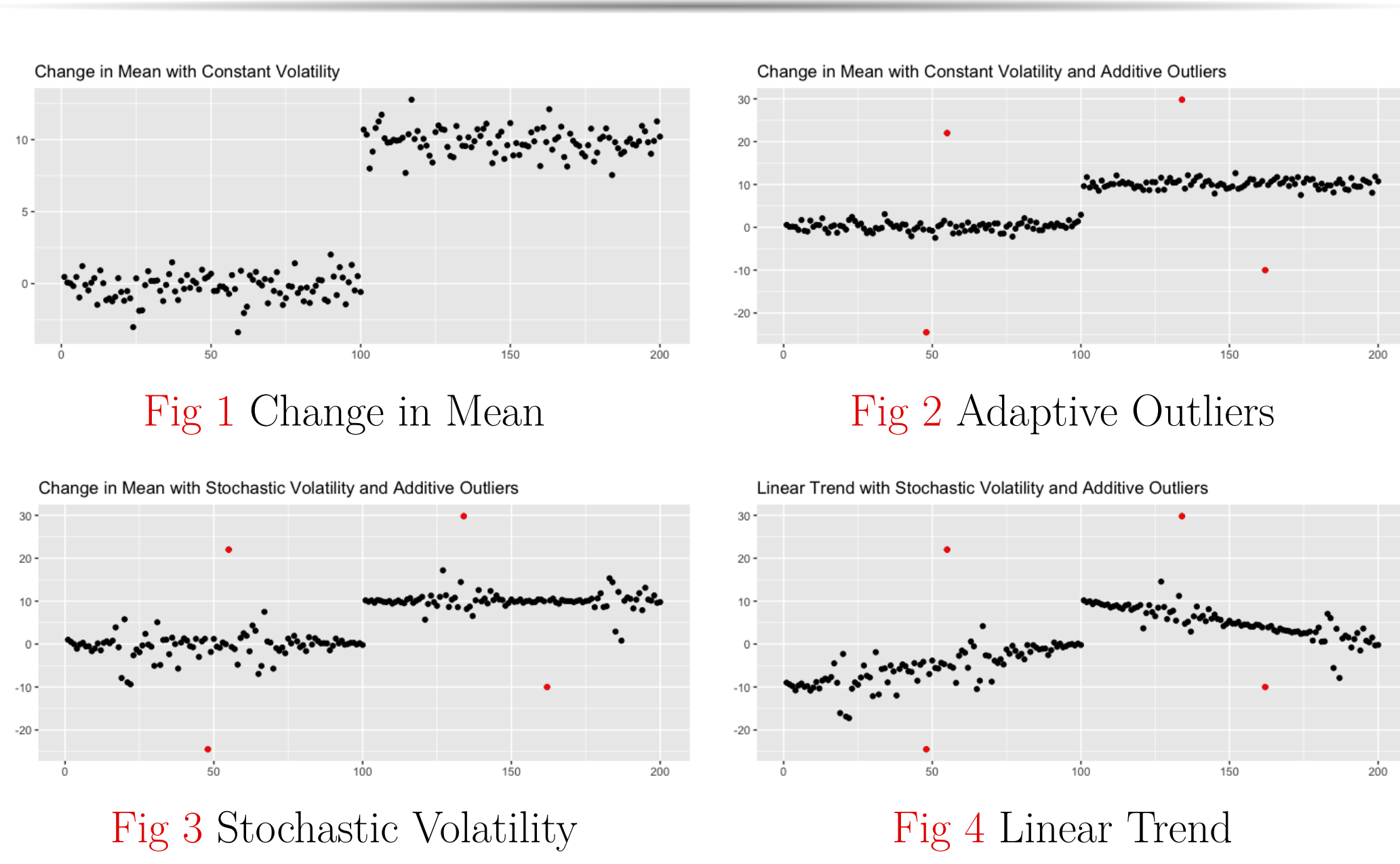
Introduction

- Goal:** Distinguish global/macro patterns from local/micro fluctuations
- 'Drift' describes the micro-level evolution of a process. This may appear as variation about gradual trends.
- 'Shifts' refer to discontinuities, rapid changes, or major breaks in trend. These represent macro-level changes in a process.
- Both might be mechanically or stochastically generated and/or modeled. However, causes of shifts are typically different from those of drift.
- While understanding such differences is a prime objective, this first requires distinguishing: **Drift vs Shift**.

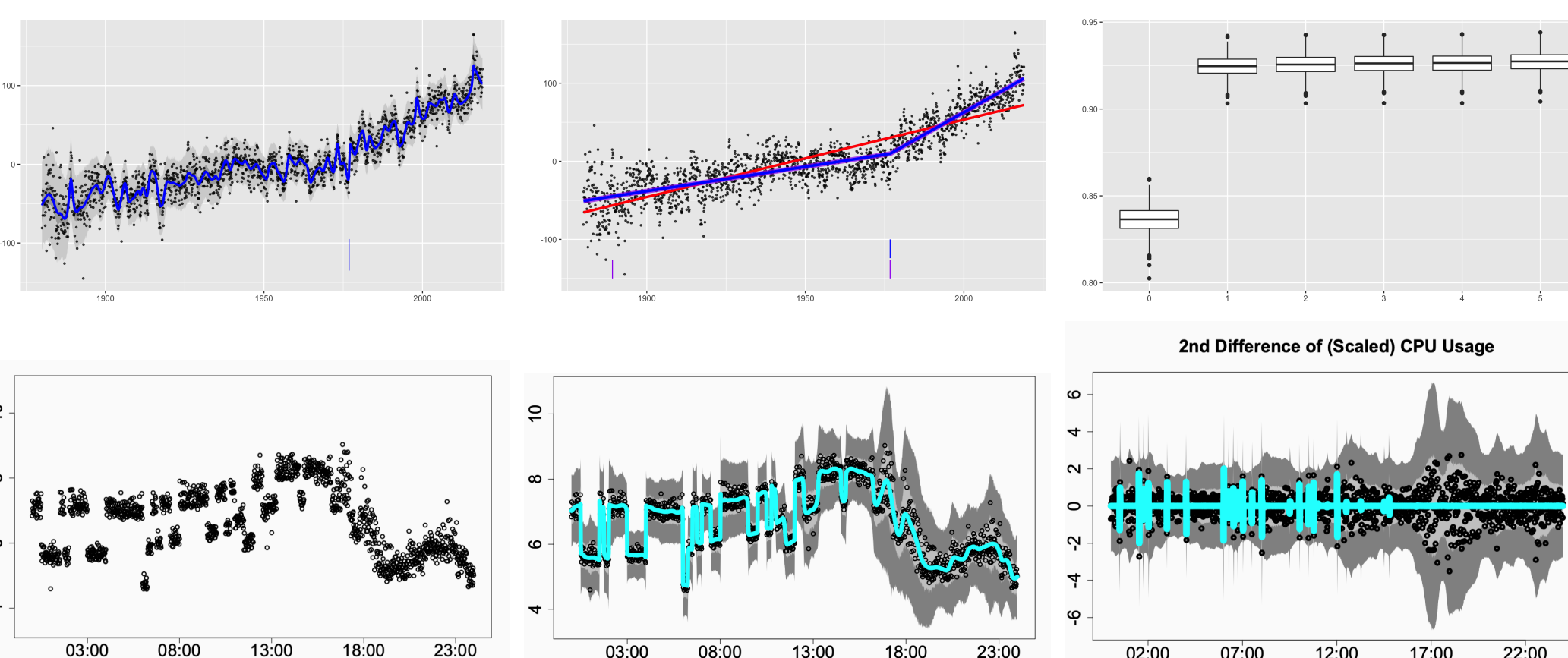
Tools include:

- Trend Filtering
- Stochastic Volatility
- Outlier Detection
- Dynamic/Adaptive Shrinkage
- Dynamic Linear Models (DLM)
- Change Point Analysis
- Bayesian (Time Series) Analysis
- Machine Learning (Regularization)

Challenges



- Outliers** violate common Gaussian noise assumptions.
- Heterogeneity** leads to over-prediction of changepoints.



- Real world data has complex patterns and trends.
- Outliers and heterogeneity are the norm.
- Nature of changepoints ambiguous.

Solutions

- Model based **ABCO**: Adaptive Bayesian Changepoints w/ Outliers[1].
- A two-step Bayesian 'decoupling' method developed via DLM[2].

ABCO Model

Given a time series $\{y_t\}$, ABCO supposes the decomposition:

$$y_t = \underbrace{\beta_t}_{\text{mean signal}} + \underbrace{\zeta_t}_{\text{additive outlier}} + \underbrace{\epsilon_t}_{\text{heteroskedastic noise}}$$

Trend Signal $\{\beta_t\}$

Ref 'Dynamic Shrinkage Process' [3], ABCO uses **global-local shrinkage priors** on the D th order difference (Δ^D , $D = 1, 2$) on the state variable $\{\beta_t\}$:

$$\Delta^D \beta_t = \omega_t, \quad \omega_t \sim N(0, \tau_\omega^2 \lambda_{\omega,t}^2 = e^{h_t}),$$

$$h_{t+1} = \mu + (\phi_1 + \phi_2 s_t)(h_t - \mu) + \eta_{t+1}, \quad \eta_{t+1} \sim Z(\alpha, \beta, 0, 1).$$

Z-distribution: log inverted Beta; heavy left-tail

- **Changepoint:** threshold γ & indicator $s_t = \begin{cases} 1 & \text{if } \log(\omega_t^2) > \gamma \\ 0 & \text{if } \log(\omega_t^2) \leq \gamma \end{cases}$

Additive Outlier $\{\zeta_t\}$

The outlier term $\{\zeta_t\}$ follows a 'horseshoe+' shrinkage prior:

$$\zeta_t | \sigma_{\zeta,t} \sim N(0, \sigma_{\zeta,t}^2)$$

$$(\sigma_{\zeta,t} | \tau_\zeta, \eta_{\zeta,t}) \sim C^+(0, \tau_\zeta \eta_{\zeta,t})$$

$$\tau_\zeta \sim C^+(0, \sigma_{\tau,\zeta})$$

$$\eta_{\zeta,t} \sim C^+(0, \sigma_{\eta,\zeta})$$

with half-Cauchy $C^+(\cdot)$ and prior shrinkage hyper-parameters $\sigma_{\tau,\zeta}, \sigma_{\eta,\zeta}$.

- **Outlier:** custom cutoff & locally adaptive score $o_t := \tilde{E} \left(\frac{\sigma_{\zeta,t}^2}{\sigma_{\zeta,t}^2 + \sigma_{\epsilon,t}^2} \right)$

Heteroskedastic Noise $\{\epsilon_t, \sigma_{\epsilon,t}^2\}$

The noise $\{\sigma_{\epsilon,t}^2\}$ follows a stochastic volatility model of order 1.

$$y_t = \beta_t + \zeta_t + \epsilon_t, \quad \epsilon_t \sim N(0, \sigma_{\epsilon,t}^2),$$

$$\log(\sigma_{\epsilon,t}^2) = \mu_\epsilon + \phi_\epsilon (\log(\sigma_{\epsilon,t-1}^2) - \mu_\epsilon) + \xi_{\epsilon,t}, \quad \xi_{\epsilon,t} \sim N(0, \sigma_\xi^2).$$

Dynamic Regression Generalizations

Set $\mathbf{x}_t = (x_{1,t}, \dots, x_{p,t})$ as p predictors at time t , and $\omega, \mathbf{h}, \mu, \phi, \eta$ analogously.

$$y_t = \mathbf{x}_t' \beta_t + \zeta_t + \epsilon_t \quad \Delta^D \beta_{t+1} = \omega_t$$

$$\omega_{j,t} \sim N(0, \tau_{\omega,j}^2 \sigma_{\omega,j}^2 \lambda_{\omega,j,t}^2 = e^{h_{j,t}}) \quad \mathbf{h}_{t+1} = \mu + (\phi_1 + \phi_2 \mathbf{s}_t)(\mathbf{h}_t - \mu) + \eta_{t+1}$$

ABCO Simulations

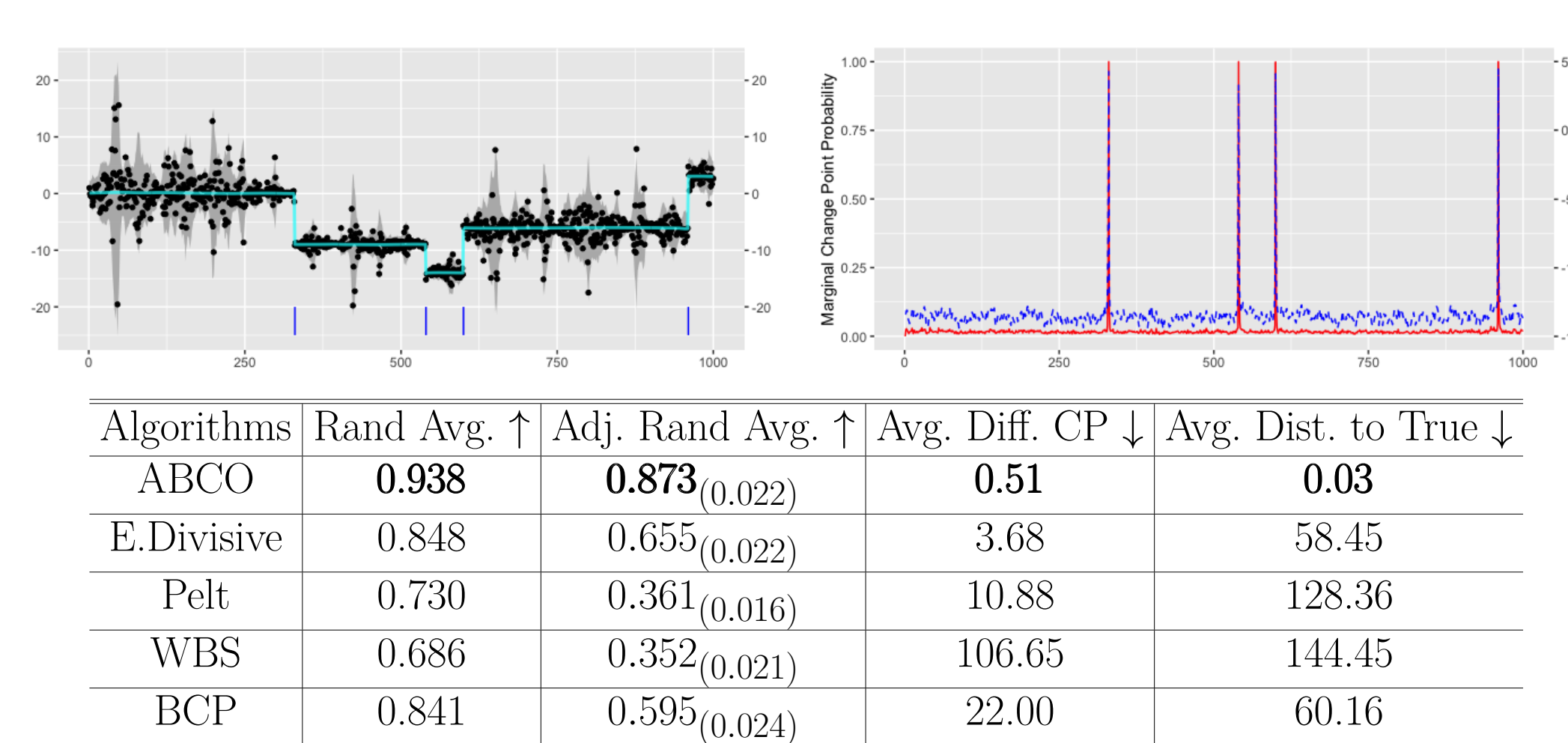
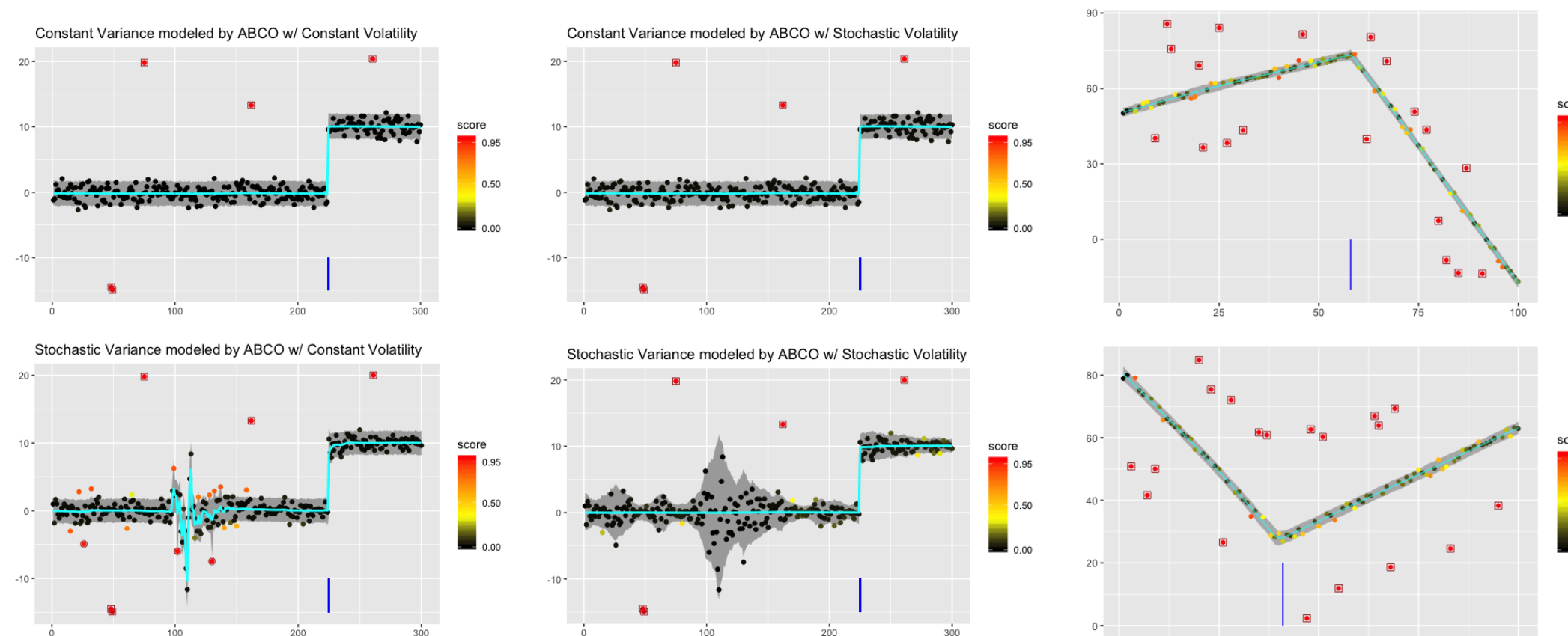


Fig 6 Length-1000 time series with shift and SV(1) variance model $\log(\sigma_{\epsilon,t}^2) = \phi_\epsilon \log(\sigma_{\epsilon,t-1}^2) + \alpha_t$, $\alpha_t \sim N(0, \sigma_\alpha^2)$, $\phi_\epsilon = 0.9$, $\sigma_\alpha = 0.4$.



ABCO Applications

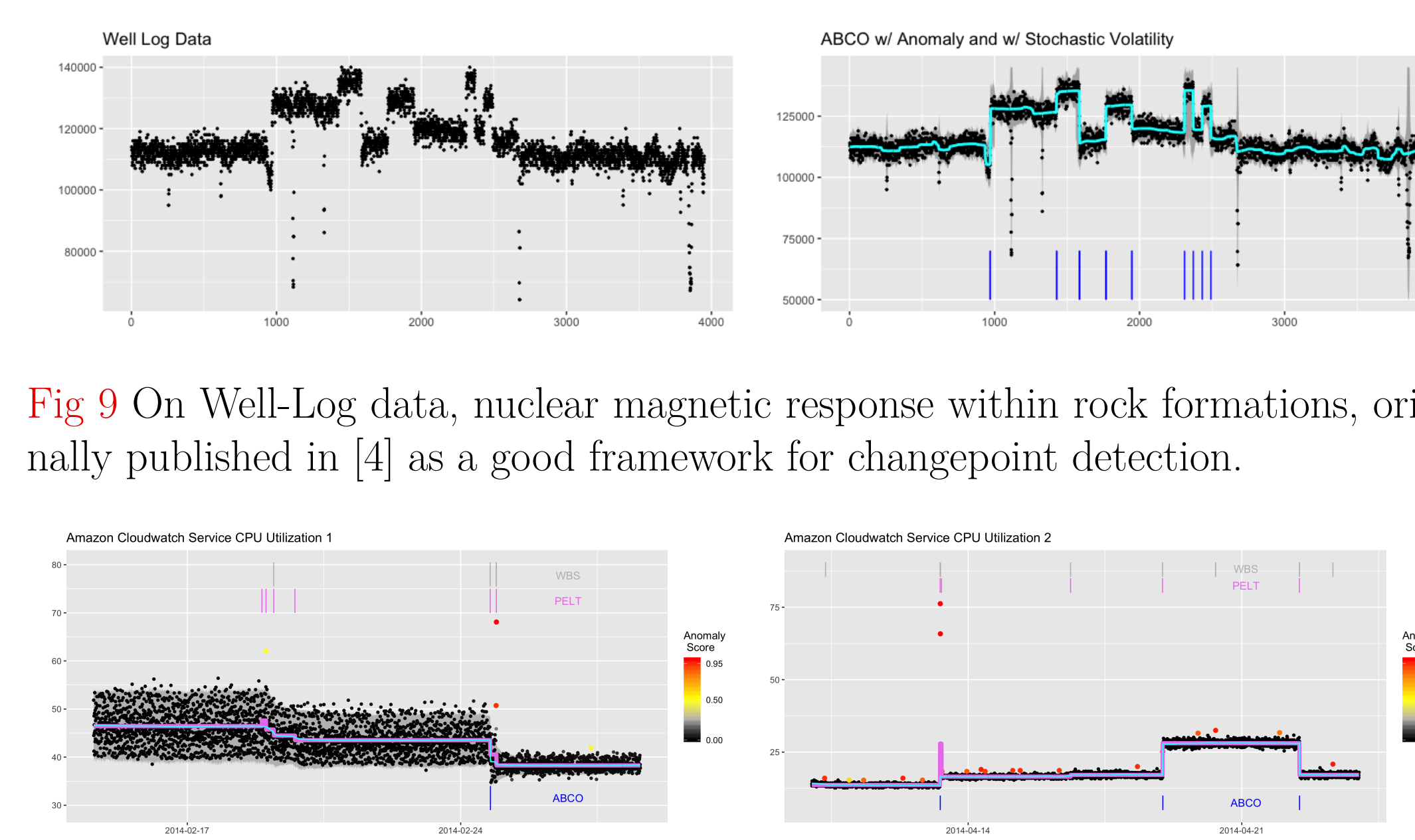


Fig 10 On Amazon Cloudwatch Service CPU Utilization data.

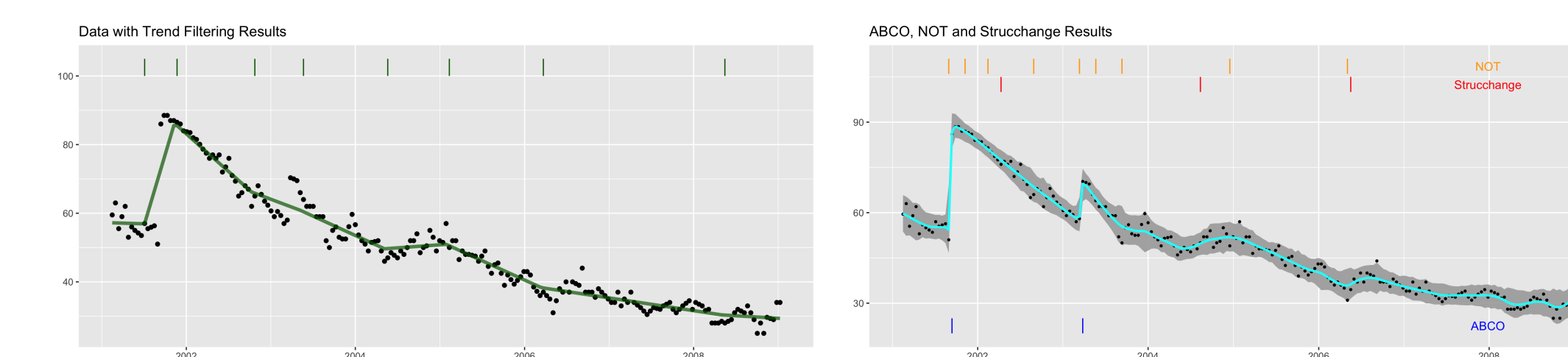


Fig 11 On George W. Bush Approval Rating data.

Decoupling Approach

Dynamic Linear Models (DLM)

Given a time series $\mathbf{Y} = (y_1, \dots, y_n)'$, a predictor series $\mathbf{X} = (x_1, \dots, x_n)'$,

$$y_t = x_t \beta_t + \epsilon_t, \quad \epsilon_t \sim N(0, \sigma_{\epsilon,t}^2),$$

$$\Delta^D \beta_t = \omega_t, \quad \omega_t \sim N(0, \sigma_\omega^2).$$

Decoupled Regularized Loss

Denote $\tilde{\beta}$ as the posterior mean of k MCMC draws $\{\beta^{(i)}, i = 1, \dots, k\}$ of $\{\beta_t\}$.

Decoupled loss: $L_\lambda(\tilde{\beta}) = \|\mathbf{W}^{1/2}(\mathbf{X} \circ \tilde{\beta} - \mathbf{X} \circ \tilde{\beta})\|_2^2 + q_\lambda(\tilde{\beta})$.

- $\mathbf{W} = \text{diag}(w_1, \dots, w_n)$ is diagonal with weights for each measurement being $w_i = 1/\sigma_{\epsilon,i}^2$, for $i = 1, \dots, n$.
- Penalty function $q_\lambda(\cdot)$ induces sparsity into $\tilde{\beta}$ with form

$$q_\lambda(\tilde{\beta}) = \lambda \sum_t \frac{1}{|\psi_t|} |\Delta^D \tilde{\beta}_t|,$$

where $\psi_t = \frac{1}{k} \sum_{i=1}^k \Delta^D \beta_t^{(i)}$ and $D = 1, 2$ controls the type of change.

Changepoint Selection

Given λ , denote η_λ as the time indices which $\{\Delta^D \tilde{\beta}_t \neq 0\}$.

$$\text{Diagnostic tool: } R_\lambda^2 = \frac{1}{k} \sum_{i=1}^k \frac{\|\beta^{(i)} - \tilde{\beta}\|_2^2}{\|\beta^{(i)} - \tilde{\beta}^{(i)}\|_2^2}$$

where $\tilde{\beta}^{(i)} = \frac{1}{n} \sum_{t=1}^n \beta_t^{(i)}$, and the optimal λ determined by least changepoints given $E[R_\lambda^2]$ exceeds a certain threshold.

* Multiple Predictors & Covariates

Set predictors $\mathbf{X} = \text{blockdiag}(\mathbf{x}_1', \dots, \mathbf{x}_n')$ with $\mathbf{x}_i = (x_{i,1}, \dots, x_{i,p})'$, and

covariates $\mathbf{Z} = (\mathbf{z}_1, \dots, \mathbf{z}_n)$ with $\mathbf{z}_i = (z_{i,1}, \dots, z_{i,l})'$.

$$y_t = \mathbf{x}_t' \beta_t + \alpha' \mathbf{z}_t + \epsilon_t, \quad \epsilon_t \sim N(0, \sigma_{\epsilon,t}^2), \quad \Delta^D \beta_t = \omega_t, \quad \omega_t \sim N(0, \Sigma_{\omega,t}).$$

Extend the model with the decoupled loss:

$$L_\lambda(\tilde{\beta}, \tilde{\alpha}) = \|\mathbf{W}^{1/2}(\mathbf{X} \tilde{\beta} + \mathbf{Z} \tilde{\alpha} - \mathbf{X} \tilde{\beta} - \mathbf{Z} \tilde{\alpha})\|_2^2 + q_\lambda(\tilde{\beta}),$$

$$q_\lambda(\tilde{\beta}) = \lambda \sum_{t=1}^n \sum_{g=1}^G \frac{1}{|\psi_{g,t}|} |\Delta^D \beta_{g,t}|.$$

Decoupling Simulations

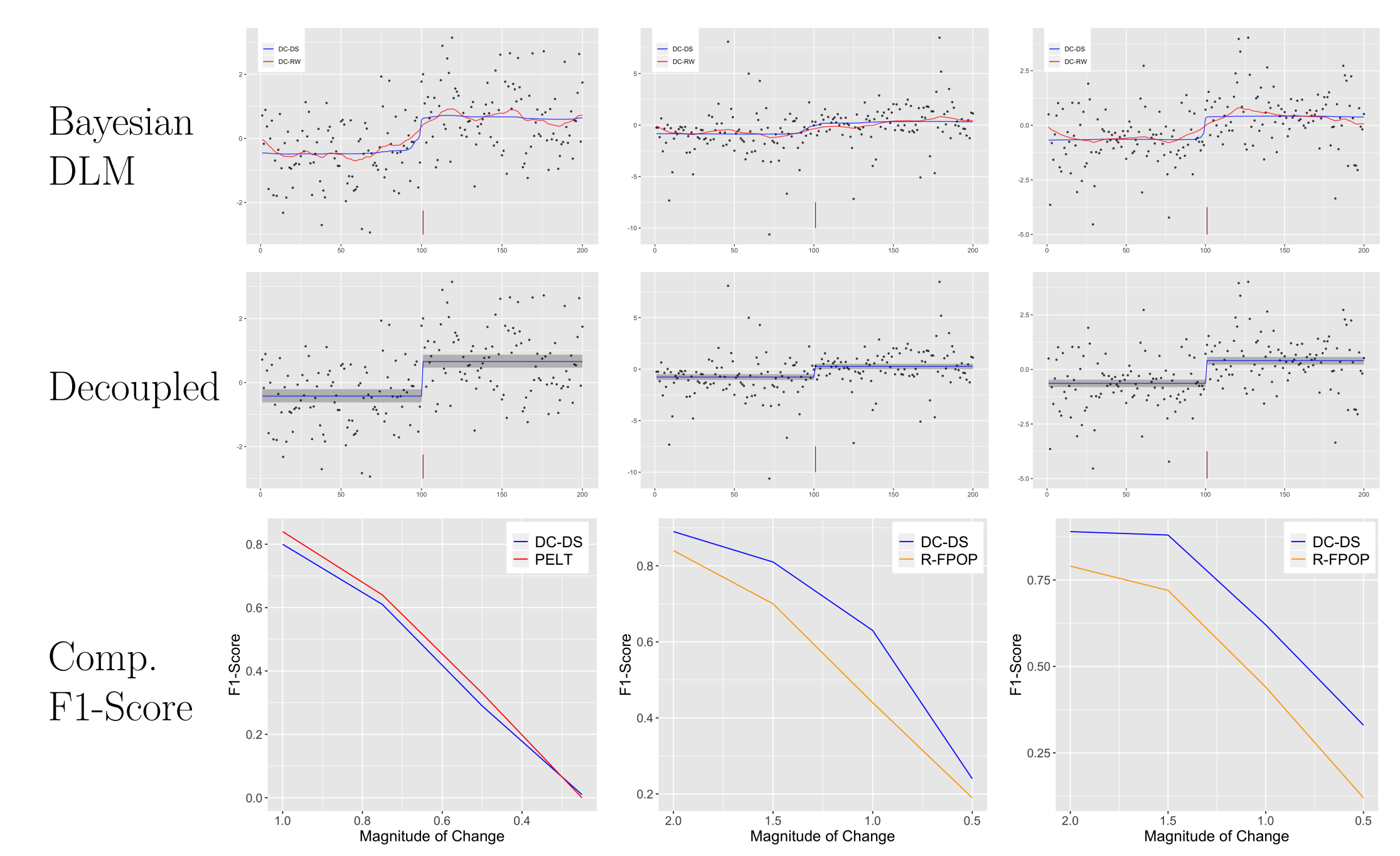


Fig 12 Gaussian Noise (left); Outliers (middle); Stochastic Volatility (right). DC-DS: decoupled results with shrinkage. DC-RW: decoupled with random walk.

Decoupling Applications

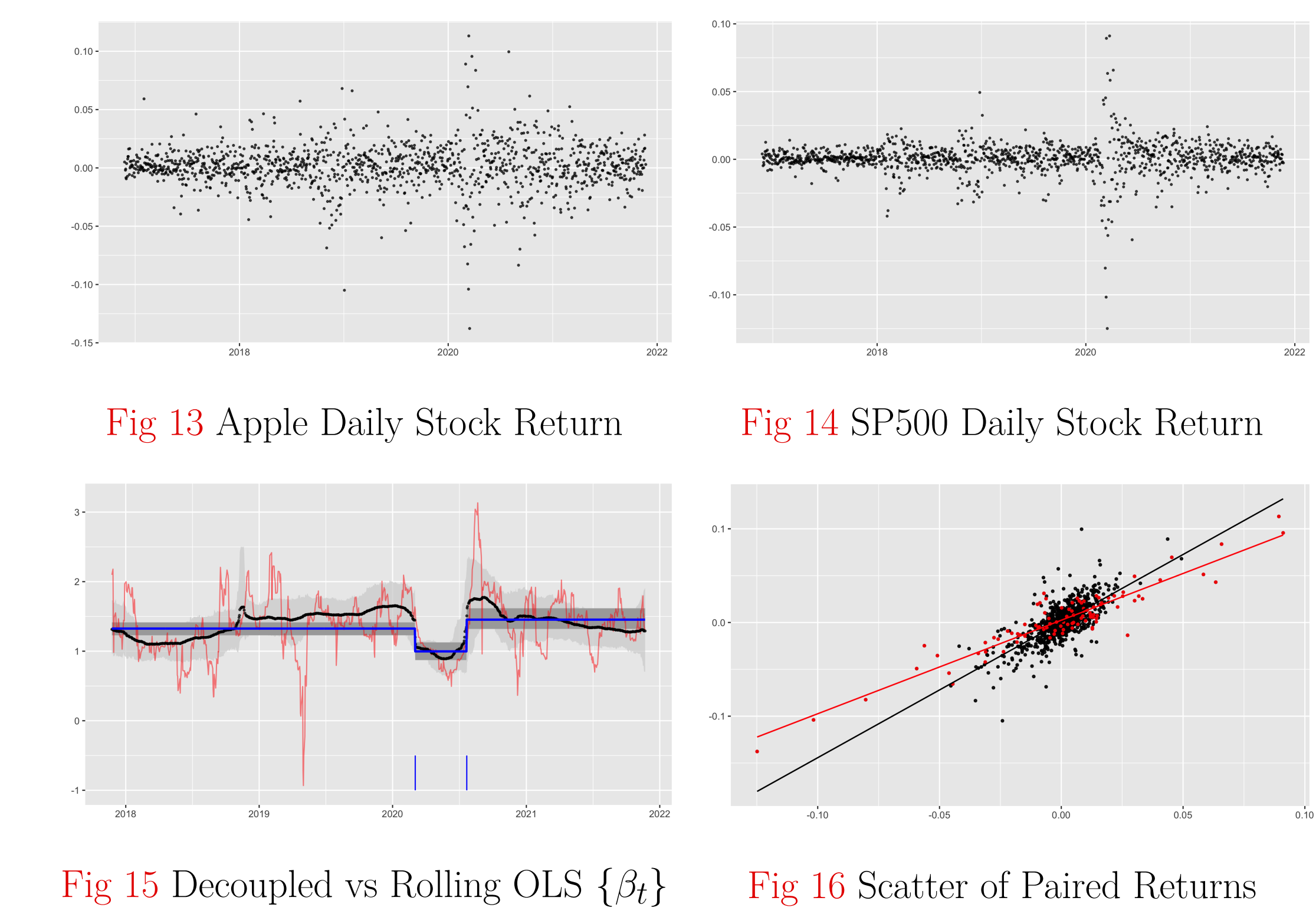


Fig 15 Decoupled vs Rolling OLS $\{\beta_t\}$

Fig 16 Scatter of Paired Returns

Conclusions

- A framework for inferring changepoints from posteriors produced by Bayesian time-varying parameter models.
- By decoupling trend modeling and changepoint analysis, we allow fitting an arbitrarily complex model to deal with intricacies inherent in data.
- Extensions: higher order trend changes, regression coefficients, multivariate.

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