Next generation cosmological analysis with nested sampling

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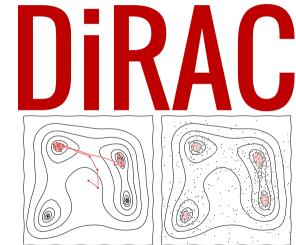






Overview

- DiRAC 2020 RAC allocation of 30MCPUh
- Main goal: Planck Legacy Archive equivalent
- ▶ Parameter estimation → Model comparison
- ► MCMC → Nested sampling
- $\blacktriangleright \mathsf{Planck} \rightarrow \{\mathsf{Planck}, \mathsf{DESY1}, \mathsf{BAO}, \ldots\}$
- Pairwise combinations
- Suite of tools for processing these
 - anesthetic 2.0
 - unimpeded 1.0
 - zenodo archive
- MCMC chains also available.
- Work in progress, but beta testers requested (email wh260@cam.ac.uk)



The three pillars of Bayesian inference

 $\begin{array}{l} \text{Posterior} = \frac{\text{Likelihood} \times \text{Prior}}{\text{Evidence}} \quad \text{Posterior} = \frac{\text{Evidence} \times \text{Prior}}{\text{Normalisation}}. \end{array}$

Parameter estimation

What do the data tell us about the parameters of a model?

e.g. the size or age of a ΛCDM universe $P(\theta|D,M) = \frac{P(D|\theta,M)P(\theta|M)}{P(D|M)}, \quad P(M|D) = \frac{P(D|M)P(M)}{P(D)},$

Model comparison

 $\frac{\mathcal{Z}_{\mathcal{M}}\Pi_{\mathcal{M}}}{\sum_{m}Z_{m}\Pi_{m}},$

How much does the data support a particular model? *e.g.* ΛCDM vs a dynamic dark energy cosmology

Tension quantification

Do different datasets make consistent predictions from the same model? *e.g. CMB vs Type IA supernovae data*

$$\mathcal{R} = rac{\mathcal{Z}_{AB}}{\mathcal{Z}_{A}\mathcal{Z}_{B}},$$

$$\begin{split} \log \mathcal{S} &= \langle \log \mathcal{L}_{AB} \rangle_{\mathcal{P}_{AB}} \\ &- \langle \log \mathcal{L}_{A} \rangle_{\mathcal{P}_{A}} \\ &- \langle \log \mathcal{L}_{B} \rangle_{\mathcal{P}_{B}} \end{split}$$

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 $\mathcal{P} = \frac{\mathcal{L} \times \pi}{\mathcal{Z}},$

Occam's Razor [2102.11511]

Bayesian inference quantifies Occam's Razor:

- "Entities are not to be multiplied without necessity"
 - "Everything should be kept as simple as possible, but not simpler" "Albert Einstein"
- Properties of the evidence: rearrange Bayes' theorem for parameter estimation

$$\mathcal{P}(\theta) = rac{\mathcal{L}(\theta)\pi(\theta)}{\mathcal{Z}} \implies \log \mathcal{Z} = \log \mathcal{L}(\theta) - \log rac{\mathcal{P}(\theta)}{\pi(\theta)}$$

Evidence is composed of a "goodness of fit" term and "Occam Penalty"

- RHS true for all θ. Take max likelihood value θ_{*}:
 - $\log \mathcal{Z} = -\chi^2_{
 m min} {
 m Mackay}$ penalty
- Be more Bayesian and take posterior average to get the "Occam's razor equation"

- William of Occam

$$\boxed{\log \mathcal{Z} = \langle \log \mathcal{L} \rangle_{\mathcal{P}} - \mathcal{D}_{\mathrm{KL}}}$$

Natural regularisation which penalises models with too many parameters.

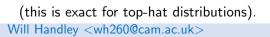
Kullback Liebler divergence

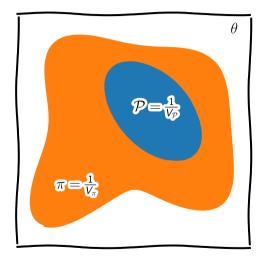
The KL divergence between prior π and posterior P is is defined as:

$$\mathcal{D}_{ ext{KL}} = \left\langle \log rac{\mathcal{P}}{\pi}
ight
angle_{\mathcal{P}} = \int \mathcal{P}(heta) \log rac{\mathcal{P}(heta)}{\pi(heta)} d heta.$$

- Whilst not a distance, $\mathcal{D} = 0$ when $\mathcal{P} = \pi$.
- Occurs in the context of machine learning as an objective function for training functions.
- In Bayesian inference it can be understood as a log-ratio of "volumes":

$$\mathcal{D}_{ ext{KL}}pprox \log rac{m{V_{\pi}}}{m{V_{\mathcal{P}}}}$$



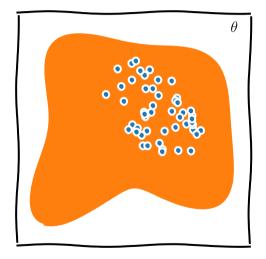


Why do sampling?

- The cornerstone of numerical Bayesian inference is working with samples.
- Generate a set of representative parameters drawn in proportion to the posterior θ~ P.
- ► The magic of marginalisation ⇒ perform usual analysis on each sample in turn.
- The golden rule is stay in samples until the last moment before computing summary statistics/triangle plots because

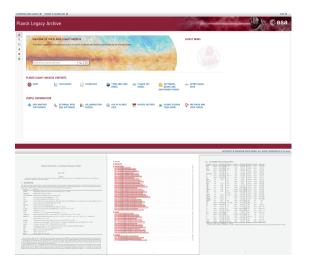
$f(\langle X\rangle) \neq \langle f(X)\rangle$

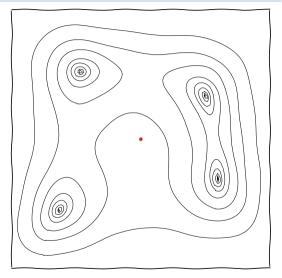
- Generally need ~ O(12) independent samples to compute a value and error bar.
- Will Handley <wh260@cam.ac.uk>

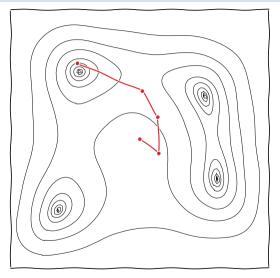


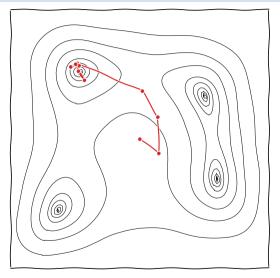
The Planck legacy archive

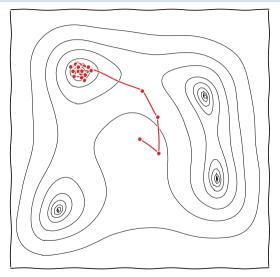
- Planck collaboration science products
- distributed cosmology inference results as MCMC chains
- Across a grid of:
 - subsets/combinations of *Planck* data
 - TT, lowl, lowE, lensing
 - ACDM extensions
 - 🕨 base, mnu, nrun, omegak, r
- importance sampling across some other likelihoods (BAO, JLA,...)
- Cannot compute evidences in high dimensions from MCMC chains
 - Only parameter estimation
 - no model comparison

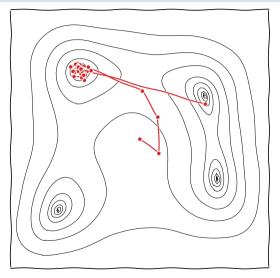


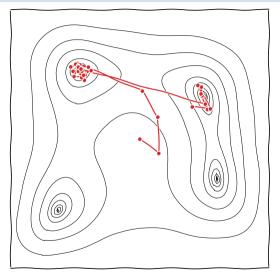


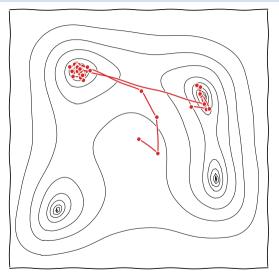






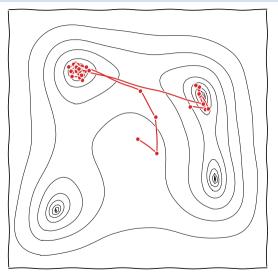


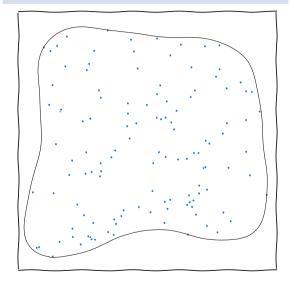


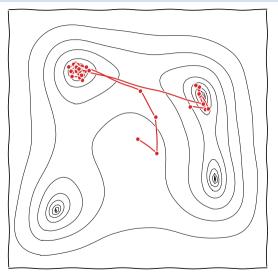


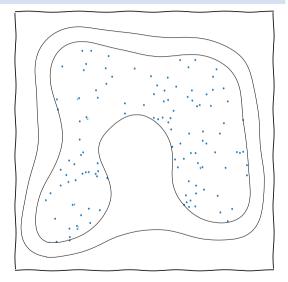
Nested sampling

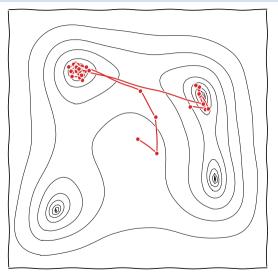
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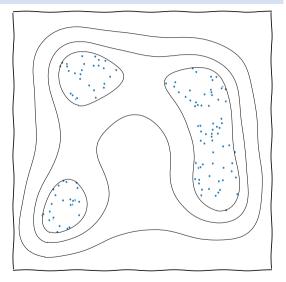


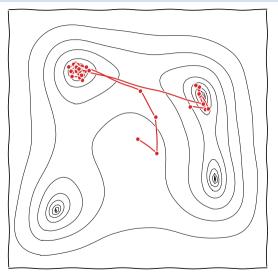


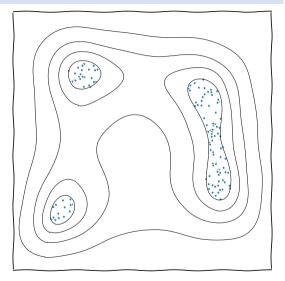


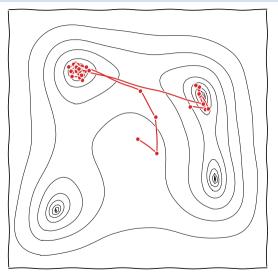


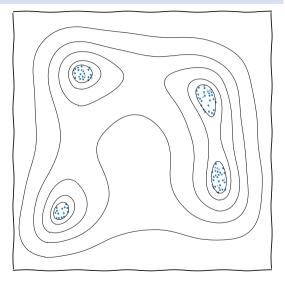


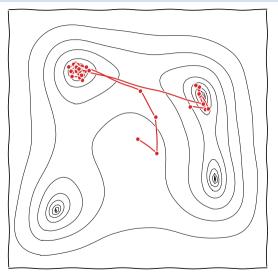


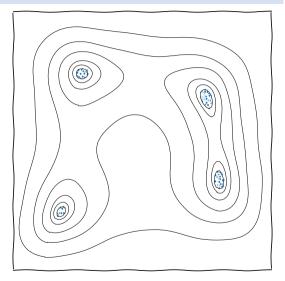


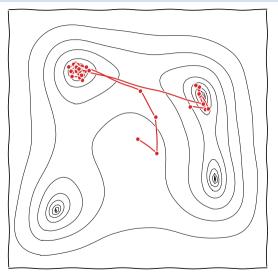


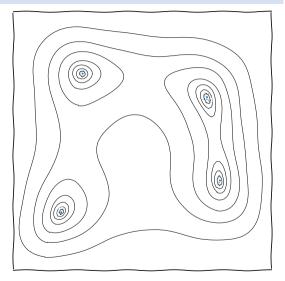


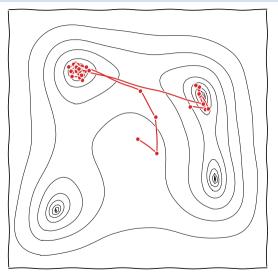




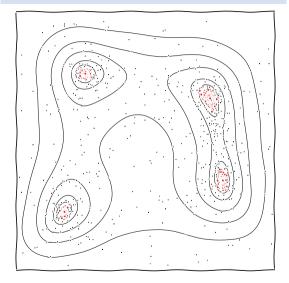




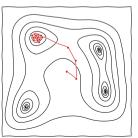




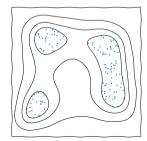
Nested sampling



- Single "walker"
- Explores posterior
- Fast, if proposal matrix is tuned
- Parameter estimation, suspiciousness calculation
- Channel capacity optimised for generating posterior samples



- Ensemble of "live points"
- Scans from prior to peak of likelihood
- Slower, no tuning required
- Parameter estimation, model comparison, tension quantification
- Channel capacity optimised for computing partition function



The grid (so far)

- Models: [Λ CDM, Ω_K , ν , r, w, w(a)]
- Data: [plik, camspec, DESY1, bicep+keck, BAO(DR16), pantheon]
- Pairwise combinations of datasets
- Breakdown of Planck & BAO data
- Samplers: [Metropolis Hastings MCMC, Nested Sampling]
- These exhaust what is currently available by default in cobaya
- Wide priors to allow for importance readjustment as desired
- roughly halfway through computational allocation.
- Feedback desirable as to what extensions to the grid would be of community interest (email wh260@cam.ac.uk).
- Further checking needed before first release by end of this year.

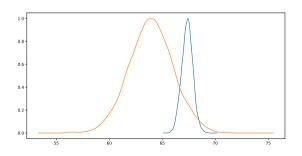
unimpeded

Universal Model comparison and Parameter Estimation Distributed over Every Dataset

- Python tool for seamlessly downloading and cacheing chains
- Data stored on zenodo
- hdf5 storage for fast & reliable download & storage
- Library of trained bijectors to be used as priors/emulators [2102.12478]/nuisance marginalised likelihoods [2207.11457]
- anesthetic compatible for processing of chains [1905.04768]
- α-testers wanted! (email wh260@cam.ac.uk)
- End goal community library which everyone contributes to so expensive runs reusable.

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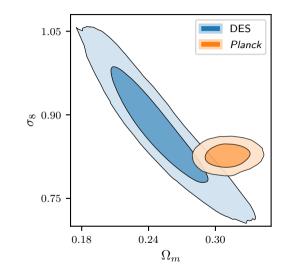
from unimpeded import Unimpeded
store = Unimpeded(cache='data.hdf5')
samps = store('planck')
samps.H0.plot.kde_ld()
samps = store('planck', model='klcdm')
samps.H0.plot.kde_ld()



10

The importance of global measures of tension

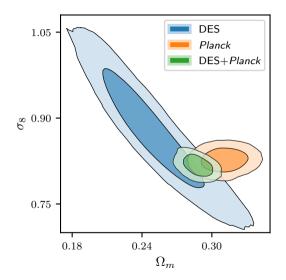
- Hubble tension [1907.10625]
 - ▶ *Planck*: $H_0 = 67.4 \pm 0.5$
 - ▶ $SH_0ES: H_0 = 74.0 \pm 1.4$
- In other situations the discrepancy doesn't exist in a single interpretable parameter
- ► For example: DES+*Planck* [1902.04029]
- Are these two datasets in tension?
- There are a lot more parameters are we sure that tensions aren't hiding? Are we sure we've chosen the best ones to reveal the tension?
- Should use "Suspiciousness" statistic S, or Bayes ratio R to determine global tension.



The importance of global measures of tension

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The DES evidence ratio R

The Dark Energy Survey [1708.01530] quantifies tension between two datasets A and B using the Bayes ratio:

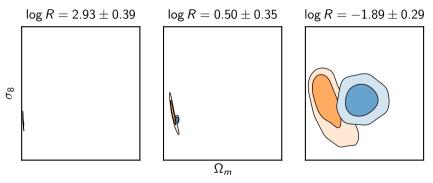
$$R = \frac{\mathcal{Z}_{AB}}{\mathcal{Z}_{A}\mathcal{Z}_{B}} = \frac{P(A \cap B)}{P(A)P(B)} = \frac{P(A|B)}{P(A)} = \frac{P(B|A)}{P(B)}$$

where \mathcal{Z} is the Bayesian evidence.

- Many attractive properties:
 - Symmetry
 - Parameterisation independence
 - Dimensional consistency
 - Use of well-defined Bayesian quantities
- R gives the relative change in our confidence in data A in light of having seen B (and vice-versa).

- R > 1 implies we have more confidence in A having received B.
- ► Like evidences, it is prior-dependent from D in log Z = ⟨log L⟩_P - D
- Increasing prior widths ⇒ decreasing evidence.
- Increasing prior widths ⇒ increasing confidence.

The DES evidence ratio *R*: Prior dependency



- What does it mean if increasing prior widths \Rightarrow increasing confidence?
- Wide priors mean *a-priori* the parameters could land anywhere.
- We should be proportionally more reassured when they land close to one another if the priors are wide

How do we deal with the prior dependency in R?

Option 1 Take the Bayesian route, accept the prior dependency, and spend time trying to justify why a given set of priors are "physical".

Option 2 Try to find a principled way of removing this prior dependency

• Decompose ratio using Occam's Razor equation $\log \mathcal{Z} = \langle \log \mathcal{L} \rangle_{\mathcal{P}} - \mathcal{D}$

$$egin{aligned} & \operatorname{og} \mathcal{R} = \log \mathcal{Z}_{AB} - \log \mathcal{Z}_{A} - \log \mathcal{Z}_{A} \ & = \langle \log \mathcal{L}_{AB}
angle_{\mathcal{P}_{AB}} - \langle \log \mathcal{L}_{A}
angle_{\mathcal{P}_{A}} - \langle \log \mathcal{L}_{B}
angle_{\mathcal{P}_{B}} - \mathcal{D}_{AB} + \mathcal{D}_{A} + \mathcal{D}_{B} \ & = \log \mathcal{S} + \log \mathcal{I} \end{aligned}$$

where we have defined the suspiciousness S, which is prior independent, and the information \mathcal{I} , which depends on the parameter compression of the shared space

- Focussing on the prior-independent portion S gives R for the "Narrowest reasonable priors" which do not impinge on the posterior
- One of the critical observations is that one can only hide tension by widening priors. Narrowing them will only ever show tension if it is present.

Suspiciousness S

► For a Gaussian set of posteriors:

$$\log \mathcal{S} = rac{d}{2} - rac{1}{2}(\mu_A - \mu_B)(\Sigma_A + \Sigma_B)^{-1}(\mu_A - \mu_B).$$

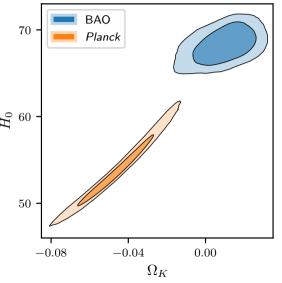
- ► The Malhanobis term is suggestive, so we can use this to calibrate a "sigma" level of tension using a χ^2 distribution for $\chi^2_d = d 2 \log S$, or a tension probability.
- ► S is composed of evidences Z and KL divergences D, which are Gaussian-independent concepts, so the only thing to determine is d, the "number of shared parameters".
- Can do this with Gaussian dimensionality $\frac{d}{2} = \operatorname{var}_{\mathcal{P}}(\log \mathcal{L})$ [1903.06682]

| Planck vs BAO : | $p=42\pm4\%$ |
|----------------------------|-------------------|
| Planck vs DESY1 : | $p=3.2\pm1.0\%$ |
| Planck vs S <i>H</i> 0ES : | $p=0.25\pm0.17\%$ |

• Under this metric, SH₀ES is unambiguously inconsistent, although not quite as brutal as $> 4\sigma$. BAO is consistent, and DESY1 is inconsistent, but only just. This is pleasingly similar to ones intuition.

Curvature tension? [1908.09139]

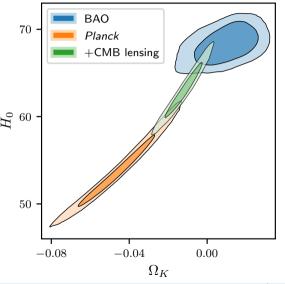
- ▶ If you allow $\Omega_K \neq 0$, *Planck* (plikTTTEEE) has a moderate preference for closed universes (50:1 betting odds on), $\Omega_K = -4.5 \pm 1.5\%$ [1911.02087]
- *Planck*+lens+BAO strongly prefer $\Omega_{\mathcal{K}} = 0$.
- But, *Planck* vs lensing is 2.5σ in tension, and Planck vs BAO is 3σ.
- Reduced if plik \rightarrow camspec [2002.06892]
- BAO and lensing summary assume ΛCDM.
- Doing this properly with BAO retains preference for closed universe (though closer to flat Ω_K = -0.4 ± 0.2%) [2205.05892]
- Present-day curvature has profound consequences for inflation [2205.07374]
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16 / 17

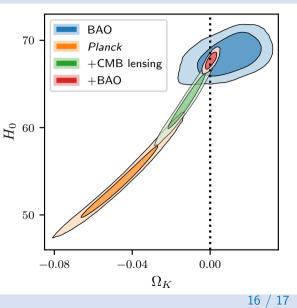
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Conclusions

DiRAC RAC allocation for building a legacy grid of

- MCMC & Nested sampling chains
- gridded over (pairwise) up-to-date datasets
- gridded over extensions to ΛCDM
- Bijectors & emulators for fast re-use
- Importance sampling toolkit via anesthetic for (re)processing
- Long-term goal: community repository of chains to share model comparison compute resource

Looking for:

- α -testers for unimpeded
- Suggestions for more datasets (and their incorporation into cobaya)