# **CTEQ-TEA update, studies and tools to understand PDF uncertainties**

## **Pavel Nadolsky**

Southern Methodist University, USA

With CTEQ-TEA (Tung Et. Al.) working group

China: A. Ablat, S. Dulat, J. Gao, T.-J. Hou, I. Sitiwaldi, M. Yan, and collaborators Mexico: A. Courtoy USA: T.J. Hobbs, M. Guzzi, X. Jing,

J. Huston, H.-W. Lin, D. Stump, C. Schmidt, K. Xie, C.-P. Yuan





## CT18 parton distributions

#### PRD 103 (2021) 014013

Four PDF ensembles: CT18 (default), A, X, and Z



New CT18 NNLO grids for precision calculations

- Soon to appear in the LHAPDF library
- Contain more x and Q points improved interpolation at the expense of slightly slower evaluation
- Crossing of quark mass thresholds implemented with multiple *Q* grids
- Complement the published (less dense) CT18 grids that remain sufficient for most applications

# Toward a new generation of CT202X PDFs

See detailed presentations at DIS'2023 workshop

- 1. Identify sensitive, mutually consistent new experimental data sets using preliminary fits and fast techniques ( $L_2$  sensitivities and ePump)
- 2. Implement N3LO QCD and NLO EW contributions as they become available. N3LO accuracy is reached only when N3LO terms are **fully** implemented.
  - Meanwhile, "NNLO+" PDFs: e.g., include theoretical uncertainty due to QCD scale dependence for key processes as has been done in CT18/CT18X NNLO PDFs
- 3. Explore quark sea flavor dependence:  $s \bar{s}$  (CT18As), fitted charm (CT18FC),...
- 4. Include lattice QCD constraints (CT18As\_Lat)
- 5. Next-generation PDF uncertainty quantification: META PDFs, Bézier curves, MC sampling, multi-Gaussian combination, ...

# From talk by M. Boonekamp and CERN-LPCC-2022-06





The fitted experiments are not perfectly consistent and may have a non-negligible scale dependence.

CT18+CT18Z PDF uncertainties **together** account for the crucial tension in the fitted data (ATLAS 7 TeV W/Z vs. (SI)DIS) and for QCD scale variations in DIS,  $Z p_T$ , jet production.

PDF set	Chi2/ndf	PDF set	Chi2/ndf		may
Cteq66	231/126	CT18NNLO	163/126	CT18Z	
CT10	179/126	CT18ANNLO	170/126		CT1
NNPDF31	200/126	MSHT20	270/126		
NNPDF40	195/126	ABMP16	236/126		$Z p_T$

## New post-CT18 LHC Drell-Yan data

Boson	$\sqrt{s}$	Lumi	Observable	Ref.	
ATLAS					
W, Z	2.76	$4.0 \text{ pb}^{-1}$	$\sigma^{ m fid,tot}$	1907.03567	
W, Z	13	81.0 pb <sup>-1</sup>	$\sigma^{ m fid}$	1603.09222	
W, Z	5.02	25.0 pb <sup>-1</sup>	$(oldsymbol{\eta}_\ell, y_{\ell\ell})$	1810.08424	
Z	8	$20.2 \ {\rm fb}^{-1}$	$(m_{\ell\ell},y_{\ell\ell})$	1710.05167	
$W \rightarrow \mu \nu$	8	20.2 fb <sup>-1</sup>	$\eta_{\mu}$	1904.05631	
Z	13	<b>36</b> .1 fb <sup>-1</sup>	$p_T^{\ell\ell}$	1912.02844	
CMS					
Z	13	$2.8 { m ~fb^{-1}}$	$m_{\ell\ell}$	1812.10529	
Z	13	35.9 fb <sup>-1</sup>	$(y, p_T, \phi^*)$	1909.04133	
W	13	35.9 fb <sup>-1</sup>	$oldsymbol{\sigma}^{ ext{fid}}$ , $y_W, (oldsymbol{\eta}_\ell, p_T^\ell)$	2008.04174	
LHCb					
$W \rightarrow e \mathbf{v}$	8	2.0 fb <sup>-1</sup>	$\eta_e$	1608.01484	
Z	13	<b>294</b> pb <sup>-1</sup>	$oldsymbol{\sigma}^{ ext{fid}}$ , $(y, p_T, oldsymbol{\phi}^*)$	1607.06495	
$Z \rightarrow \mu \mu$	13	$5.1 { m ~fb^{-1}}$	$oldsymbol{\sigma}^{ ext{fid}}$ , $(y, p_T, oldsymbol{\phi}^*)$	2112.07458	

We mainly focus on (pseudo)rapidity distributions in this work.

#### K. Xie et al., in progress

Multiple candidate fits to explore the impact of 8 and 13 TeV Drell-Yan data using NNLO and resummed N3LL-NNLO cross sections



## Post-CT18 LHC Drell-Yan data [See K. Xie's talk for the details.]





- Most of the post-CT18 LHC Drell-Yan data are consistent with the ATLAS 7 TeV W, Z precision measurement, which enhance the strangeness (CT18A).
- Exceptions for ATLAS and LHCb 8 TeV W data, which push the  $d(\bar{d})$  PDFs to the opposite direction.
- The post-CT18 LHC Drell-Yan data shrink the error bands.
- The joint impact of these new data sets pull the PDFs and predictions from CT18 to CT18Z direction.

# CT18As\_Lat NNLO: Strangeness asymmetry with a lattice QCD constraint



## Sensitivity of experiments to the strangeness asymmetry



Preference for  $s - \bar{s} \neq 0$  at x > 0.1 emerges from competing  $\chi^2$  pulls of NuTeV dimuon, LHCb W/Z, BCDMS and E866 fixed-target cross sections. We estimated it using the  $L_2$  sensitivity fast technique [*T. Hobbs et al., arXiv:1904.00022*]. The lattice prediction by *R. Zhang et al., 2005.01124* is consistent with  $s - \bar{s} = 0$  at x > 0.3.

## New CT18 Fitted Charm analysis

moments of the FC PDFs often used to characterize magnitude, asymmetry

$$\langle x^n \rangle_{c^{\pm}} = \int_0^1 dx \, x^n (c \pm \bar{c})[x, Q]$$

$\langle x \rangle_{\rm FC} \equiv \langle x \rangle_{\rm c^+} [$	$Q_0 = 1.27 \mathrm{GeV}$ ]at NNLO.		
$= 0.0048 {}^{+0.006}_{-0.004}$	$^{3}_{3} \left( \frac{+0.0090}{-0.0048} \right), \text{CT18 (BHPS3)}$		
$= 0.0041  {}^{+0.004}_{-0.004}$	$^{9}_{1}$ $\binom{+0.0091}{-0.0041}$ , CT18X (BHPS3)		
$= 0.0057 {}^{+0.004}_{-0.004}$	$_{5}^{8} \left(\frac{+0.0084}{-0.0057}\right), \text{CT18} (\text{MBMC})$		
$= 0.0061  {}^{+0.003}_{-0.003}$	$\binom{0}{8} \binom{+0.0064}{-0.0061}$ , CT18 (MBME)		
$\Delta\chi^2 \le 10$	$\Delta\chi^2 \le 30$		
(restrictive tolerance) (~CT standard tolerance)			



M. Guzzi, T. Hobbs, K. Xie et al., arXiv:2211.01387

#### possible charm-anticharm asymmetries

pQCD only very weakly breaks  $c = \bar{c}$  through HO corrections

- → large(r) charm asymmetry would signal nonpert dynamics, IC
- $\rightarrow$  MBM breaks  $c = \overline{c}$  through hadronic interactions



## A Lagrange Multiplier scan



A slow method inside the global fit to compute the  $\chi^2$  dependence on the quantity of interest (here the momentum fraction carried by the fitted charm in CT18 FC NNLO).

## An L<sub>2</sub> sensitivity



A fast approximation to the LM scan to estimate  $\Delta \chi^2$  of a fitted experiment (here ATLAS 7 TeV W/Z) when the PDF increases by  $1\sigma$  for a given tolerance  $T^2$ . Needs only published error PDFs and  $\chi^2$  tables. Can be combined with TRExFitter for PDF errors on  $M_W$ .

## Epistemic PDF uncertainty in PDF fits

"Hopscotch scans" to quantify epistemic uncertainty on MC replicas

Can be applied to understand the PDF uncertainty on  $M_W$  using open-source programs

Based on numerical results from

A. Courtoy, J. Huston, P. N., K. Xie, M. Yan, C.-P. Yuan, **Phys. Rev. D 107, (2023) 034008** 

[full comparisons in arXiv:2205.10444

and at <a href="https://ct.hepforge.org/PDFs/2022hopscotch/">https://ct.hepforge.org/PDFs/2022hopscotch/</a>





#### **Representative sampling**



## Epistemic PDF uncertainty...

...reflects **methodological choices** such as PDF functional forms or NN architecture and hyperparameters.

... can dominate the full uncertainty when experimental and theoretical uncertainties are small.

... is associated with the prior probability.

... can be estimated by **representative sampling** of the PDF solutions obtained with acceptable methodologies.

 $\Rightarrow$  sampling over choices of experiments, PDF/NN functional space, models of correlated uncertainties...

 $\Rightarrow$  in addition to sampling over data fluctuations



#### **Components of PDF uncertainty**



In each category, one must maximize

**PDF fitting accuracy** (accuracy of experimental, theoretical and other inputs)

PDF sampling accuracy

(adequacy of sampling in space of possible solutions)

#### Fitting/sampling classification is borrowed

from the statistics of large-scale surveys [Xiao-Li Meng, *The Annals of Applied Statistics*, Vol. 12 (2018), p. 685]

## HEP is not alone

#### Various domains contend with **multi-dimensional non-probability samples**

Forecasting: presidential elections, financial markets, weather and climate, ... Meng, The Annals of Applied Statistics, 12(2), 685; Isakov and Kuriwaki, Harvard Data Science Review, 2(4), 2020

Political polling
M. R. Elliott, R. Valliant, Statistical Science, 32(2), 249 (2017)
M. A. Bailey, Polling at a Crossroads – Rethinking Modern Survey Research. Cambridge University Press, 2023

COVID-19 vaccination assessments and epidemiological studies *Bradley et al.*, <u>https://doi.org/10.1038/s41586-021-04198-4</u> *W. Dempsey, arXiv:2005.10425* 

Clinical trials of medical treatments P. Msaouel, <u>https://doi.org/10.1080/07357907.2022.2084621</u>

Studies of biodiversity R. Boyd et al., <u>https://doi.org/10.1016/j.tree.2023.01.001</u>

• • •

#### AI/ML techniques are superb for finding an excellent fit to data. Are these techniques adequate for uncertainty estimation [exploring all good fits]?

A common resampling procedure used by experimentalists and theorists:

- 1. Train a neural network model  $T_i$  with N<sub>par</sub> (hyper)parameters on a randomly fluctuated replica of discrete data  $D_i$ . Repeat N<sub>rep</sub> times. In a typical application: N<sub>par</sub> > 10<sup>2</sup>, N<sub>rep</sub> < 10<sup>4</sup>.
- 2. Out of N<sub>rep</sub> replicas  $T_i$  with "good" description of data [i.e., with a high likelihood  $P(D_i|T_i) \propto e^{-\chi^2(D_i,T_i)/2}$ ], discard "badly behaving" (overfitted, not smooth, ...) replicas
- 3. Estimate the uncertainties of  $T_i$  using the remaining "well-behaved" replicas

## Is this procedure rigorous? How many $N_{rep}$ replicas does one need?

## A likelihood-ratio test of NN models $T_1$ and $T_2$

From Bayes theorem, it follows that

 $\frac{P(T_2|D)}{P(T_1|D)} = \frac{P(D|T_2)}{P(D|T_1)} \times \frac{P(T_2)}{P(T_1)}$  $\equiv r_{\text{posterior}} \equiv r_{\text{likelihood}} \equiv r_{\text{prior}}$  $= a \text{leatory} \quad \text{epistemic + aleatory} \quad \text{probabilities}$ 

Suppose replicas  $T_1$  and  $T_2$  have the same  $\chi^2 [r_{\text{likelihood}} = \exp\left(\frac{\chi_1^2 - \chi_2^2}{2}\right) = 1]$ , but  $T_2$  is disfavored compared to  $T_1 [r_{\text{posterior}} \ll 1]$ .

This only happens if  $r_{\text{prior}} \ll 1 : T_2$  is discarded based on its **prior** probability.

# Epistemic PDF uncertainty is important in W boson mass and $\alpha_s$ measurements

#### ATLAS-CONF-2023-004

PDF-Set	$p_{\mathrm{T}}^{\ell}$ [MeV ]	$m_{\rm T}$ [MeV ]	combined [MeV ]
CT10	$80355.6^{+15.8}_{-15.7}$	$80378.1^{+24.4}_{-24.8}$	80355.8 <sup>+15.7</sup> -15.7
CT14	$80358.0^{+16.3}_{-16.3}$	80388.8 <sup>+25.2</sup> -25.5	$80358.4^{+16.3}_{-16.3}$
CT18	$80360.1^{+16.3}_{-16.3}$	80382.2 <sup>+25.3</sup> -25.3	80360.4+16.3
MMHT2014	80360.3 <sup>+15.9</sup> -15.9	$80386.2^{+23.9}_{-24.4}$	$80361.0^{+15.9}_{-15.9}$
MSHT20	80358.9 <sup>+13.0</sup> -16.3	$80379.4^{+24.6}_{-25.1}$	80356.3 <sup>+14.6</sup>
NNPDF3.1	$80344.7^{+15.6}_{-15.5}$	80354.3 <sup>+23.6</sup> -23.7	80345.0 <sup>+15.5</sup> _15.5
NNPDF4.0	$80342.2^{+15.3}_{-15.3}$	80354.3 <sup>+22.3</sup> -22.4	$80342.9^{+15.3}_{-15.3}$

Table 2: Overview of fitted values of the *W* boson mass for different PDF sets. The reported uncertainties are the total uncertainties.

#### ATLAS-CONF-2023-015

The statistical analysis for the determination of  $\alpha_s(m_Z)$  is performed with the xFitter framework [60]. The value of  $\alpha_s(m_Z)$  is determined by minimising a  $\chi^2$  function which includes both the experimental uncertainties and the theoretical uncertainties arising from PDF variations:

$$\chi^{2}(\beta_{\exp},\beta_{th}) = \frac{\sum_{i=1}^{N_{data}} \left(\sigma_{i}^{\exp} + \sum_{j} \Gamma_{ij}^{\exp} \beta_{j,\exp} - \sigma_{i}^{th} - \sum_{k} \Gamma_{ik}^{th} \beta_{k,th}\right)^{2}}{\Delta_{i}^{2}} + \sum_{j} \beta_{j,\exp}^{2} + \sum_{k} \beta_{k,th}^{2}.$$

profiling of CT and MSHT PDFs requires to include a tolerance factor  $T^2 > 10$  as in the ePump code

[T.J. Hou et al., <u>1912.10053</u>, Appendix F]

Also the next slide.

(1)

## Augmented likelihood for PDFs with global tolerance

1. Start by defining the correspondence between  $\Delta \chi^2$  and cumulative probability level: 68% c.l.  $\Leftrightarrow \Delta \chi^2 = T^2$ . 2. Write the **augmented** likelihood density for this definition:

 $P(D_i|T_i) \propto e^{-\chi^2/(2T^2)}$ 

3. When profiling 1 new experiment with the prior imposed on PDF nuisance parameters  $\lambda_{\alpha,th}$ :

$$\chi^{2}(\vec{\lambda}_{exp},\vec{\lambda}_{th}) = \sum_{i=1}^{N_{pt}} \frac{\left[D_{i} + \sum_{\alpha} \beta_{i,\alpha}^{exp} \lambda_{\alpha,exp} - T_{i} - \sum_{\alpha} \beta_{i,\alpha}^{th} \lambda_{\alpha,th}\right]^{2}}{s_{i}^{2}} + \sum_{\alpha} \lambda_{\alpha,exp}^{2} + \sum_{\alpha} T^{2} \lambda_{\alpha,th}^{2}. \qquad \beta_{i,\alpha}^{th} = \frac{T_{i}(f_{\alpha}^{+}) - T_{i}(f_{\alpha}^{-})}{2},$$

$$new \text{ experiment} \qquad priors \text{ on expt. systematics} and PDF \text{ params}$$
4. Alternatively, we can reparametrize  $\chi^{2'} \equiv \chi^{2}/T^{2}$ , so that 68% c.l.  $\Leftrightarrow \Delta \chi^{2'} = 1$ . We have
$$P(D_{i}|T_{i}) \propto e^{-\chi^{2'/2}}$$

$$\chi^{2'}(\vec{\lambda}_{exp}, \vec{\lambda}_{th}) = \sum_{i=1}^{N_{pt}} \frac{\left[D_{i} + \sum_{\alpha} \beta_{i,\alpha}^{exp} \lambda_{\alpha,exp} - T_{i} - \sum_{\alpha} \beta_{i,\alpha}^{th} \lambda_{\alpha,th}\right]^{2}}{s_{i}^{2} T^{2}} + \sum_{\alpha} \lambda_{\alpha,exp}^{2} + \sum_{\alpha} \lambda_{\alpha,th}^{2}.$$

5. Inconsistent redefinitions:

$$\chi^{2}(\vec{\lambda}_{exp},\vec{\lambda}_{th}) = \sum_{i=1}^{N_{pt}} \frac{\left[D_i + \sum_{\alpha} \beta_{i,\alpha}^{exp} \lambda_{\alpha,exp} - T_i - \sum_{\alpha} \beta_{i,\alpha}^{th} \lambda_{\alpha,th}\right]^2}{s_i^2} + \sum_{\alpha} \lambda_{\alpha,exp}^2 + \sum_{\alpha} \lambda_{\alpha,th}^2. \qquad \text{and } P(D_i|T_i) \propto e^{-\chi^2/2} + \sum_{\alpha} \lambda_{\alpha,th}^2 + \sum_{\alpha} \lambda_{$$

## Why augmented likelihood?

The term is accepted in lattice QCD to indicate that the log-likelihood contains quadratic prior terms

$$\chi^{2}(\vec{\lambda}_{\exp},\vec{\lambda}_{th}) = \sum_{i=1}^{N_{pt}} \frac{\left[D_{i} + \sum_{\alpha} \beta_{i,\alpha}^{\exp} \lambda_{\alpha,\exp} - T_{i} - \sum_{\alpha} \beta_{i,\alpha}^{th} \lambda_{\alpha,th}\right]^{2}}{s_{i}^{2}} + \sum_{\alpha} \lambda_{\alpha,\exp}^{2} + \sum_{\alpha} T^{2} \lambda_{\alpha,th}^{2}.$$
new experiment
new experiment
priors on expt. systematics
and PDF params

After minimization w.r.t. to  $\lambda_{\alpha,exp}$ ,  $\lambda_{\alpha,th}$ , the prior terms are **hidden** inside the covariance matrix:

$$\chi^{2} = \sum_{i,j}^{N_{pt}} (T_{i} - D_{i}) (\text{cov}^{-1})_{ij} (T_{j} - D_{j})$$

The usual  $\chi^2$  definition therefore contains a **prior** component, which may be handled differently by the various groups

## **Tolerances explained by epistemic uncertainties**

Relative PDF uncertainties on the *gg* luminosity at 14 TeV in three PDF4LHC21 fits to the **identical** reduced global data set arXiv:2203.05506



While the fitted data sets are identical or similar in several such analyses, the differences in uncertainties can be explained by methodological choices adopted by the PDF fitting groups.

NNPDF3.1' and especially 4.0 (based on the NN's+ MC technique) tend to give smaller nominal uncertainties in data-constrained regions than CT18 or MSHT20

# Epistemic uncertainties explain some of these differences.

- 1. Inclusion of multiple parametric forms in the CT18 uncertainty
- 2. Constraints from the effective prior in the NNPDF4.0 uncertainty
- 3. Parametrization uncertainty in xFittter/JAM PDF fits, lattice QCD PDFs...

### CT18: the uncertainty reflects multiple PDF parametrizations



**Upper figure:** A large part of the CT18 PDF uncertainty accounts for the sampling over 250-350 parametrization forms, possible choices of fitted experiments and fitting parameters, definitions of  $\chi^2$ 

**Lower figure:** this approach sometimes enlarges the uncertainties compared to the other groups, reflecting the chosen goodness-of-fit (tolerance) criterion more than the strength of experimental constraints

However, more restrictive tolerance criteria elevate the risk of sampling biases.

A more advanced CT tolerance prescription is under development.

Easier to examine these issues for specific QCD observables than in abstract

## NNPDF4.0: hopscotch scans suggest enlarged uncertainties

NNPDF replicas sample **aleatory** data fluctuations for a fixed training methodology (called "importance sampling" by NNPDF)

Representative sampling of **epistemic** uncertainty is challenging because of the large NN (hyper)parameter space

- Curse of dimensionality
- Big-data paradox [X.-L. Meng, Ann. App. Stat., 12 (2018) 685; F. Hickernell, MCQMC 2016, 1702.01487]

A **hopscotch scan** is a technique to densely sample a few PDF parameter combinations relevant for the QCD observable of interest by using NNPDF4.0 **Hessian PDFs** and NNPDF4.0 fitting code

The hopscotch scan relies on **dimensionality reduction** 



Figure 3.9. The neural network architecture adopted for NNPDF4.0. A single network is used, whose eight output values are the PDFs in the evolution (red) or the flavor basis (blue box). The architecture displayed corresponds to the optimal choice in the evolution basis; the optimal architecture in the flavor basis is different as indicated by Table 3.3).

R. Ball et al., arXiv:2109.02653



# How the hopscotch solutions are found

- 1. Examine the quasi-Gaussian  $\chi^2$ dependence along 50 Hessian EV directions
- 2. Perform high-density MC sampling of a span of a few EV directions that drive the specific PDF uncertainty





### Monte-Carlo sampling of PDF parametrizations

Using the public NNPDF4.0 fitting code, we find well-behaving PDF solutions to the NN4.0 fit that have better  $\chi^2$  with respect to central data values (by as much as 35-80 units depending on the  $\chi^2$  definition) than the published replica 0. These replicas follow a regular pattern. They lie outside of the nominal (red) NN4.0 uncertainties in the 50-dimensional PDF parameter space.

## The hopscotch scans: NNPDF4.0 vs CT18 uncertainties



The ellipses are projections of 68% c.l. ellipsoids in  $N_{par}$ -dim. spaces

 $N_{par} = 28$  and 50 for CT18 and NNPDF4.0 Hessian PDFs

#### Monte-Carlo sampling of PDF parametrizations



Nominal NN4.0 Hessian or MC 68%cl

least as large as shown

#### Hopscotch scans realize the likelihood-ratio test



According to the LR test, the NN4.0 analysis discards PDFs in the green and blue regions based on the prior probabilities and differences in the likelihood definitions – both associated with prior terms

The allowed regions will change for the other acceptable  $\chi^2$  definitions, which exist in reflection of the biasvariance dilemma

## Goodness-of-fit functions in PDF analyses

Analysis	χ <sup>2</sup> prescription to fit PDFs	$\chi^2$ prescription to compare PDFs	Comments
HERAPDF	HERA	HERA	
СТ	Extended $T$ +prior	Extended <i>T</i> , Experimental	
MSHT'20	Т	Т	
NNPDF4.0	t <sub>0</sub> + prior with fluctuated cross-sampled data	Experimental or t <sub>0</sub> with unfluctuated full data	<i>t</i> <sub>0</sub> prescription has pre- and post-NNPDF3.0 versions
Hopscotch'2022	N/A	Experimental or $t_0$ [2022] with unfluctuated data	

Different prescriptions reflect modeling of additive and multiplicative systematic errors in covariance matrices



Search docs

Getting started

Fitting code: n3fit

Code for data: validphys

/ Chi square figures of merit

#### Chi square figures of merit

Within the NNPDF methodology various figures of merit are used, each of which can be used in different situations. To avoid confusion, it is important to understand the differences between the various figures of merit, and to understand which definition we are referring to in a given context. In particular, it is worth stressing that whenever a figure of merit is discussed, the  $t_0$  method (discussed below) applies.

#### Note

From NNPDF2.0 onwards the  $t_0$  formalism has been used to define the figure of merit used during the fitting of the PDFs.

#### Note

The  $t_0$  method is **not** used by default in other validphys applications, and instead the default is to compute the experimental  $\chi^2$ . To compute  $\chi^2_{t_0}$ , users need to specify

use\_t0: True
t0pdfset: <Some LHAPDF set>

in the relevant namespace. This will instruct actions such as validphys.results.dataset chi2 table() to compute the  $t_0$  estimator.

https://docs.nnpdf.science/figuresofmerit/index.html, accessed on 2023-03-28

## Systematic uncertainties and the bias-variance dilemma

$$\chi^{2} = \sum_{i,j}^{N_{pt}} (T_{i} - D_{i})(\text{cov}^{-1})_{ij} (T_{j} - D_{j}) \qquad (\text{cov})_{ij} = s_{i}^{2} \delta_{ij} + \sum_{\alpha=1}^{N_{\lambda}} \beta_{i,\alpha} \beta_{j,\alpha}$$

$$\beta_{i,\alpha} = \sigma_{i,\alpha} X_i$$

 $D_i$ ,  $T_i$ ,  $s_i$  are the central data, theory, uncorrelated error

 $\beta_{i,\alpha} \equiv \sigma_{i,\alpha} \hat{X}_i$  is the correlation matrix for  $N_{\lambda}$  nuisance parameters. Experiments publish  $\sigma_{i,\alpha}$ .

The "truth" normalizations  $\hat{X}_i$  in the experiment are of order  $T_i$  or  $D_i$ . { $\hat{X}_i$ } are learned as a model { $X_i$ } together with PDFs f and theory { $T_i(f)$ }. For example, we can sample as  $X_i = a_i D_i + b_i T_i$ , with free  $0 \le a_i, b_i \le 1$ .

**Mean variation \delta\_X^2 of the model from truth** on an ensemble of replicas, for data point *i*:

$$\delta_X^2 \equiv \left\langle \left(X_i - \hat{X}_i\right)^2 \right\rangle = \underbrace{\left\langle \left(\hat{X}_i - \langle X_i \rangle\right)^2 \right\rangle}_{\text{model bias}} + \underbrace{\left\langle (X_i - \langle X_i \rangle)^2 \right\rangle}_{\text{variance}} = \underbrace{\left\langle \left(\hat{X}_i - \langle X_i \rangle\right)^2 \right\rangle}_{\text{model bias}} - \underbrace{\left\langle (D_i - \langle X_i \rangle)^2 \right\rangle}_{\text{data bias}} + \underbrace{\left\langle (D_i - X_i)^2 \right\rangle}_{\chi^2(D_i, T_i)}$$

Experimental definition,  $X_i = D_i$ :  $\langle (X_i - \hat{X}_i)^2 \rangle = (\hat{X}_i - D_i)^2 \equiv \delta_D^2$ 

$$t_0$$
 definition,  $X_i = t_{0i}$ :  $\left\langle \left(X_i - \hat{X}_i\right)^2 \right\rangle = \left(\hat{X}_i - t_{0i}\right)^2 \equiv \delta_{t_0}^2$ 

In general, not enough information to compare  $\delta_D$  and  $\delta_{t_0}$ 

## Smoothing of *K*-factors

An analogous **bias-variance tradeoff** arises during smoothing of MC integration errors for *K*-factor tables

A smoother curve reduces the  $\chi^2$  for the data, but the best-fit result retains some dependence on the fitted functional form

This dependence can be conservatively estimated by including an uncorrelated MC integration error



## Possible criticisms [see R. Ball et al., arXiv:2211.12961] and our detailed response [arXiv: 2205.10444, version 5]

 Criticism: hopscotch solutions are improbable according to the random resampling ("importance sampling") of fitted data with the fixed NNPDF4.0 training methodology.
 Our response: Hopscotch solutions will be likely if the NN training methodology is varied. Experimental data resampling does not account for methodology variations.

2. **Criticism:** hopscotch solutions fail smoothness conditions during NN4.0 replica training and are discarded. **Our response:** Unclear how many of 2330+50 hopscotch solutions were tested by NNPDF. Most of hopscotch solutions are sufficiently smooth upon a typical CTEQ-TEA examination and largely fall within NNPDF4.0 uncertainty bands. Smoothness is not a sharply defined criterion, cf. the bias-variance dilemma.

3. **Criticism:** among the various prescriptions for approximating correlated systematic uncertainties in  $\chi^2$ , only  $t_0$  prescription used for NNPDF replica training should be used for exploring the PDF uncertainty. **Our response:** beyond relatively simple examples of D'Agostini's bias explored by NNPDF [arXiv:0912.2276] and others, there is no rigorous demonstration that a particular  $\chi^2$  prescription is preferable. Counterexamples exist. A variety of other  $\chi^2$  prescriptions are used, cf. the bias-variance dilemma. NNPDF continues to use the experimental  $\chi^2$  prescription for PDF comparisons in the NN4.0 publication and NN4.0 validphys code [except during NN training].

## The hopscotch scan counterbalances the bias of the nominal replica ensemble

#### 6.2 Creating a less biased sub-sample

The basic idea is to use such partial information about the selection bias to design a *biased* subsampling scheme to *counterbalance* the bias in the original sample, such that the resulting sub-samples have a *high likelihood* to be less biased than the original sample from our target population. That is, we create a sub-sampling indicator  $S_I$ , such that with high likelihood, the correlation between  $S_IR_I$  and  $G_I$ is reduced, compared to the original  $\rho_{R,G}$ , to such a degree that it will compensate for the loss of sample size and hence reduce the MSE of our estimator (e.g., the sample average). We say with *high likelihood*, in its non-technical meaning, because without full information on the response/recording mechanism, we can never guarantee such a counterbalance sub-sampling (CBS) would always do better. However, with judicious execution, we can reduce the likelihood of making serious mistakes.

X.-L. Meng, Survey Methodology, Catalogue 12-001-X, vol. 48 (2022), #2

## Hopscotch NN4.0 replicas

LHAPDF6 grids available at <a href="https://ct.hepforge.org/PDFs/2022hopscotch/">https://ct.hepforge.org/PDFs/2022hopscotch/</a>





## Scans of the log-likelihood in EV directions 25 and 33



## Hopscotch replicas enlarge the error bands



FIG. 9. Solid bands indicate the nominal 68% NNPDF4.0 uncertainties for strangeness asymmetry (left) and charm PDF (right) at Q = 1.7 GeV. The alternative EV sets with  $\Delta \chi^2_{t_0} = 0$  are plotted as dashed lines.

At x > 0.2,  $Q \approx Q_0 = 1.51$  GeV, the hopscotch replicas reduce significance of  $(s - \bar{s})/(s + \bar{s}) \approx 50\%$  (left) and  $c(x, Q) \neq 0$  (right). This washes out the  $3\sigma$  evidence for the "intrinsic charm" stated in R. Ball et al., Nature 608 no. 7923, (2022) 483.

### Epistemic PDF uncertainty:

*Epistemic uncertainty* (due to parametrization, methodology, parametrization/NN architecture, smoothness, data tensions, model for syst. errors, ...) is increasingly important in NNLO global fits as experimental and theoretical uncertainties decrease

Nominal PDF uncertainties in high-stake measurements (ATLAS W mass, Higgs cross sections...) thus should be tested for *robustness of sampling over acceptable methodologies* and demonstrate *absence of biases* in this sampling.

This is also necessary for combination of PDFs including data correlations [LHC EW, Jet & Vector boson WGs, <u>https://tinyurl.com/4wcnd8xn</u>; <u>https://tinyurl.com/2p8d8ba3</u>; <u>https://tinyurl.com/2p8tcn5b</u>; Ball, Forte, Stegeman, arXiv:2110.08274].

Such tests can be done outside of the PDF fits using hopscotch scans. [arXiv: 2205.10444, Sec. 2.].

The ambiguity due to the  $\chi^2$  definition is significant. Publication of full likelihoods for experimental systematic errors [Cranmer, Prosper, et al., arXiv:2109.04981] will suppress this ambiguity.

- Hopscotch scans were illustrated using the NNPDF4.0 public code and LHAPDF grids, and mp4lhc program.
- Impact on the uncertainties at small and large x, PDF ratios, fitted charm, ...
- Insights applicable to other analyses using a large parameter space CT/MSHT tolerance, polarized PDFs, etc.

Uncertainty quantification, a challenge for AI, As we try to analyze PDFs and understand why. With machine learning methods we strive To make sense of the data and derive.

But uncertainty presents a hurdle As we try to make predictions and be certain. It's a challenge that we must face As we work to improve our models with grace.

Parton distributions, oh how they vex As we try to understand their complex effects. But still we persist, for we must know The secrets that uncertainty has yet to show.

#### **Microsoft Bing**

## Backup

## Computing uncertainty $\Delta X$

 By unweighted averaging of predictions for 100 (or 1000) MC replicas:

$$\langle X \rangle = \frac{1}{N_{rep}} \sum_{i=1}^{N_{rep}} X_i \, ; \quad \Delta X^2 = \langle (X - \langle X \rangle)^2 \rangle$$

(NNPDF calls it "**importance sampling**". The MC replicas are distributed according to the fluctuated data [Ball:2011gg] using the same training algorithm).



Replica 0 is the mean of 1000 MC replicas; has better unfluctuated  $\chi^2$  than MC replicas.

2. Using  $N_{eig} = 50$  Hessian PDFs.

$$\Delta X^2 = \sum_{i=1}^{N_{eig}} (X_i - X_0)^2 \,.$$

NNPDF4.0 MC and Hessian uncertainties are in a good agreement.

## Figures of merit in the NNPDF4.0 analysis I

1.  $\chi^2$  with respect to the central experimental values

$$\chi^{2} = \sum_{i,j}^{N_{pt}} (T_{i} - D_{i})(\operatorname{cov}^{-1})_{ij} (T_{j} - D_{j})$$
$$(\operatorname{cov})_{ij} \equiv s_{i}^{2} \delta_{ij} + \sum_{\alpha=1}^{N_{\lambda}} \beta_{i,\alpha} \beta_{j,\alpha}, \qquad \qquad \beta_{i,\alpha} = \sigma_{i,\alpha} X_{i},$$

 $D_i$ ,  $T_i$ ,  $s_i$  are the central data, theory, uncorrelated error  $\beta_{i,\alpha}$  is the correlation matrix for  $N_{\lambda}$  nuisance parameters.

Experiments publish  $\sigma_{i,\alpha}$ . To reconstruct  $\beta_{i,\alpha}$ , we need to decide on the normalizations  $X_i$ .

NNPDF4.0 use:

*a.*  $X_i = D_i$  : "**exp**erimental scheme"; can result in a bias *b.*  $X_i = \text{fixed } T_i$  : " $t_0$  scheme"; can result in a (different) bias

## Figures of merit in the NNPDF4.0 analysis II

$$(\operatorname{cov})_{ij} \equiv s_i^2 \delta_{ij} + \sum_{\alpha=1}^{N_{\lambda}} \beta_{i,\alpha} \beta_{j,\alpha},$$

$$\beta_{i,\alpha} = \sigma_{i,\alpha} X_i,$$

NNPDF4.0 use:

*a.*  $X_i = D_i$  : **exp**erimental scheme; can result in a bias *b.*  $X_i = \text{fixed } T_i : t_0$  scheme; can result in a (different) bias

The conventions are neither complete nor unique. Ambiguity affects all groups. See Appendix in <u>1211.5142</u>.

2. NNPDF4.0 trains MC replicas with  $\chi^2$  for fluctuated  $D_i$ ,  $t_0$  scheme, and replica selection (prior) conditions:

$$Cost = \chi_{t_0}^2(T_i, D_i^{fluctuated}) + \chi_{prior}^2$$

3. NNPDF4.0 quotes the final unfluctuated  $\chi^2$  in the "exp" scheme.

 $t_0$  scheme:  $\chi^2_{tot}/N_{pt} = 1.233$ .

 $\chi^2_{tot}/N_{pt} = 1.160$ .

$$\chi^{2}(\exp) - \chi^{2}(t_{0}) = -340$$
 for 4618 data points

## **PRIOR PROBABILITY IN PDF FITS**

✓ PDF fitting example of inverse problem: aim to find a posterior probability of **f** given the data **D**.

✓ Parametrization of PDFs: finite-dimensional problem.

 $f(x) \approx \tilde{f}(x, \theta) \in \mathcal{F}$ 

✓ The posterior probability for the parametrization depends on both the figure of merit that maximises the data likelihood given the parameters and on prior probability *H*.

(M. Ubiali, HP2 2022 workshop, Durham, 2022-09-22)

### Why doesn't NNPDF4.0 find HS solutions?



NNPDF authors find that some HS replicas fail the initial-stage overfitting test (M. Ubiali, HP2 2022 workshop, Durham, 2022-09-22) -10 = 10 = 10 -20 = 10 -20 = 10 -4 = -3 = -2 = -1 = 0 xg (x,Q) at Q = 1.7 GeV (sym. err)  $NNPDF4.0 \text{ NNLO 68\% (solid), alt. } (\Delta \chi^2)_{t0} = 0 \text{ (dashed)}$ 

10 0 NN40full t0 EV



HS solutions have much lower  $\chi^2$  than NN MC replicas. HS PDFs are outside the 50-dim neighborhood of NN replica 0. We do not see evidence of "overfitting" according to CT18 criteria. Collaborations with other groups

## Snowmass'21 whitepaper: Proton structrure at the precision frontier

S. Amoroso et al., Acta Physica Polonica B 53 (2022) 12, A1

#### A summary of recent trends in the global analysis of proton PDFs

- 1. Status of modern NNLO PDFs and their applications
- 2. Future experiments to constrain PDFs
- 3. Theory of PDF analysis at N2LO and N3LO
- 4. New methodological advancements
  - Experimental systematic uncertainties in PDF fits
  - Theoretical uncertainties in PDF fits
  - Machine learning/AI connections
- 5. Delivery of PDFs; PDF ensemble correlations in critical applications
- 6. PDFs and QCD coupling strength on the lattice
- 7. Nuclear, meson, transverse-momentum dependent PDFs
- 8. Public PDF fitting codes
- 9. Fast (N)NLO interfaces

10. PDF4LHC21 recommendation and PDF4LHC21 PDFs for the LHC analyses



# Progress in PDF analysis



# Snowmass 2021 whitepaper: Proton structure at the precision frontier

S. Amoroso et al., Acta Physica Polonica B 53 (2022) 12, A1



# An ATLAS, CTEQ-TEA, and MSHT comparative study of NNLO PDF sensitivities



- Comparisons of strengths of constraints from individual data sets in 8 PDF analyses using the common L<sub>2</sub> sensitivity metric
- An interactive website to plot such comparisons [2070 figures in total]

Preview