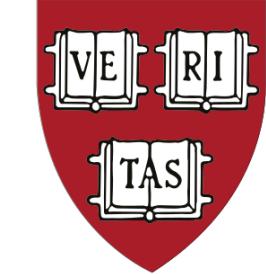




FI



HARVARD
UNIVERSITY



Massachusetts
Institute of
Technology

Boosting Statistical Anomaly Detection via *multiple test* with NPLM

Gaia Grosso^{1,2,3,*}, Marco Letizia^{4,5}, Phil Harris^{1,2}

¹ NSF Institute for Artificial Intelligence and Fundamental Interactions (IAIFI)

²Massachusetts Inst. of Technology, ³Harvard University

⁴MaLGa Center - DIBRIS, University di Genova, ⁵INFN, Sezione di Genova

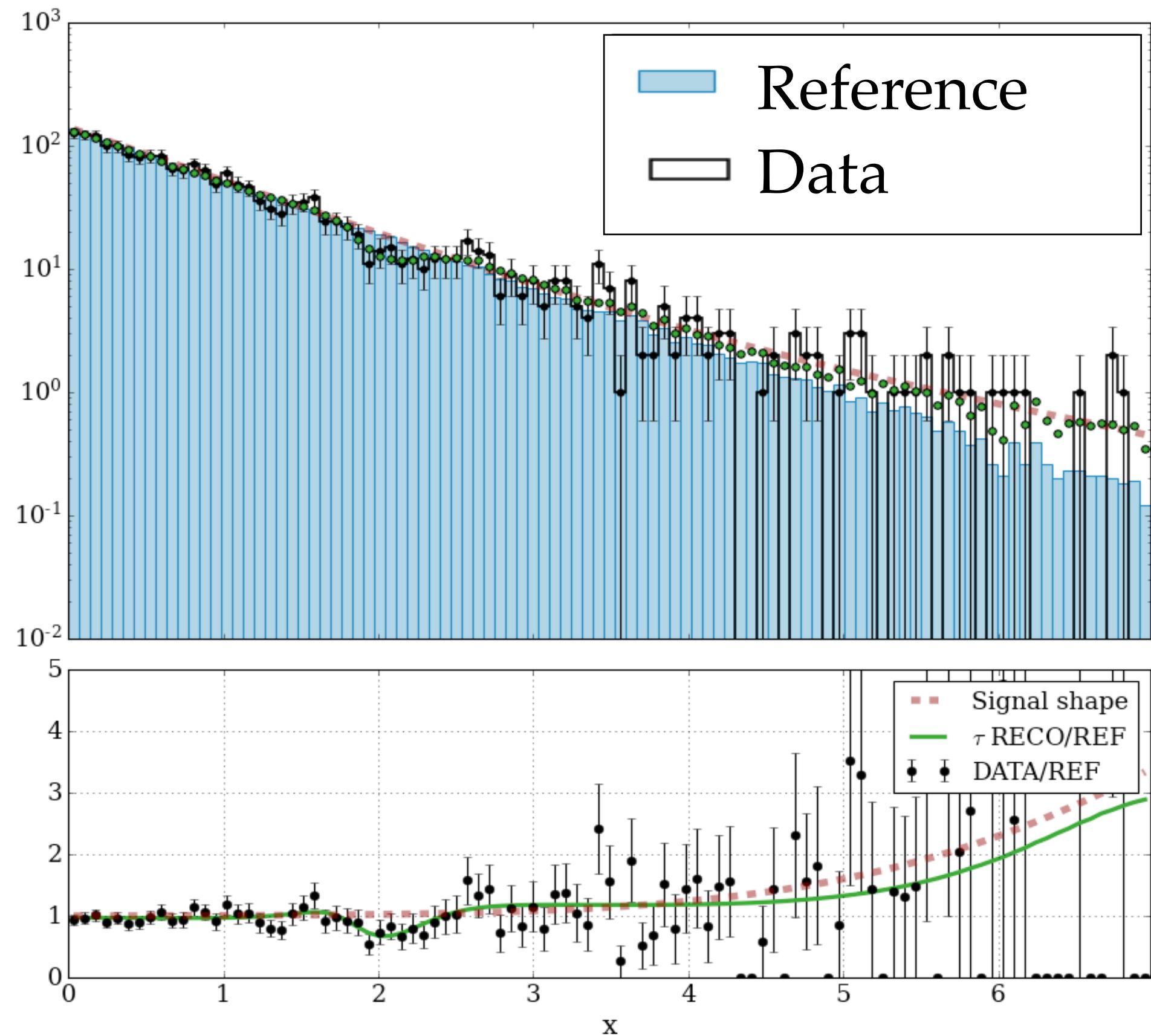
*gaia.grosso@cern.ch

March 13th, 2023



ACAT 2024

Statistical anomaly detection as a goodness of fit

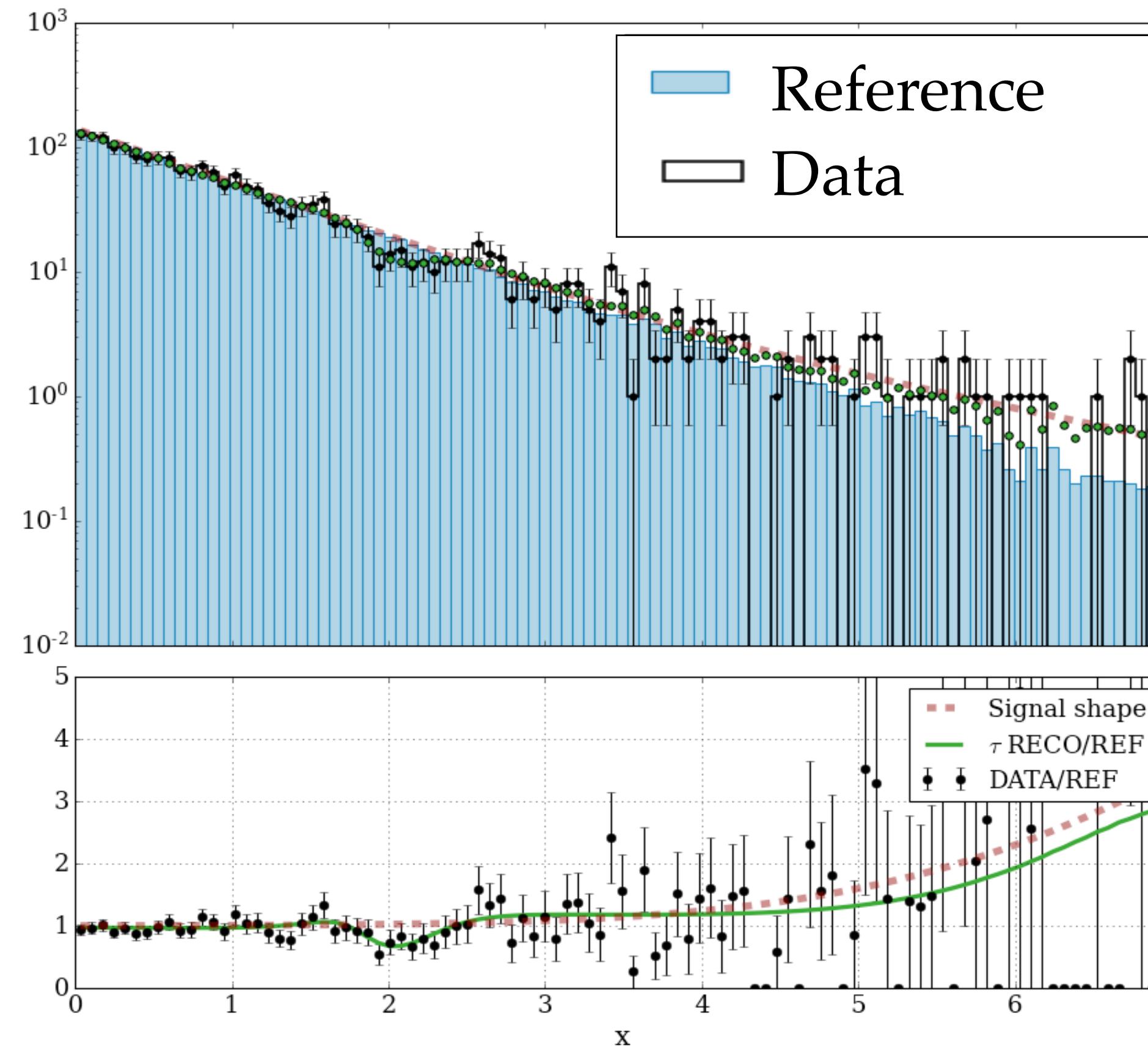


Problems defined by:

- **Data:** experimental measurements of the natural process
- **Reference model:** expected nominal behavior of the data

Is the Reference model a *good* description of the data?

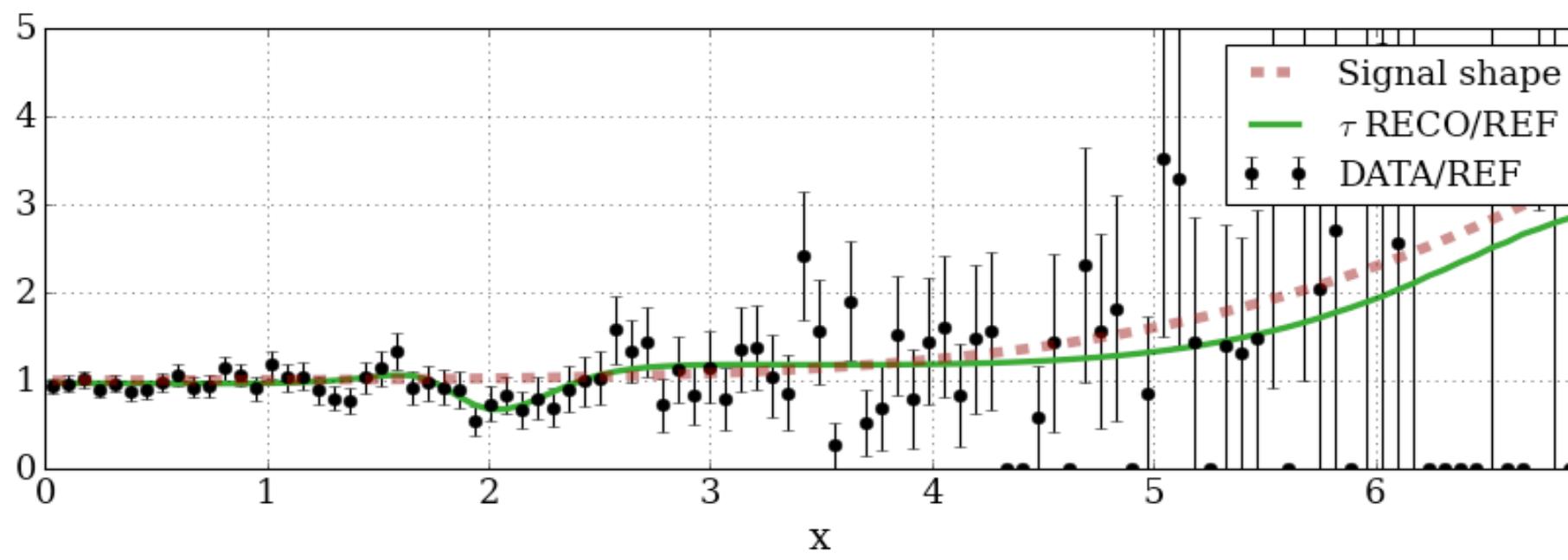
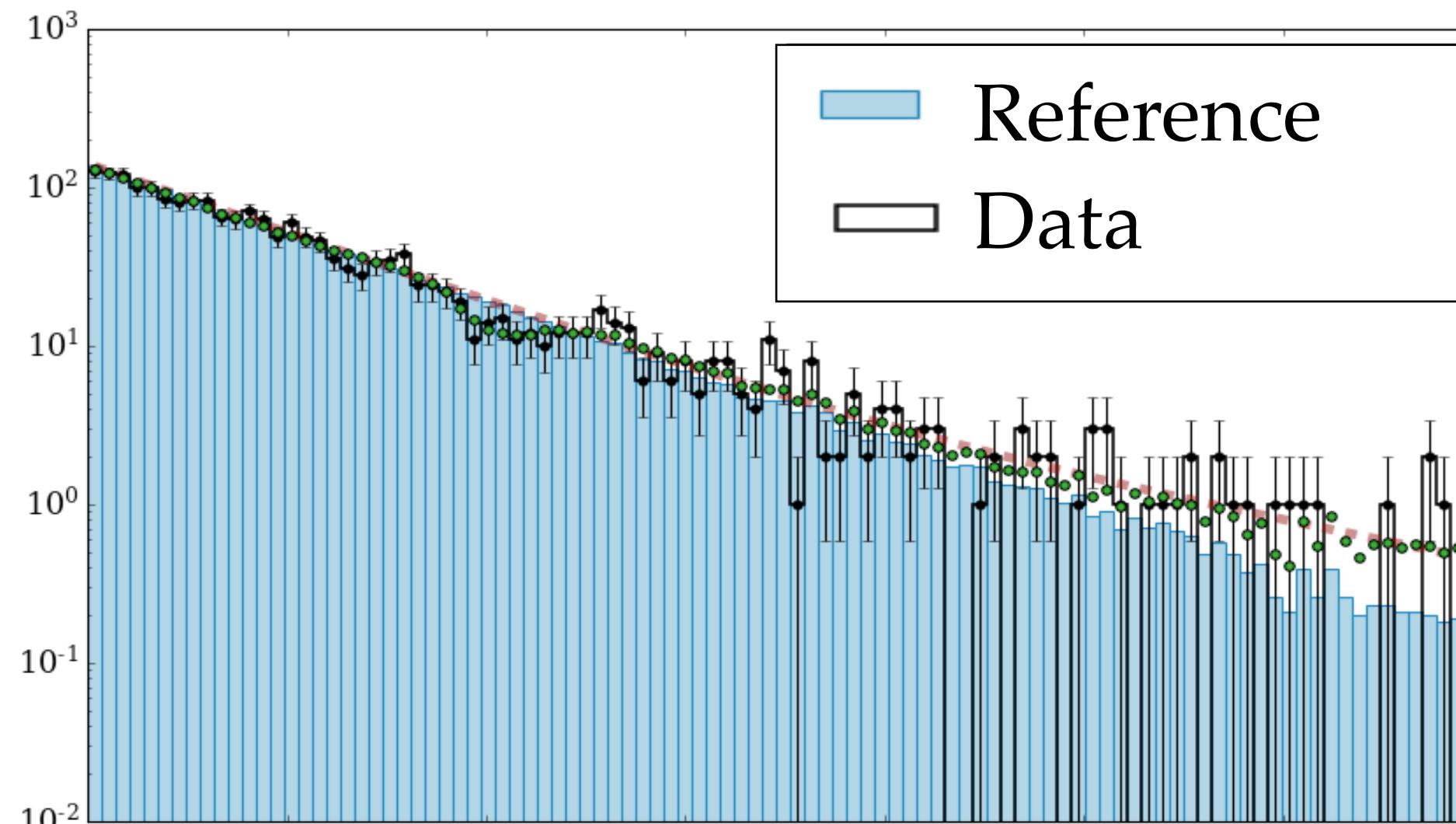
Statistical anomaly detection as a goodness of fit



Many use cases:

- Experimental system monitoring (DQM)
 - Signal agnostic New Physics searches
 - Data validation
- assessing the *goodness of Generative models*

Statistical anomaly detection as a **goodness of fit**



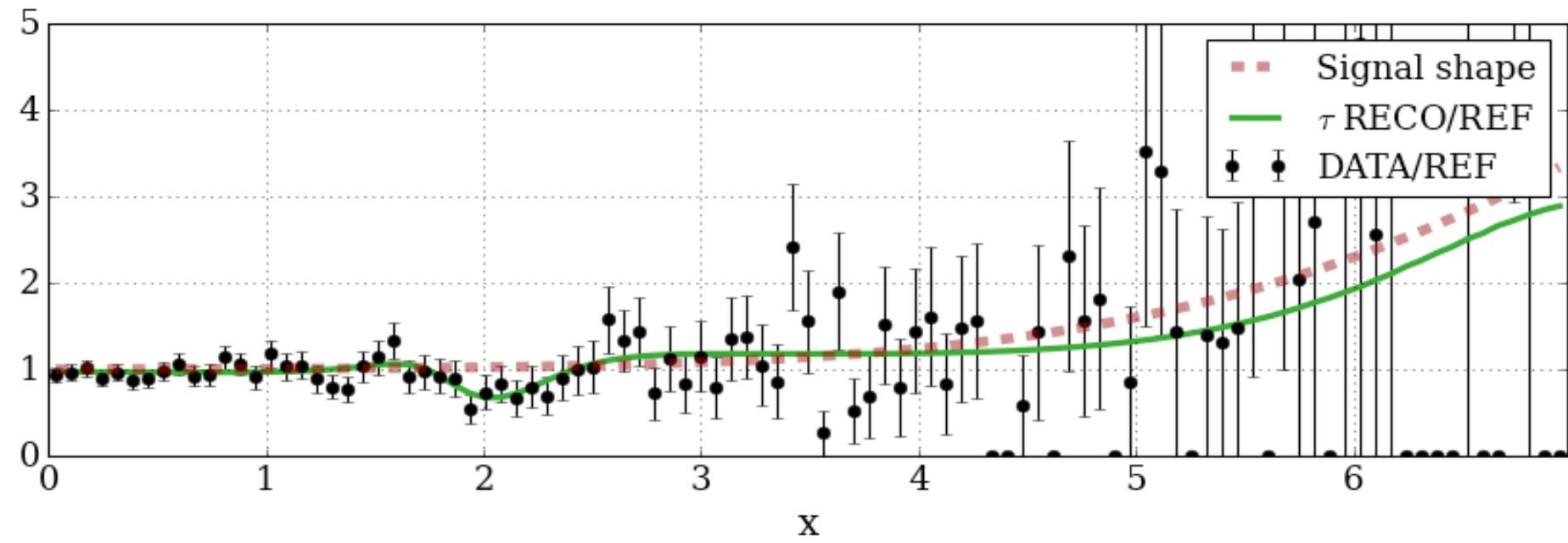
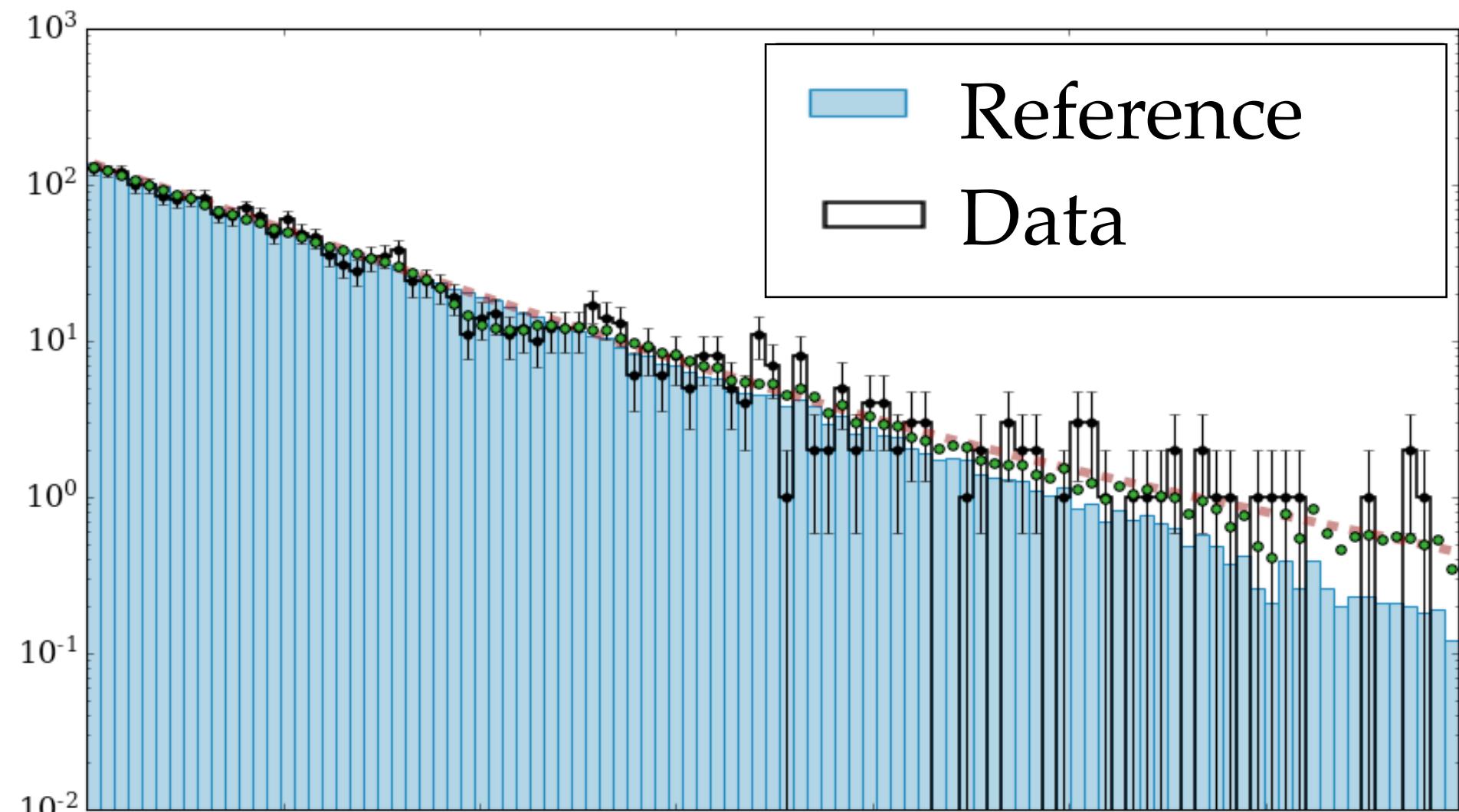
Challenges:

- Anomalies are *rare!*
- Anomalies are *unexpected!*

- *large statistics samples* to reach sensitivity
- *high dimensional raw data* for inclusive test
- the ideal statistical model of the data
(aka *inductive bias*) is unknown

Motivation for ML based
solutions

Statistical anomaly detection as a goodness of fit



How to exploit large datasets?

How to work around inductive biases?

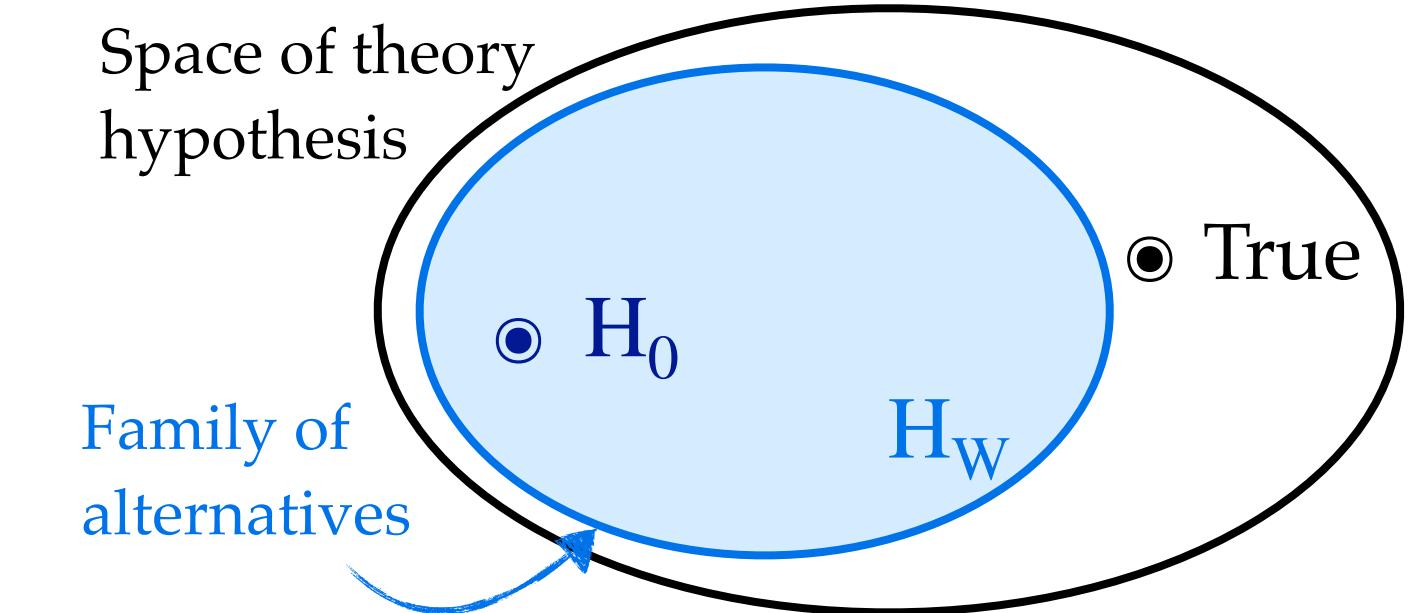
→ *multiple testing* in the context of the **Neyman-Pearson GoF test**
[GG, Letizia, Wulzer, Pierini [2305.14137](#)]

ML-based Neyman-Pearson GoF test

Compare the Reference hypothesis with an *alternative*.

Inductive bias: definition of the family of alternatives

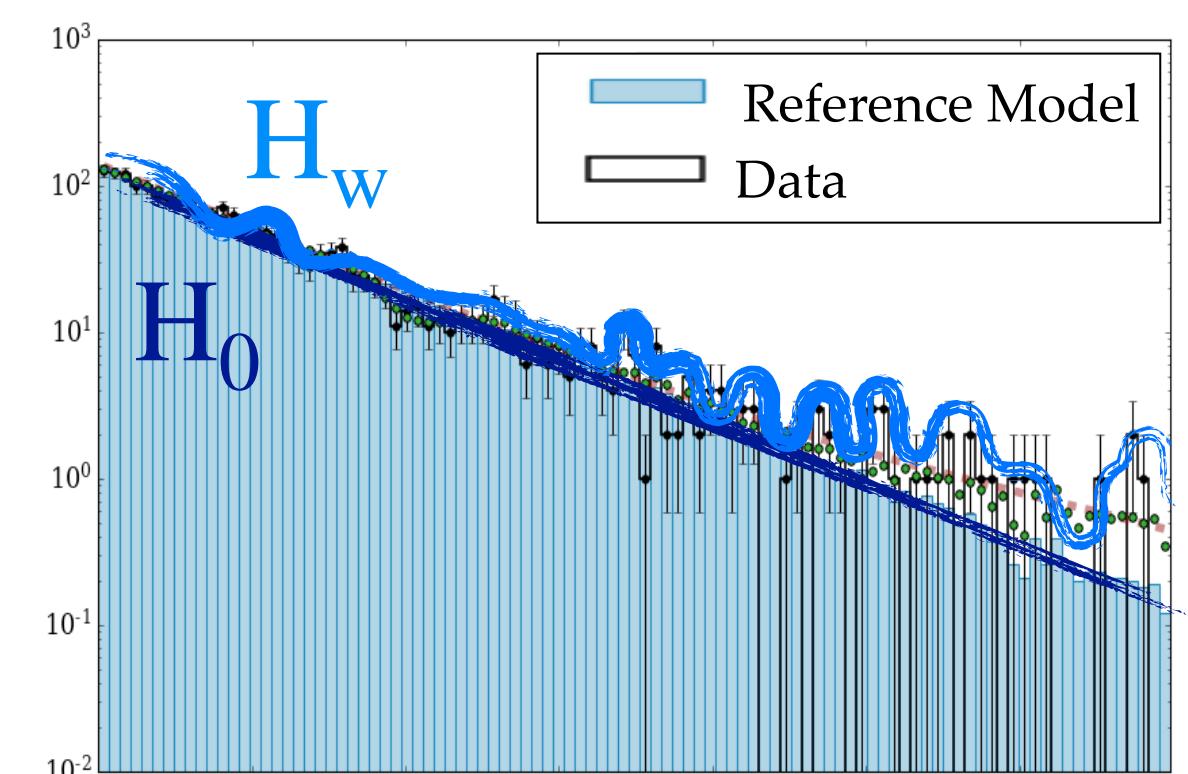
$$t(\mathcal{D}) = \max_{\mathbf{w}} \left[2 \log \frac{\mathcal{L}(\mathcal{D} | H_w)}{\mathcal{L}(\mathcal{D} | H_0)} \right]$$



New physics Learning Machine (NPLM)

Universal approximator
(NN, kernel methods, ...)

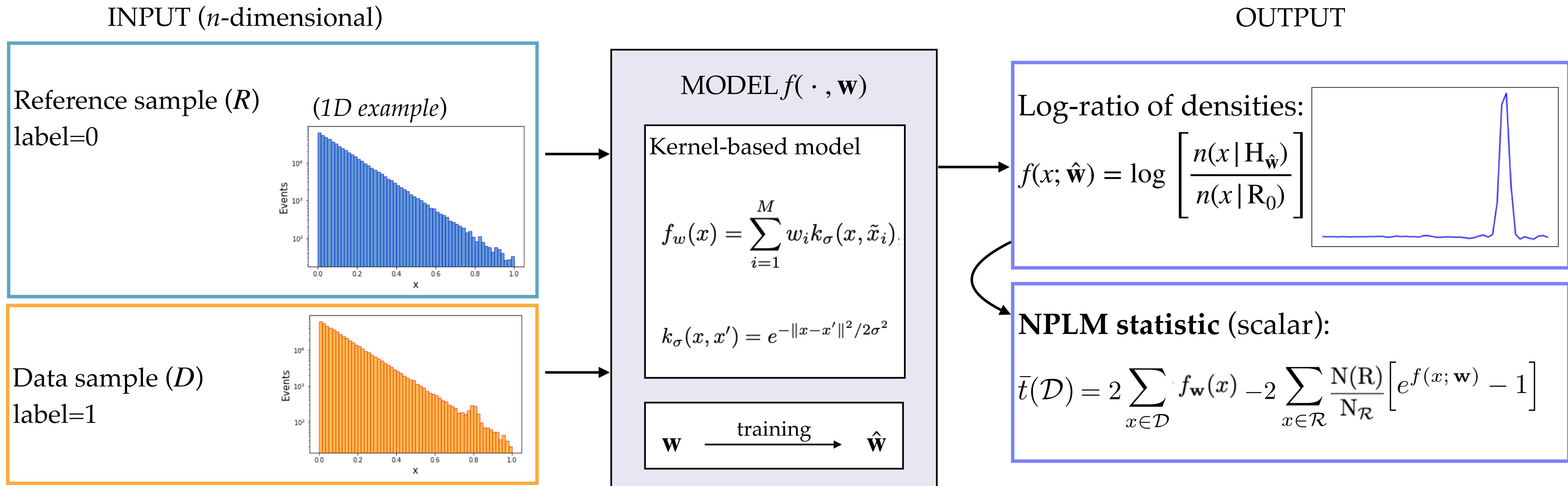
$$n(x|H_w) = e^{f(x; \mathbf{w})} n(x|R_0)$$



"Learning New Physics from a Machine" [Phys. Rev. D](#)

ML-based Neyman-Pearson GoF test

NPLM: implementation

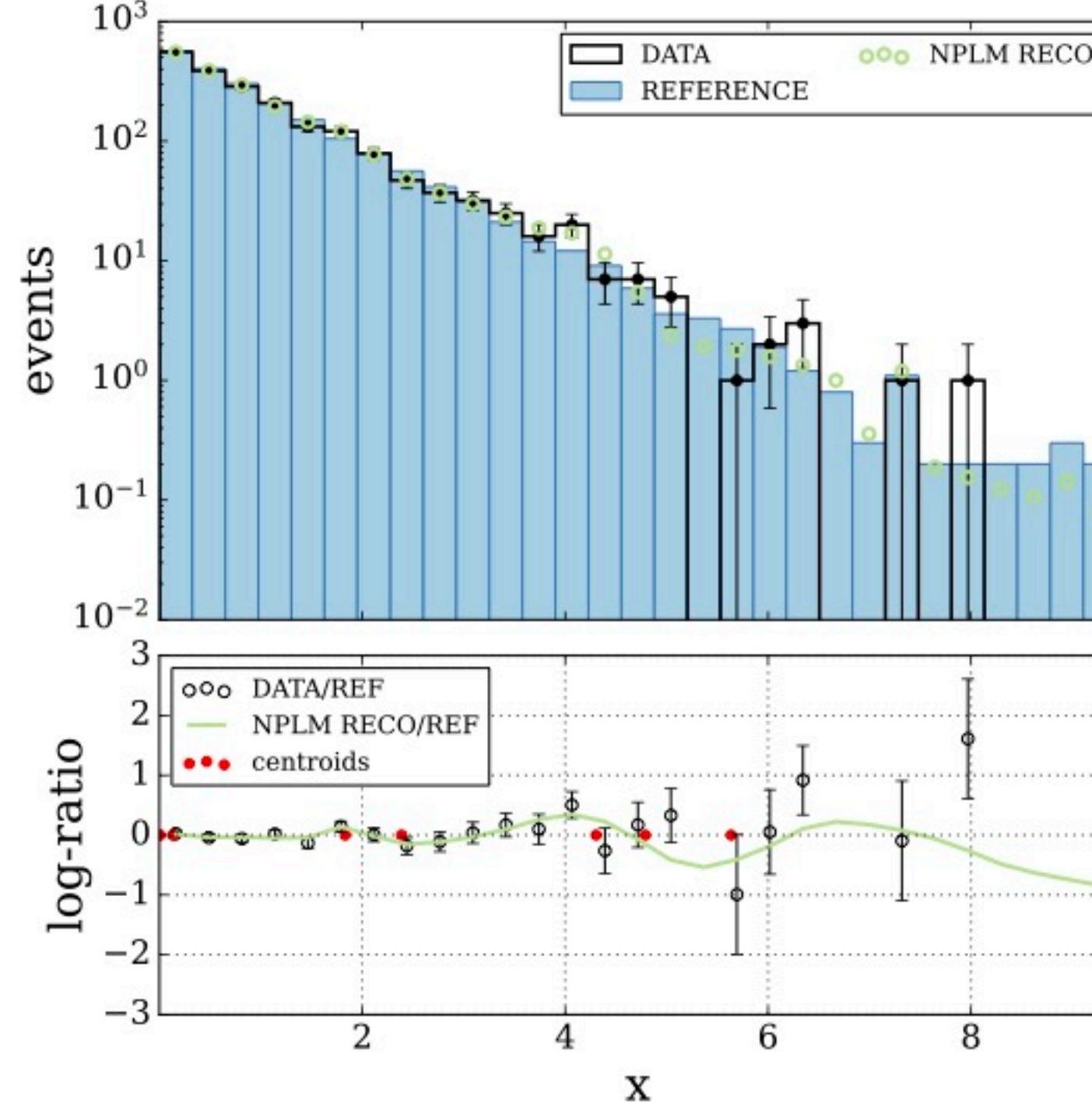


"Learning New Physics from a Machine" [Phys. Rev. D](#)

How to mitigate wrong inductive biases?

Inductive bias from model selection

Kernel Methods model hyper-parameter choice



Kernel-based model: $f_w(x) = \sum_{i=1}^M w_i k_\sigma(x, \tilde{x}_i)$ $k_\sigma(x, x') = e^{-\|x-x'\|^2/2\sigma^2}$
(gaussian kernel)

Loss:

$$\hat{L}(f_w) + \lambda R(f_w)$$

Weighted binary cross entropy:

$$\hat{L}(f_w) = \frac{1}{N} \sum_{i=1}^N a_0(1-y) \log(1+e^{f(x)}) + a_1 y \log(1+e^{-f(x)})$$

Regularization term:

$$R(f_w) = \sum_{ij} w_i w_j k_\sigma(x_i, x_j)$$

Hyperparameters:

M : number of kernels

σ : kernel width

λ : L2 regularization

The hyper parameters M, σ, λ define the family of alternatives

Aggregation of multiple tests

Aggregation rule for the p -value:

$$p_{\text{aggreg}} = \min_{\sigma \in [\sigma_1, \dots, \sigma_n]} [p_\sigma]$$

$$[\sigma_1, \dots, \sigma_n] = [5\%, 25\%, 50\%, 75\%, 95\%]$$

quantiles of the pairwise distance between reference-distributed data points (after standardization).

Strategy:

1. **Test toys under the null hypothesis** by sampling background events:

$$\{t_\sigma(D_{\text{pseudo}}), \sigma \in [\sigma_1, \dots, \sigma_n]\}_{i=0}^{N_{\text{toys}}}.$$

2. **Test the data of interest** $\forall \sigma: \{t_\sigma(D), \sigma \in [\sigma_1, \dots, \sigma_n]\}$
3. Compute the empirical p-values $\forall \sigma: \{p_\sigma(D), \sigma \in [\sigma_1, \dots, \sigma_n]\}$
4. Select the **minimum p-value**: $p_{\text{aggreg}} = \min_{\sigma \in [\sigma_1, \dots, \sigma_n]} [p_\sigma]$

Aggregation of multiple tests

1D proof-of-concept

Signal benchmarks: gaussian resonances with various width (σ_{NP}) and locations (\bar{x}_{NP}).
[N(S) signal injection over N(B) = 2000 events]

NPLM simple tests
with different
kernel widths (σ)

Aggregation of
simple tests

N(S)	7	18	13	10	90
\bar{x}_{NP}	4	4	4	6.4	1.6
σ_{NP}	0.01	0.16	0.64	0.16	0.16
$\sigma = 0.1$	0.008 ± 0.003	0.032 ± 0.006	0.002 ± 0.001	0.026 ± 0.005	0.30 ± 0.02
$\sigma = 0.3$	0.001 ± 0.001	0.056 ± 0.007	0.001 ± 0.001	0.14 ± 0.01	0.49 ± 0.02
$\sigma = 0.7$	0	0.059 ± 0.008	0.003 ± 0.002	0.21 ± 0.01	0.53 ± 0.02
$\sigma = 1.4$	0	0.045 ± 0.007	0.005 ± 0.002	0.19 ± 0.01	0.41 ± 0.02
$\sigma = 3.0$	0	0.020 ± 0.004	0.008 ± 0.003	0.11 ± 0.01	0.23 ± 0.02
aggregation	0.009 ± 0.003	0.11 ± 0.01	0.013 ± 0.004	0.27 ± 0.02	0.62 ± 0.02

Table 1: 1D experiments: probability of observing $Z \geq 3$

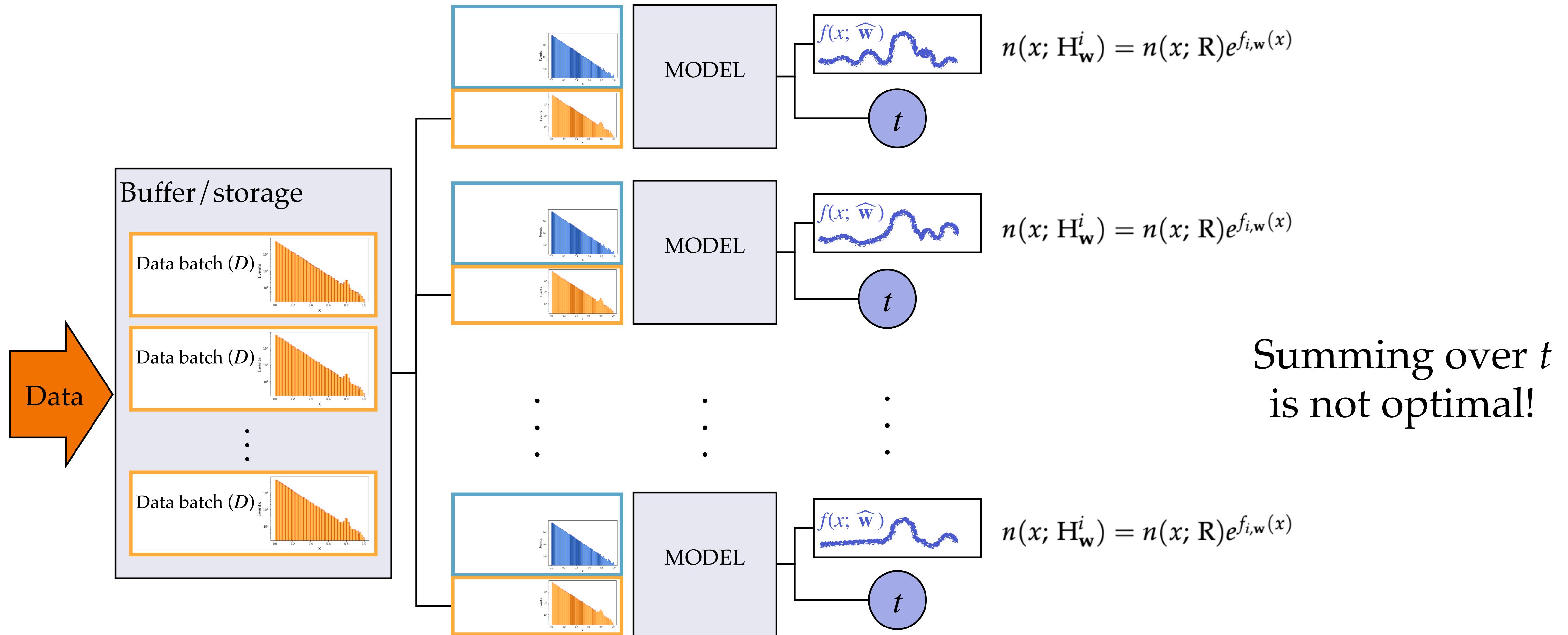
Single tests have different power over the signal benchmarks.
The aggregation shows **uniform enhanced power** over the range of benchmarks.

Preliminary multi-dimensional tests confirm these findings

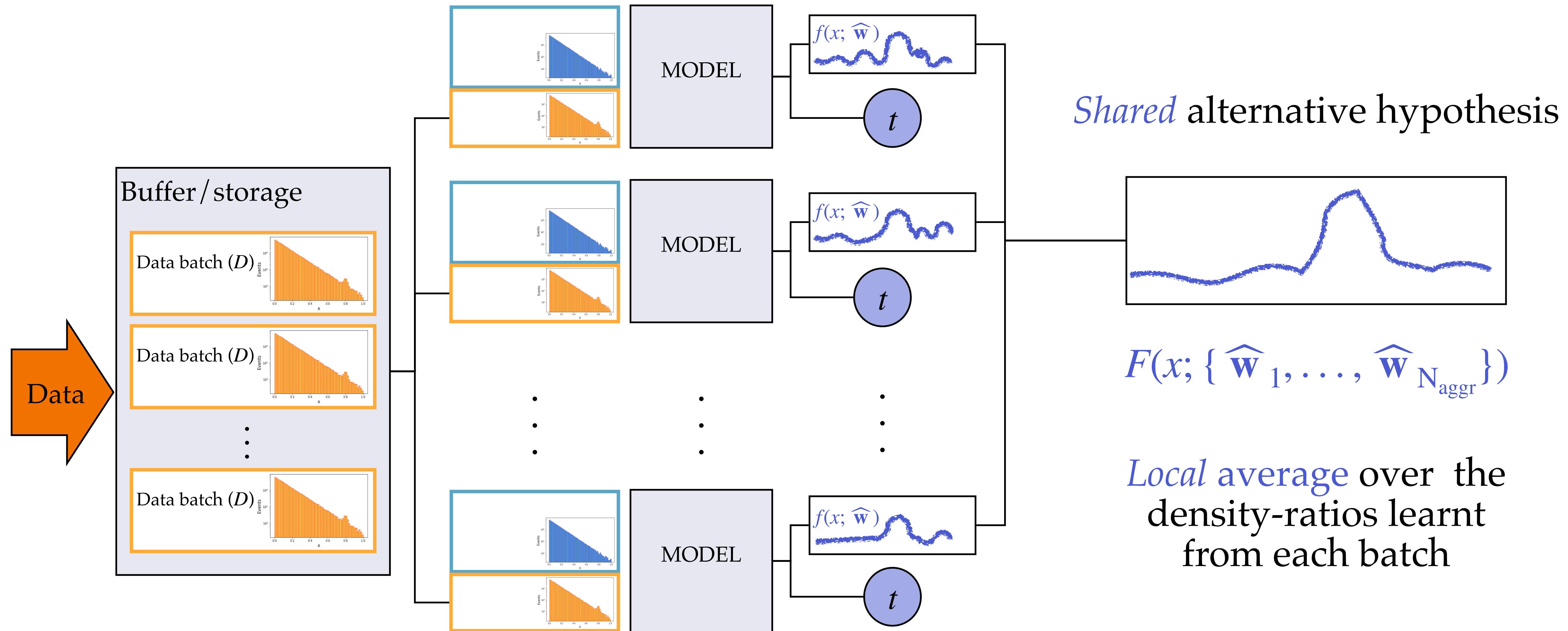
How to deal with large samples?



Combining NPLM over *multiple batches*



Combining NPLM over *multiple batches*



Combining NPLM over *multiple batches*

Shared alternative hypothesis

$$\begin{aligned}
 F_{\mathbf{W}}^{N_{\text{aggr}}}(x) &= \log \frac{n(x; \mathbf{H}_{\mathbf{w}}^{N_{\text{aggr}}})}{n(x; \mathbf{R})} \\
 &= \log \left[\frac{1}{N_{\text{aggr}}} \sum_{i=1}^{N_{\text{aggr}}} e^{f_{i,\mathbf{w}}(x)} \right]
 \end{aligned}$$

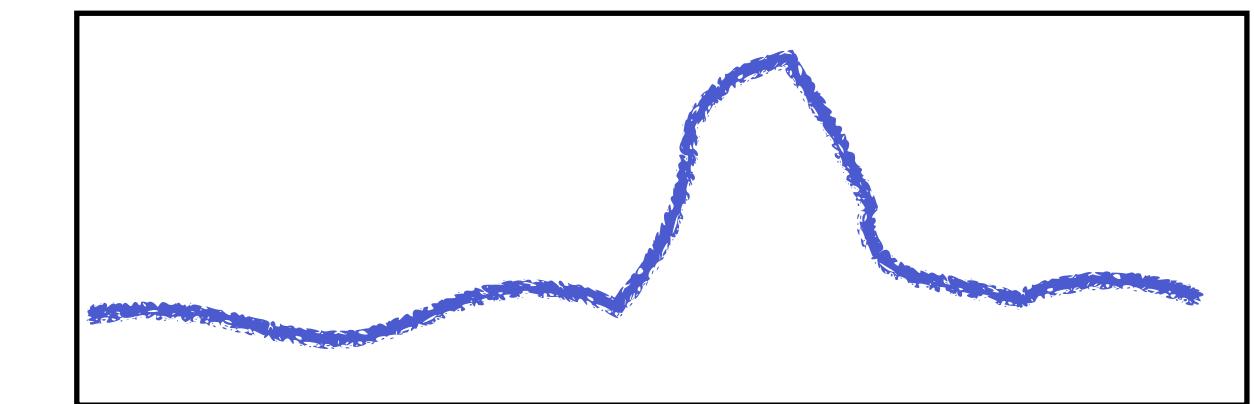
N_{aggr} : # of aggregated batches

$n(x; \mathbf{H}_{\mathbf{w}}^i) = n(x; \mathbf{R})e^{f_{i,\mathbf{w}}(x)}$

$$\begin{aligned}
 t_{\text{AGGR}}^{N_{\text{aggr}}, N_{\text{test}}}(\mathcal{D}) &= 2 \sum_{i=1}^{N_{\text{test}}} \log \frac{\mathcal{L}(\mathcal{D}_i | \mathbf{H}_{\mathbf{w}}^{N_{\text{aggr}}})}{\mathcal{L}(\mathcal{D}_i | \mathbf{R})} \\
 &= 2 \sum_{i=1}^{N_{\text{test}}} \left[\sum_{x \in \mathcal{R}} w_{\mathcal{R}}(x) (1 - e^{F_{\mathbf{W}}^{N_{\text{aggr}}}(x)}) + \sum_{x \in \mathcal{D}_i} F_{\mathbf{W}}^{N_{\text{aggr}}}(x) \right]
 \end{aligned}$$

N_{aggr} : # of aggregated batches

N_{test} : # of tested batches



$$F(x; \{ \widehat{\mathbf{w}}_1, \dots, \widehat{\mathbf{w}}_{N_{\text{aggr}}} \})$$

Local average over the density-ratios learnt from each batch

Combining NPLM over *multiple batches*

1D proof-of-concept

Signal benchmarks:

- Broad peak: $\bar{x} = 4, \sigma = 0.64$
- Narrow peak: $\bar{x} = 4, \sigma = 0.01$

TESTS:

Single batch jobs:

1) $N(B) = 2\,000$

- Narrow peak: $N(S) = 3 (Z_{\text{ideal}} = 1.7)$
- Broad peak: $N(S) = 13 (Z_{\text{ideal}} = 1.5)$

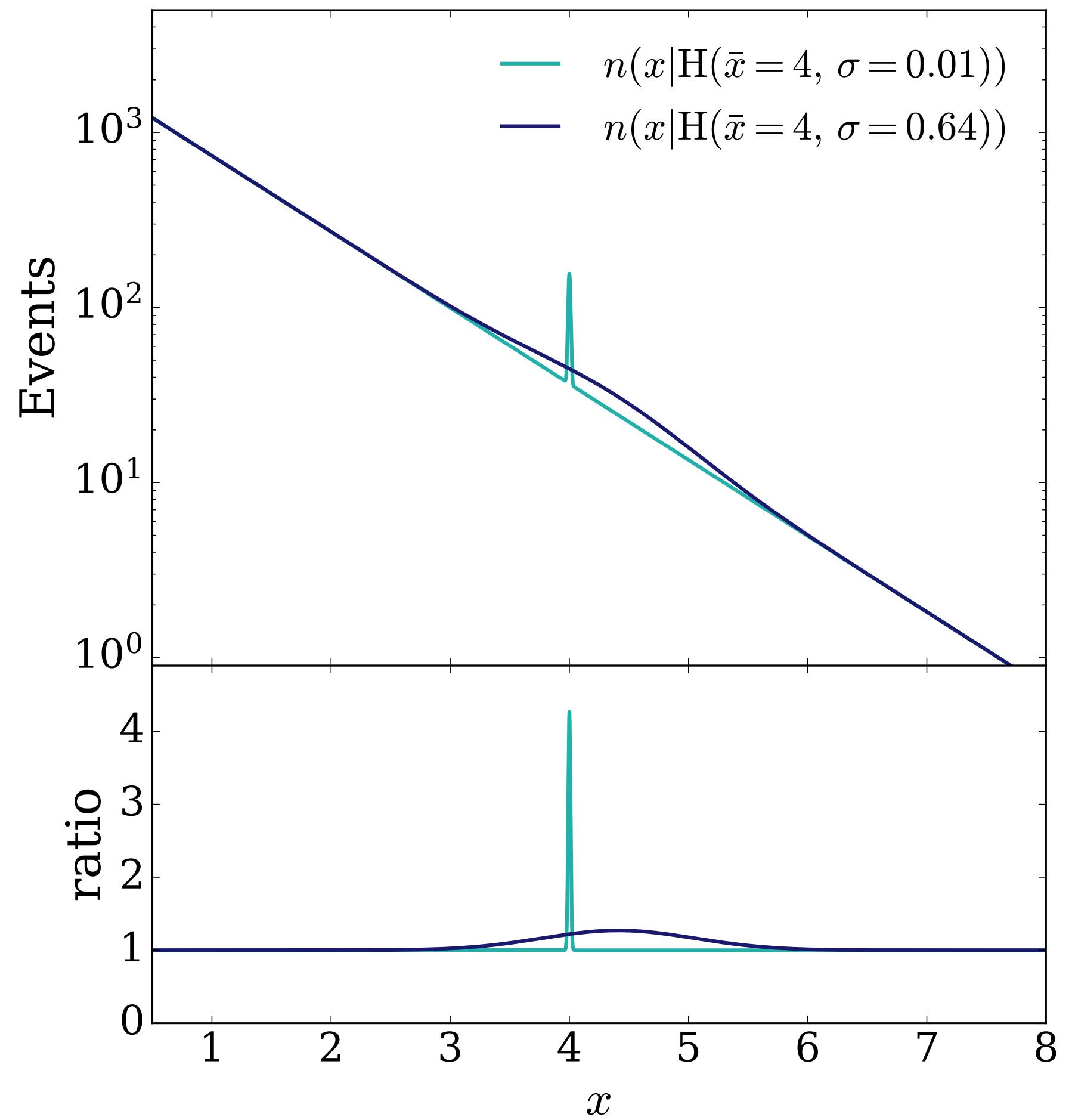
2) $N(B) = 16\,000$

- Narrow peak: $N(S) = 24 (Z_{\text{ideal}} = 4.8)$
- Broad peak: $N(S) = 104 (Z_{\text{ideal}} = 4.2)$

Aggregation over 8 batches:

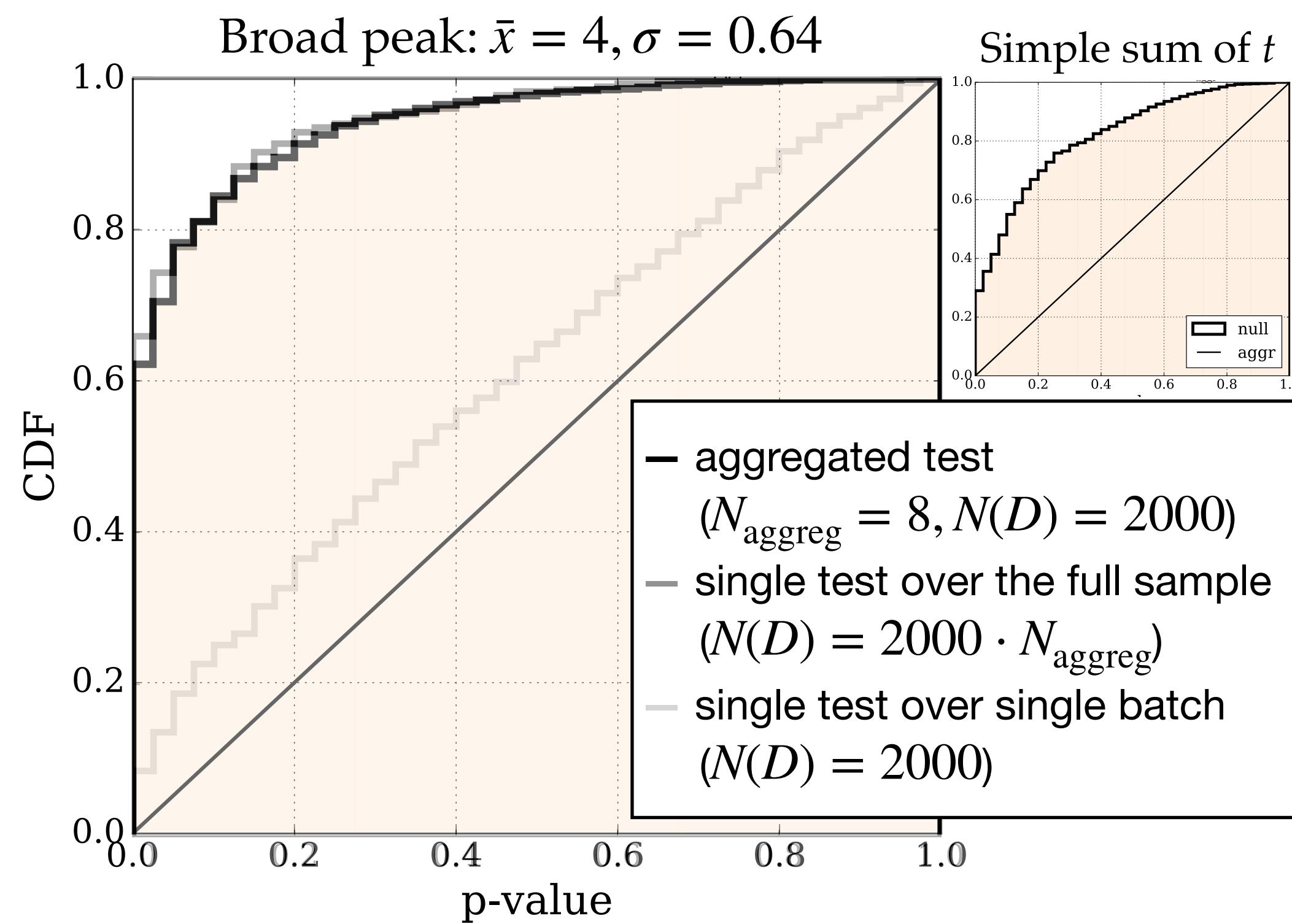
$N(B) = 2\,000, N_{\text{aggr}} = 8$

- $N_{\text{test}} = 1$
- $N_{\text{test}} = 8$

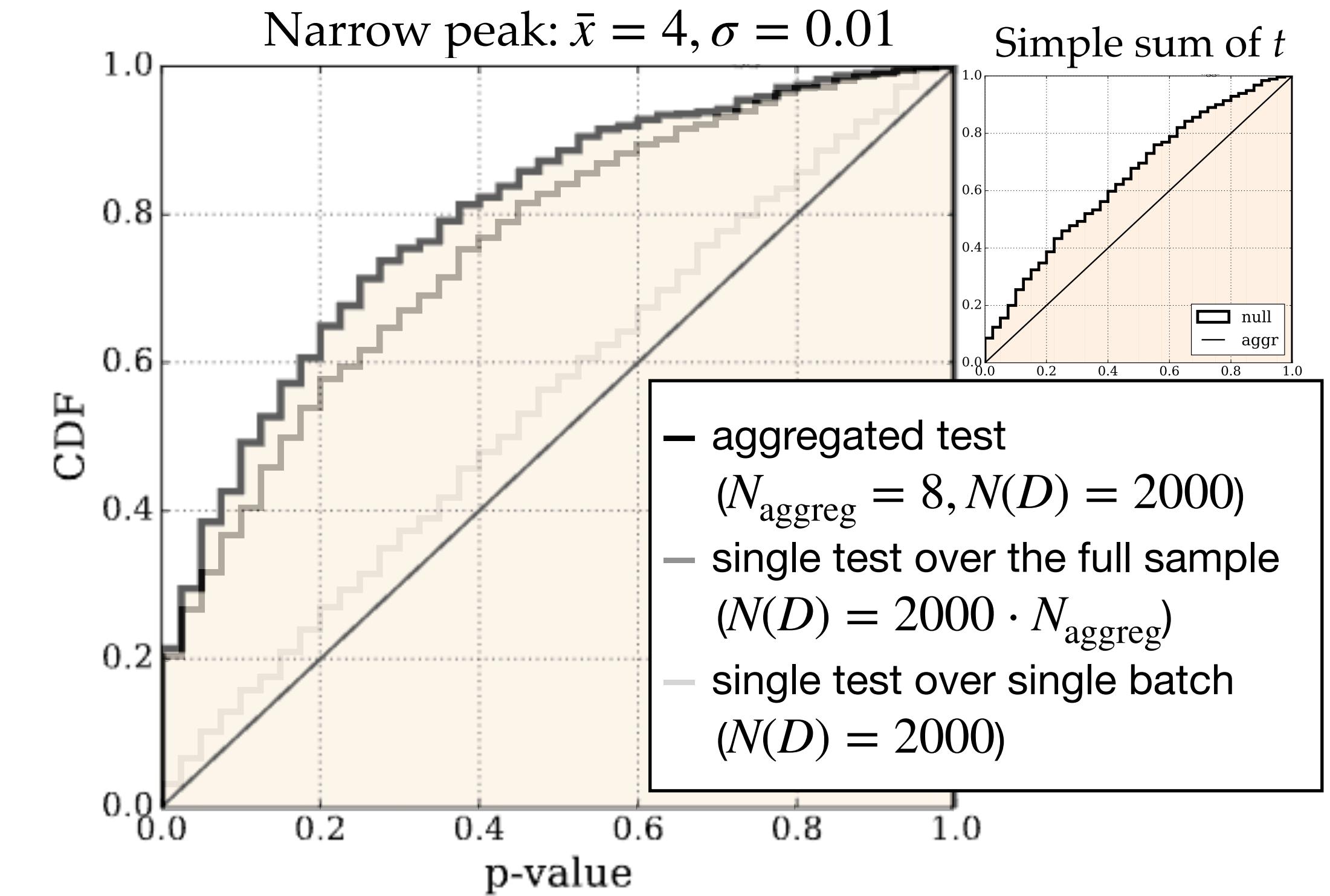


Combining NPLM over *multiple batches*

1D proof-of-concept



The proposed combination lead to performances that are comparable to the ones obtained with the full statistics!
Clear gain with respect to simple sum of tests



Physics benchmark	$\Pr(\text{p-value} < 0.001) (Z > 3) [\%]$		$\Pr(\text{p-value} < 0.02) (Z > 2) [\%]$	
	single train (N(R)=16000)	8 splits (N(R)=2000)	single train (N(R)=16000)	8 splits (N(R)=2000)
narrow resonance	8.0 ± 0.9	7.3 ± 0.9	19 ± 1	20 ± 1
broad resonance	33 ± 2	54 ± 2	64 ± 3	78 ± 3

Summary and next steps

We addressed two outstanding questions of agnostic goodness of fit with multiple testing:

- How to mitigate induced biases due to model selection
- How to exploit large statistics in a fast and efficient way

Relevant to:

- Quasi-online monitoring
- New Physics searches
- Data validation

Towards a resource *efficient*,
automatized, and *powerful* GoF tool

Ongoing efforts:

- Multiple testing in presence of *systematic uncertainties*
- Comparison with state-of-the-art GoF approaches

Backup slides

Inductive bias from model selection

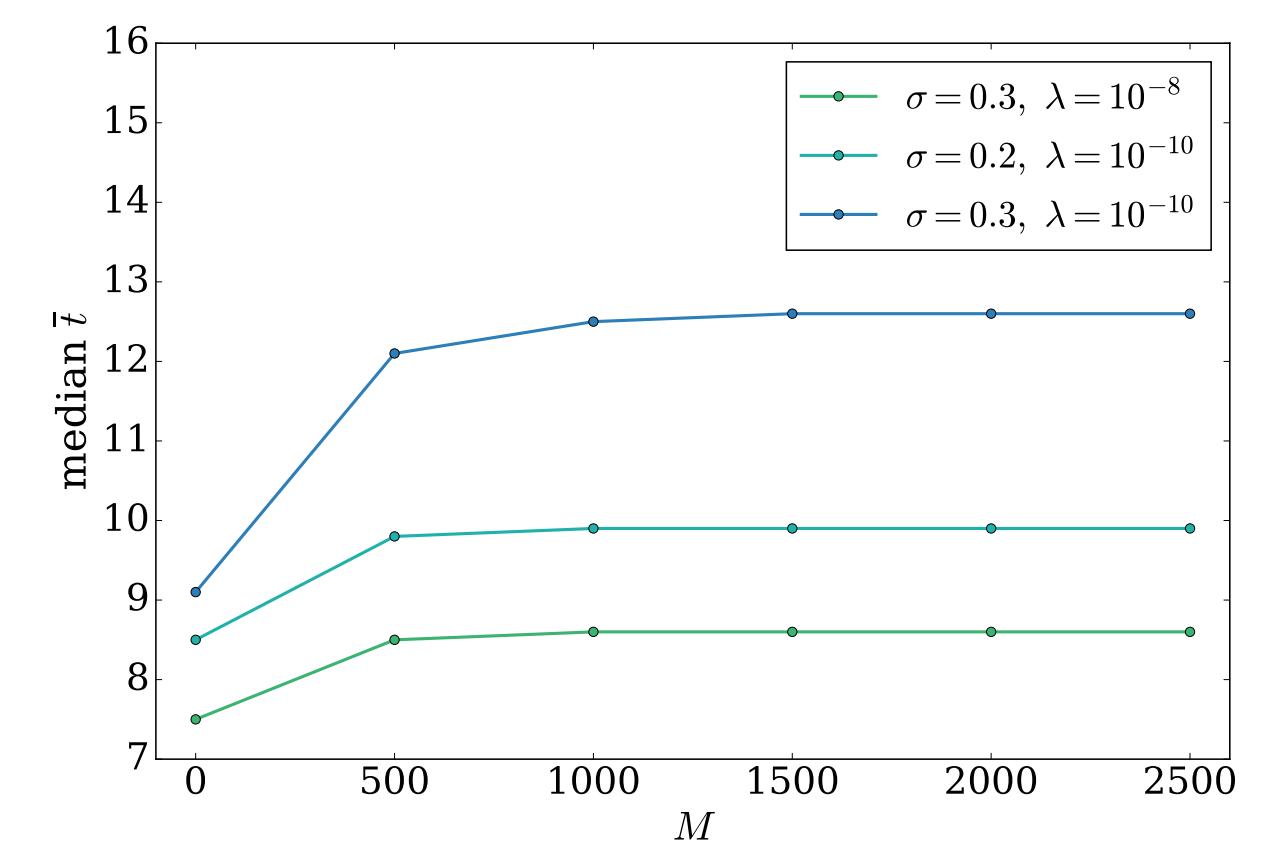
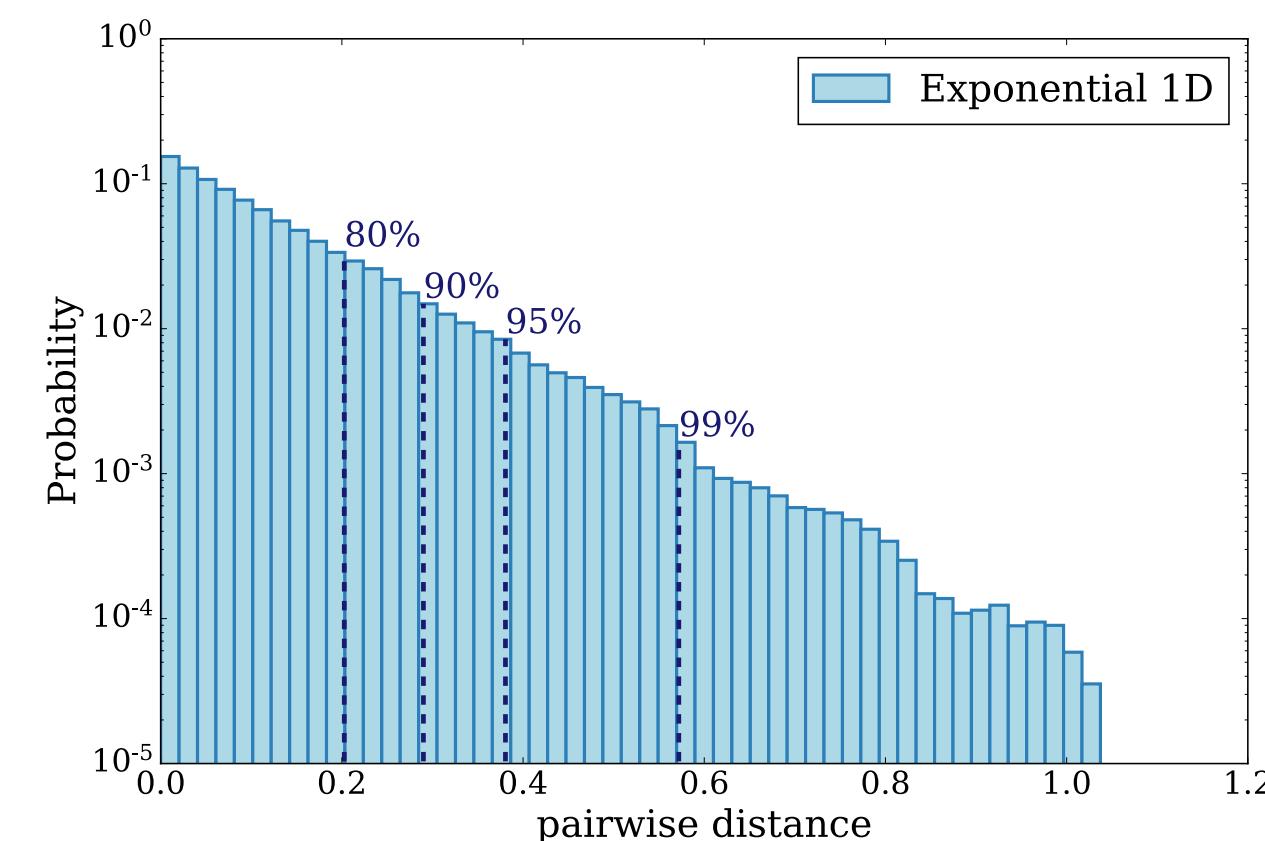
Kernel Methods model hyper-parameter choice

Asymptotic χ^2 behavior is observed for each choice of (M, σ, λ) , provided that $N_R \gg N_D$.

How to choose $(M, \sigma, \lambda)??$

Heuristics:

- **Number of centers M :** at least as large as \sqrt{N} to achieve statistically optimal bounds of the training convergence).
- **Gaussian width σ :** we select it as the 90th percentile of the pairwise distance between reference-distributed data points (after standardisation).
- **Regularisation parameter λ :** is kept as small as possible while keeping training stable



The hyper parameter choice impact the sensitivity to different signal patterns. How to mitigate this effect?

ML-based Neyman-Pearson GoF test

NPLM: Likelihood ratio from weighted Binary Cross Entropy

Test statistic (unbinned extended likelihood ratio)

$$\begin{aligned}\bar{t}(\mathcal{D}) &= 2 \max_{\mathbf{w}} \log \left[\frac{\mathcal{L}(\mathcal{D} | H_{\mathbf{w}})}{\mathcal{L}(\mathcal{D} | R_0)} \right] = 2 \max_{\mathbf{w}} \left\{ \log \left[\frac{e^{-N(\mathbf{v}) N_D V_{\mathcal{D}}}}{e^{-N(R)}} \prod_{i=1}^{N_D} \frac{n(x_i | \mathbf{w})}{n(x_i | R)} \right] \right\} \\ &= 2 \sum_{x \in \mathcal{D}} f_{\mathbf{w}}(x) - 2 \sum_{x \in \mathcal{R}} \frac{N(R)}{N_{\mathcal{R}}} \left[e^{f(x; \mathbf{w})} - 1 \right] \\ &\quad \text{(f(x; \widehat{\mathbf{w}}) = log } \left[\frac{n(x | H_{\widehat{\mathbf{w}}})}{n(x | R)} \right]\end{aligned}$$

Loss function

$$\bar{L} [f(x; \mathbf{w})] = - \sum_{x \in \mathcal{D}} \log \left[1 + e^{-f_{\mathbf{w}}(x)} \right] + \sum_{x \in \mathcal{R}} \frac{N(R)}{N_{\mathcal{R}}} \log \left[1 + e^{f_{\mathbf{w}}(x)} \right]$$

\mathbf{w} : trainable parameters on the NN model

D : data sample

R : reference sample (built according to the R_0 hypothesis); could be weighted (w)

Assumptions:

- $N_R \gg N_D$ the statistical fluctuations of the reference sample are negligible.
- the weights of the reference sample (w) are such that the reference sample is normalised to match the data sample luminosity

Aggregation of multiple tests

Preliminary results: 1D, 5D, 21D benchmarks

N(S)	7	18	13	10	90
\bar{x}_{NP}	4	4	4	6.4	1.6
σ_{NP}	0.01	0.16	0.64	0.16	0.16
$\sigma = 0.1$	0.008 ± 0.003	0.032 ± 0.006	0.002 ± 0.001	0.026 ± 0.005	0.30 ± 0.02
$\sigma = 0.3$	0.001 ± 0.001	0.056 ± 0.007	0.001 ± 0.001	0.14 ± 0.01	0.49 ± 0.02
$\sigma = 0.7$	0	0.059 ± 0.008	0.003 ± 0.002	0.21 ± 0.01	0.53 ± 0.02
$\sigma = 1.4$	0	0.045 ± 0.007	0.005 ± 0.002	0.19 ± 0.01	0.41 ± 0.02
$\sigma = 3.0$	0	0.020 ± 0.004	0.008 ± 0.003	0.11 ± 0.01	0.23 ± 0.02
aggregation	0.009 ± 0.003	0.11 ± 0.01	0.013 ± 0.004	0.27 ± 0.02	0.62 ± 0.02

Table 1: 1D experiments: probability of observing $Z \geq 3$

N(S)	1000	2500
$\sigma = 4.3$	0.003 ± 0.002	0.11 ± 0.01
$\sigma = 5.3$	0.006 ± 0.002	0.19 ± 0.01
$\sigma = 6.0$	0.007 ± 0.003	0.25 ± 0.02
$\sigma = 6.6$	0.007 ± 0.003	0.36 ± 0.02
$\sigma = 7.5$	0.008 ± 0.003	0.49 ± 0.02
aggregation	0.009 ± 0.003	0.49 ± 0.02

Table 3: HIGGS dataset: probability of observing $Z \geq 3$

test	Z' M = 180 GeV, W = 0.02 GeV	Z' M = 180 GeV, W = 2 GeV	Z' M = 200 GeV, W = 10 GeV	Z' M = 300 GeV, W = 15 GeV	Z' M = 600 GeV, W = 30 GeV	EFT $c_w = 10^{-6}$
$\sigma = 0.57$	0.04 ± 0.02	0.04 ± 0.02	0.06 ± 0.02	0.01 ± 0.01	0.01 ± 0.01	0.005 ± 0.005
$\sigma = 1.19$	0.03 ± 0.02	0.05 ± 0.02	0.06 ± 0.02	0.01 ± 0.01	0.02 ± 0.01	0.02 ± 0.007
$\sigma = 1.79$	0.03 ± 0.02	0.04 ± 0.02	0.05 ± 0.01	0.05 ± 0.02	0.04 ± 0.02	0.03 ± 0.01
$\sigma = 2.49$	0.01 ± 0.01	0	0.02 ± 0.01	0.02 ± 0.01	0.02 ± 0.01	0.005 ± 0.005
$\sigma = 3.57$	0.01 ± 0.01	0.02 ± 0.01	0.02 ± 0.01	0.02 ± 0.01	0.02 ± 0.01	0.02 ± 0.01
aggregation	0.08 ± 0.03	0.11 ± 0.03	0.12 ± 0.02	0.08 ± 0.03	0.08 ± 0.02	0.05 ± 0.02

Table 2: 5D MUMU experiments: probability of observing $Z \geq 3$