

SUSY 2024

Theory meets Experiment

Madrid, 10 – 14 June 2024

Modular tool for bayesian analysis with Machine Learning in DM Direct Detection

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arXiv: 2406.XXXXX



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Teórica
UAM-CSIC



Universidad Autónoma
de Madrid

Outline

Reconstruct the DM parameters with DD experiments

Combine different data in a fast and simple way using
Machine Learning.

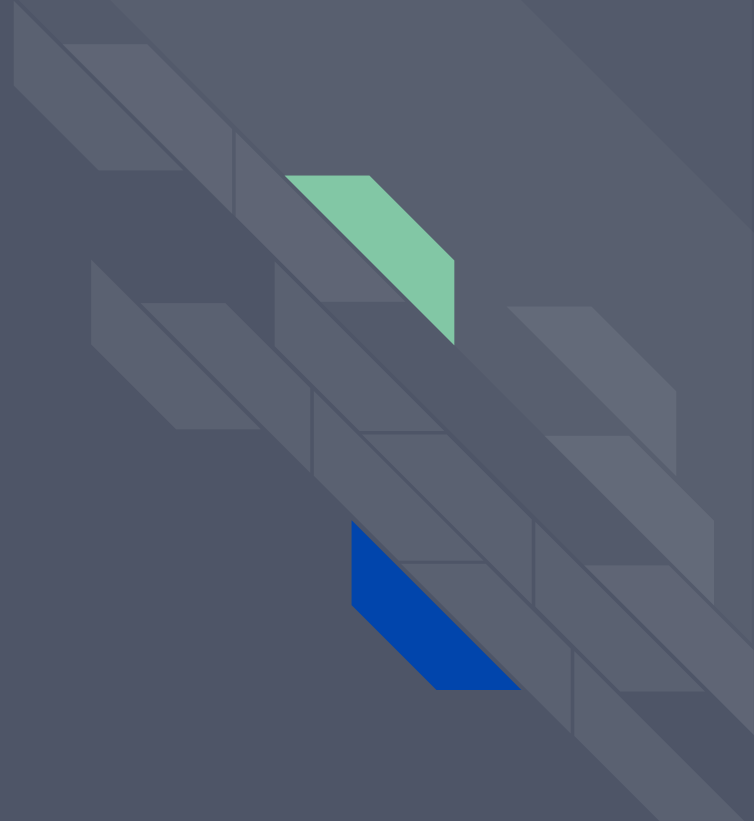
Outline

Reconstruct the DM parameters with DD experiments

Combine different data in a fast and simple way using
Machine Learning.

- Bayesian analysis with ML analysis to obtain posteriors
- Data sample generation:
 - DM-nucleon interaction with NR-EFT
 - Simulate the expected signal
- Parameter space that can be reconstructed

Bayesian Analysis



Bayesian Analysis

Bayes theorem:

Posterior probability of the parameters Θ of interest given the data X

Probability of the data X given the parameters Θ (likelihood)

Prior probability of the parameters Θ

$$P(\Theta|X) = \frac{P(X|\Theta)P(\Theta)}{P(X)}$$

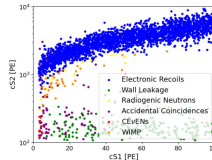
Parameters

Data

e.g. $(\sigma_i^{SI}, \theta_i, m_{DM})$

Number of events

or



Probability of the data X also call evidence

Bayesian Analysis

Bayes theorem:

Traditional methods:
need to assume a
likelihood function

Posterior probability of the
parameters Θ of interest
given the data X

Probability of the data X
given the parameters Θ
(likelihood)

Prior probability of
the parameters Θ

$$P(\Theta|X) = \frac{P(X|\Theta)P(\Theta)}{P(X)}$$

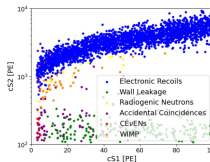
Parameters

Data

e.g. $(\sigma_i^{SI}, \theta_i, m_{DM})$

Number of
events

or



Probability of the data X
also call evidence

Bayesian Analysis

Likelihood-to-evidence ratio:

Conditional probabilities

$$P(\mathbf{X}|\Theta) = \frac{P(\mathbf{X}, \Theta)}{P(\Theta)}$$

$$P(\Theta|\mathbf{X}) = \frac{P(\mathbf{X}, \Theta)}{P(\mathbf{X})}$$

$$r(\mathbf{X}, \Theta) := \frac{P(\mathbf{X}|\Theta)}{P(\mathbf{X})} = \frac{P(\Theta|\mathbf{X})}{P(\Theta)} = \frac{P(\mathbf{X}, \Theta)}{P(\Theta)P(\mathbf{X})}$$

Conditional probabilities

$$P(\mathbf{X}|\Theta) = \frac{P(\mathbf{X}, \Theta)}{P(\Theta)}$$

$$P(\Theta|\mathbf{X}) = \frac{P(\mathbf{X}, \Theta)}{P(\mathbf{X})}$$

Bayesian Analysis

Likelihood-to-evidence ratio:

$$r(\mathbf{X}, \Theta) := \frac{P(\mathbf{X}|\Theta)}{P(\mathbf{X})} = \frac{P(\Theta|\mathbf{X})}{P(\Theta)} = \frac{P(\mathbf{X}, \Theta)}{P(\Theta)P(\mathbf{X})}$$

compute the
posterior



$$P(\Theta|\mathbf{X}) = r(\mathbf{X}, \Theta) P(\Theta)$$

Bayesian Analysis with SWYFT

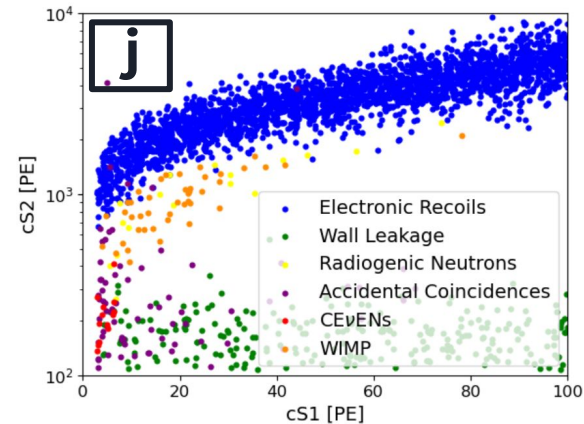
SWYFT → Sampling-based inference tool that estimates *likelihood-to-evidence ratio* with ML algorithms to obtain marginal and joint posteriors

Parameters

$(\sigma^j, \theta^j, m_{\text{DM}}^j)$



Data

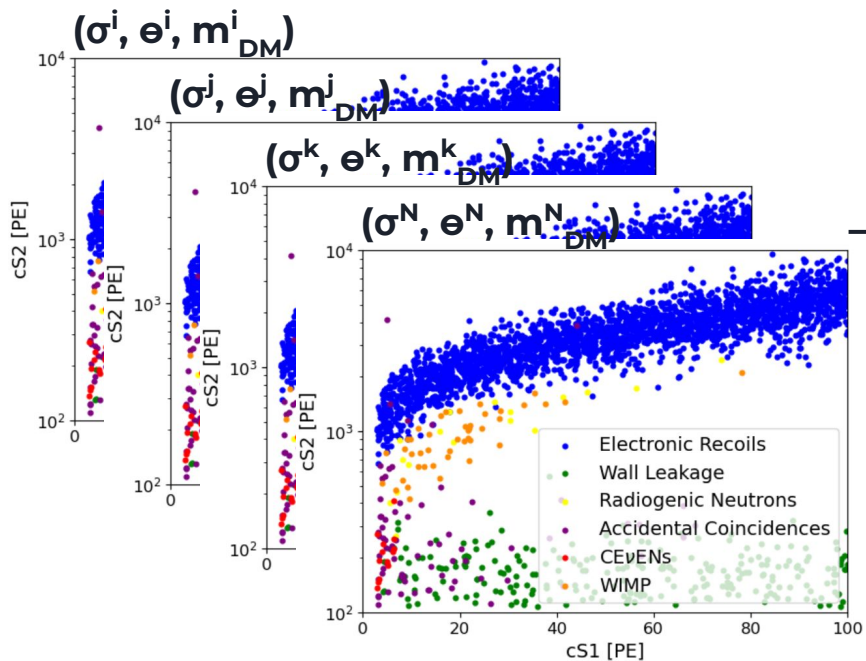


You can consider uncertainties:

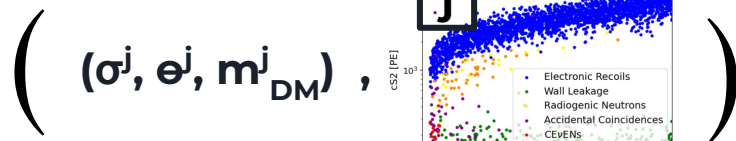
- DM halo model,
- Nuclear response functions,
- Detector response ...

Bayesian Analysis with SWYFT

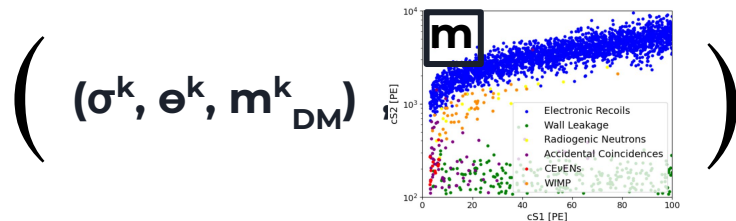
SWYFT → Sampling-based inference tool that estimates *likelihood-to-evidence ratio* with ML algorithms to obtain marginal and joint posteriors



Matching (parameter, data) → **label 1**

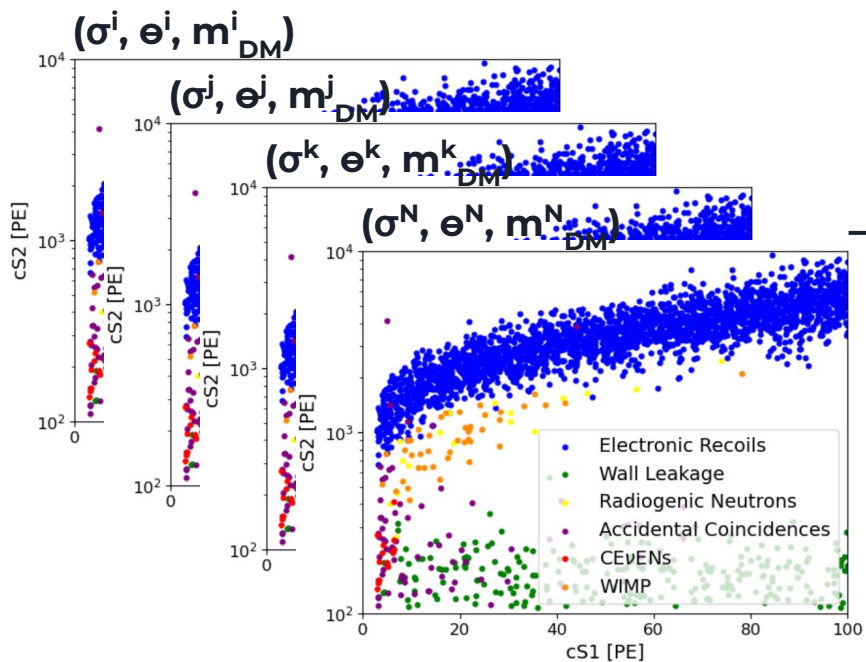


Scrambled (parameter, data) → **label 0**



Bayesian Analysis with SWYFT

SWYFT → Sampling-based inference tool that estimates *likelihood-to-evidence ratio* with ML algorithms to obtain marginal and joint posteriors



Matching (parameter, data) → **label 1**

Sampled from $P(\mathbf{X}, \Theta)$

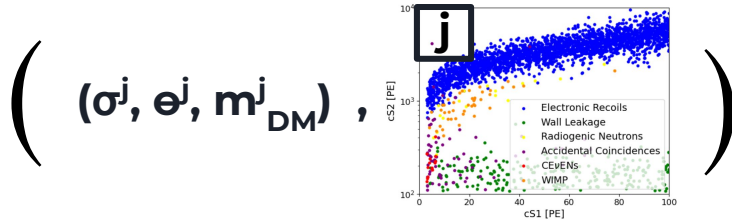
Scrambled (parameter, data) → **label 0**

Sampled from $P(\Theta)P(\mathbf{X})$

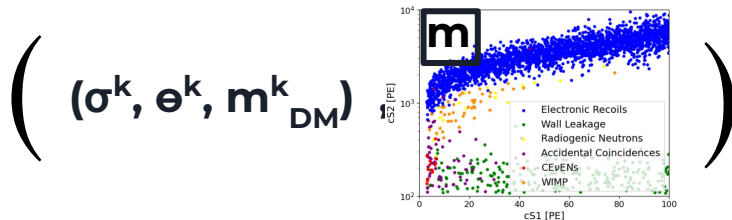
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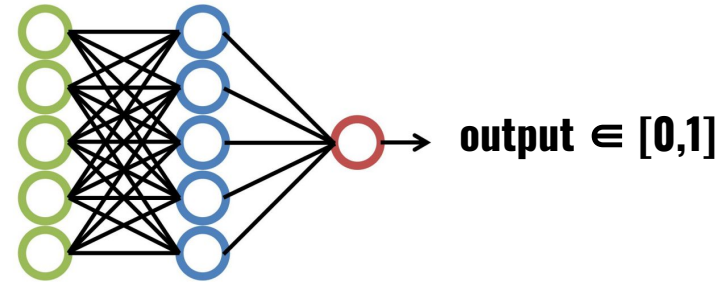
Matching (parameter, data) → **label 1**



Scrambled (parameter, data) → **label 0**



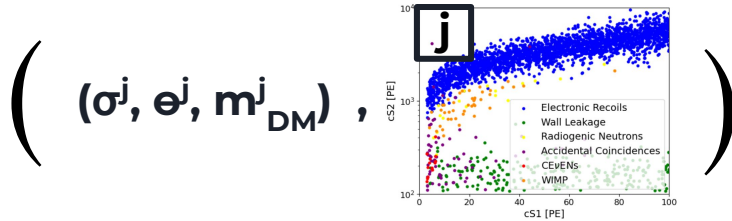
Binary classifier (DNN, CNN, ...)



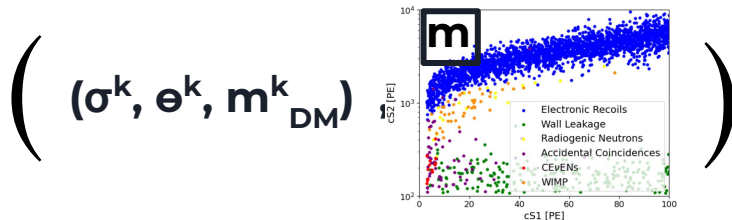
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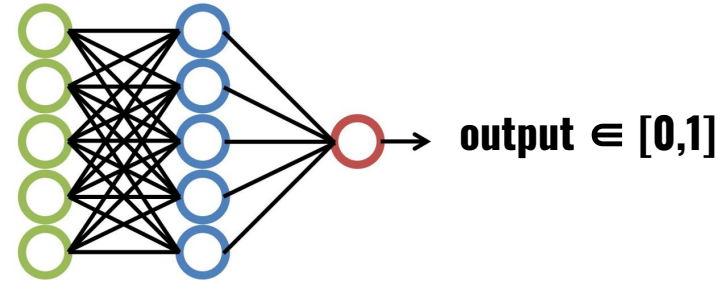
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Scrambled (parameter, data) → **label 0**



Binary classifier (DNN, CNN, ...)

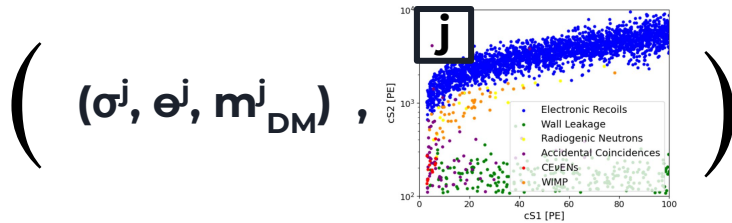


label 1
output = $P(k = 1 | \mathbf{X}, \Theta)$ → $P(\mathbf{X}, \Theta)$

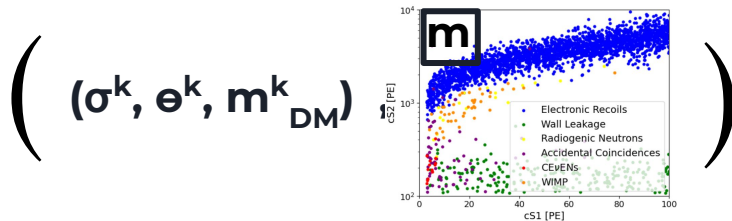
Bayesian Analysis with SWYFT

SWYFT → Sampling-based inference tool that estimates *likelihood-to-evidence ratio* with ML algorithms to obtain marginal and joint posteriors

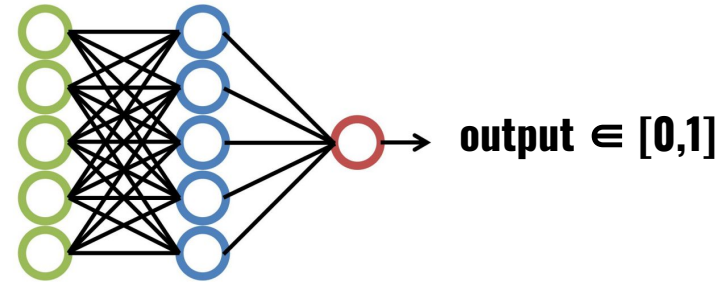
Matching (parameter, data) → **label 1**



Scrambled (parameter, data) → **label 0**



Binary classifier (DNN, CNN, ...)



label 1
output = $P(k = 1 | \mathbf{X}, \Theta)$ → $P(\mathbf{X}, \Theta)$

label 0
1 - output = $P(k = 0 | \mathbf{X}, \Theta)$ → $P(\Theta)P(\mathbf{X})$

Bayesian Analysis with SWYFT

SWYFT → Sampling-based inference tool that estimates *likelihood-to-evidence ratio* with ML algorithms to obtain marginal and joint posteriors

$$r(\mathbf{X}, \Theta) = \frac{\text{output}}{1 - \text{output}}$$

compute the
posterior

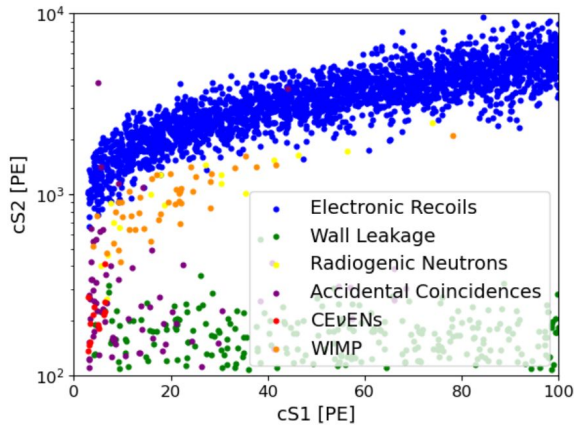


$$P(\Theta | \mathbf{X}) = r(\mathbf{X}, \Theta) P(\Theta)$$

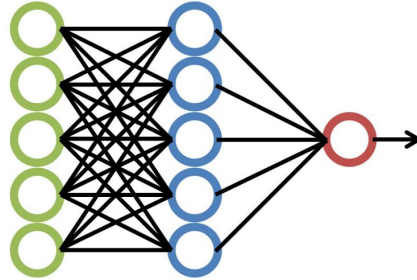
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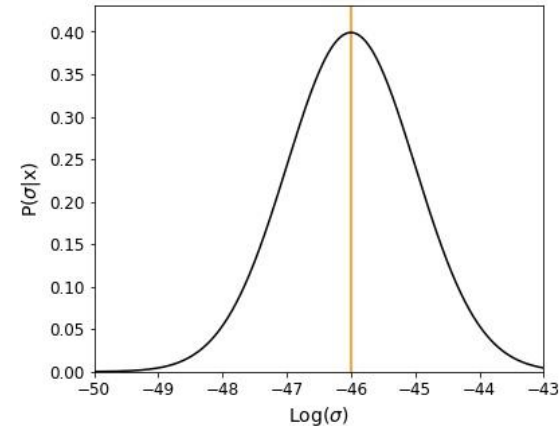
New data sample x^{new}



Trained binary classifier



Posterior $P(\sigma|x^{\text{new}})$



For another data sample → we do **not** need to train everything again, use the same classifier

Combined Bayesian Analysis

Posterior using an experimental dataset \mathbf{X}_1 (e.g. images)

$$P(\Theta|\mathbf{X}_1) = \frac{P(\mathbf{X}_1|\Theta)P(\Theta)}{P(\mathbf{X}_1)} = r(\mathbf{X}_1, \Theta)P(\Theta)$$

non-informative
prior



Combined Bayesian Analysis

Posterior using an experimental dataset \mathbf{X}_1 (e.g. images)

$$P(\Theta|\mathbf{X}_1) = \frac{P(\mathbf{X}_1|\Theta)P(\Theta)}{P(\mathbf{X}_1)} = r(\mathbf{X}_1, \Theta)P(\Theta)$$

non-informative
prior



Posterior using an experimental dataset \mathbf{X}_2 (e.g. spectra)

$$P(\Theta|\mathbf{X}_2) = \frac{P(\mathbf{X}_2|\Theta)P(\Theta)}{P(\mathbf{X}_2)} = r(\mathbf{X}_2, \Theta)P(\Theta) \rightarrow r(\mathbf{X}_2, \Theta)P(\Theta|\mathbf{X}_1)$$

use the previous
posterior as prior



Combined Bayesian Analysis

Posterior using an experimental dataset \mathbf{X}_1 (e.g. images)

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non-informative
prior



Posterior using an experimental dataset \mathbf{X}_2 (e.g. spectra)

$$P(\Theta|\mathbf{X}_2) = \frac{P(\mathbf{X}_2|\Theta)P(\Theta)}{P(\mathbf{X}_2)} = r(\mathbf{X}_2, \Theta)P(\Theta) \rightarrow r(\mathbf{X}_2, \Theta)P(\Theta|\mathbf{X}_1)$$

use the previous
posterior as prior



Product of likelihood-to-evidence ratios

$$P(\Theta|\mathbf{X}_2) = r(\mathbf{X}_2, \Theta) r(\mathbf{X}_1, \Theta) P(\Theta)$$

Combined Bayesian Analysis

Product of likelihood-to-evidence ratios

$$P(\Theta | \mathbf{X}_n) = \prod_{i=1}^n r(\mathbf{X}_i, \Theta) P(\Theta)$$

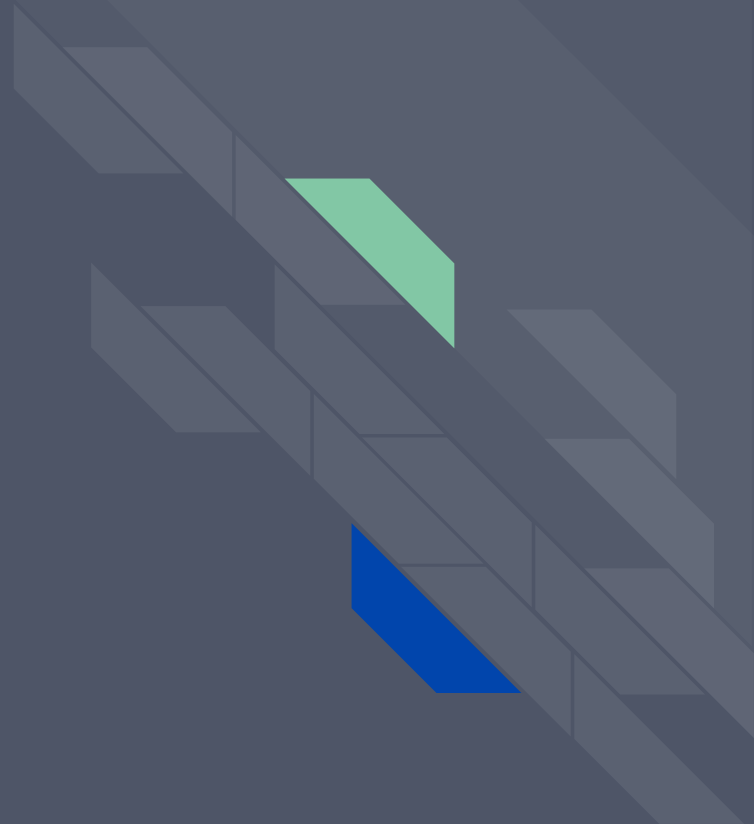


each one obtained training SWYFT
with only the \mathbf{X}_i dataset

To include new data you do not need to re-train everything. Train with your new data and include $r(\mathbf{X}_{\text{new}}, \Theta)$

To remove a dataset do not include $r(\mathbf{X}_{\text{old}}, \Theta)$

Data sample generation



DM-nucleon effective field theory (NR-EFT)

For DM particles with spin up to $\frac{1}{2}$, the effective DM-nucleon scattering interaction Lagrangian

$$\mathcal{L}_{\text{EFT}} = \sum_{\tau} \sum_i c_i^{\tau} \mathcal{O}_i \bar{\chi} \chi \bar{\tau} \tau$$

i=14 possible interactions

\mathcal{O}_1 : spin-independent (SI)

\mathcal{O}_4 : spin-dependent (SD)


Change to polar coordinates:

$$c_i^0 = \frac{1}{2} (c_i^p + c_i^n) = A_i \sin(\theta_i)$$

$$c_i^1 = \frac{1}{2} (c_i^p - c_i^n) = A_i \cos(\theta_i)$$

isospin basis
 c^0 : isoscalar
 c^1 : isovector

nucleon basis
 c^p : proton
 c^n : neutron


$$\sigma_i \propto A_i^2$$

usually shown assuming isoscalar interactions

$$c^p = c^n \quad c^0 = 1 \text{ and } c^1 = 0$$

DM-nucleon effective field theory (NR-EFT)

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nucleon basis
 c^p : proton
 c^n : neutron



$$\sigma_i \propto A_i^2$$

For each operator

2 parameters:

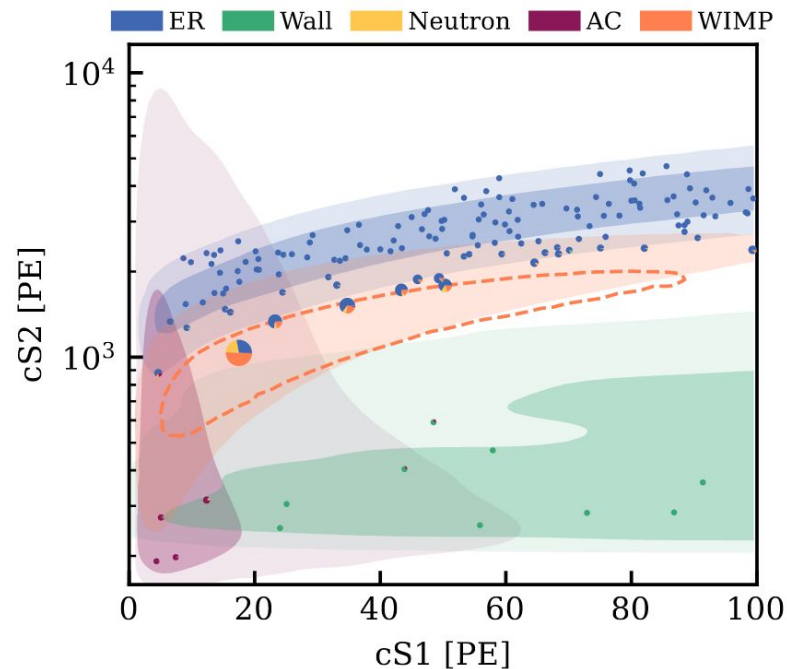
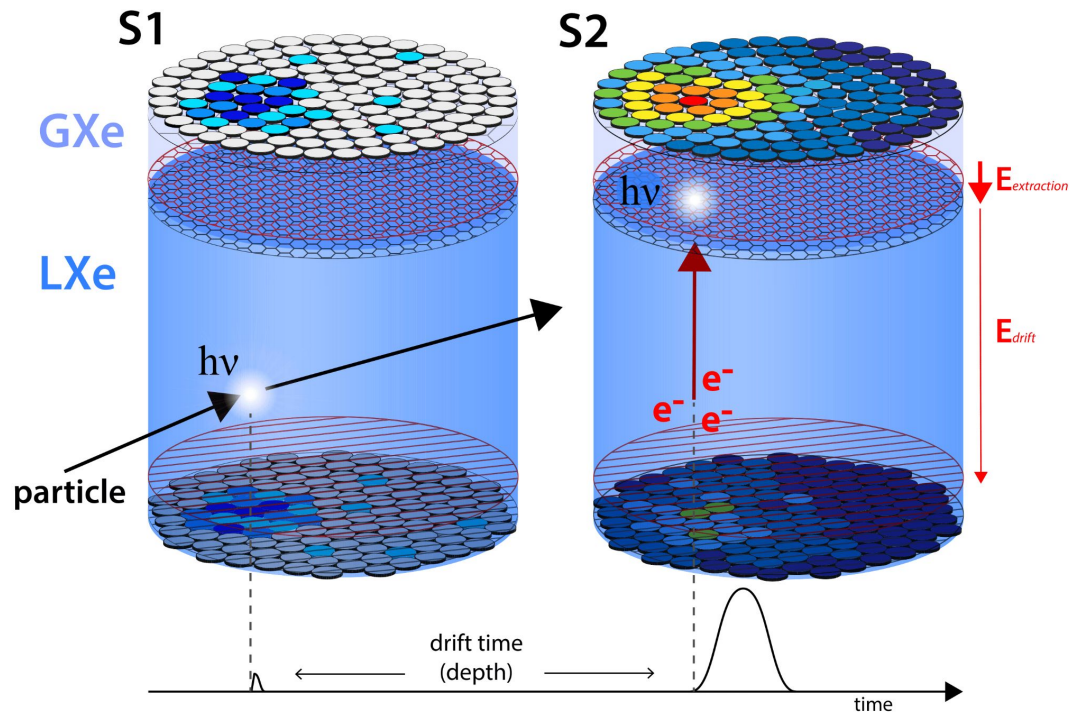
- amplitude (cross-section)
- phase

+ DM mass

$(\sigma_i, \theta_i, m_{\text{DM}})$

DM signal

XENONnT

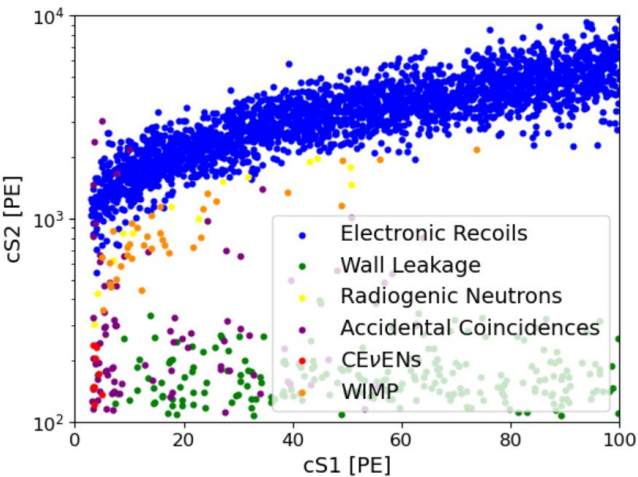


Data Representation:

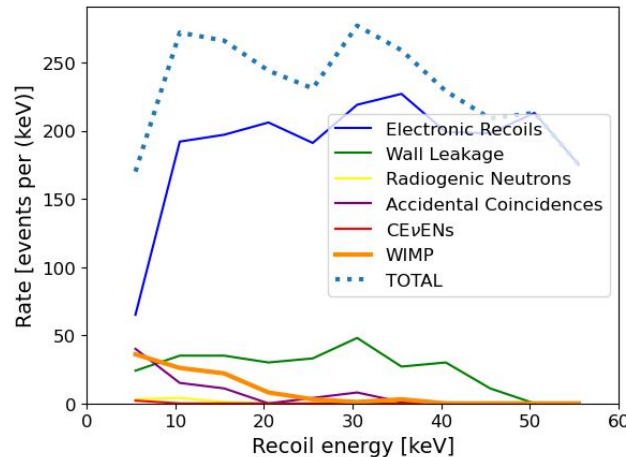
XENONnT 20ty simulator

We specify background and signal characteristics

cS1 vs cS2 plane



differential rate



number of events

	name	pseudo_exp_events
0	er	2459
1	radiogenics	17
2	ac	71
3	wall	246
4	WIMP	43
5	CEVNS-SM	13

WIMPs: differential rate compute with WimPyDD for a particular operator, (σ_i, e_i, m_{DM}) , and standard DM halo model

O1: $\sigma=10^{-47}\text{cm}^2$
 $\theta=\pi/2$ ($c^b=c^n$)
 $m=50\text{GeV}$

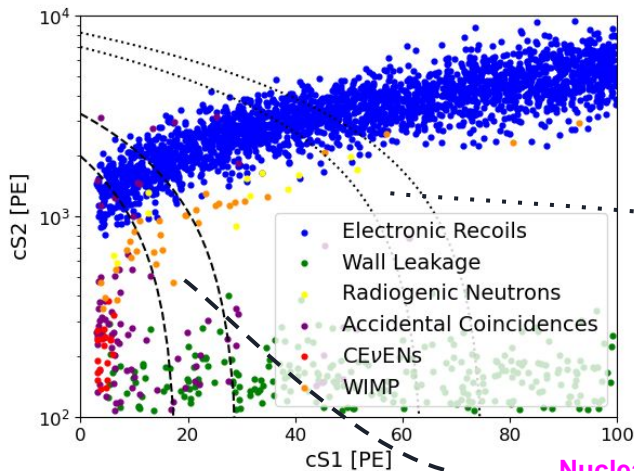
CEvENS: differential rate compute with SNUDD

Data Representation:

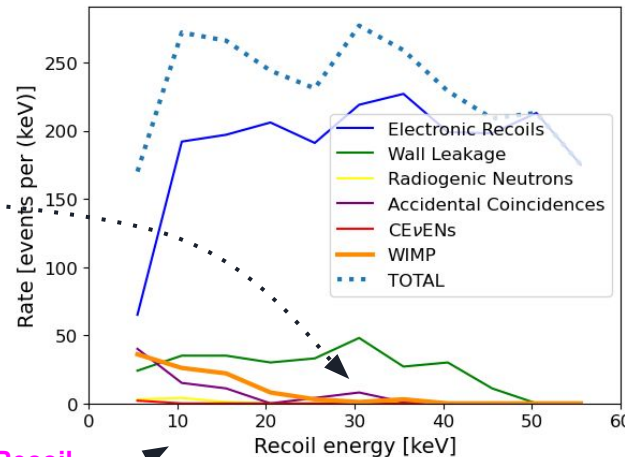
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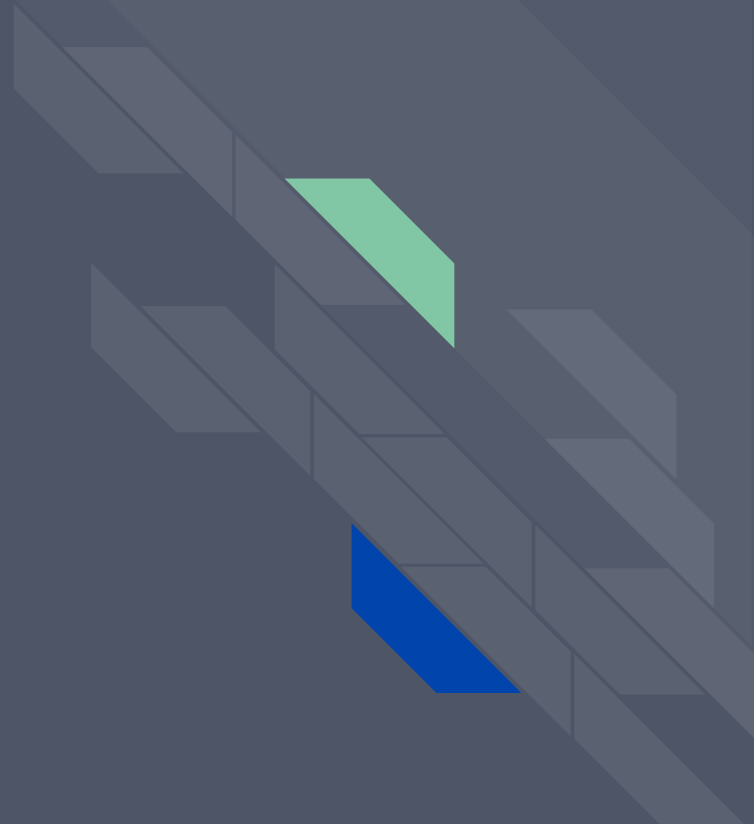
Nuclear Recoil isoenergy curves

WIMPs: differential rate compute with WimPyDD for a particular operator, (σ_i, e_i, m_{DM}) , and standard DM halo model

O1: $\sigma = 10^{-47} \text{cm}^2$
 $\theta = \pi/2$ ($c^b = c^n$)
 $m = 50 \text{GeV}$

CEvENS: differential rate compute with SNUDD

Results



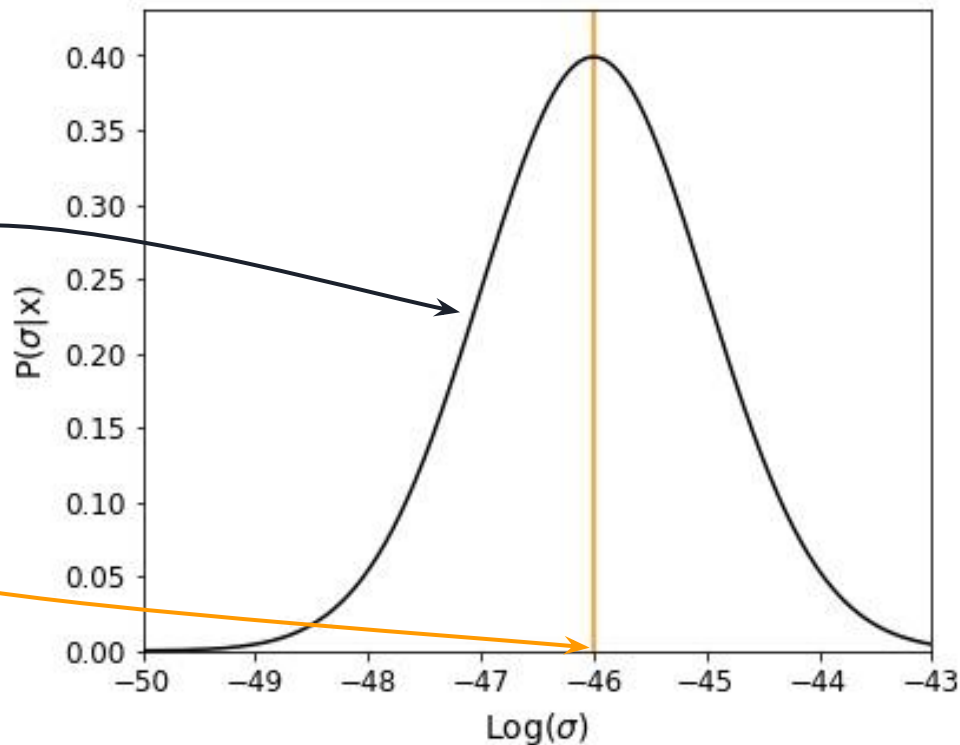
Posterior

Compute the posterior for a new pseudo experiment

For example:

$$P(\sigma|x)$$

$$\sigma^{\text{true}} = 10^{-46} \text{cm}^2$$



**this is a gaussian as
an example, not the
actual posterior!**

Posterior

Compute the posterior for a new pseudo experiment

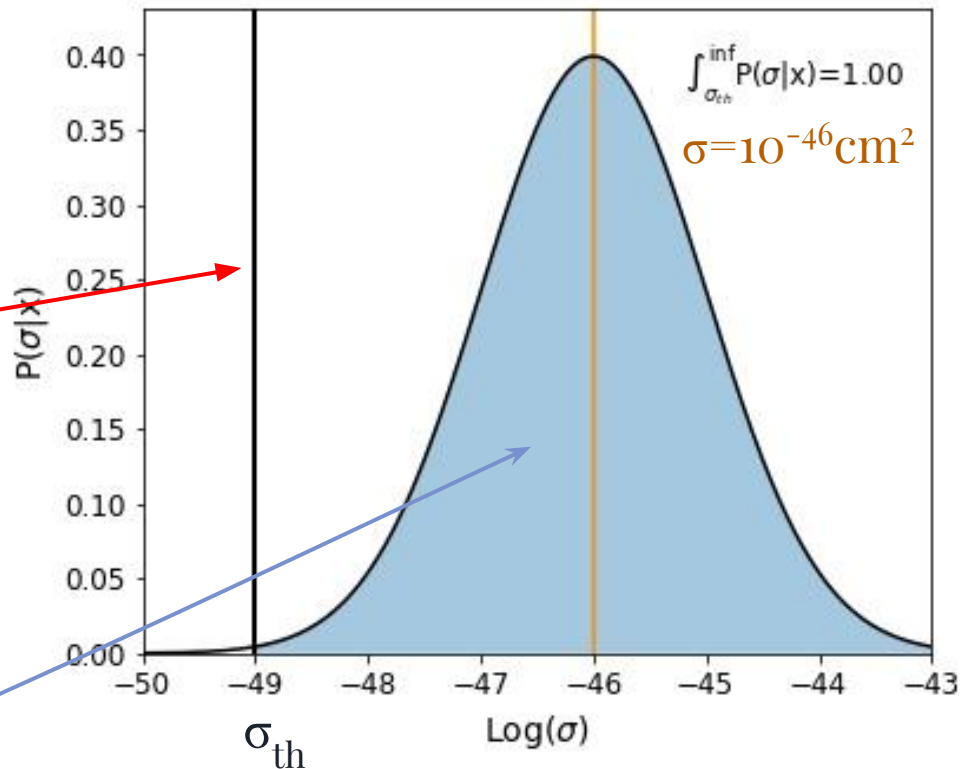
We define a σ_{th} threshold:

$\sigma_{th} = 10^{-49} \text{cm}^2 \rightarrow$ **NO SIGNAL!**

Then, we can **reconstruct** σ if:

$$\int_{\sigma_{th}}^{\infty} P(\sigma|x) > 0.90$$

this is a gaussian as an example, not the actual posterior!



Posterior

Compute the posterior for a new pseudo experiment

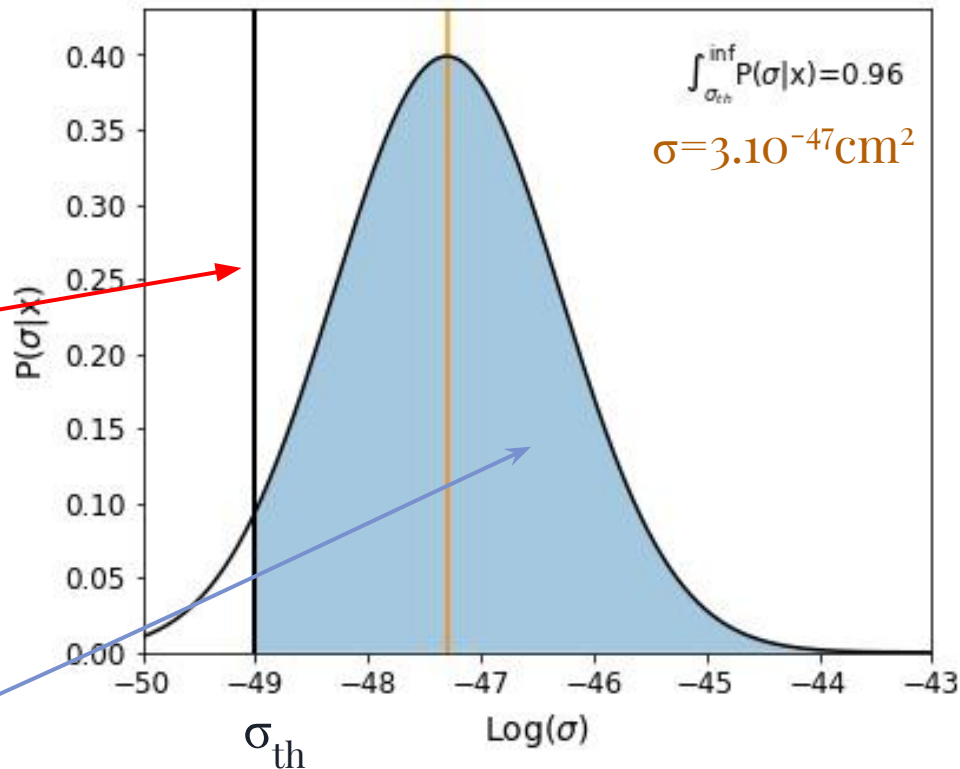
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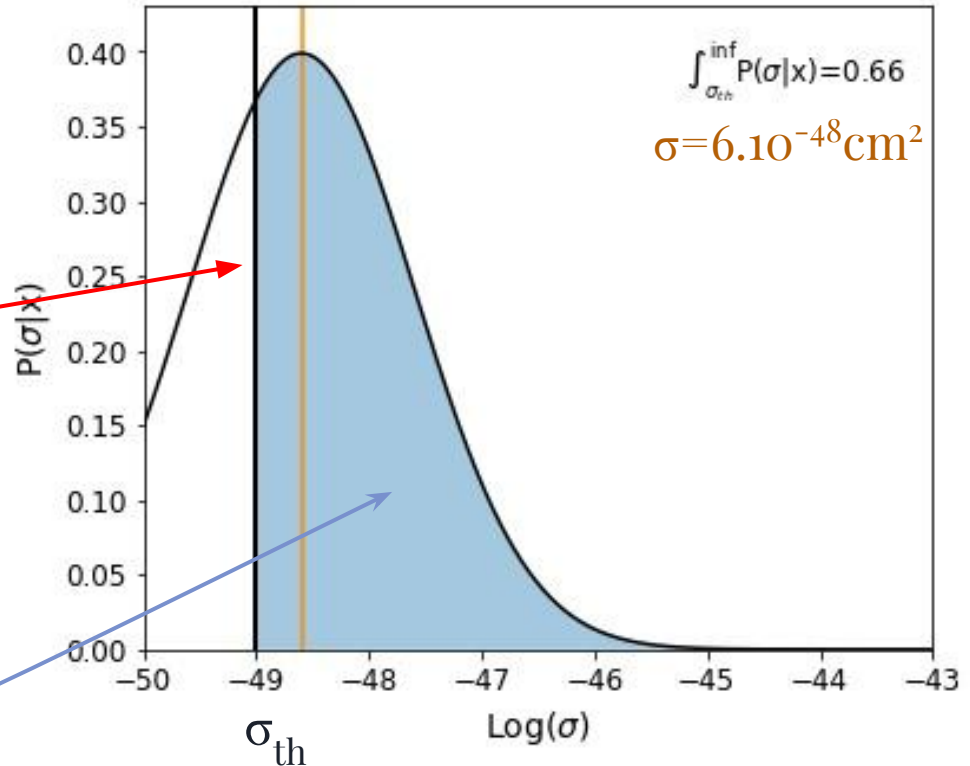
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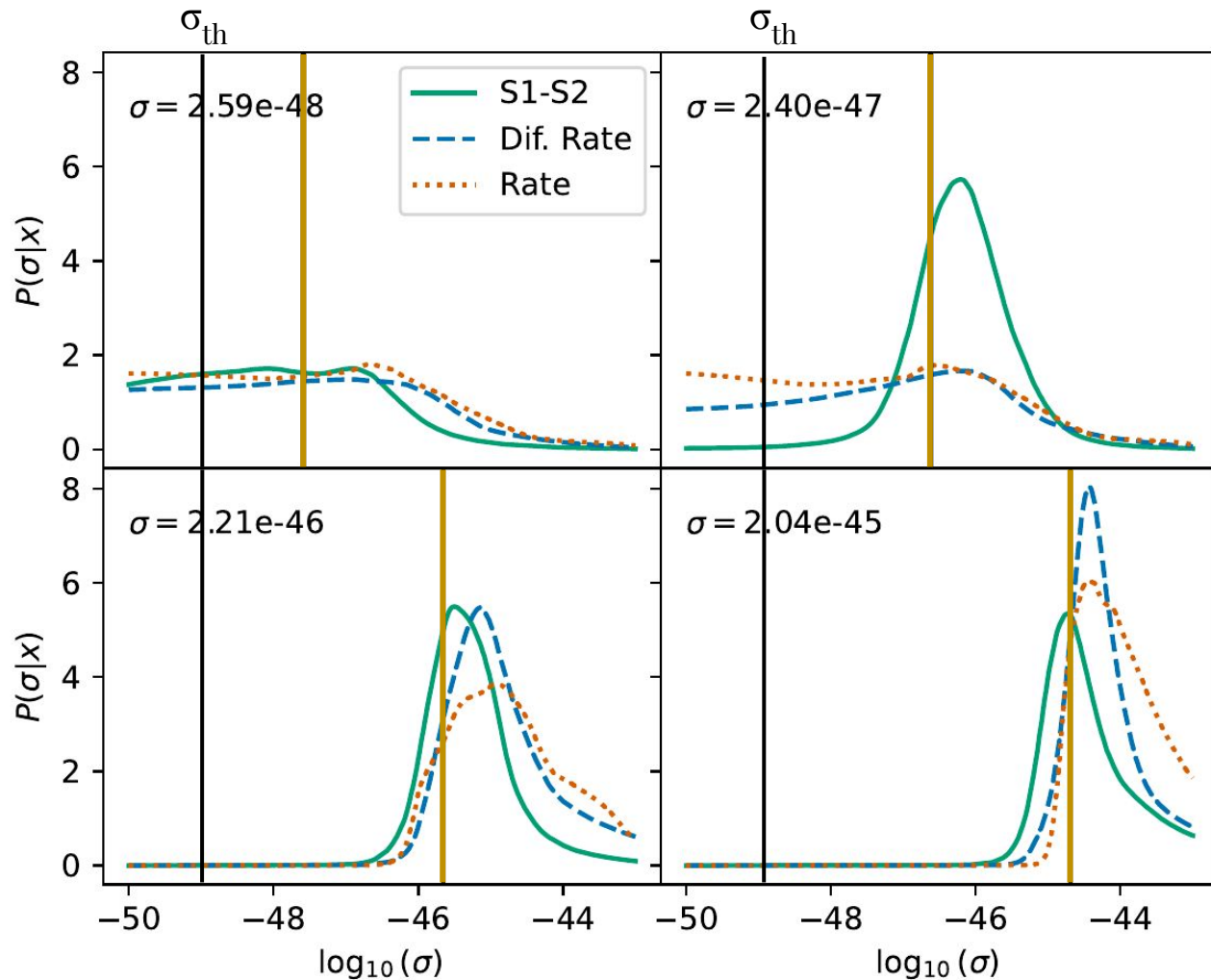
Results

Data:

cS1 vs cS2 plane
differential rate
total number of
events

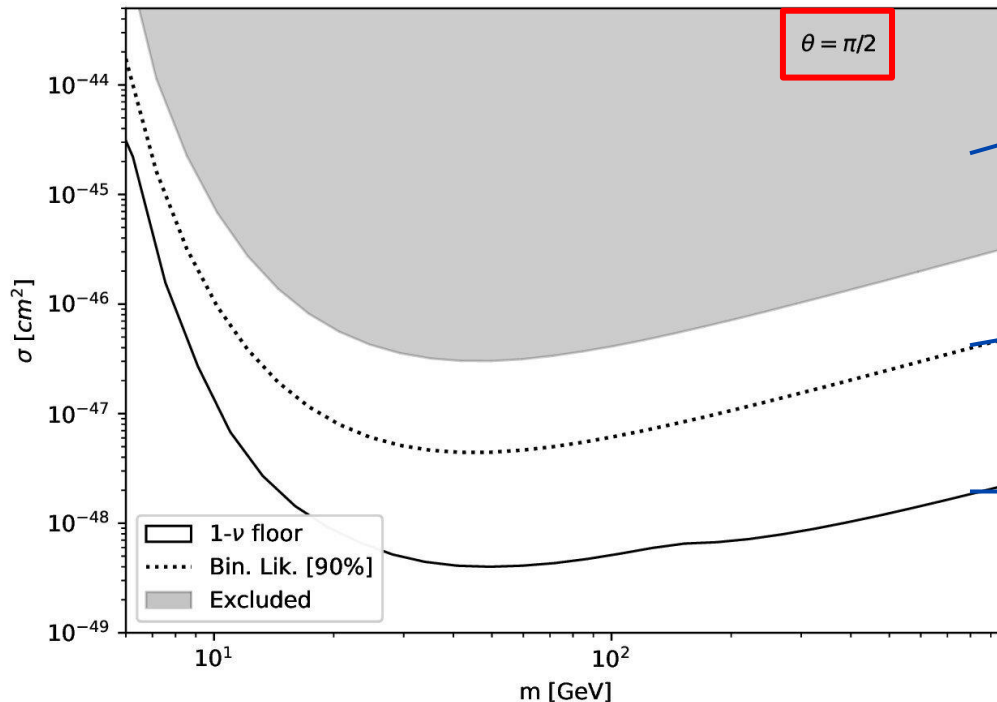
Posteriors for:

O_1 (SI)
 $m_{DM} \approx 100\text{GeV} \rightarrow$ **fixed**
 $\theta = \pi/2 \rightarrow$ **fixed**



Results: parameter reconstruction

For O_1 with fixed $\theta = \pi/2$



Exclusion:

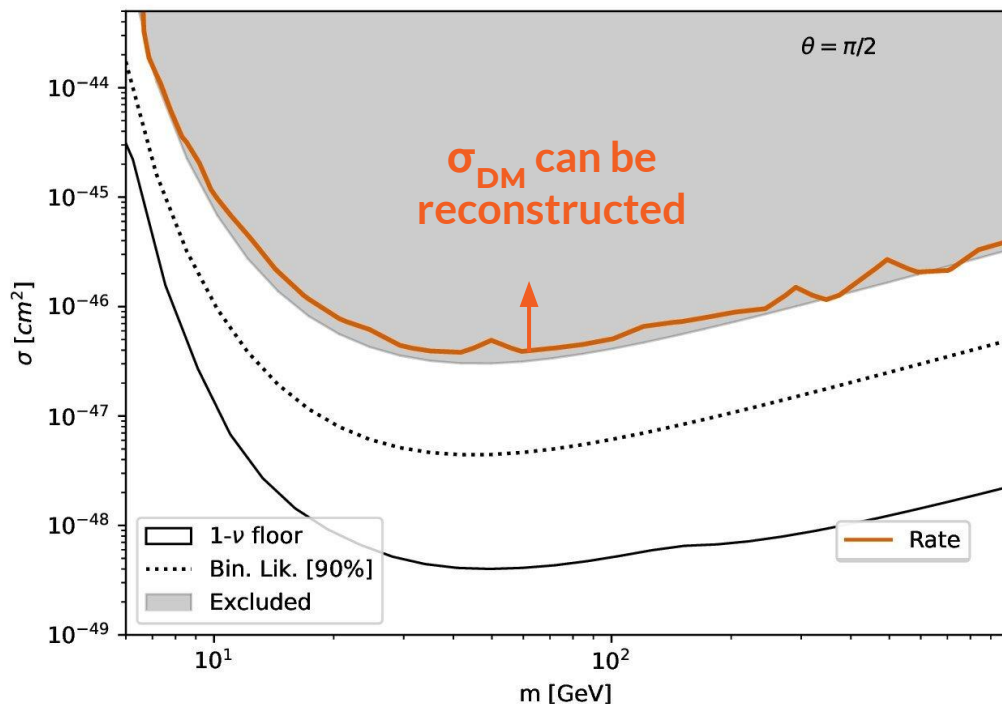
XENONnT 1 ton year
expected **exclusion** limit
(90% C.L.)

XENONnT 20 ton year
projected expected
exclusion limit (90% C.L.)

1ν neutrino floor

Results: parameter reconstruction

For O_1 with fixed $\theta = \pi/2$



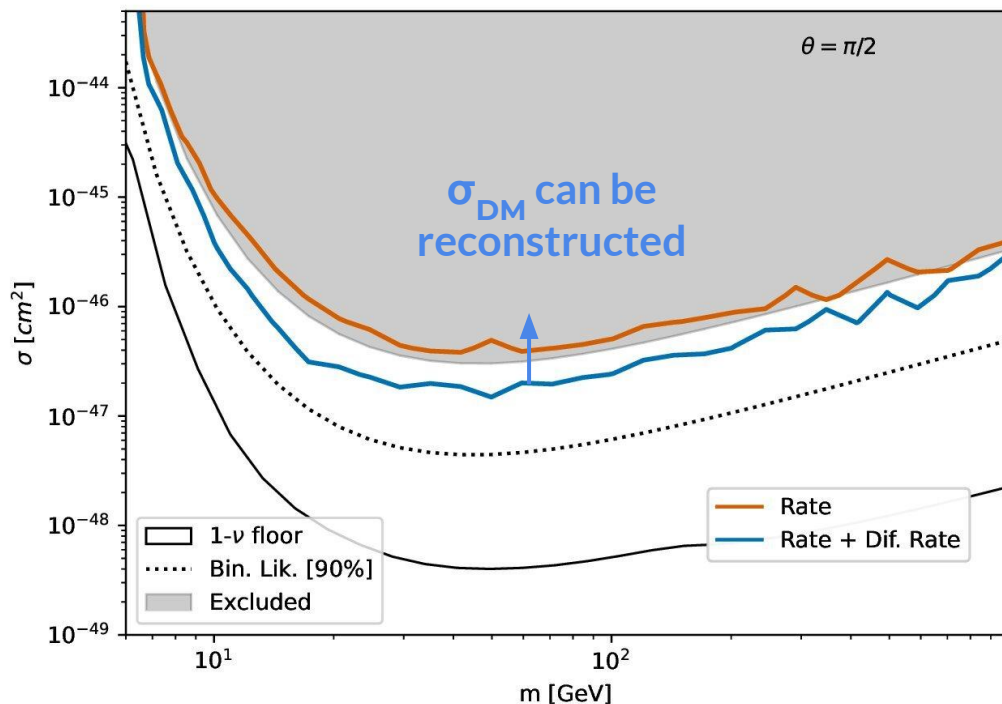
Reconstruction of the cross-section σ_{DM}

ML (SWYFT):

— **rate**

Results: parameter reconstruction

For O_1 with fixed $\theta = \pi/2$



Reconstruction of the cross-section σ_{DM}

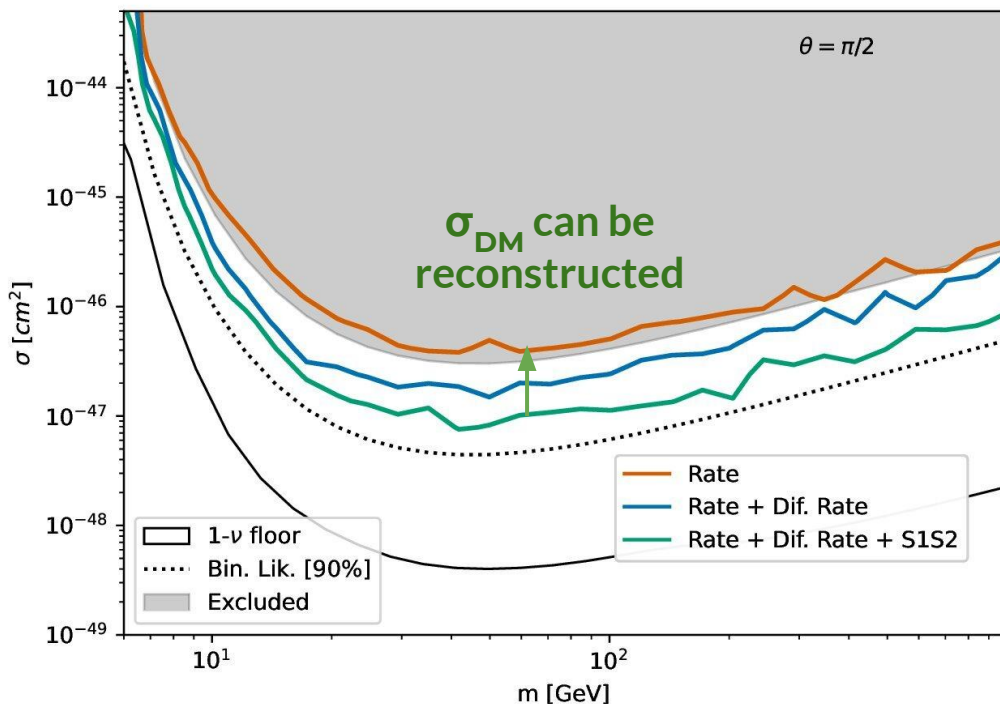
ML (SWYFT):

— **rate**

— **rate + diff. rate**

Results: parameter reconstruction

For O_1 with fixed $\theta = \pi/2$



Reconstruction of the cross-section σ_{DM}

ML (SWYFT):

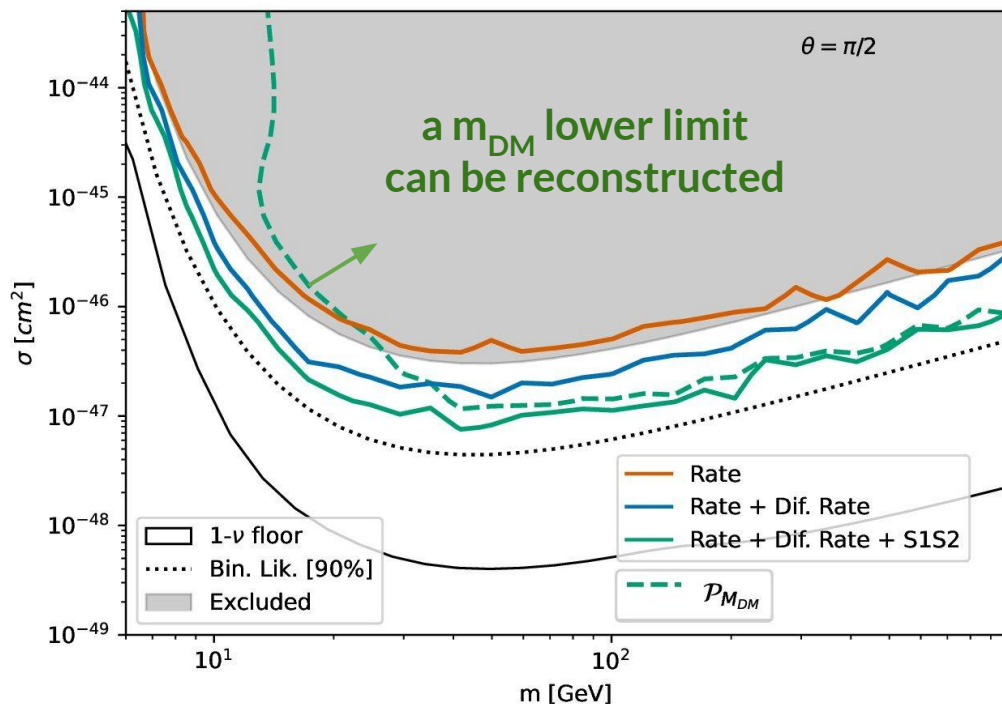
— **rate**

— **rate + diff. rate**

— **rate + diff. rate + cS1-cS2**

Results: parameter reconstruction

For O_1 with fixed $\theta = \pi/2$



Reconstruction of the cross-section σ_{DM}

ML (SWYFT):

— **rate**

— **rate + diff. rate**

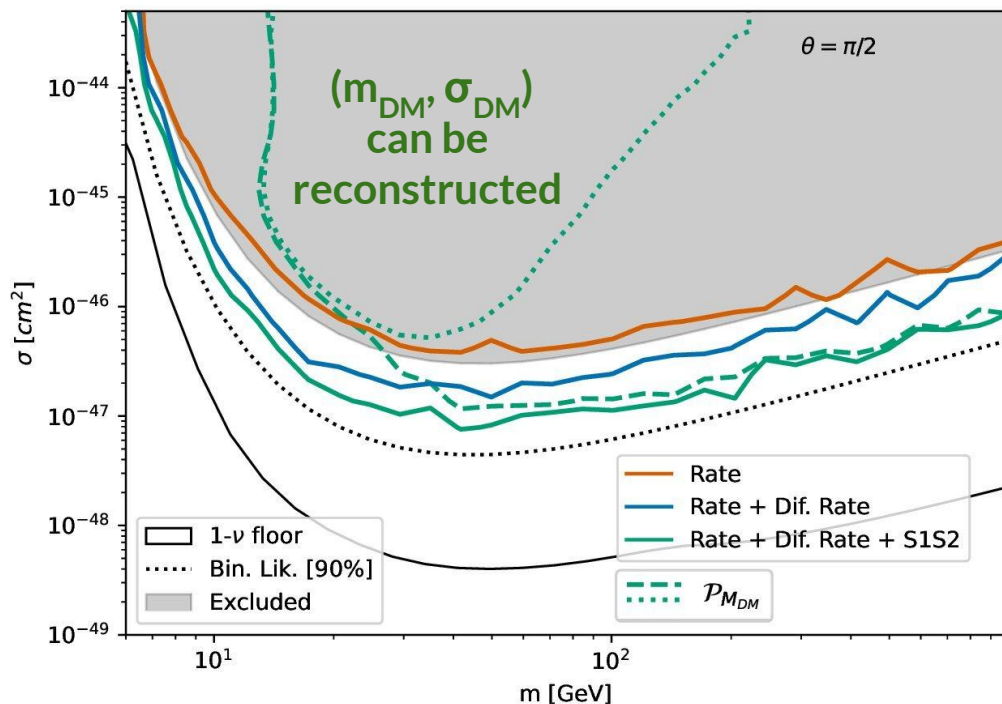
— **rate + diff. rate + cS1-cS2**

Reconstruction of the mass m_{DM}

-- **rate + diff. rate + cS1-cS2**

Results: parameter reconstruction

For O_1 with fixed $\theta = \pi/2$



Reconstruction of
the cross-section σ_{DM}

ML (SWYFT):

— **rate**

— **rate + diff. rate**

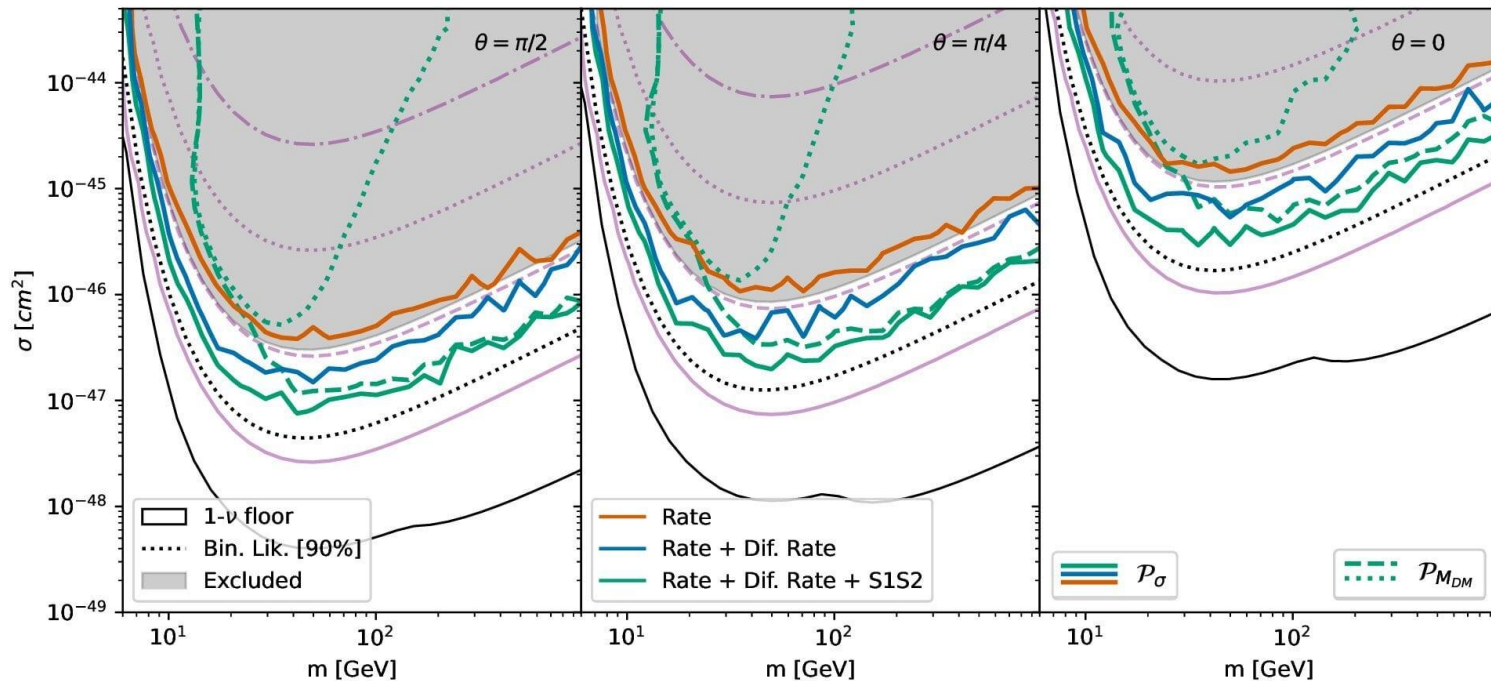
— **rate + diff. rate + cS1-cS2**

Reconstruction of
the mass m_{DM}

-- **rate + diff. rate + cS1-cS2**

Results: parameter reconstruction

O_1 operator
XENONnT 20ty



This panel is the usually shown **SI** parameter space

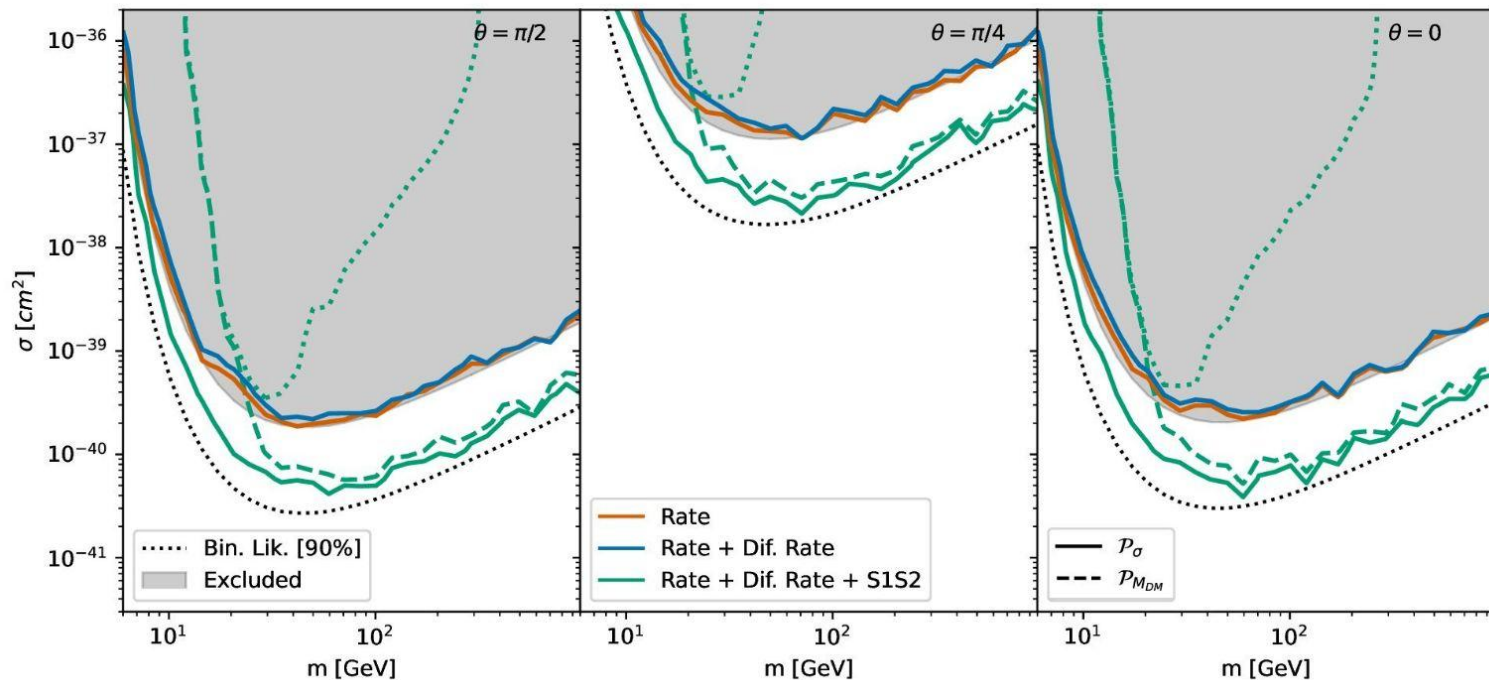
total number of events

Data: + differential rate

+ cS1 vs cS2 plane

Results: parameter reconstruction

O_4 operator
XENONnT 20ty



total number of events

Data: + differential rate

+ cS1 vs cS2 plane

Python package



CADDENA → <https://github.com/Martindelosrios/CADDENA>

```
In [6]: # Now let's load the dataset with XENON nT simulations.
# This data was create with .....

ref = files("CADDENA") / "dataset/"
DATA_PATH = str(ref)
with h5py.File(DATA_PATH + "/testset.h5", "r") as data:
    x_norm_rate = data['x_norm_rate'][()]
    x_norm_drate = data['x_norm_drate'][()]
    x_norm_s1s2 = data['x_norm_s1s2'][()]
    pars_norm = data['pars_norm'][()]
```

Load the dataset

```
In [7]: # Let's pick some random synthetic observation from the simulated dataset.

i = np.random.randint(len(pars_norm))
print('nobs = ' + str(i))

pars_true = pars_norm[i,:]
x_obs_rate = x_norm_rate[i,:]
x_obs_drate = x_norm_drate[i,:]
x_obs_s1s2 = x_norm_s1s2[i,:].reshape(1,96,96)
```

**Pick one as the
'observation'**

Python package



CADDENA → <https://github.com/Martindelosrios/CADDENA>

In [11]:

```
models.XENONnT_01_rate.load_weights()  
models.XENONnT_01_drate.load_weights()  
models.XENONnT_01_s1s2.load_weights()
```

**Load the already
trained ML algorithms**

In [13]:

```
# Let's sample the parameters from a flat prior  
  
pars_prior = np.random.uniform(low = 0, high = 1, size = (100_000, 3))
```

Set a prior

Python package

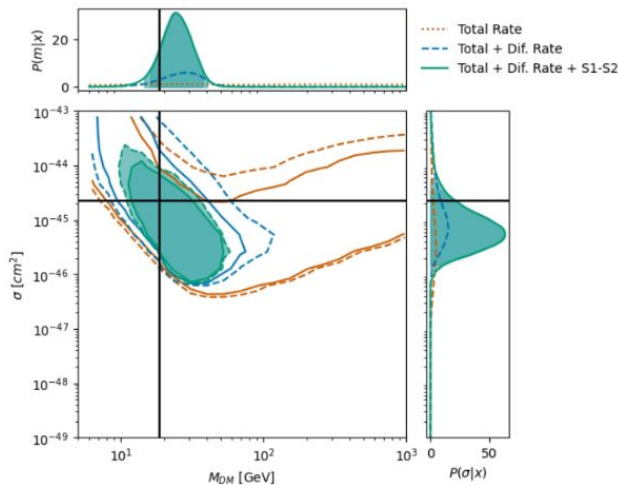


CADDENA → <https://github.com/Martindelosrios/CADDENA>

In [11]:

```
1 # Let's analyze the each type of data with the corresponding model
2
3 # The first argument is a list with the observation that will be analyzed.
4 # The second argument is the sample of paramters that will be paired with the observations.
5 # The third argument is a list with the models that will be used.
6
7 logratios1D, logratios2D = caddena.ratio_estimation([x_obs_rate, x_obs_drates, x_obs_sls2],
8                                                    pars_prior,
9                                                    [models.XENONnT_01_rate, models.XENONnT_01_drates,
10                                                     models.XENONnT_01_sls2])
```

Compute the posteriors



Plot the posterior distributions

Conclusions

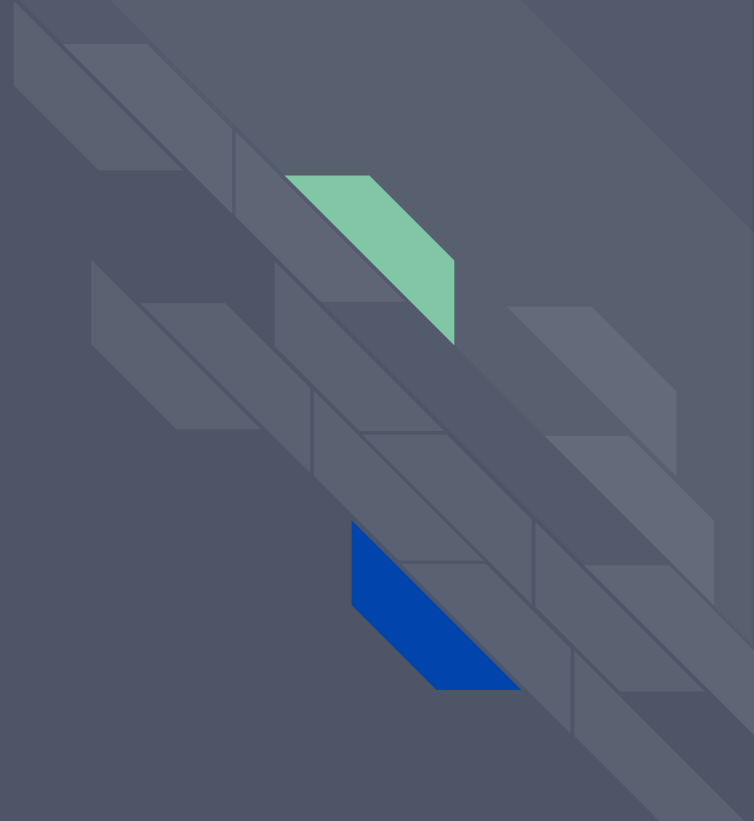
- We developed a Bayesian analysis to explore the reach of direct detection experiments that can be applied to any DM model (NR-EFT)
- The ML implementation (SWYFT) is fundamental:
 - **fast** estimation of posteriors,
 - simple way to **include-combine-remove** data.
- O_1 (SI) and O_4 (SD) presented here as examples
- We computed the parameter space where \mathbf{m}_{DM} and $\boldsymbol{\sigma}$ that can be **reconstructed**, using: total number of events - differential rate - full cS1,cS2 space.
- **Next:**
 - Apply to other NR-EFT operators → combine operators
 - Different DD experiments → combine experiments



Thank you!



Back-up



Bayesian Analysis

Marginalization (traditional method):

one needs to assume
a functional form for
the likelihood

$$P(\Theta_0 | \mathbf{X}) = \int \frac{P(\mathbf{X} | \Theta) P(\Theta) \prod_j d\Theta_{j \neq 0}}{P(\mathbf{X})}$$

If you want to
study only one
parameter Θ_0

The full posterior
has to be
computed

Bayesian Analysis with SWYFT

SWYFT → Sampling-based inference tool that estimates *likelihood-to-evidence ratio* with ML algorithms to obtain marginal and joint posteriors

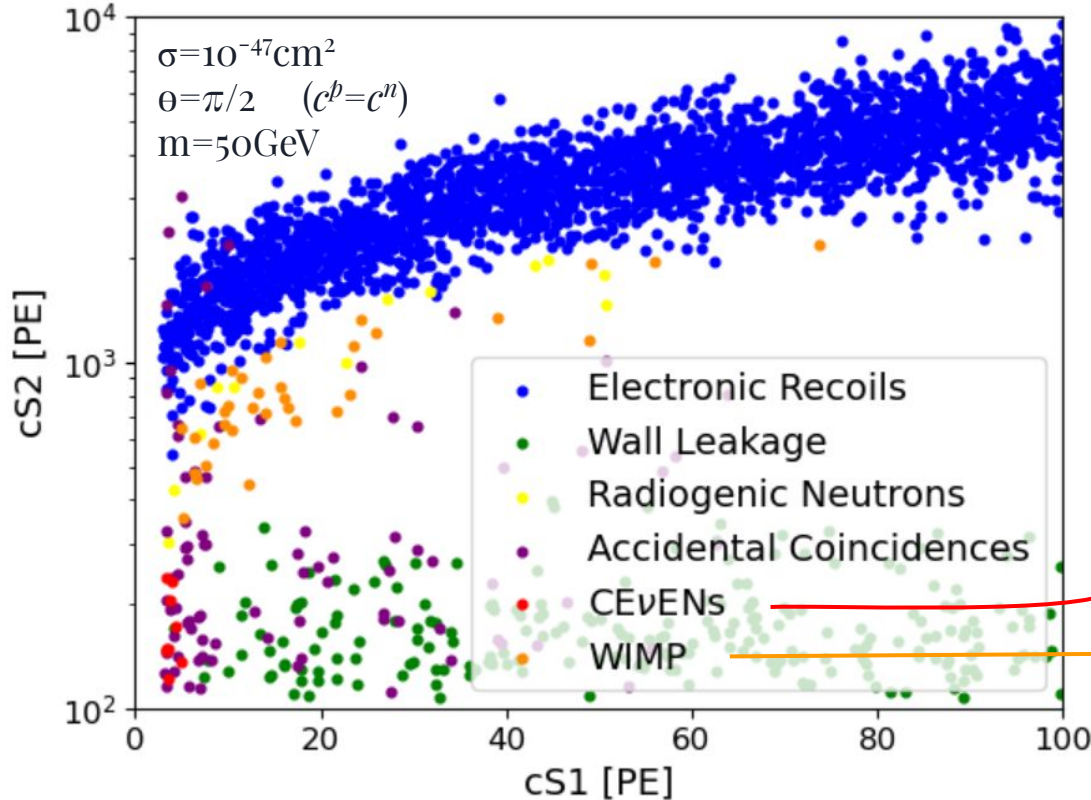
$$r(\mathbf{X}, \Theta) = \frac{\text{output}}{1 - \text{output}} \quad \xrightarrow{\text{compute the posterior}} \quad P(\Theta | \mathbf{X}) = r(\mathbf{X}, \Theta) P(\Theta)$$

$$\begin{aligned} r(\mathbf{X}, \Theta) &= \frac{P(\mathbf{X}, \Theta)}{P(\Theta)P(\mathbf{X})} = \frac{P(\mathbf{X}, \Theta | k = 1)}{P(\mathbf{X}, \Theta | k = 0)} \\ &= \frac{P(\mathbf{X}, \Theta, k = 1)}{P(\mathbf{X}, \Theta, k = 0)} = \frac{P(k = 1 | \mathbf{X}, \Theta)}{P(k = 0 | \mathbf{X}, \Theta)} \\ &= \frac{P(k = 1 | \mathbf{X}, \Theta)}{1 - P(k = 1 | \mathbf{X}, \Theta)} = \frac{\text{output}}{1 - \text{output}} \end{aligned}$$

DM signal

XENONnT 20ty

NR-EFT: O_1



XENONnT simulator

We specify background and signal characteristics

differential rate
compute with SNUDD

differential rate compute with
WimPyDD for a particular
operator, $(\sigma_i, \theta_i, m_{\text{DM}})$, and
standard DM halo model

Data Representation: cS1 vs cS2 plane

XENONnT 20ty

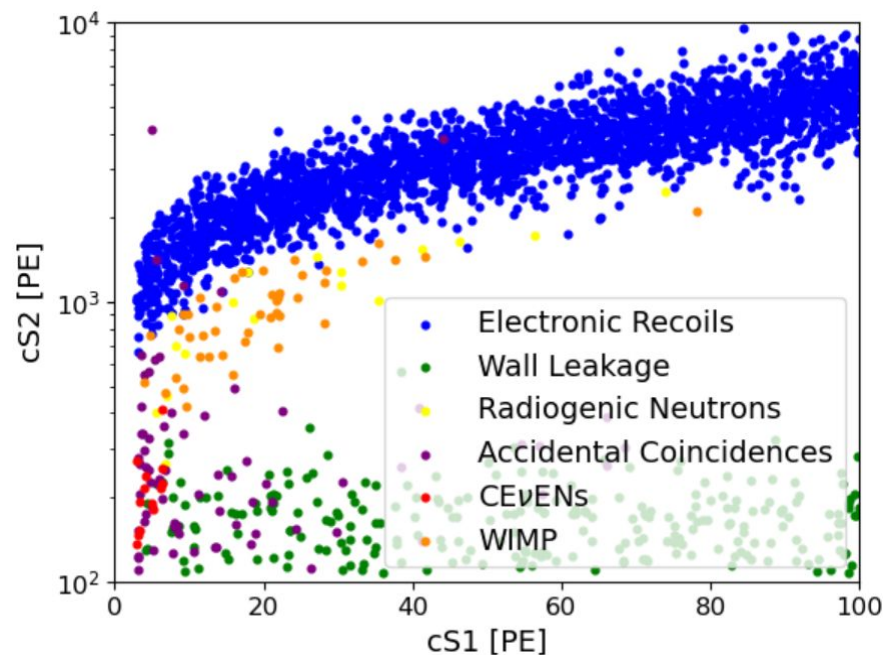
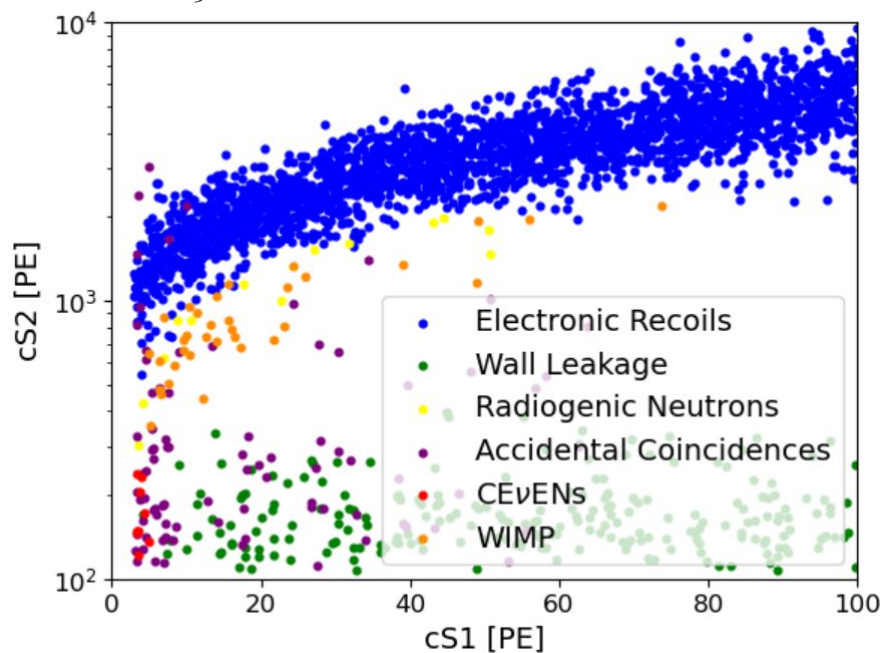
NR-EFT: O₁

$$\sigma = 10^{-47} \text{cm}^2$$

$$\theta = \pi/2 \quad (c^b = c^n)$$

$$m = 50 \text{GeV}$$

We generate a 10k pseudo experiments per operator varying σ , θ , and m_{DM}



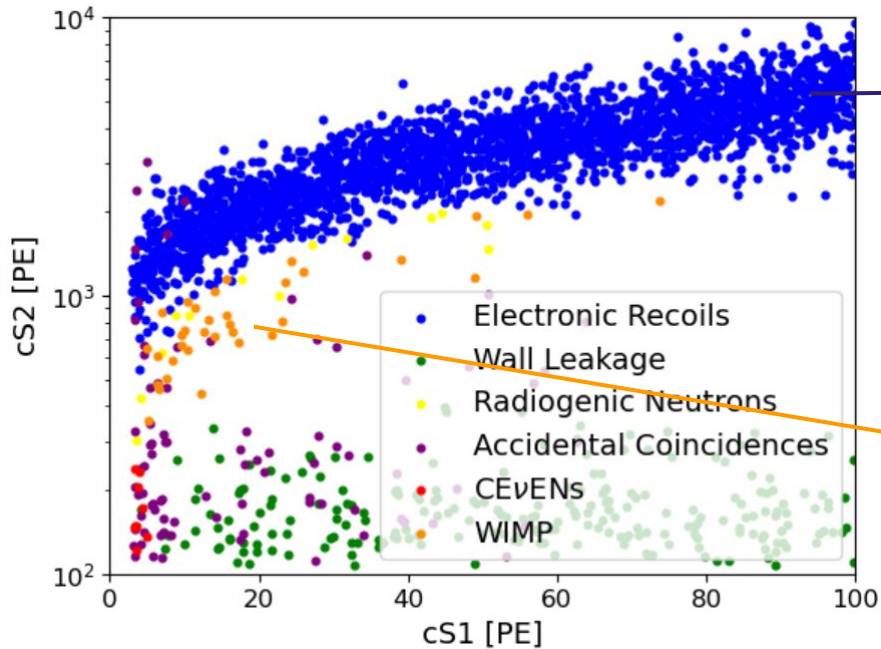
Data Representation: number of events

NR-EFT: O₁

$$\sigma = 10^{-47} \text{cm}^2$$

$$\theta = \pi/2 \quad (c^b = c^n)$$

$$m = 50 \text{GeV}$$



	name	pseudo_exp_events
0	er	2459
1	radiogenics	17
2	ac	71
3	wall	246
4	WIMP	43
5	CEVNS-SM	13

Data Representation: differential rate

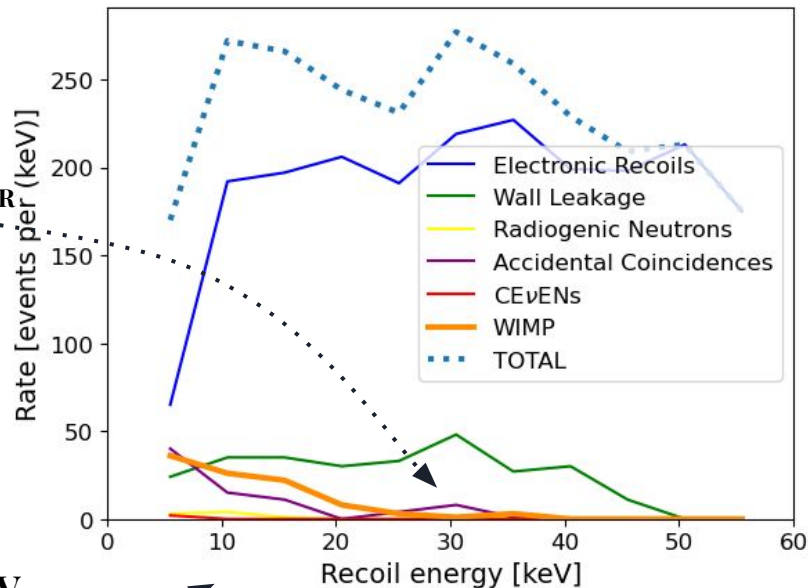
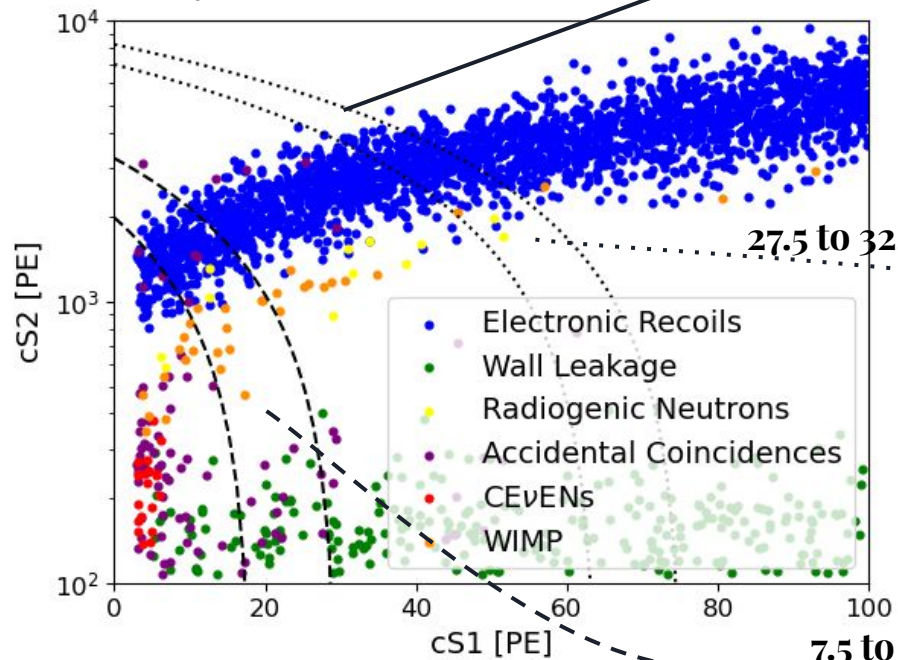
NR-EFT: O₁

$$\sigma = 10^{-47} \text{cm}^2$$

$$\theta = \pi/2 \quad (c^b = c^n)$$

$$m = 50 \text{GeV}$$

Nuclear Recoil
isoenergy
curves



Results

Data:

cS1 vs cS2 plane

Posteriors for:

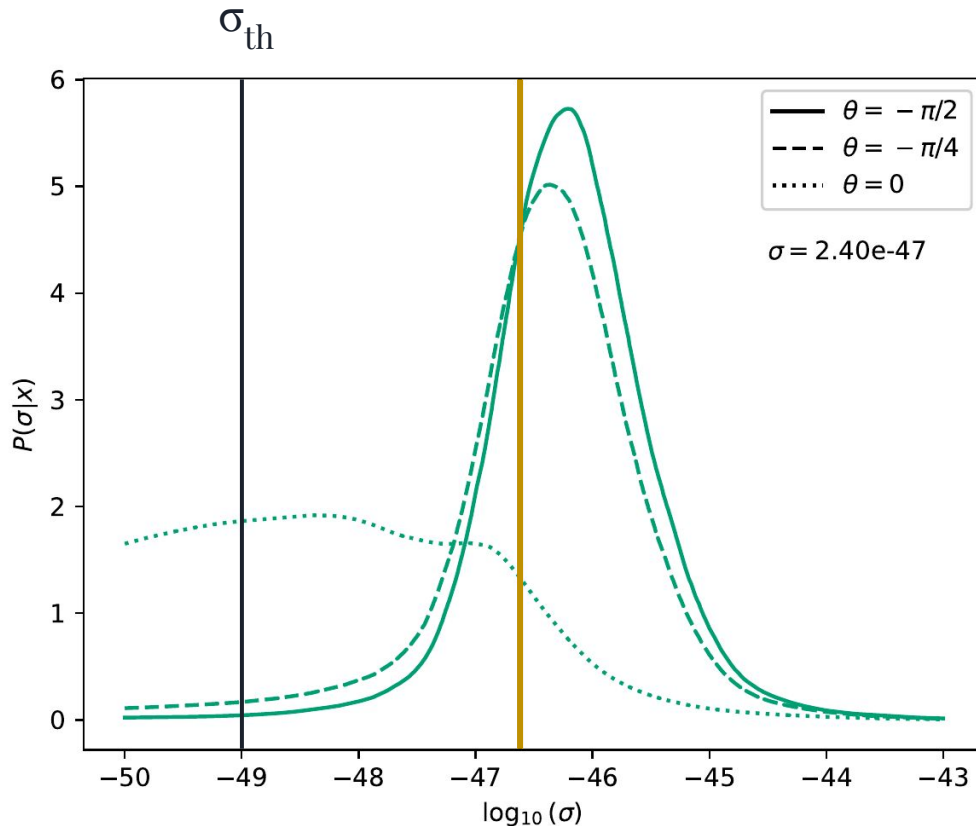
O_1 (SI)

$m_{\text{DM}} \approx 100 \text{ GeV} \rightarrow$ **fixed**

$\sigma = 2.4 \cdot 10^{-47} \text{ cm}^2 \rightarrow$ **fixed**

reconstruction
depends on θ

cS1 vs cS2



Comparison with Multinest

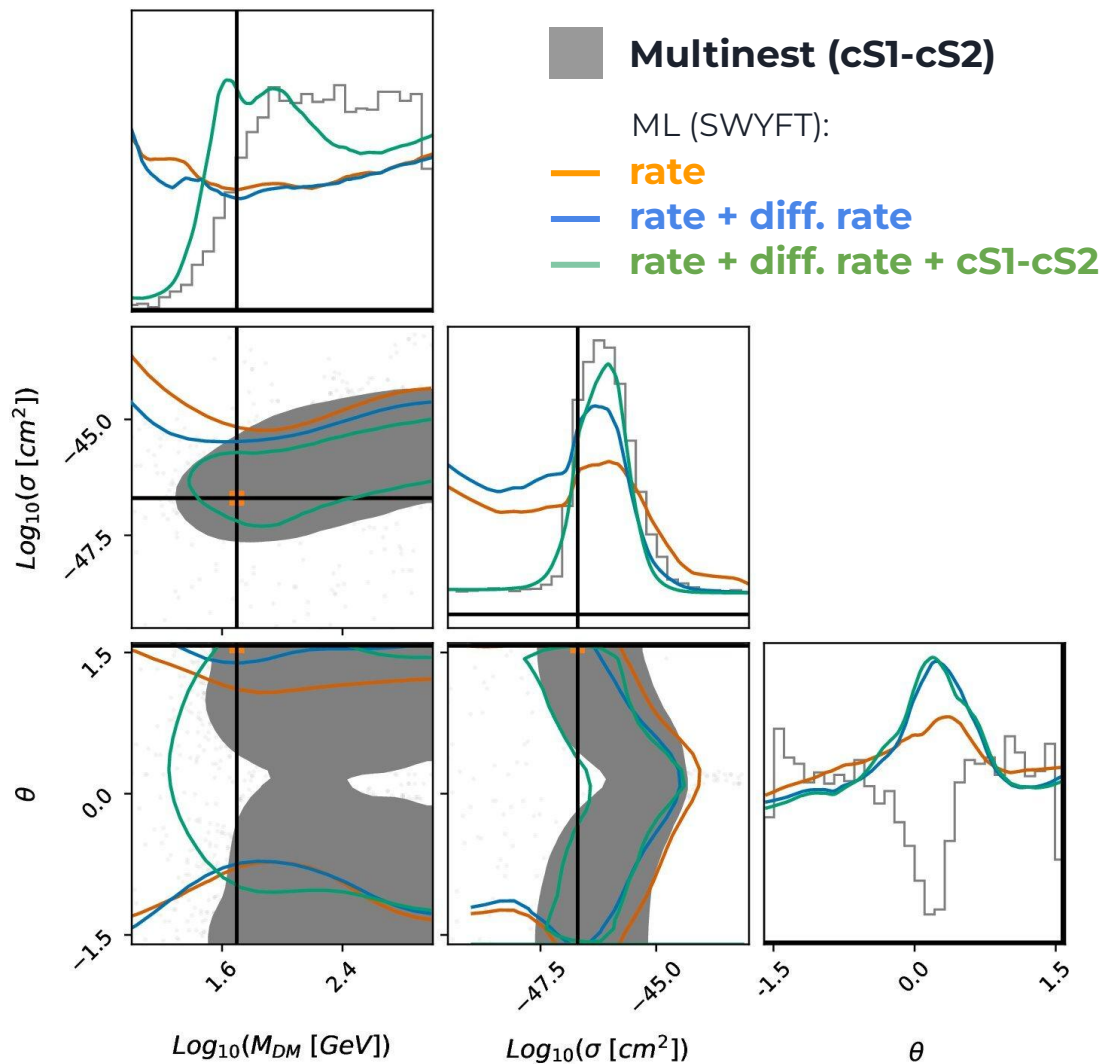
Posterior distributions for:

O_1 (SI)

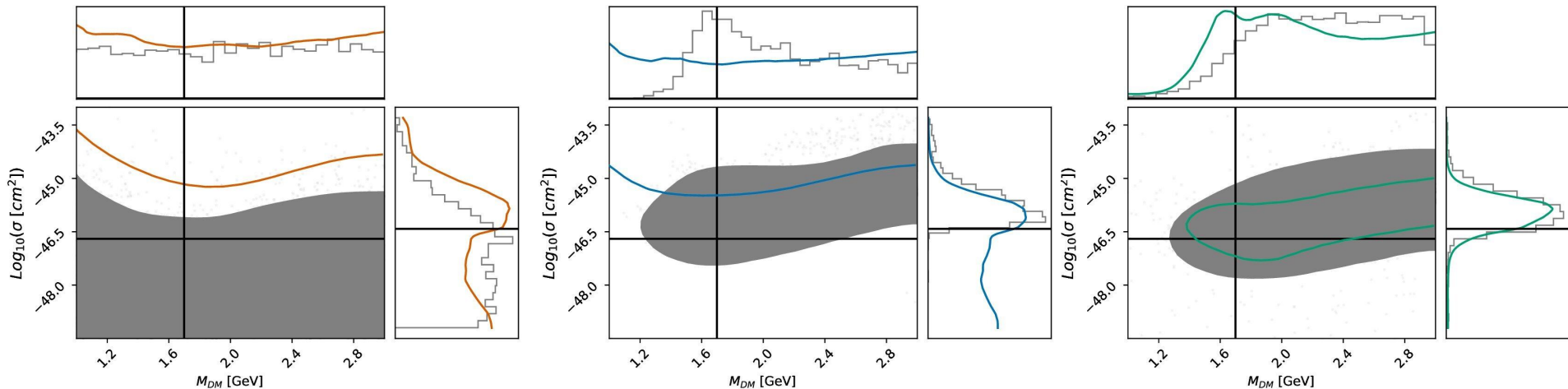
$m_{DM} \approx 50 \text{ GeV}$ → **fixed**

$\sigma = 2.10^{-47} \text{ cm}^2$ → **fixed**

$\theta = \pi/2$ → **fixed**



Results: parameter reconstruction



Posterior distributions for:

O_1 (SI)

$m_{\text{DM}} \approx 50 \text{ GeV}$

$\sigma = 2 \cdot 10^{-47} \text{ cm}^2$

$\theta = \pi/2$

→ **fixed**

→ **fixed**

→ **fixed**

■ **Multinest (cS1-cS2)**

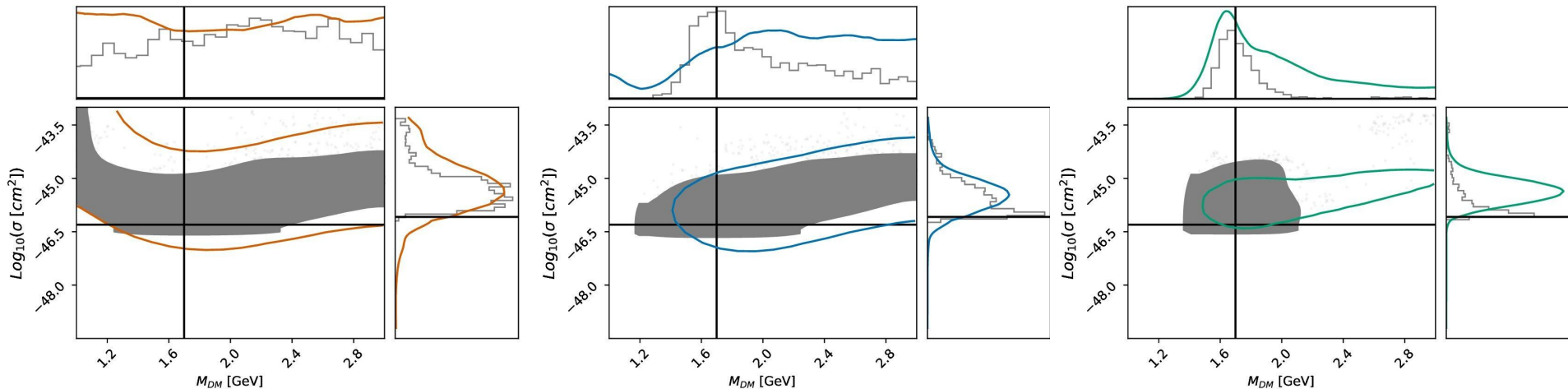
ML (SWYFT):

— **rate**

— **rate + diff. rate**

— **rate + diff. rate + cS1-cS2**

Results: parameter reconstruction



Posterior distributions for:

O_1 (SI)

$m_{DM} \approx 50 \text{ GeV}$

$\sigma = 5 \cdot 10^{-47} \text{ cm}^2$

$\theta = \pi/2$

→ **fixed**

→ **fixed**

→ **fixed**

■ **Multinest (cS1-cS2)**

ML (SWYFT):

— **rate**

— **rate + diff. rate**

— **rate + diff. rate + cS1-cS2**

DM-nucleon non-relativistic effective field theory (NR-EFT)

For DM particles with spin up to $\frac{1}{2}$, the effective DM-nucleon scattering interaction Lagrangian

$$\mathcal{L}_{\text{EFT}} = \sum_{\tau} \sum_i c_i^{\tau} \mathcal{O}_i \bar{\chi} \chi \bar{\tau} \tau \quad \text{i=14 possible interactions}$$

$$i \frac{\vec{q}}{m_N}, \quad \vec{v}^{\perp}, \quad \vec{S}_{\chi}, \quad \vec{S}_N.$$

$$\vec{v} \equiv \vec{v}_{\chi, \text{in}} - \vec{v}_{N, \text{in}}$$

$$\vec{v}^{\perp} = \vec{v} + \frac{\vec{q}}{2\mu_N}$$

momentum transfer, spin operators, relative velocity

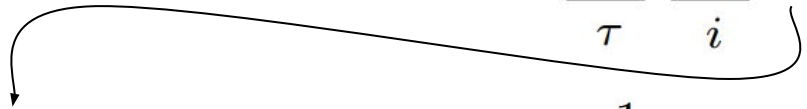
$\mathcal{O}_1 = 1_{\chi} 1_N$	$\mathcal{O}_9 = i \vec{S}_{\chi} \cdot (\vec{S}_N \times \frac{\vec{q}}{m_N})$
$\mathcal{O}_3 = i \vec{S}_N \cdot (\frac{\vec{q}}{m_N} \times \vec{v}^{\perp})$	$\mathcal{O}_{10} = i \vec{S}_N \cdot \frac{\vec{q}}{m_N}$
$\mathcal{O}_4 = \vec{S}_{\chi} \cdot \vec{S}_N$	$\mathcal{O}_{11} = i \vec{S}_{\chi} \cdot \frac{\vec{q}}{m_N}$
$\mathcal{O}_5 = i \vec{S}_{\chi} \cdot (\frac{\vec{q}}{m_N} \times \vec{v}^{\perp})$	$\mathcal{O}_{12} = \vec{S}_{\chi} \cdot (\vec{S}_N \times \vec{v}^{\perp})$
$\mathcal{O}_6 = (\vec{S}_{\chi} \cdot \frac{\vec{q}}{m_N})(\vec{S}_N \cdot \frac{\vec{q}}{m_N})$	$\mathcal{O}_{13} = i(\vec{S}_{\chi} \cdot \vec{v}^{\perp})(\vec{S}_N \cdot \frac{\vec{q}}{m_N})$
$\mathcal{O}_7 = \vec{S}_N \cdot \vec{v}^{\perp}$	$\mathcal{O}_{14} = i(\vec{S}_{\chi} \cdot \frac{\vec{q}}{m_N})(\vec{S}_N \cdot \vec{v}^{\perp})$
$\mathcal{O}_8 = \vec{S}_{\chi} \cdot \vec{v}^{\perp}$	$\mathcal{O}_{15} = -(\vec{S}_{\chi} \cdot \frac{\vec{q}}{m_N})((\vec{S}_N \times \vec{v}^{\perp}) \cdot \frac{\vec{q}}{m_N})$

DM-nucleon non-relativistic effective field theory (NR-EFT)

For DM particles with spin up to $\frac{1}{2}$, the effective DM-nucleon scattering interaction Lagrangian

$$\mathcal{L}_{\text{EFT}} = \sum_{\tau} \sum_i c_i^{\tau} \mathcal{O}_i \bar{\chi} \chi \bar{\tau} \tau$$

$i=14$ possible interactions


$$c_i^0 \mathbb{1}_{2 \times 2} + c_i^1 \tau_3$$

isospin basis
 c^0 : isoscalar
 c^1 : isovector

$$c_i^0 = \frac{1}{2} (c_i^p + c_i^n)$$

$$c_i^1 = \frac{1}{2} (c_i^p - c_i^n)$$

nucleon basis
 c^p : proton
 c^n : neutron

\mathcal{O}_1 : spin-independent (SI)
 \mathcal{O}_4 : spin-dependent (SD)

usually shown assuming isoscalar interactions

$$c^p = c^n \quad c^0 = 1 \text{ and } c^1 = 0$$

DM-nucleon non-relativistic effective field theory (NR-EFT)

For DM particles with spin up to $\frac{1}{2}$, the effective DM-nucleon scattering interaction Lagrangian

$$\mathcal{L}_{\text{EFT}} = \sum_{\tau} \sum_i c_i^{\tau} \mathcal{O}_i \bar{\chi} \chi \bar{\tau} \tau \quad \text{i=14 possible interactions}$$

Change to polar coordinates:

$$c_i^0 = \frac{1}{2} (c_i^p + c_i^n) = A_i \sin(\theta_i)$$

$$c_i^1 = \frac{1}{2} (c_i^p - c_i^n) = A_i \cos(\theta_i)$$

Natural choice for the EFT parameter space because the interaction cross section:

$$\sigma_i \propto A_i^2$$

For SI (O1) $\sigma_{\chi\mathcal{N}}^{\text{SI}} = \frac{A_1^2 \mu_{\chi\mathcal{N}}^2}{\pi}$ ↗ DM-nucleon reduced mass

DM-nucleon non-relativistic effective field theory (NR-EFT)

Contact interaction between a spin $\frac{1}{2}$ DM and nucleon

$$\mathcal{L}_{\text{int}}^{\text{SI}}(\vec{x}) = c_1 \bar{\Psi}_\chi(\vec{x})\Psi_\chi(\vec{x}) \bar{\Psi}_N(\vec{x})\Psi_N(\vec{x})$$

$$U_\chi(p) = \sqrt{\frac{E+m}{2m}} \begin{pmatrix} \xi_\chi \\ \frac{\vec{\sigma} \cdot \vec{p}}{E+m_\chi} \xi_\chi \end{pmatrix} \sim \begin{pmatrix} \xi_\chi \\ \frac{\vec{\sigma} \cdot \vec{p}}{2m_\chi} \xi_\chi \end{pmatrix}$$

at low momenta.

Idem for the nucleon spinor

ξ Pauli spinors

at leading order in p/m $c_1 \mathbf{1}_\chi \mathbf{1}_N \equiv c_1 \mathcal{O}_1$

DM-nucleon non-relativistic effective field theory (NR-EFT)

Another interaction

$$\mathcal{L}_{\text{int}}^{\text{SD}} = c_4 \bar{\chi} \gamma^\mu \gamma^5 \chi \bar{N} \gamma_\mu \gamma^5 N$$

the dominant contribution in
the non-relativistic limit
comes from the spatial indices

$$\bar{\chi} \gamma^i \gamma^5 \chi \sim \xi_\chi^\dagger \sigma^i \xi_\chi$$

$$\text{Since } \hat{S}^i = \sigma^i / 2 \quad -4c_4 \vec{S}_\chi \cdot \vec{S}_N \equiv -4c_4 \mathcal{O}_4$$

Data analysis to obtain posteriors

SWYFT → Sampling-based inference tool that estimates likelihood to evidence ratio with ML algorithms to obtain marginal and joint posteriors

	MCMC	SWYFT
Forward Model	$x=f(\text{parameters})$	$x=f(\text{parameters})$
Likelihood	$L(x, f(\text{parameters}))$	Data Driven
Samples	All parameters space > # samples	Only Interesting parameters
Amortization	NO	YES

Motivation

Bayes' Rule: determine a probability distribution over model parameters θ given an observation \mathbf{x}

$$p(\theta | \mathbf{x}) = \frac{p(\mathbf{x} | \theta)}{p(\mathbf{x})} p(\theta)$$

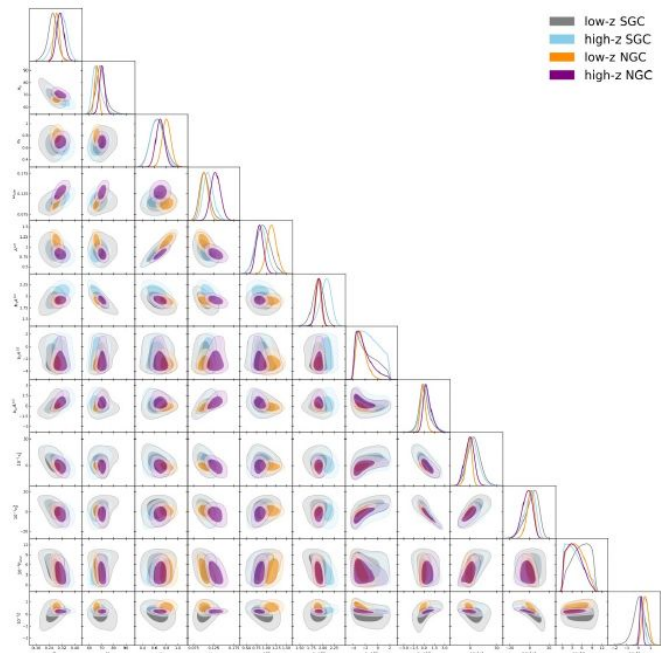
Posterior

Likelihood of \mathbf{x} given θ

Evidence of the data

Prior

Samples typically generated with *Markov Chain Monte Carlo (MCMC)* or *Nested sampling*



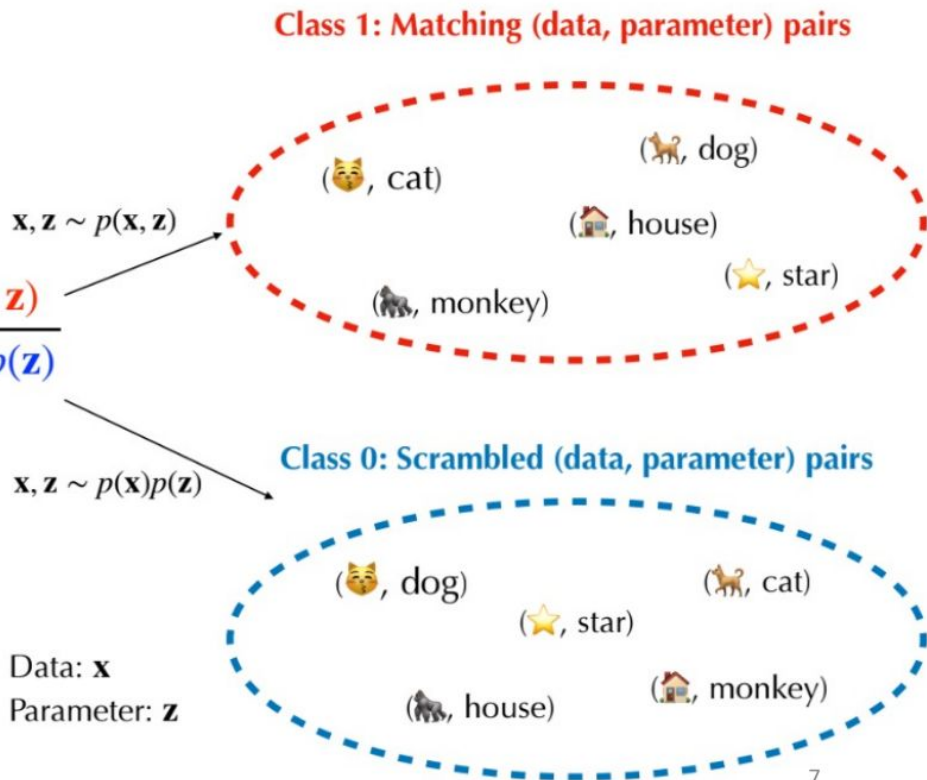
BOSS, Ivanov+ 1909.05277

Neural Ratio Estimation (NRE)

Approximate density ratios.

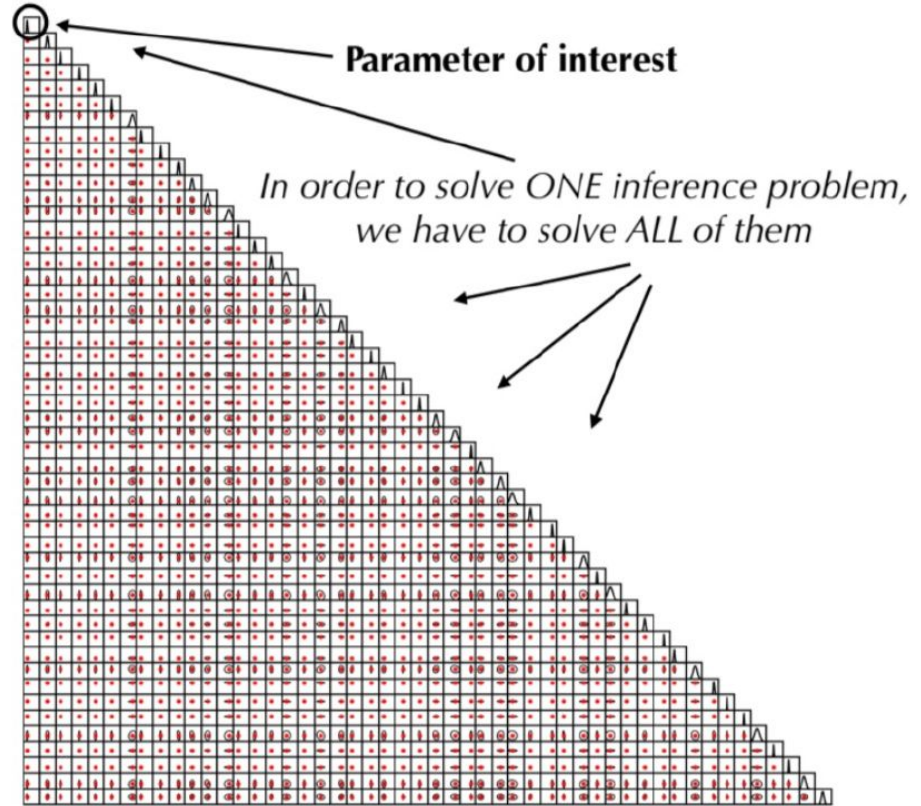
$$r(\mathbf{x}; \mathbf{z}) \equiv \frac{p(\mathbf{x} | \mathbf{z})}{p(\mathbf{x})} = \frac{p(\mathbf{z} | \mathbf{x})}{p(\mathbf{z})} = \frac{p(\mathbf{x}, \mathbf{z})}{p(\mathbf{x})p(\mathbf{z})}$$

Strategy: We estimate posteriors-to-prior ratio by training a binary classifier to discriminate between matching and scrambled (data, parameter) pairs.

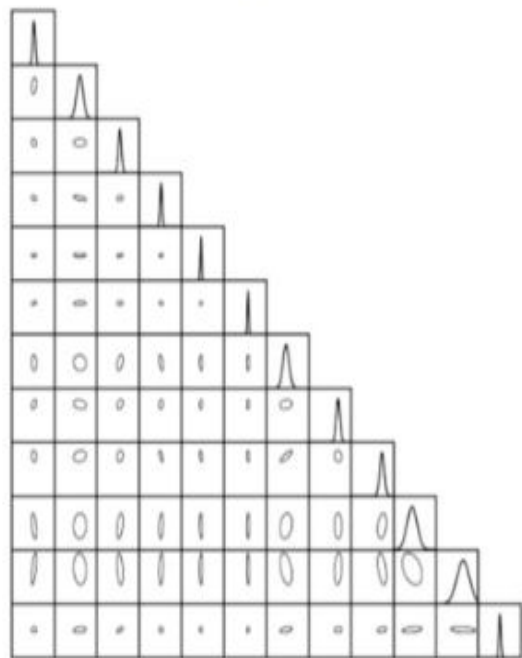


MARGINAL

- MCMC or Nested sampling methods produce samples from the **posterior distribution**.
- Classical methods require sampling the **full joint posterior**, so that they are slow to converge.
- Novel approaches in the field of *simulation-based inference* (SBI) are starting to overcome these obstacles.

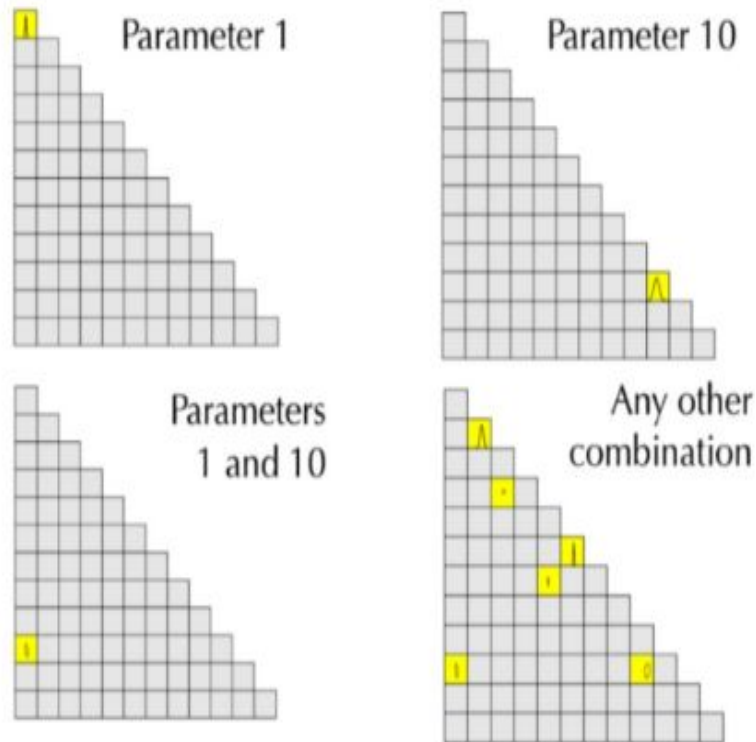


Instead of estimating all parameters...



50 parameters ~ 100 Million simulations

...we can choose what we care about



Depending on which parameter is scrambled

Results

Data:

entire **cS1 vs cS2** plane

These are all the posteriors for

$m_{\text{DM}} \approx 100\text{GeV} \rightarrow \text{fixed}$
 $\theta = \pi/2 \rightarrow \text{fixed}$

red $\rightarrow \sigma$ not reconstructed
~no signal,
~similar posteriors

black $\rightarrow \sigma$ reconstructed

