Frequentist Uncertainties on Density Ratios with $w_i f_i$ **Ensembles**

Sean Benevedes ML4Jets2024 - November 6 2024

Work in collaboration with Jesse Thaler





Most talks

- 1. Introduction
- Methods 2.
- Results 3.
- 4. Discussion



Organization



- 1. Introduction 2.
 - Results
- 3. Methods
- 4. Discussion



Density ratio estimation





What: Given samples from two distributions p(x)and q(x), estimate $\log r(x) \equiv \log \frac{q(x)}{p(x)}$







How: Machine learning!

- by any monotonic function of r
- Cranmer et al., 1506.02169)

But... how do we use a point estimate for r? If the estimate is poor, downstream applications will suffer. Need a measure of uncertainty in the estimate!

See Hermans et al. 2110.06581 for more on problems caused by neglecting these uncertainties



Neyman-Pearson Lemma: The best classifier between q and p is given

Idea: Train a classifier on samples from q and p, solve for r (see e.g.





$\log r(x \mid \{$

where w_i are scalar weights and $f_i(x)$ are frozen networks

Pros

- Fully frequentist, no prior dependence
- Intuitive: just like nuisance parameters!
- Computationally efficient; no bootstrapping

$w_i f_i$ ensembles

$$w_i\}) = w_i f_i(x)$$

Cons
• Does not assess uncertainty from training of the f_i



$\log r(x \mid \{$

where w_i are scalar weights and $f_i(x)$ are frozen networks

Pros

- Fully frequentist, no prior dependence
- Intuitive: just like nuisance parameters!
- Computationally efficient; no bootstrapping

$w_i f_i$ ensembles

$$w_i\}) = w_i f_i(x)$$





Where do The lay of the Bootstrap Pareto exclude Neyman construc (e.g. https://cds.cern.ch/recor

Doesn't include training uncertainties

Includes training uncertainties

these fit in? (frequentist) land		
	No bootstrap	
b	KLIEP (Sugiyama et al. 200 $w_i f_i$ ensembles	
ction d/2915316)	Hopefully next year's talk!	







The big surprise: for the problems we've checked, this actually works!







- Let q be $\mathcal{N}(\mu, \sigma^2)$ and p be
 - $\mathcal{N}(-\mu,\sigma^2)$, then $\log r = 2\mu x/\sigma^2$
- Evaluation:
 - Directly check coverage of r with pseudoexperiments
 - "Mixture fraction task" gives us a measure of performance in downstream task



1D toy problem: Gaussians



Sean Benevedes

 $\hat{f} = \operatorname{argmax}_{f} \sum \log \frac{p(x_{\alpha}|f)}{p(x_{\alpha})} = \operatorname{argmax}_{f} \sum_{\alpha} \log(fr(x_{\alpha}) + (1-f))$



Mixture fraction task

Estimate the parameter f given a set of samples D drawn i.i.d. from p(x | f) = fq(x) + (1 - f)p(x)

r(x) suffices to estimate f with maximum likelihood estimation:







Sean Benevedes









ML4Jets: Quark/gluon likelihood

- Now, $q \sim$ Pythia 8.226 quark jets, $p \sim$ Pythia 8.226 gluon jets We use the dataset included in the EnergyFlow package (you can find the jets here: <u>https://zenodo.org/records/3164691</u>)
- Since I don't know the exact answer for the likelihood ratio (let me know if you do!), we only report coverage for the mixture fraction task

an IRC-safe observable



Technical detail: We construct an IRC-safe ensemble with energy-flow networks, so the objective function is actually the projection of log r to











Ok, what did we actually do?





$w_i f_i$ ensembles: the recipe

Ingredients

- *M* randomly initialized networks
- A training dataset with N_{train} samples each from q and p
- (Optional) A dataset of size N_{test} for downstream applications
- 1. Train the f_i on the training dataset
- 2. Find \hat{w}_i using the MLC loss (introduced in Tito D'Agnolo et al. 1806.02350)
- 3. Compute $C_{ij} \equiv \text{Cov}(\hat{w}_i, \hat{w}_j)$ with analytic formulae 4. Propagate uncertainties for downstream tasks





Training the f_i

- Objective: Train f_i such that there is some set of w_i for which $w_i f_i$ is a good model of log r... that leaves a lot of freedom!
- (At least) two families of approaches:
 - 1. Train one distinguished f_0 to model log r as well as possible (just like before!), use the remaining f_i to model the residuals
 - " $W_i = (1, 0, 0, 0, ...)_i$ "
 - 2. Train the f_i on equal footing (conventional ensembling)
 - " $W_i = (1/M, 1/M, \dots)_i$ "







Training the f_i

- Objective: Train f_i such that there is some set of w_i for which $w_i f_i$ is a good model of log r... that leaves a lot of freedom!
- (At least) two families of approaches: 1. Train one distinguished f_0 to model log r as well as possible (just like before!), use the remaining f_i to model the residuals
 - " $w_i = (1, 0, 0, 0, ...)_i$ "
 - 2. Train the f_i on equal footing (conventional ensembling) • " $w_i = (1/M, 1/M, \dots)_i$ "







Training the f_i

- 1. Train f_0 as a classifier between q and p (using e.g. the MLC loss) 2. Train $\{f_1, f_2, \ldots, f_{M-1}\}$ to be uncorrelated with each other and f_0 , but not as classifiers (ask me why!)









Fitting the *W*;

Now, find the \hat{w}_i minimizing the symmetrized MLC loss:

$$\mathscr{L}(w) = -\left(\sum_{x_{\alpha} \sim q}^{N_{train}} [w_i f_i(x_{\alpha}) - (e^{-w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - (e^{w_i f_i(x_{\alpha})} - 1)] + \sum_{x_{\alpha} \sim p}^{N_{train}} [-w_i f_i(x_{\alpha}) - 1] + \sum_{x_{\alpha} \sim$$

Estimators of this form are called *M*-estimators (Hubert 1964) in the statistics literature, and their covariances are known!

$$C_{ij} = V_{ik}^{-1} U_{kl} V_{lj}^{-1} \quad \left(\text{where } V_{ik} \equiv \text{Cov}\left(\frac{\partial \mathscr{L}}{\partial w_i}, \frac{\partial \mathscr{L}}{\partial w_j}\right), U_{kl} \equiv \frac{\partial^2 \mathscr{L}}{\partial w_i \partial w_j} \right)$$







Uncertainty propagation

fraction task

(Reminder, given r: $\hat{f} = argr$

 Gong-Samaniego theorem (Gong and Samaniego 1981): With a into the MLE is asymptotically consistent with known variance!

$$\sigma_{\hat{f}}^2 = \sigma_{\hat{f},\mathsf{MLE}}^2 \left(1 + A_i C_{ij} A_j\right), A_i \equiv \frac{\partial \mathscr{L}}{\partial w_i \partial f}$$



• As an example of a downstream application, consider the mixture

$$\max_{f} \sum_{i=1}^{N_{test}} \log(fr(x_{\alpha}) + (1 - f)))$$

consistent estimator of the \hat{w}_i , and therefore of $\log r$ (under the wellspecified assumption), the estimator for f that we get by just plugging \hat{r}







That's it!



Sean Benevedes



$w_i f_i$ ensembles: the recipe

Ingredients

- *M* randomly initialized networks
- A training dataset with N_{train} samples each from q and p • (Optional) A dataset of size N_{test} for downstream applications
- 1. Train the f_i on the training dataset
- 2. Find \hat{w}_i using the MLC loss (introduced in Tito D'Agnolo et al. 1806.02350)
- 3. Compute $C_{ij} \equiv \text{Cov}(\hat{w}_i, \hat{w}_j)$ with analytic formulae (M-estimator) 4. Propagate uncertainties for downstream tasks (e.g. GS)











- Fully frequentist, no prior dependence
- Intuitive: just like nuisance parameters!
- Computationally efficient; no bootstrapping

$$w_i\}) = w_i f_i(x)$$





Backup slide: Bias

Two possible causes of bias in inference: 1. Model misspecification (M too small/networks poorly trained) 2. Breakdown of asymptotic regime (N too small and/or M too large)







Sean Benevedes





Backup slide: Why decorrelation?

- Better reason: In vanilla ensembles, uncorrelated ensemble members decrease variance for prediction (see e.g. Mehta et al. 2019); no reason to expect different behavior here
- Worse reason: Numerics! Minimization of the symMLC loss is way easier when covariance is diagonal



 $\frac{\partial \mathscr{L}}{\partial w_i \partial w_j} \propto \operatorname{Cov}\left(f_i, f_j\right)$



