Makov Chain Monte Carlo for PDF uncertainties

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Introduction

- We would like to have access to the underlying distribution of PDFs (PDF parameters) the **posterior distribution**
 - ▶ this would allow for reliable compution of errors for any PDF dependent quanties
- Posterior distribution too complicated to sample directly
 - need clever way to choose Monte Carlo samples
- Solution: construct the Monte Carlo samples via a Markov chain



Markov Chain Monte Carlo representation of the likelihood



with: $\pi(\mathbf{c}|\text{data})$ - posterior, $l(\text{data}|\mathbf{c})$ - likelihood, $p(\mathbf{c})$ - prior distribution, $\mathcal{N} = \int d\mathbf{c} \ l(\text{data}|\mathbf{c}) \ p(\mathbf{c})$ - normalization

• Likelihood given by the χ^2 function:

$$l(\text{data}|\mathbf{c}) \propto \exp\left(-rac{1}{2}\chi^2(\mathbf{c},D)
ight)$$

- Draw random samples from the posterior function: $\pi(\mathbf{c}|\text{data}) \rightarrow \{\mathbf{c_1}, \mathbf{c_2}, \dots, \mathbf{c_n}\}$
- Samples have to reproduce the expectation value and higher modes:

$$E\{\mathcal{O}(\mathbf{c})\} = \frac{1}{n} \sum_{i=1}^{n} \mathcal{O}(\mathbf{c}_{i})$$

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4 Use normal random walk Metropolis-Hastings

- set initial values of the parameters c₀
- ▶ propose set of new parameters: $\tilde{\mathbf{c}}_{i+1}$ proposal distribution: $q(\tilde{\mathbf{c}}_{i+1}, \mathbf{c}_i) = \mathcal{N}(\mathbf{c}_i, C_0)$
- compute acceptance probablity: $\pi(\tilde{a}, \dots, data)$

$$\alpha = \min\left(\frac{\pi(\mathbf{c}_i+1|\text{data})}{\pi(\mathbf{c}_i|\text{data})}, 1\right)$$

if $\mathcal{U}(0,1) < \min(1,\alpha)$: $\mathbf{c}_{i+1} = \tilde{\mathbf{c}}_{i+1}$ else: $\mathbf{c}_{i+1} = \mathbf{c}_i$



2 After N_0 samples switch to a **self learning proposal distribution**

$$\tilde{\mathbf{c}}_{i+1}$$
 proposed from $q(\tilde{\mathbf{c}}_{i+1}, \mathbf{c}_i) = (1-\beta)\mathcal{N}\left(\mathbf{c}_i, \text{scale} \cdot \overline{C}_i\right) + \beta \mathcal{N}(\mathbf{c}_i, C_0)$

with self learned covariance \overline{C}_i

- ▶ $0 \le \beta \le 1$ controls the impact of the 'learned' proposal
- **③** Reset self learned proposal distribution to boost convergence
 - this reduces the impact of the starting point

[H. Haario et al.: "An adaptive Metropolis algorithm", Bernoulli 7.2 (Apr. 2001)]

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- Because of the dependence on the previous step MCMC samples are correlated.
- Hence, standard Monte Carlo error estimates do not work $\sigma_{MC}^2 = \frac{1}{n-1} \sum_{k=1}^n (X_k - \hat{\mu})^2$
 - these severely underestimate the true uncertainties
- Since every new sample depends on the previous one the gain in information is reduced
- This is what is called **autocorrelation**
 - twice the autocorrelation-time τ estimates the number of links in the chain until the next independent sample is drawn



severe substructure

- Lattice QCD has several methods to deal with this problem
- One example is the $\Gamma\text{-method}$
 - allows to estimate the autocorrelation time, τ_{int} , directly from the chain
 - used to enlarge error estimates to eliminate bias: $\sigma_{MCMC}^2 = 2\tau_{int}\sigma_{MC}^2$
 - or filter/thin the time series to get uncorrelated samples

Monte Carlo errors with less errors.

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Abstract

We explain in detail how to estimate mean values and assess statistical errors for arbitrary functions of elementary observables in Monte Carlo simulations. The method is to estimate and sum the relevant autocorrelation functions, which is argued to produce more certain error estimates than binning techniques and hence to help toward a bitrary ambidution of manimic simulations. As off-arbitrary interaction

[arXiv:hep-lat/0306017]



using 300 samples directly

thinning 10^4 samples to a total of 300

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Proton PDF analysis

Experimental data

- DIS: 1660 points
 - ▶ HERA NC/CC
 - \blacktriangleright NMC F_2
 - \blacktriangleright BCDMS F_2

- DY: 324 points
 - ► CDF & DØ
 - ► CMS
 - ► ATLAS
 - ► LHCb

• Total: 1984 points





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PDF parameters

$$\begin{aligned} xf_i(x,Q_0) &= \mathbf{c_0} x^{\mathbf{c_1}} (1-x)^{\mathbf{c_2}} (1+\mathbf{c_3}\sqrt{x}+\mathbf{c_4}x) \\ & \mathbf{u_v} & \to & c_1 \quad c_2 \quad c_4 \\ & \mathbf{d_v} & \to & c_1 \quad c_2 \quad c_4 \\ & \overline{\mathbf{u}} + \overline{\mathbf{d}} & \to & c_1 \quad c_2 \quad c_4 \\ & \mathbf{s} + \overline{\mathbf{s}} & \to & c_0 \\ & \mathbf{g} & \to & c_0 \quad c_1 \quad c_2 \quad c_3 \quad c_4 \end{aligned}$$
Total: 15 parameters

PDF parameters

$$\begin{split} xf_i(x,Q_0) &= \mathbf{c_0} x^{\mathbf{c_1}} (1-x)^{\mathbf{c_2}} (1+\mathbf{c_3}\sqrt{x}+\mathbf{c_4}x) \\ & \mathbf{u_v} \quad \rightarrow \quad c_1 \quad c_2 \quad c_4 \\ & \mathbf{d_v} \quad \rightarrow \quad c_1 \quad c_2 \quad c_4 \\ & \mathbf{\overline{u}} + \mathbf{\overline{d}} \quad \rightarrow \quad c_1 \quad c_2 \quad c_4 \\ & \mathbf{s} + \mathbf{\overline{s}} \quad \rightarrow \quad c_0 \\ & \mathbf{g} \quad \rightarrow \quad c_0 \quad c_1 \quad c_2 \quad c_3 \quad c_4 \end{split}$$
Total: 15 parameters

 $\bullet\,$ Down-valence distribution becomes independent of c_4

$$\lim_{\mathbf{c_4} \to \infty} x d_v(x, Q_0) = \lim_{\mathbf{c_4} \to \infty} c_0 x^{\mathbf{c_1}} (1-x)^{\mathbf{c_2}} [\mathbf{c_4} x]$$
$$= \tilde{\mathbf{c}}_0 x^{\mathbf{c_1}+1} (1-x)^{\mathbf{c_2}}$$

• Need to onstrain c_4 by Uniform Prior:

$$-1000 \le \mathbf{c_4} \le 10.000$$

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PDF parameters

$$xf_{i}(x,Q_{0}) = \mathbf{c_{0}}x^{\mathbf{c_{1}}}(1-x)^{\mathbf{c_{2}}}(1+\mathbf{c_{3}}\sqrt{x}+\mathbf{c_{4}}x)$$

$$\mathbf{u_{v}} \rightarrow c_{1} \quad c_{2} \quad c_{4}$$

$$\mathbf{d_{v}} \rightarrow c_{1} \quad c_{2} \quad c_{4}$$

$$\mathbf{\overline{u}}+\mathbf{\overline{d}} \rightarrow c_{1} \quad c_{2} \quad c_{4}$$

$$\mathbf{s}+\mathbf{\overline{s}} \rightarrow c_{0}$$

$$\mathbf{g} \rightarrow c_{0} \quad c_{1} \quad c_{2} \quad c_{3}$$
Total: 15 parameters

Hyperparameters

- Proposals: Adaptive Metropolis Hastings
- 36 independent chains with 479k samples each
 - burn-in phase: 140k samples
 - Total: 17 million samples
- removing autocorrelation (thinning rate: 3000) and burn-in:

Total: 4068 uncorrelated samples

 χ^2 /d.o.f. = 2380.25/1969 = 1.20

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Thermalization



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Thermalization



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χ^2 /DATA DATA SET Ref. DATA POINTS DIS HERA σ_{red} neutral current [54]10391.26HERA σ_{red} charged current [54]81 1.08BCDMS F_2 proton [135]3391.09NCM F_2 proton [136]2011.54DIS total 16601.25DY CDF Z-rapidity [137]281.10DØ Z-rapidity [138]280.60ATLAS Z p_T 8 TeV (M_{ll}) [139]44 1.06ATLAS Z p_T 8 TeV (y_Z) [139]48 0.65CMS $Z p_T 8$ TeV [140]280.46CMS double diff. 2011 7 TeV [141]88 1.02LHCb $W^{\pm}, Z \rightarrow \mu$ 7 TeV [142]291.07LHCb $W^{\pm}, Z \rightarrow \mu 8$ TeV [143]311.18DY total 3240.91Total 1984**1.20** (per dof)

Description of Experimental Data

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Pairwise correlations

Definition of Uncertainties

From Samples to PDF-Uncertainties

We want to define **confidence interval** for observable $\mathcal{O}(c)$

$$\mathcal{O}_{-} \leq \mathcal{O} \leq \mathcal{O}_{+}$$

- α %-symmetric (MCSE):
 - central value: mean $\frac{1}{N} \sum_{i}^{N} O_{i}$
 - ▶ lower (upper) bound: mean ± standard deviation
- α %-asymmetric (Percentile/NNPDF-like):
 - central value: samle with minimal χ^2 value or 50th percentile
 - lower (upper) bound: 16th (84th) percentile of distribution of samples
- Cumulative χ^2 :
 - central value: sample with minimal χ^2 value
 - ▶ lower (upper) bound: keep only the samles including lowest α % of samples







Comparison with Hessian – Gaussian parameters

- Cumulative χ^2 : $\Delta \chi^2_{90\%} = 22$
- Hessian Method $\Delta \chi^2 = 22$



Comparison with Hessian – non-Gaussian parameters

- Cumulative χ^2 : $\Delta \chi^2_{90\%} = 22$
- Hessian Method $\Delta \chi^2 = 22$



Markov Chain Monte Carlo

• Access uncertainties without approximations

• Γ-method to deal with autocorrelation

Proton PDF extraction

- 15 parameters, DIS & DY data
- χ^2 /d.o.f. = 2380.25/1969 = 1.20
- Generated samples: 17 million
- Result: 4068 uncorrelated samples

Definition of Uncertainties

- Confidence limits using χ^2 -values
- Tolerance estim. from χ^2 -samples
- comparison with Hessian
 - agreement for Gaussian params.
 - differences for non-Gaussian

Nuclear PDF analysis (only Pb)

Data & parameterization

- Data:
 - W and Z boson from pPb LHC
 - heavy quark(onia) from pPb LHC
 - Neutrino DIS
 - Total no data points: 1488
- Parametrization:
 - $xf_i(x,Q_0) = \mathbf{c_0} x^{\mathbf{c_1}} (1-x)^{\mathbf{c_2}} (1+\mathbf{c_3}\sqrt{x}+\mathbf{c_4}x)$
 - ▶ 10 open parameters (6 valence, 2 gluon, 2 sea)
- Theory prediction @ NLO
 - ▶ HQ with Crystal Ball (data driven)

Hyperparameters

- Proposals: Adaptive Metropolis Hastings
- 9 independent chains with total of 3.8 million samples
 - ▶ burn-in phase: $\approx 60k 80k$ for each chain
 - Total: 6729 uncorrelated samples

Combined chain

- After removing autocorrelation all 9 chains are thinned and combined
- No independent sample: $N_{tot} = 6729$.



Combined chain



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Combined chain: autocorrelation

• Atocorrelation-time τ for te combined chain (using the Γ -method)







Pb PDF: Different thinning rates



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Conclusion

Markov Chain Monte Carlo

• Access uncertainties without approximations

 Γ-method to deal with autocorrelation

Proton PDF extraction

- 15 parameters, DIS & DY data
- χ^2 /d.o.f. = 2380.25/1969 = 1.20

Nuclear (Pb) PDF extraction

• 10 parameters, CC DIS, W/Z, HQ data

- Generated samples: 17 million
- Result: 4068 uncorrelated samples

- Generated samples: 3.8 million
- Result: 6729 uncorrelated samples

Definition of Uncertainties

- Confidence limits using χ^2 -values
- Tolerance estim. from χ^2 -samples
- comparison with Hessian
 - agreement for Gaussian params.
 - ▶ differences for non-Gaussian

BACKUP SLIDES

Time series



