Exaggerated Certainty: One Statistician's Take on EoE

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Summary, not outline

- Uncertainty is routinely underestimated.
- EoE is a very nice idea.
- Statisticians usually analyze one experiment at a time.
- Exceptions are in Meta Analysis and in auxiliary data and paradata in Survey Sampling.
- Wilks theorem hinges on having "asymptotically unbiased" estimating equations.
- When adding terms to log-likelihoods focus on bias of the resulting estimating equations

Summary concluded

- A toy example shows importance of independence assumption.
- If you insist on frequency guarantees do you average over both data sets or condition on auxiliary data?
- Should all systematic errors be treated as if error rates could have frequentist control?
- Bayesian posteriors smoothly incorporate auxiliary information about common parameters into new analysis.
- Lots of study needed to deal with degrees of freedom.
- Some work on degrees of freedom in smoothing problems.

Simple vision

- Two experiments: main and auxiliary
- Two models: primary has parameters θ and γ .
- Auxiliary had parameters γ and say τ .
- \blacktriangleright In primary experiment θ,γ is not identifiable or poorly identified
- Or just "borrow strength" from the auxiliary experiment to reduce overall uncertainty about γ and therefore θ.
- Natural but usually impractical frequentist strategy: put the two data sets together and fit a model to the combined data.
- Log-likelihood is, assuming independence of experiments, (m, a for main and auxiliary)

$$\ell_{\text{comb}}(\theta, \gamma, \tau) = \ell_m(\theta, \gamma) + \ell_a(\gamma, \tau)$$

 Same marginal models for the two experiments, more parameters to fit.

Standard HEP alternative

- Find normal approximation to estimation error for γ in auxiliary experiment.
- Model $\hat{\gamma}_a$ as Normal (γ, σ^2) with σ known.
- Actually σ^2 estimated usually via standard MLE with inverse of negative Hessian .
- Call this estimate V.
- ► Now do "natural", less "impractical" thing and treat the data set as the primary data together with Ŷ_a.

Standard HEP alternative 2

- ▶ Joint model is original model for main expt with independent Gaussian model for $\hat{\gamma}_a$.
- Log-likelihood is replaced by

$$\ell_1(\theta,\gamma) = \ell_m(\theta,\gamma) - \frac{(\gamma - \hat{\gamma}_a)^2}{2V}$$

Estimate parameters by MLE – solve system of *likelihood* equations:

$$rac{\partial}{\partial heta} \ell_1(heta, \gamma) = 0$$

and

$$\frac{\partial}{\partial \gamma} \ell_1(\theta, \gamma) - \frac{\partial}{\partial \gamma} \frac{(\gamma - \hat{\gamma}_{a})^2}{2V} = 0$$

 Key point, fixed by Glen and Enzo, this system of equations is biased.

Estimating equations, Wilks Theorem

Real likelihood equations are

$$\mathsf{Score} = \mathsf{0}$$

 Basic ingredient of all MLE theory: these equations are unbiased

$$\mathbb{E}_{\theta,\gamma}\left\{
abla_{ heta,\gamma}\ell(heta,\gamma)
ight\} = 0$$

- Wilks theorem needs this to centre the χ^2 approximation.
- So you want for HEP strategy:

$$\mathbb{E}_{\theta,\gamma,\tau} \left\{ \ell_1(\theta,\gamma) \right\} = 0$$

which means we need

$$\mathbf{E}_{\gamma,\tau}\left\{\frac{\gamma-\hat{\gamma}_{a}}{V}\right\}=\mathbf{0}$$

• The presence of V makes this unlikely to be right.

Errors on Errors fixes the problem

- Most classical case for statisticians: γ is a mean parameter in a Gaussian linear model.
- Simplest: Y_1, \ldots, Y_n independent $N(\gamma, \tau)$. $(\tau = \sigma^2)$
- Auxiliary likelihood ($\tau = \sigma^2$, $S = \sum (Y_i \overline{Y})^2$)

$$\ell_{\mathsf{a}}(\gamma, \tau) = -rac{n}{2}\log(\tau) - rac{n(ar{Y} - \gamma)^2 + S}{2 au}$$

• The MLE of γ, τ is then

$$\hat{\gamma}_{\mathsf{a}} = ar{Y} \quad \hat{\tau} = S/n.$$

 \blacktriangleright Wilks theorem: for true value of γ

$$\Lambda(\gamma) = 2 \left\{ \ell_{a}(\hat{\gamma}, \hat{\tau}) - \ell(\gamma, \hat{\tau}(\gamma)) \right\} \approx \chi_{1}^{2}$$

where $\hat{\tau}(\gamma) = \sum (Y_i - \gamma)^2 / n$.

Enzo and Glen summary for me not audience

- ► Glen and Enzo treat the auxiliary data not as Y₁,..., Y_n but as Ŷ, Ŷ and write down the exact likelihood for this data.
- The actual distribution of $\hat{\tau}$ is given by

$$\frac{n\hat{\tau}}{\tau} \sim \chi_{\nu}^2$$

where $\nu = n - 1$ is the degrees of freedom.

- And $\hat{\tau}$ is independent of $\hat{\gamma}_{a} = \bar{Y}$
- Degrees of freedom n-1; uses unbiased variance estimate.
- Exact auxiliary likelihood is as in Enzo and Glen's work: normal(γ, τ)× independent chi-squared.
- Corresponding r has $\nu = n 1 = \frac{1}{2r^2} = \frac{1}{2\epsilon^2}$

• Largest r, smallest df (for linear model) are df = 1 and $r^2 = \frac{1}{2}$.

A very small example

- Goal: to illustrate extent to which independence assumption and gamma assumption could be flawed.
- ► To show these assumptions may not be really crucial.
- On-off. Main experiment observe Y ~Poisson(s + b).
 Unidentified model for On.
- Auxiliary experiment observe X ~Poisson(b).
- We have $\hat{b}_a = X$. Estimated variance of \hat{b} is V = X from neg Hessian.
- Three likelihoods to compare: ℓ_{comb} , ℓ_{HEP} and ℓ_{EoE} .
- Do low statistics strong (apparent) signal example: Y = 40, X = 10.



Signal

Coverage probabilities



True Signal

Slides commentary

- Even in this small example impacts not huge at coverage 95%.
- Choice of r matters to coverage.
- Discrete data gives unsmooth coverage probs
- The profile score for the background can have two local maxima in the permissible parameter space.

Meta Analysis and Auxiliary Information

- Statisticians have a long history of focusing on analyzing each individual data set as if it were the only data to be had.
- In Meta Analysis several experiments / studies are analyzed together parameters of interest in common but different nuisance parameters in each.
- Typical problem: different biases for the estimates of the common parameters of interest. Handled in a variety of somewhat unsatisfactory ways.
- Reminds me of problem of nearly 20 years ago: parton distribution function estimates differ between experimental groups by huge number of standard errors.
- In survey samples it is common to have control measurements produced from census data or to have paradata about respondents (e.g. how hard you had to work to get a response)
- ► In the medical literature use of empirical likelihood.

What needs doing; things to investigate maybe

- Models with many parameters: risk of accumulating bias.
- Maximum likelihood underestimates variance parameters.
 - In linear regression MLEs of variances are adjusted to remove bias.
 - Variance estimates are independent of mean estimates, have gamma distributions, and known degrees of freedom.
- In other models the location estimates are skewed and not independent of the variability estimates. We need more study of the impact of multiple slightly skewed estimates.
- There is work by, for example Zou, Hastie, and Tibshirani on degrees of freedom in variable selection.

Bayes has its advantages

- I am skeptical of the wisdom of treating theory uncertainty via frequentist methods.
- That said it seems the theory uncertainties are not cleanly handled any other way.
- The posterior from the auxiliary experiment captures any nuance in the distribution of the uncertainty in *γ*.
- So routine use of Bayesian methods would give sensible combinations of evidence.
- Bayes susceptible, in complex models, to hidden strengths of assumptions encoded in priors.
- Computer experiment literature full of uncertainty assignments to things like theory calculations where there is no frequentist uncertainty.

Three References

- H Zou, T Hastie, R. Tibshirani (2007) On the "degrees of freedom" of the LASSO. Annals of Statistics, 25,2173-2192.
- C Chen, P Han, F He (2022) Improving main analysis by borrowing information from auxiliary data. *Statistics in Medicine*, **41**, 567-579.
- M Borenstein, L Hedges, J Higgins, H Rothstein (2021) Introduction to Meta-Analysis, 2nd ed. Wiley,