

30/10/2024



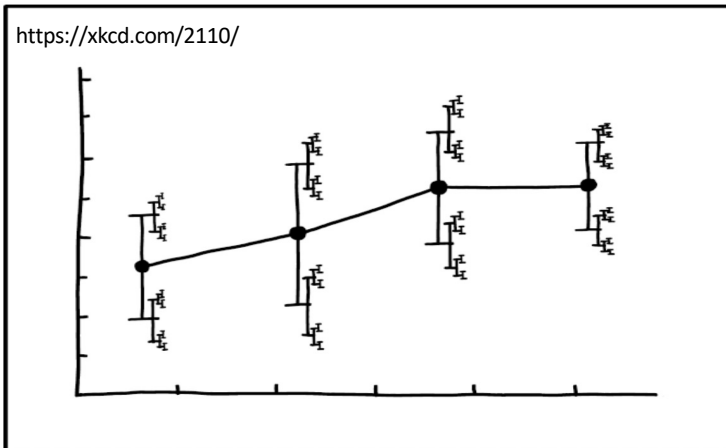
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Errors-on-errors (Gamma Variance Model)

Enzo Canonero, Glen Cowan

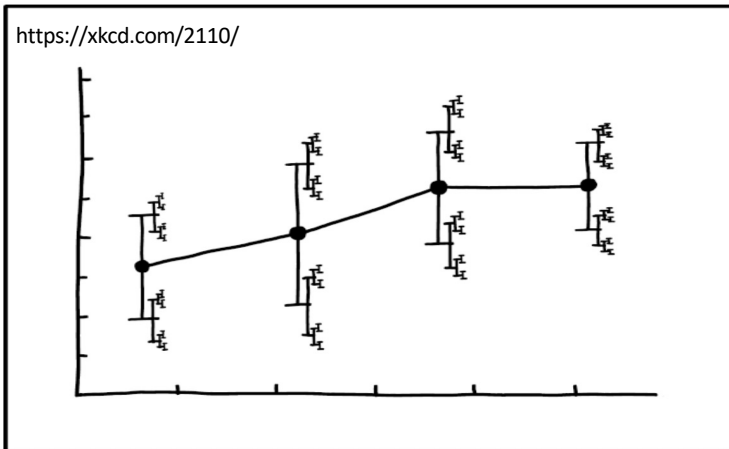
“Uncertain systematics”

“Errors-on-errors”



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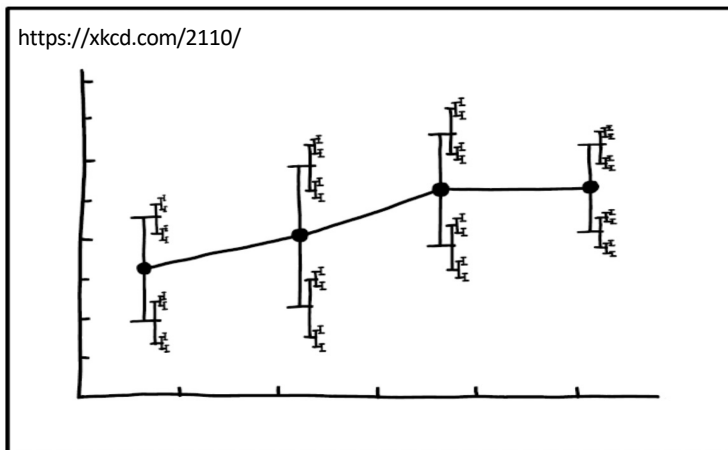
“Errors-on-errors”



- 1) Some *systematic uncertainties* can be well estimated:
 - Related to size of control measurements dataset
 - Related to size of MC sample

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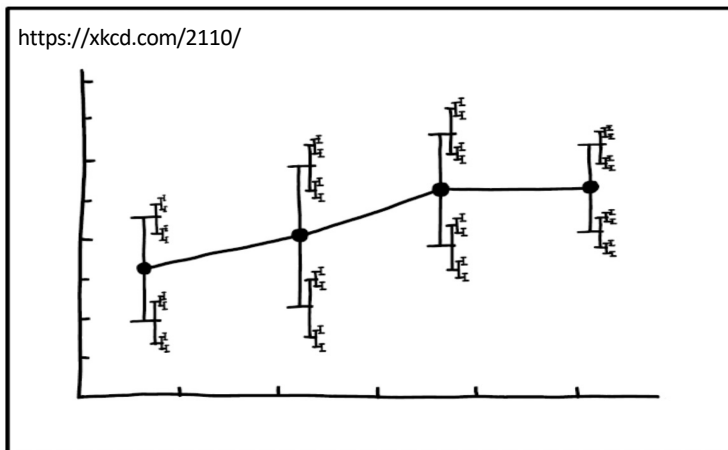
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- 1) Some **systematic uncertainties** can be well estimated:
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- 2) But they can also be **quite uncertain**:
 - Theory systematics
 - Two points systematics

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→ Non-trivial effects

Formulation of the problem



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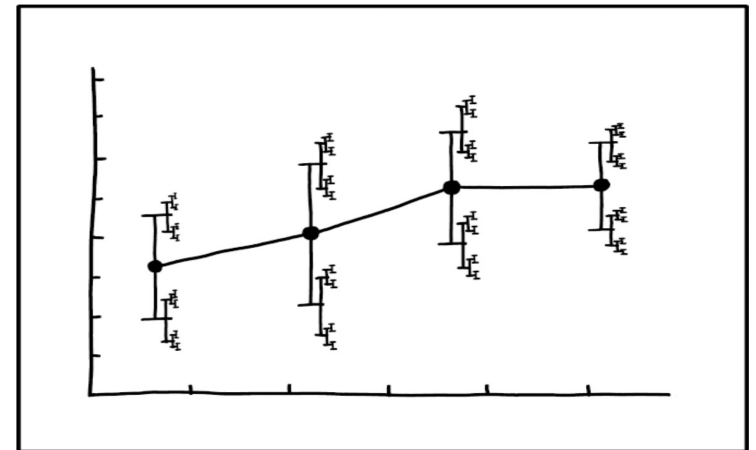
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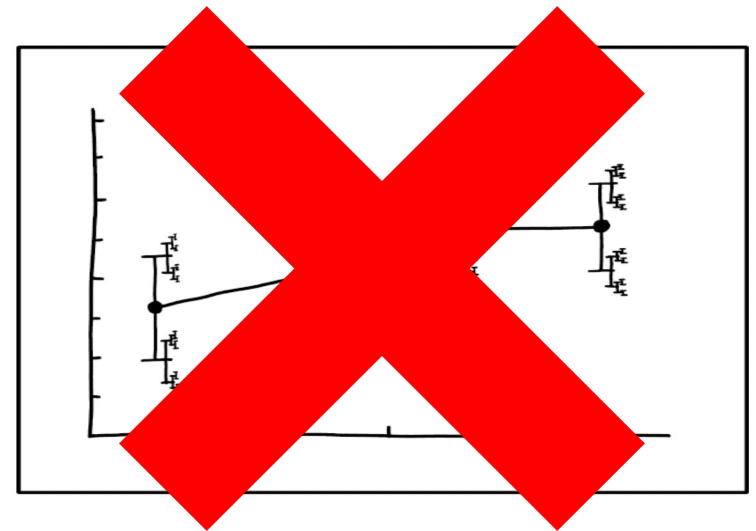


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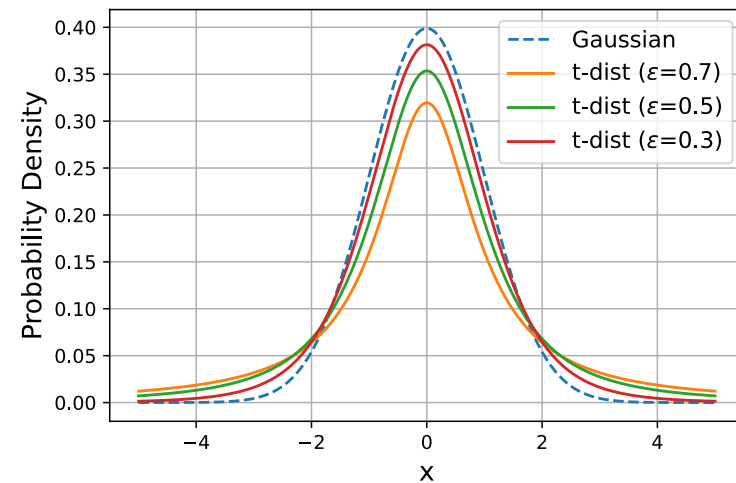
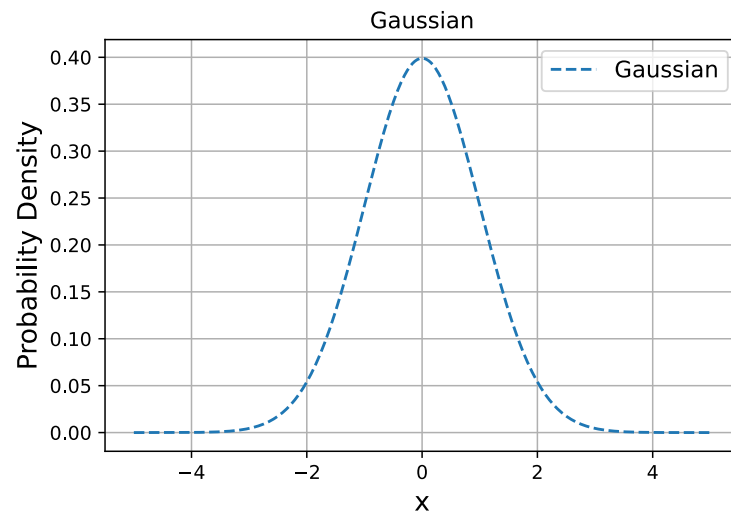
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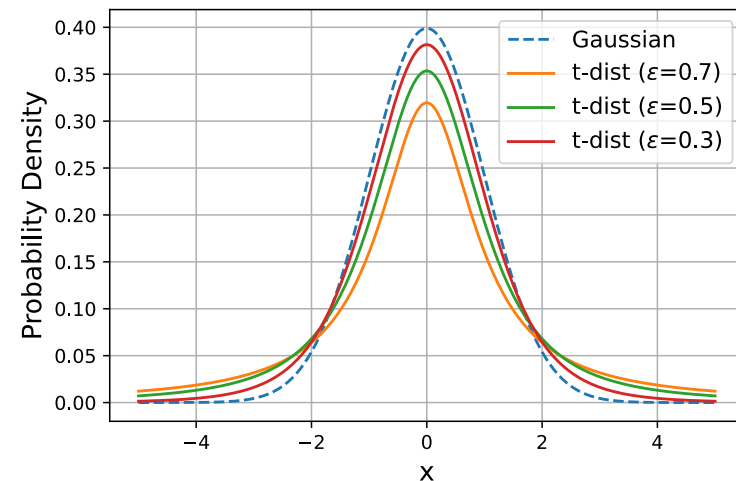
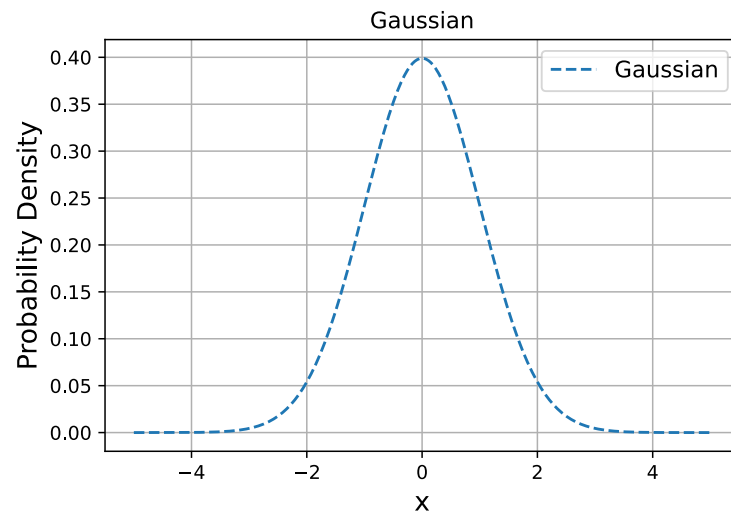
Formulation of the problem

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- We use a Student's t-distribution instead of a Gaussian distribution.

Gamma Variance Model (GVM)

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ε = error-on-error parameter

$\varepsilon = 0.3$ means 30%
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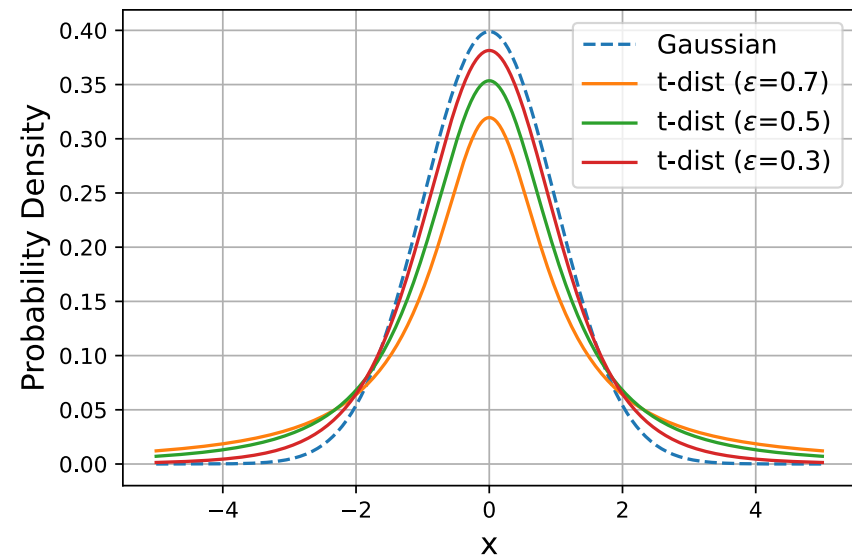
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- The student t distribution has longer tails for larger values of ε
- Outliers are **de-weighted** in the fit/combinaton



For more details: [Eur. Phys. J. C \(2019\) 79:133](#)

Formal interpretation



- The Updated log-likelihood

$$\log L_P(\boldsymbol{\mu}, \boldsymbol{\theta}) = \log P(\mathbf{y}|\boldsymbol{\mu}, \boldsymbol{\theta}) - \frac{1}{2} \sum_i \left(1 + \frac{1}{2\varepsilon_i^2} \right) \log \left(1 + 2\varepsilon_i^2 \frac{(u_i - \theta_i)^2}{v_i} \right)$$

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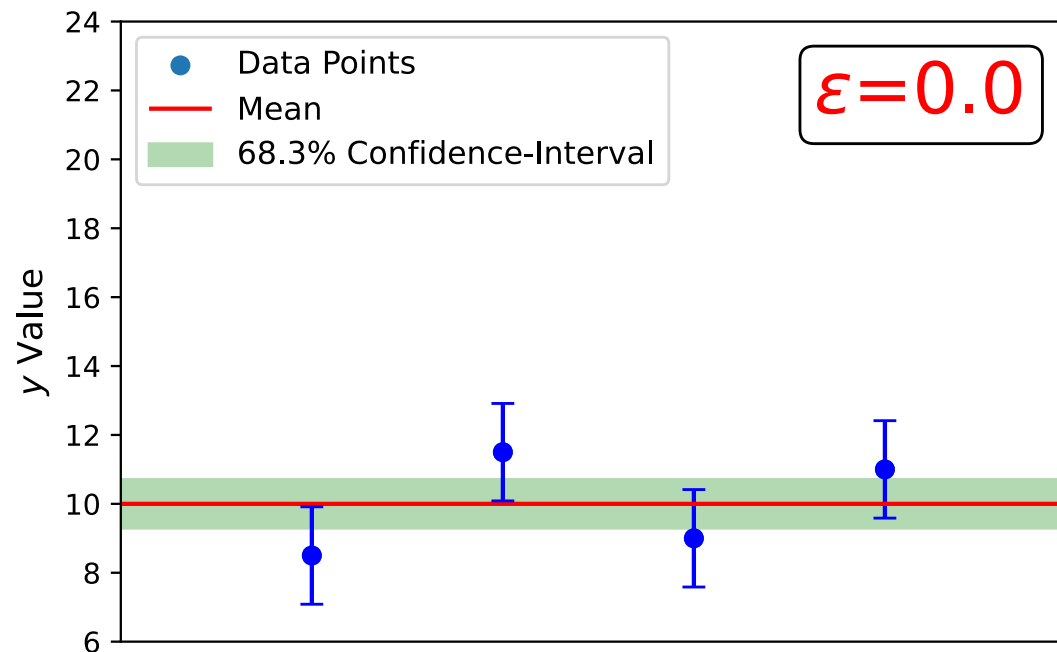
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Sensitivity to outliers

- Suppose we want to average 4 measurements
- **Syst errors** all have equal *errors-on-errors* ϵ :

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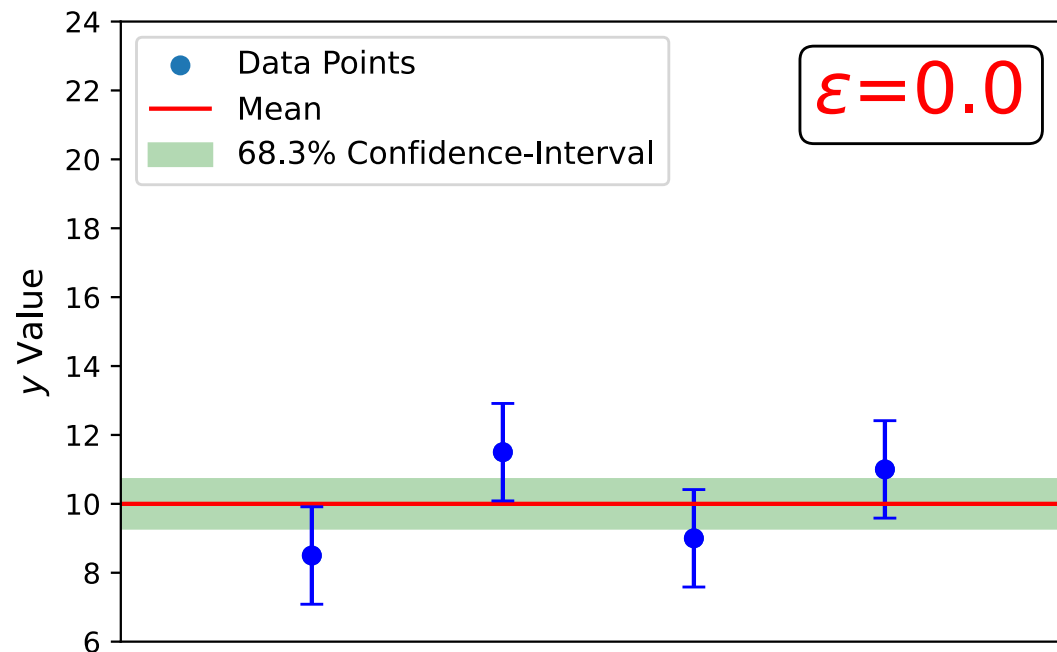


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Measurements
internally compatible

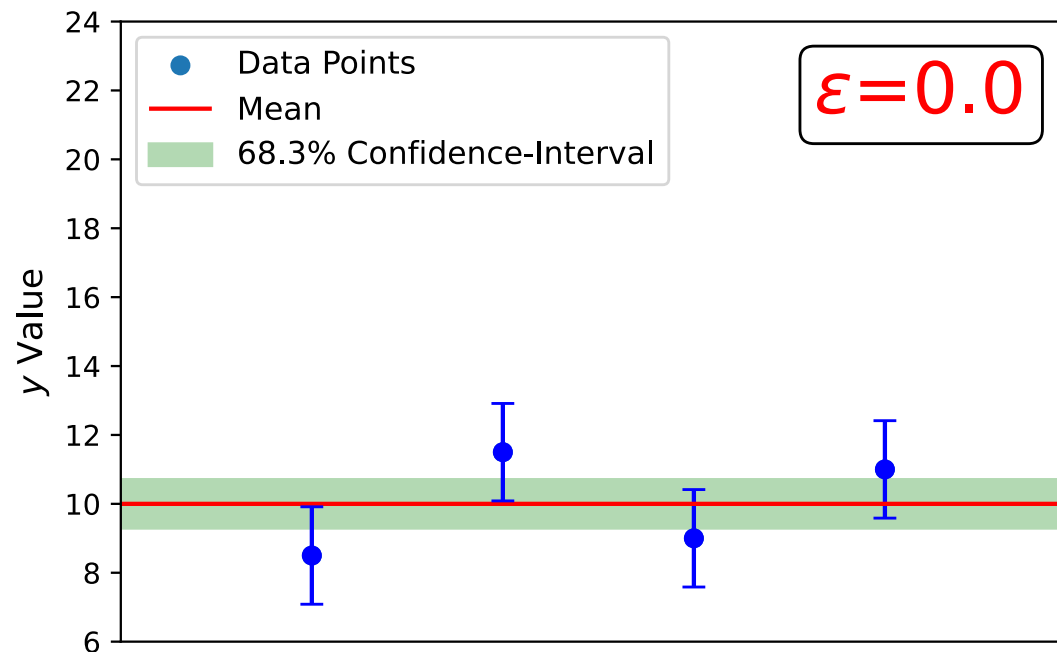


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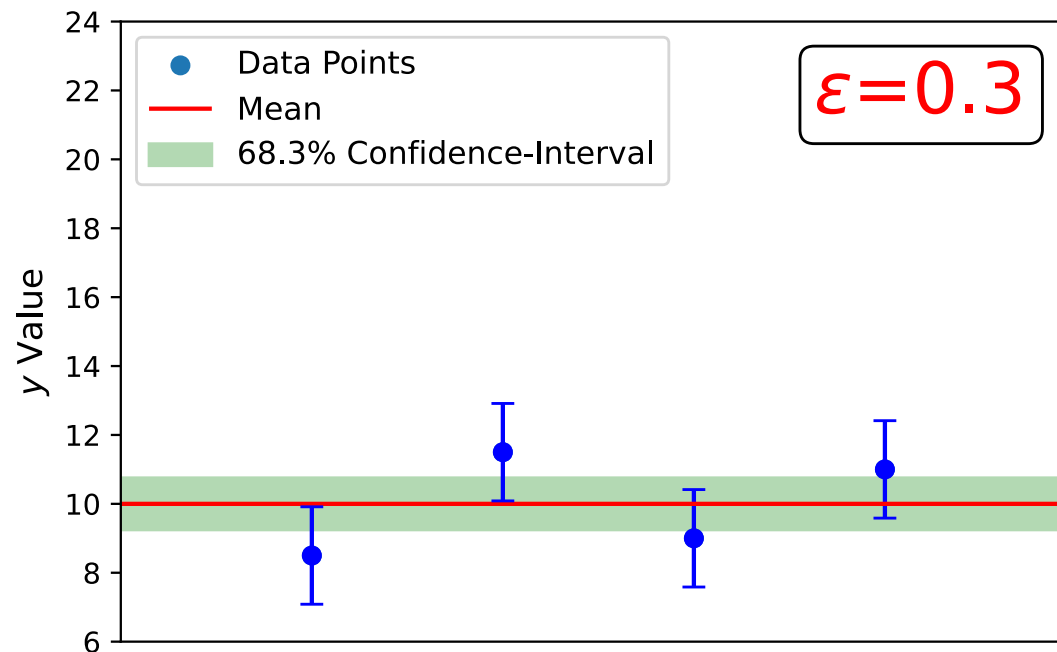
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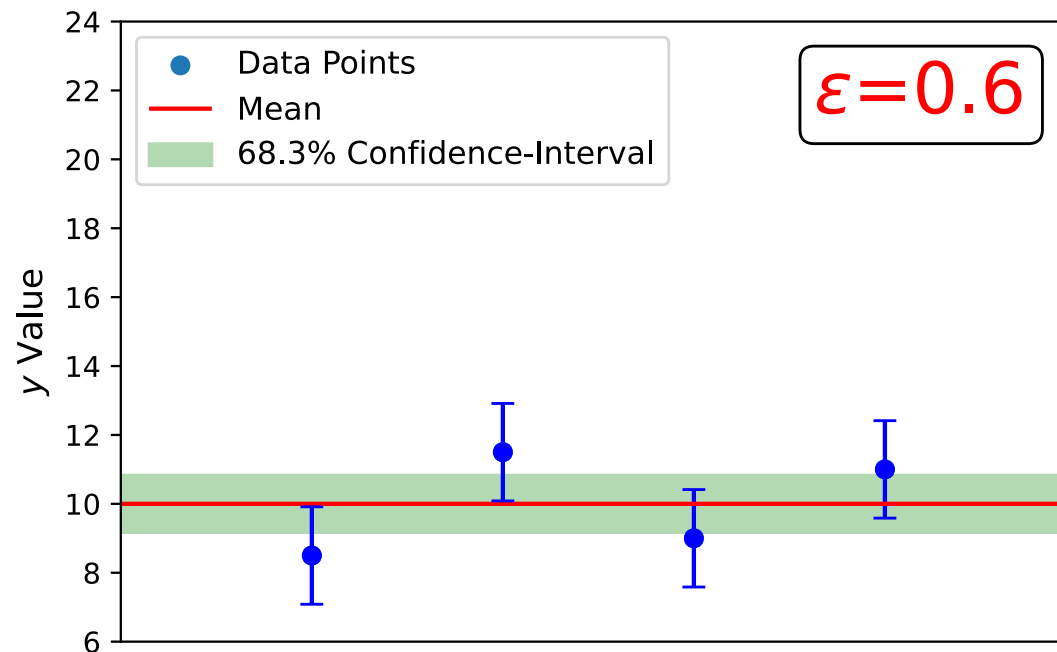
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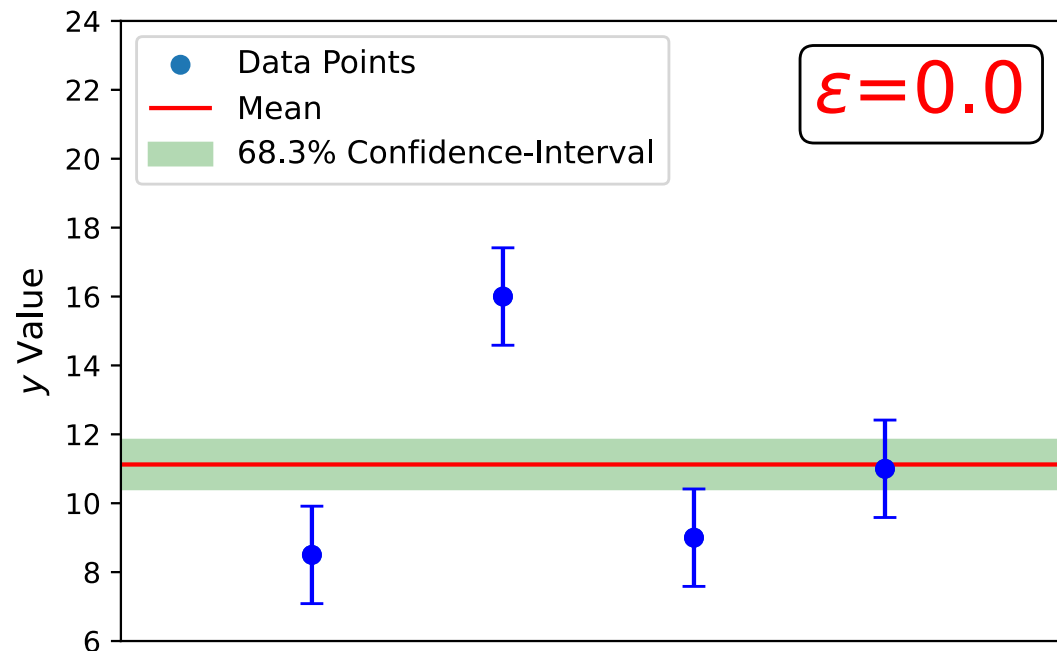
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If data are internally compatible results are only slightly modified

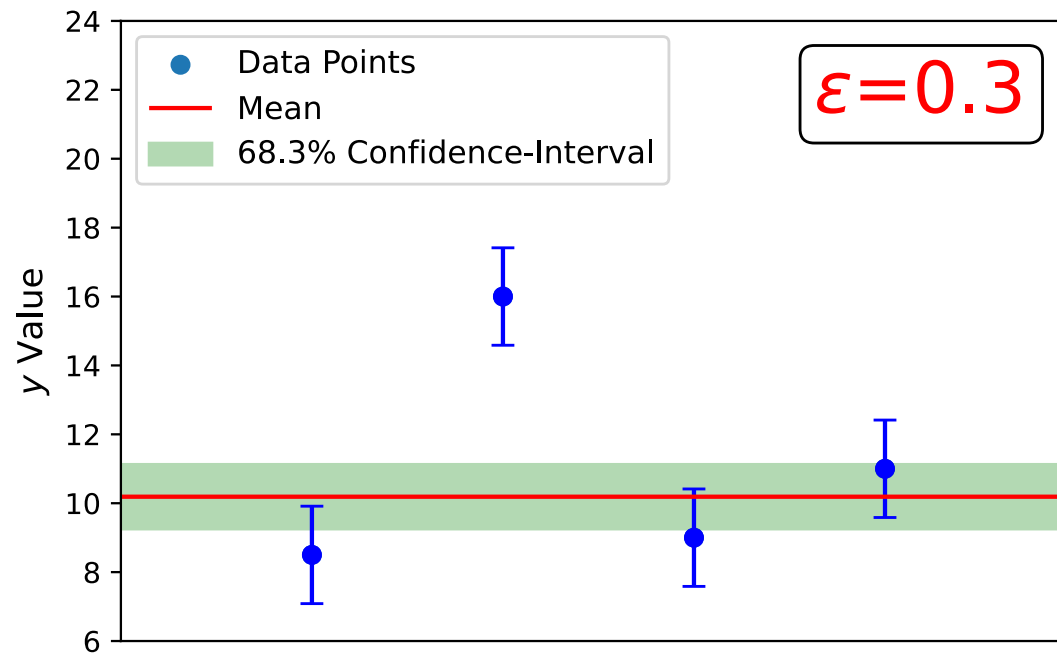
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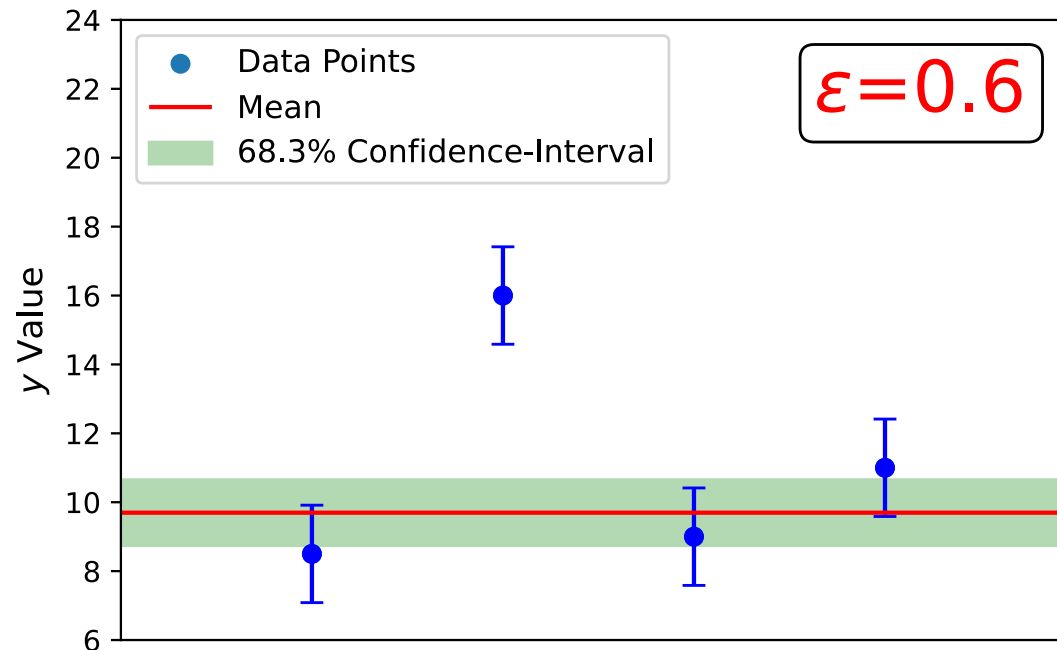
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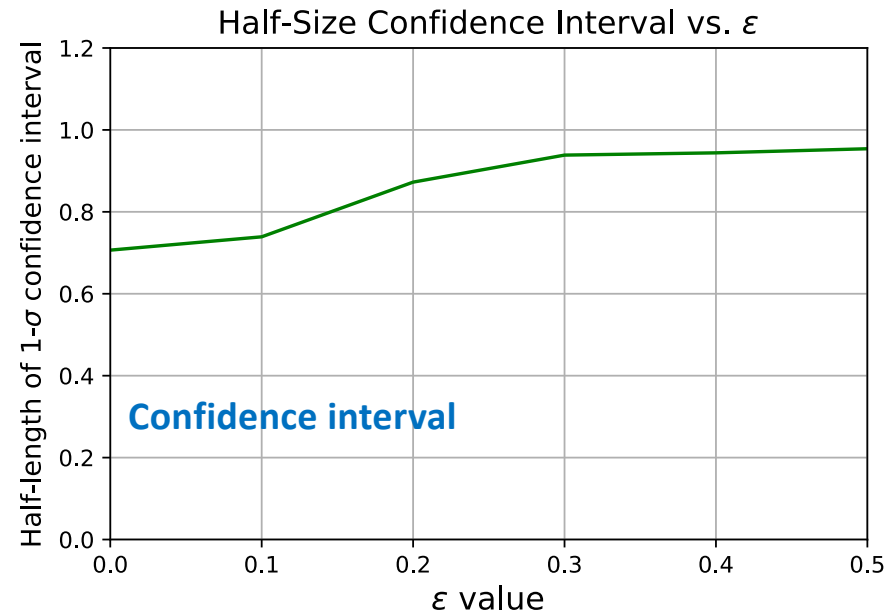
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1. The outlier is de-weighted
2. **The confidence interval grows**: the model treats internal incompatibility as an additional source of uncertainty

Conclusion: The model is sensitive to internal compatibility of the data

Beyond Wilks' theorem

- We use the likelihood to construct test statistics to compute CI (Confidence Intervals) or evaluate GOFs (Goodness-Of-Fits):
- Profile likelihood ratio (CI):

$$w_{\mu} = 2[\log L(\hat{\mu}, \hat{\theta}) - \log L(\mu, \hat{\theta})]$$

- goodness-of-fit (chi2):

$$q = -2 \log L(\hat{\mu}, \hat{\theta})$$

- To compute CIs and GOFs often asymptotics properties are used:

* P = number POIs

N = number of measurements

$$W_{\mu} \sim \chi_P^2$$

$$q \sim \chi_{N-P}^2$$

- Problem:** Our likelihood is not quadratic; test statistics deviate from asymptotic formulas by $\mathcal{O}(\epsilon^2)$, as the likelihood is not anymore quadratic because of the presence of terms like $\left(1 + \frac{1}{2\epsilon_i^2}\right) \log \left(1 + 2\epsilon_i^2 \frac{(u_i - \theta_i)^2}{\sigma_{u_i}^2}\right)$

- Equivalence:** Equivalent problem of having a small sample size: $n_{eff} = 1 + \frac{1}{2\epsilon^2}$

Effective sample size

Bartlett Corrections



Modify the test statistic t so that its distribution is closer to a χ^2 :

[Bartlett](#)

[Gauss M. Cordeiro, Francisco Cribari-Neto](#)

$$t \longrightarrow t^* = t \frac{N_{dof}}{E[t]}$$

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[Canonero, Cowan, Brazzale](#)

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**Bartlett correction originally developed for small sample sizes problems*

- All these expectation values can be computed analytically up to $\mathcal{O}(\epsilon^4)$:

$$E[q] = N_{dof} + b_q$$

$$b_q = \sum_s \left(3 - 4 \frac{\sigma_{\hat{\theta}_s}^2}{\sigma_s^2} - \frac{\sigma_{\hat{\theta}_s}^4}{\sigma_s^4} \right) \epsilon_s^2 + \mathcal{O}(\epsilon^4)$$

- We have recently computed $E[q]$ and $E[\omega_\mu]$ up to $\mathcal{O}(\epsilon^4)$

- BLUE (Best Linear Unbiased Estimators) approach to fits:

$$\chi^2 = \sum_i (y_i - \mu) W_{ij}^{-1} (y_j - \mu)$$

$$W_{ij} = \mathbf{V}_{ij} + \mathbf{U}_{ij}^{(syst)}$$

$$-2\log L = \chi^2$$

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- $\mathbf{U}_{ij}^{(syst)}$: Covariance matrix induced by systematic source.
- $\mathbf{U}_{ij}^{(syst)} = \sum_s \mathbf{U}_{ij}^{(s)}$

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- Systematic uncertainties can induce ± 1 correlations: $\rho_{ij}^{(s)} = \frac{V_{ij}^{(s)}}{\sqrt{V_{ii}^{(s)} V_{jj}^{(s)}}} = \frac{\Gamma_i^s \Gamma_j^s}{|\Gamma_i^s| |\Gamma_j^s|} = \pm 1$

Non-Trivial Correlations in Combinations



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- The new matrix $\mathbf{C}_{ij}^{(s)}$ is defined as

$$\mathbf{C}_{ij}^{(s)} = \sigma_s^2 \rho_{ij}^{(s)}$$

$$\chi^2 = \sum_{ij} \left(y_i - \mu - \Gamma_i^s \boldsymbol{\theta}_s^{(i)} \right) V_{ij}^{-1} \left(y_j - \mu - \Gamma_j^s \boldsymbol{\theta}_s^{(j)} \right) + \sum_{ij} \left(u_s^{(i)} - \boldsymbol{\theta}_s^{(i)} \right) \mathbf{C}_{ij}^{(s)-1} \left(u_s^{(j)} - \boldsymbol{\theta}_s^{(j)} \right)$$

- The new matrix $\mathbf{C}_{ij}^{(s)}$ is defined as

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- Generalization to errors-on-errors framework:

$$\chi^2 = \sum_i \left(y_i - \mu - \Gamma_i^s \boldsymbol{\theta}_s^{(i)} \right) V_{ij}^{-1} \left(y_j - \mu - \Gamma_j^s \boldsymbol{\theta}_s^{(j)} \right) + \left(N + \frac{1}{2\epsilon_s^2} \right) \log \left(1 + 2\epsilon_s^2 \sum_{ij} \left(u_s^{(i)} - \boldsymbol{\theta}_s^{(i)} \right) \mathbf{C}_{ij}^{(s)-1} \left(u_s^{(j)} - \boldsymbol{\theta}_s^{(j)} \right) \right)$$

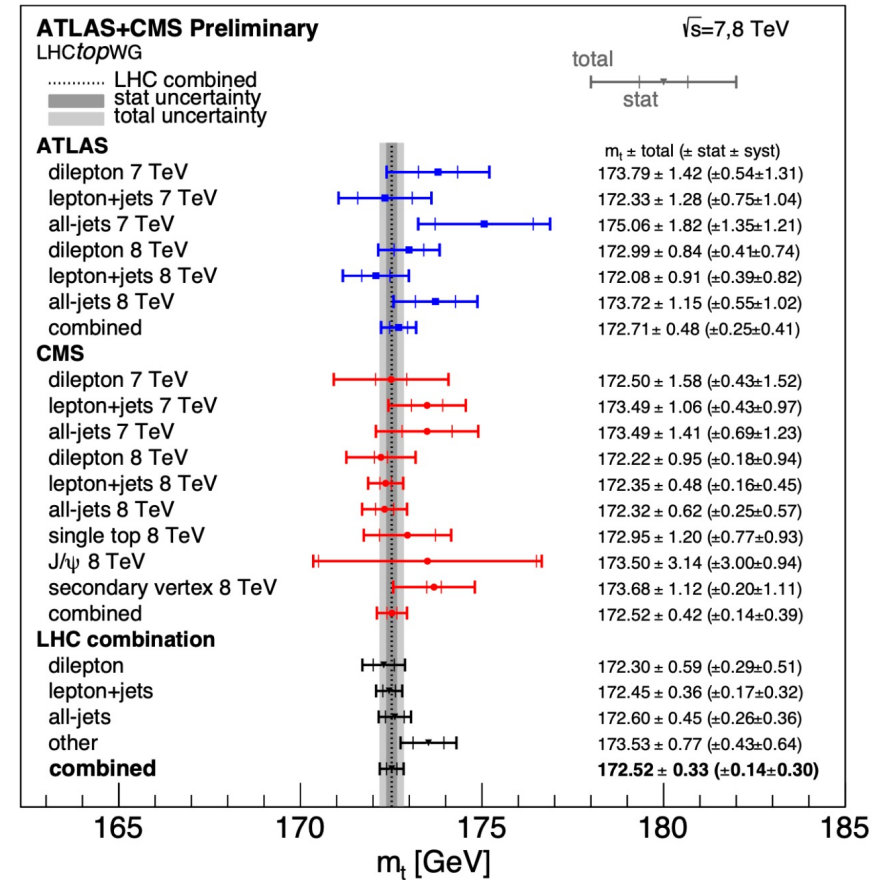
New paper for more details! [arXiv: 2407.05322](https://arxiv.org/abs/2407.05322)

7-8 TeV ATLAS-CMS top-quark mass combination

Goals

- The combination was performed with the BLUE approach:

→ Reproduce it with the nuisance parameters approach described earlier



Combination paper: [Phys. Rev. Lett. 132 \(2024\) 261902](https://arxiv.org/abs/2407.05322)

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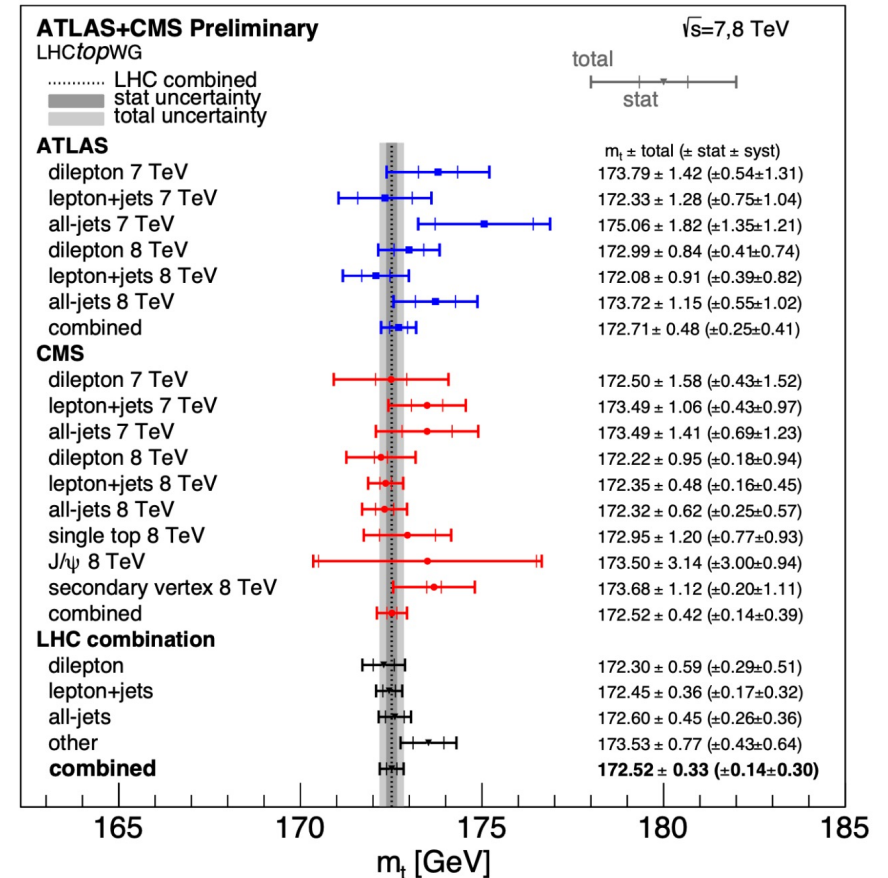
Goals

- The combination was performed with the BLUE approach:

→ Reproduce it with the nuisance parameters approach described earlier

- Top mass measurements are becoming systematics dominated
- Potentially affected by QCD modelling systematics

→ Study errors-on-errors impact on the combination



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7-8 TeV ATLAS-CMS top-quark mass combination

Robustness-test: Check if the combination is sensitive when any major systematic uncertainty is itself uncertain

Uncertainty category	Uncertainty impact [GeV]		
	LHC	ATLAS	CMS
LHC b-JES	0.18	0.17	0.25
b tagging	0.09	0.16	0.03
ME generator	0.08	0.13	0.14
LHC JES 1	0.08	0.18	0.06
LHC JES 2	0.08	0.11	0.10
Method	0.07	0.06	0.09
CMS B hadron BR	0.07	—	0.12
LHC radiation	0.06	0.07	0.10
Leptons	0.05	0.08	0.07
JER	0.05	0.09	0.02
Top quark p_T	0.05	—	0.07
Background (data)	0.05	0.04	0.06
Color reconnection	0.04	0.08	0.03
Underlying event	0.04	0.03	0.05
LHC g-JES	0.03	0.02	0.04
Background (MC)	0.03	0.07	0.01
Other	0.03	0.06	0.01
LHC l-JES	0.03	0.01	0.05
CMS JES 1	0.03	—	0.04
Pileup	0.03	0.07	0.03
LHC JES 3	0.02	0.07	0.01
LHC hadronization	0.02	0.01	0.01
p_T^{miss}	0.02	0.04	0.01
PDF	0.02	0.06	<0.01
Trigger	0.01	0.01	0.01
Total systematics	0.30	0.41	0.39
Statistical	0.14	0.25	0.14
Total	0.33	0.48	0.42

From: [arXiv:2402.08713](https://arxiv.org/abs/2402.08713)

7-8 TeV ATLAS-CMS top-quark mass combination

Robustness-test: Check if the combination is sensitive when any major systematic uncertainty is itself uncertain

1. Treat the eight largest systematic uncertainties as potentially uncertain one at a time.
2. Assign an error-on-error parameter ϵ_s to each one.
3. Study how varying each ϵ_s individually affects the central value and confidence interval.

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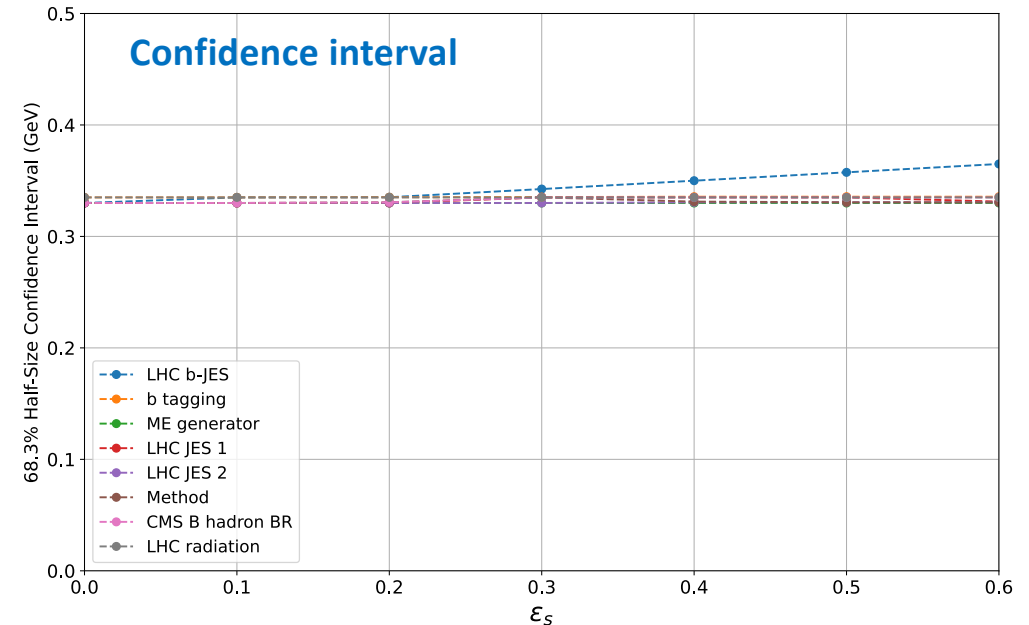
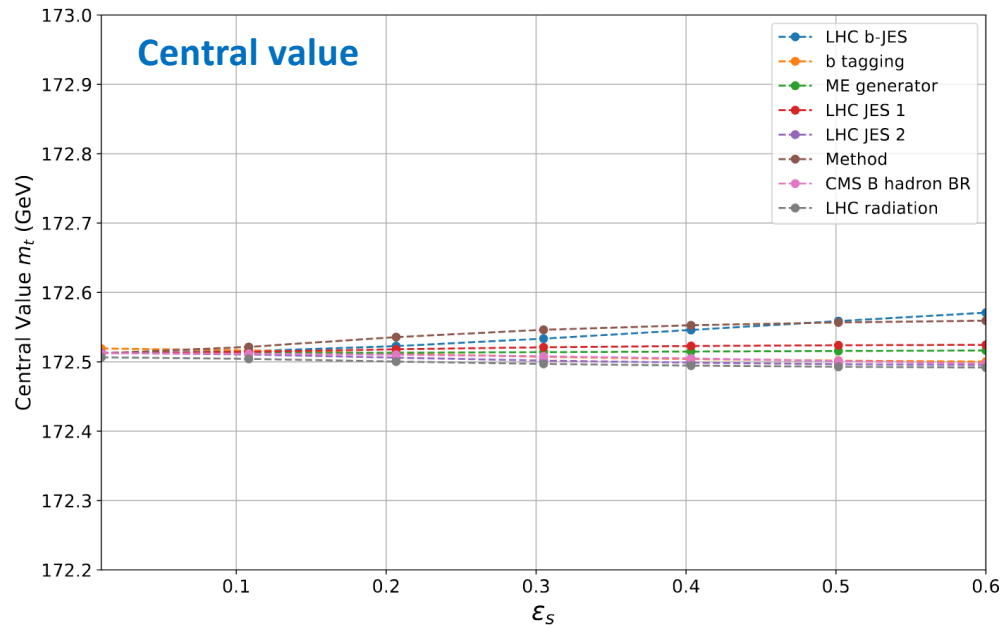
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Goal: Show a possible way to use the errors-on-errors model in an analysis

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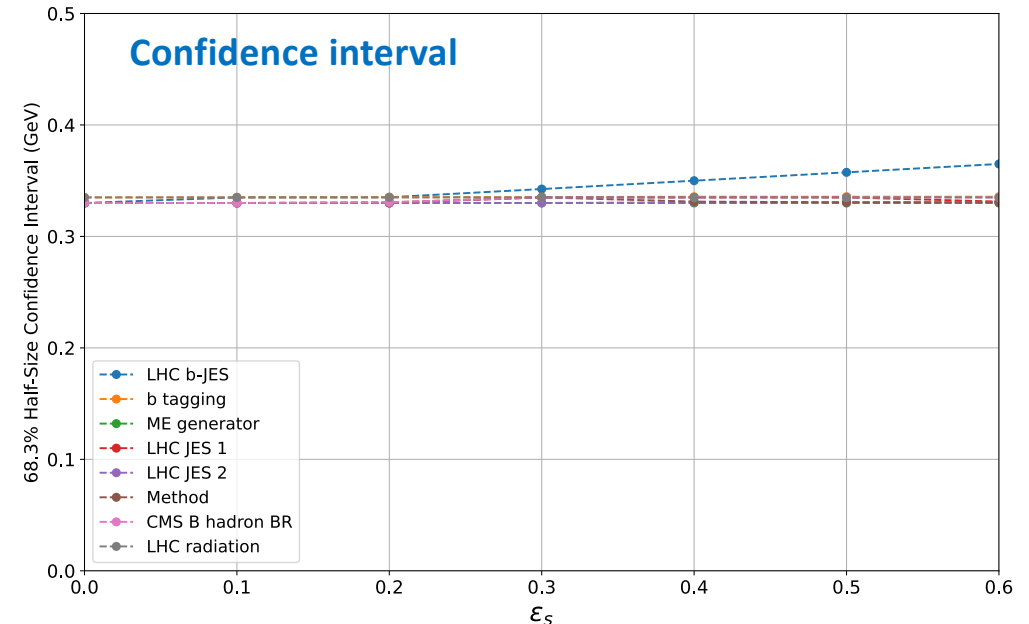
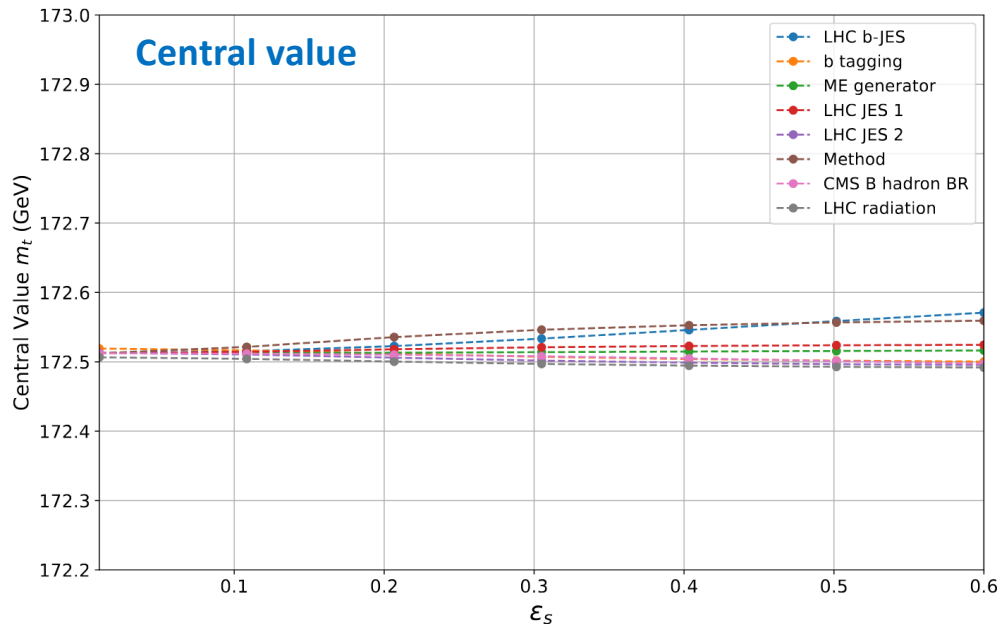
From: [arXiv:2402.08713](https://arxiv.org/abs/2402.08713)

7-8 TeV ATLAS-CMS top-quark mass combination



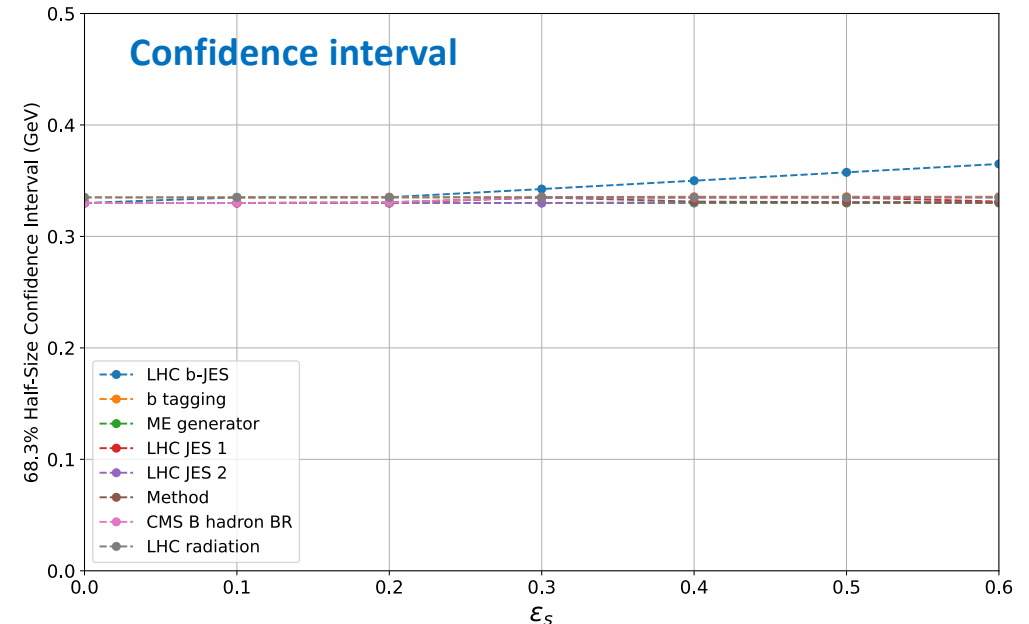
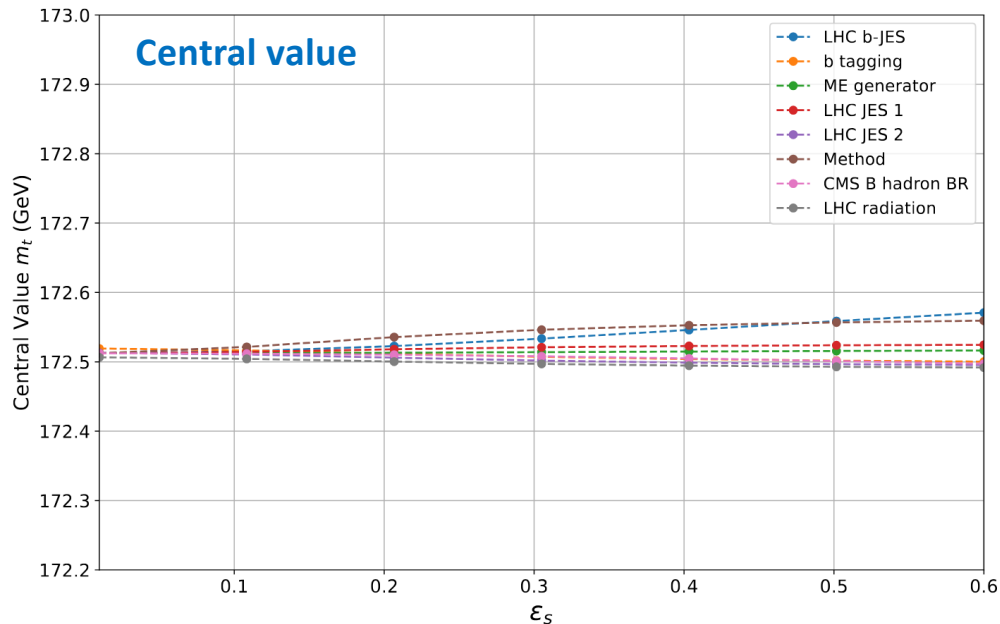
- As $\epsilon_s \rightarrow 0$ we recover the results of the combination. Treatment of correlations consistent.

7-8 TeV ATLAS-CMS top-quark mass combination



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- The central value is robust to the presence of uncertain systematic errors:
 - The change in the central value remains always within 0.1 GeV, well within the confidence interval of approximately 0.3 GeV.

7-8 TeV ATLAS-CMS top-quark mass combination



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- The central value is robust to the presence of uncertain systematic errors:
 - The change in the central value remains always within 0.1 GeV, well within the confidence interval of approximately 0.3 GeV.
- The confidence interval is also stable, though it exhibits a 10% inflation when the *LHC b-JES* uncertainty has an error the error

Sensitivity to outliers

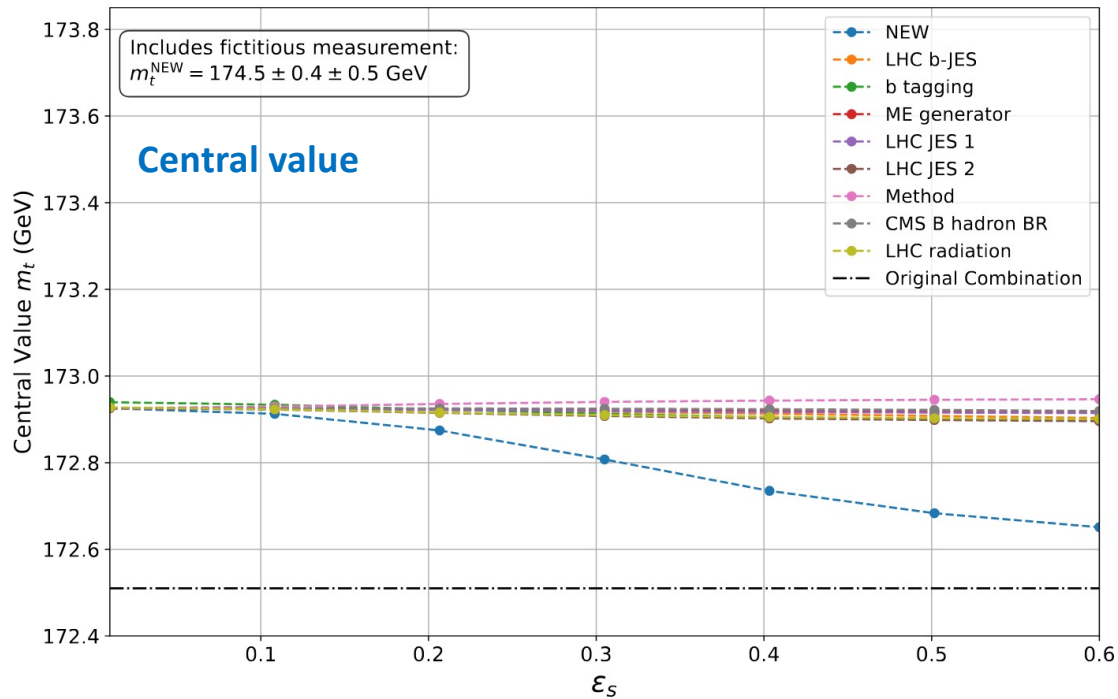


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- **Future relevance**: Relevant for future LHC–Tevatron combinations or LHC Run 2 combinations including the top mass measurement using a leptonic invariant mass ([J. High Energ. Phys. 2023, 19](#)).
- **Introduce fictitious measurement**: Add a hypothetical measurement to explore the model properties in these scenarios.

Sensitivity to outliers



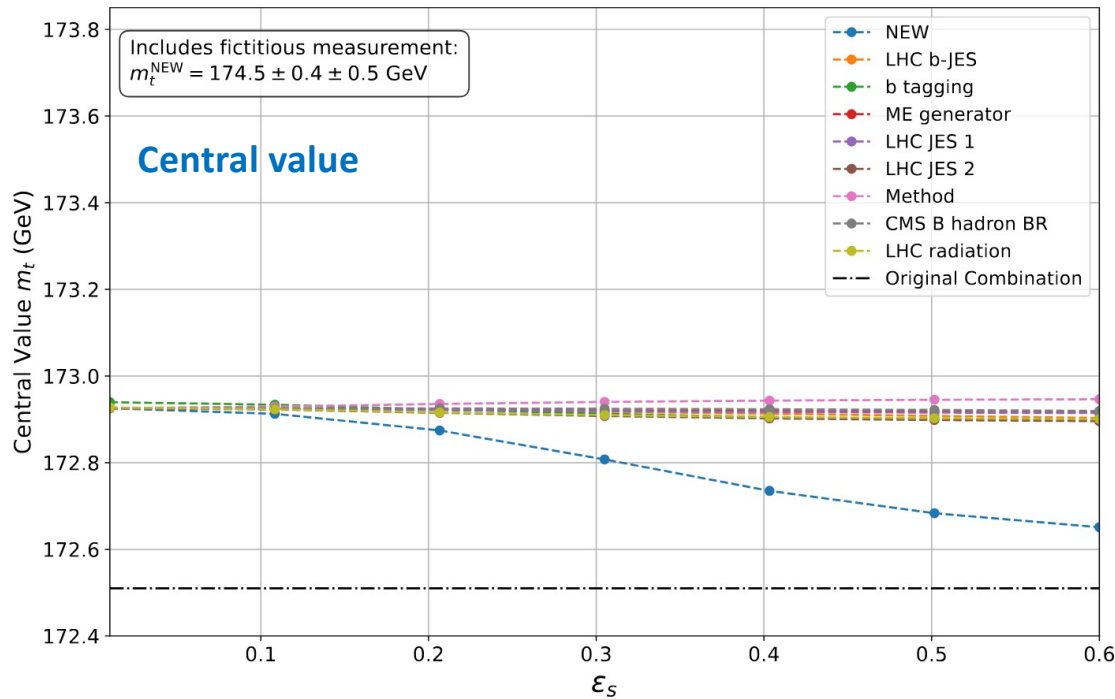
$$m_t^{NEW} = 174.5 \pm 0.4 \pm 0.5 \text{ GeV}$$

Independent systematic

- When ϵ_s is zero, the central value of the combination is pulled by the new measurement:

From **172.52 GeV** \longrightarrow To **172.91 GeV**

Sensitivity to outliers



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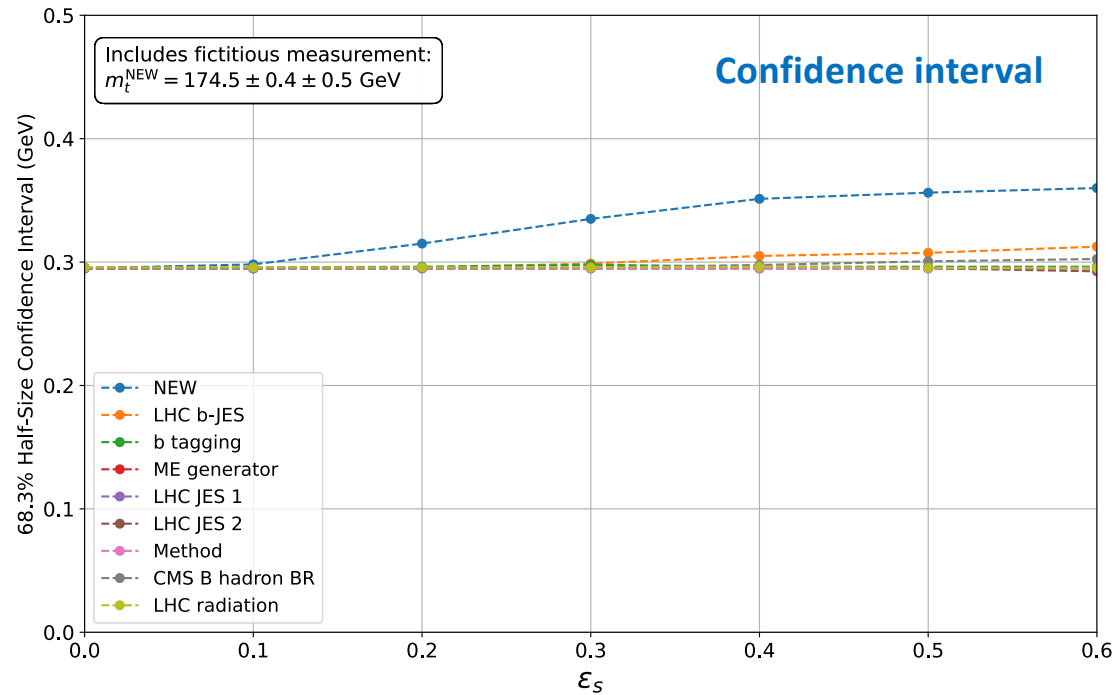
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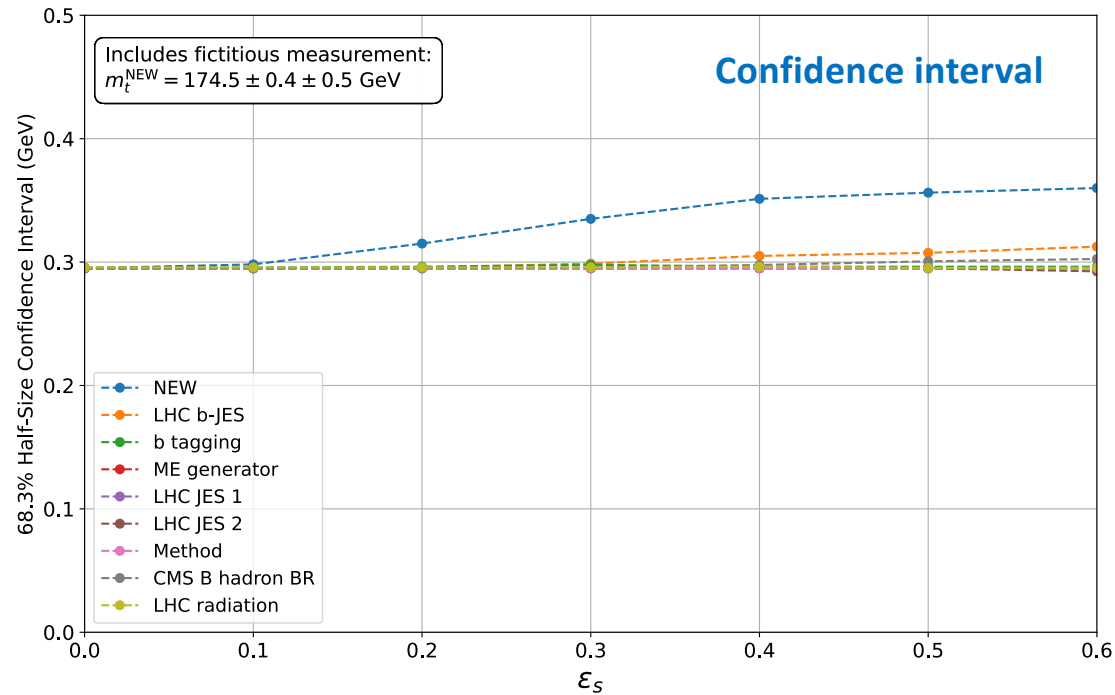


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Sensitivity to outliers

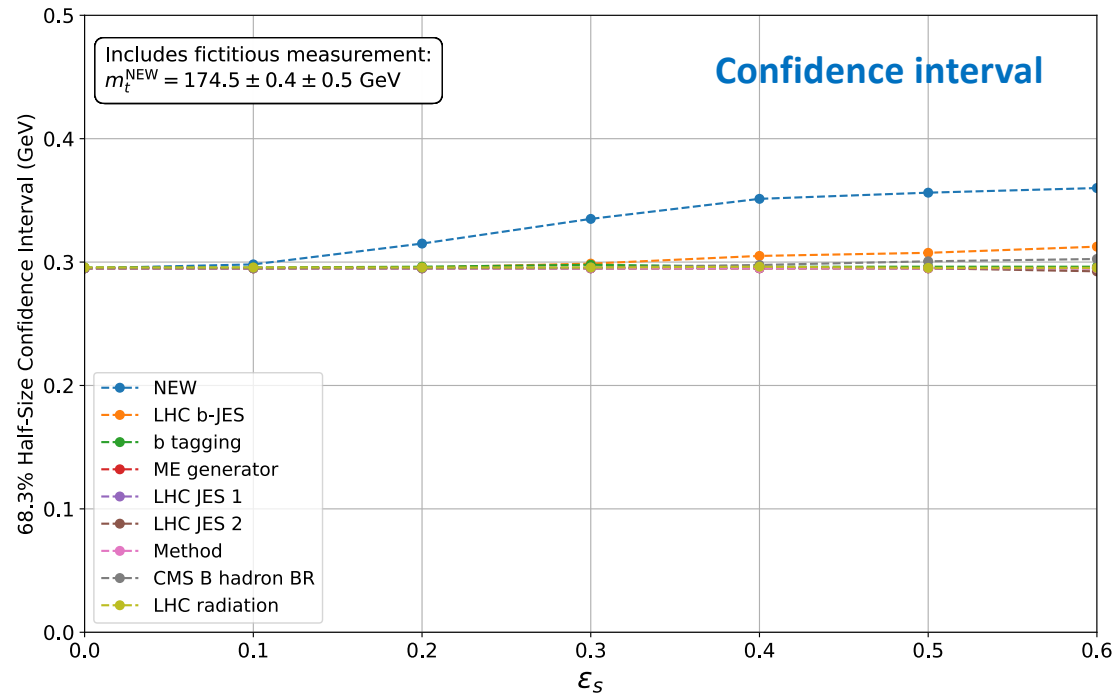


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- If the new measurement is affected by a large uncertain systematic, the CI inflates
- The tension in the dataset is treated as an additional source of uncertainty

Gamma variance model

- A model to account from theory uncertainties, two points systematics, etc..
- The primary advantage of this approach is that it reduces the sensitivity of the fits to outliers.
- The presence of incompatible data is reflected by inflated error bars on the final results.

Bartlett correction

- Used to correct for deviations from Wilks' theorem
- Useful tool to correct for small sample sizes errors

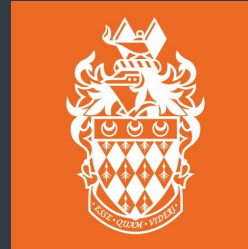
“Non-trivial” correlations

- A method to use nuisance parameters for non-trivial systematics assumptions



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Thank you for your attention



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Back-up slides

- Gamma distributions allow to parametrize distributions of positive defined variables (like estimates of variances)
- Using Gamma distributions it is possible to profile in close form over σ_i^2

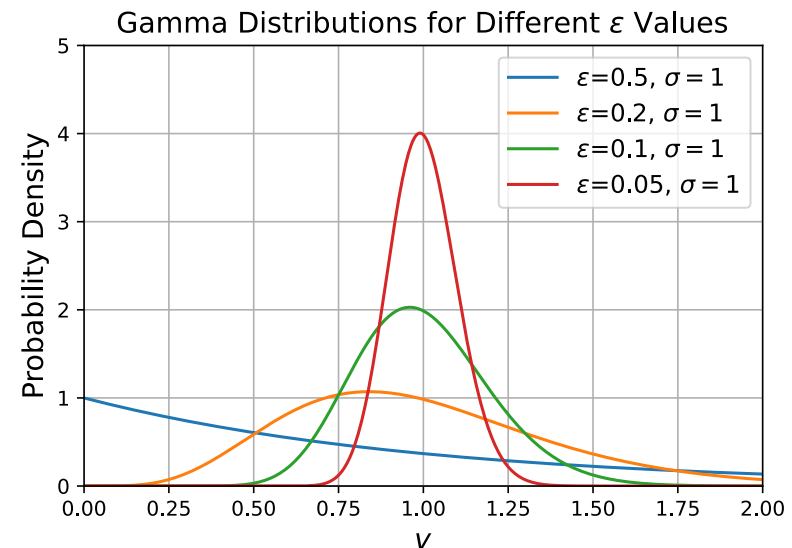
Gamma distribution

To implement “errors-on-errors” suppose the systematic variances $\sigma_{u_i}^2$ are *adjustable parameters*, and their best estimates v_i are gamma distributed:

$$v \sim \frac{\beta^\alpha}{\Gamma(\alpha)} v^{\alpha-1} e^{-\beta v}$$

$$\alpha = \frac{1}{4\varepsilon_i^2} \quad \beta = \frac{1}{4\varepsilon_i^2 \sigma_{u_i}^2}$$

- $\sigma_{u_i}^2$ Expectation value of v_i
- ε_i : relative error on σ_{u_i} : “*Error on error*”*



* ε used to be r in previous references

Gamma Variance Model (GVM)

- The likelihood is modified as follows:

$$L(\boldsymbol{\mu}, \boldsymbol{\theta}, \boldsymbol{\sigma}_{u_i}^2) = P(\mathbf{y}|\boldsymbol{\mu}, \boldsymbol{\theta}) \times \prod_i \frac{1}{\sqrt{2\pi\sigma_{u_i}}} e^{-(u_i - \theta_i)^2 / 2\sigma_{u_i}^2} \times \frac{\beta_i^{\alpha_i}}{\Gamma(\alpha_i)} v_i^{\alpha_i - 1} e^{-\beta_i v_i}$$

- One can profile over $\sigma_{u_i}^2$ in closed form:

$$\log L_P(\boldsymbol{\mu}, \boldsymbol{\theta}) = \log P(\mathbf{y}|\boldsymbol{\mu}, \boldsymbol{\theta}) - \frac{1}{2} \sum_i \left(1 + \frac{1}{2\varepsilon_i^2} \right) \log \left(1 + 2\varepsilon_i^2 \frac{(u_i - \theta_i)^2}{v_i} \right)$$

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- Profiling means computing

$$L_P(\boldsymbol{\mu}, \boldsymbol{\theta}) = L(\boldsymbol{\mu}, \boldsymbol{\theta}, \widehat{\boldsymbol{\sigma}}_{u_i}^2), \quad \widehat{\boldsymbol{\sigma}}_{u_i}^2 = \operatorname{argmax}_{\sigma_{u_i}^2} (L(\boldsymbol{\mu}, \boldsymbol{\theta}, \sigma_{u_i}^2))$$

- Gamma distributions include the case where the variance is estimate from a real dataset of control measurements:

$$v_i = \frac{1}{n_i - 1} \sum (u_{i,j} - \bar{u}_i)^2$$

- $(n - 1)v_i/\sigma_{u_i}^2$ follows a χ_{n-1}^2 distribution and v_i a Gamma distribution with:

$$\alpha_i = \frac{n_i - 1}{2}$$

$$\beta_i = \frac{n_i - 1}{2\sigma_{u_i}^2}$$

- The likelihood function can be used to construct the profile likelihood ratio test statistic:

$$w_{\mu} = -2 \ln \frac{L(\mu, \hat{\theta})}{L(\hat{\mu}, \hat{\theta})}$$

- Use the p -value:

$$p_{\mu} = \int_{w_{\mu, obs}}^{\infty} f(w_{\mu} | \mu) dw_{\mu}$$

- Include μ such that:

$$p_{\mu} < \alpha$$

- Modify the likelihood ratio w directly so that its distribution is closer to the asymptotic form:

$$w_{\mu} \longrightarrow w_{\mu}^* = w_{\mu} \frac{M}{E[w]}$$

To compute confidence intervals, rescale the results obtained with Standard methods, such as the Hessian method, by $\frac{M}{E[w]}$

$$w \sim \chi_M^2 + \mathcal{O}(n^{-1})$$

$$w^* \sim \chi_M^2 + \mathcal{O}(n^{-2})$$