

Bayesian non parametric modelling of Higgs pair production

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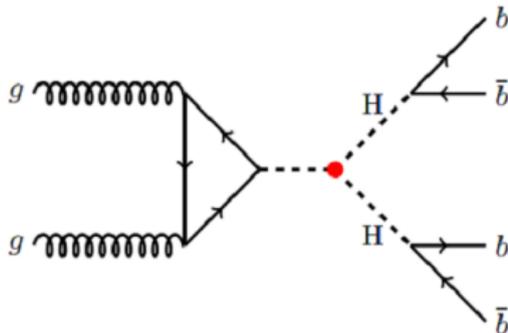


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joint work with *Annalisa Balata* (University of Padua) and *Tommaso Dorigo* (INFN)

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Isolate the signal of the Higgs boson pairs decays in the final state characterised by 4 jets of b -quark: $hh \rightarrow 4b$



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Data

- **background:** 1 259 973 observations collected by CMS during the LHC “Run 1” in 2012
(only if HLT-DiPFJet80-DiPFJet30-BTagCSVd07d05 trigger path is present)
- **signal:** 300 000 $hh \rightarrow b\bar{b}b\bar{b}$ events.
Monte Carlo simulated events (Alwall et al., 2011; Gao et al. 2014)

Events where 4 jets correspond to hadronisation of the b -quark

- 1 b -tagging algorithm CMVA (Das et al., 2013)
- 2 Selection of the first 3 jets in b -tag ranking, provided their CMVA is above the *medium cut*, 0.679
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Final dataset

At the end of preselection, we keep
68 454 MC signal events and
433 621 background CMS data

Available variables



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- Variables related to pairs of jets
- Other variables

Variables of the 4 selected jets and to the couples of dijets

Name	Variable
QP_t1	Transverse momentum related to the first jet
QP_t2	Transverse momentum related to the second jet
QP_t3	Transverse momentum related to the third jet
QP_t4	Transverse momentum related to the fourth jet
$QEta1$	Pseudorapidity related to the first jet
$QEta2$	Pseudorapidity related to the second jet
$QEta3$	Pseudorapidity related to the third jet
$QEta4$	Pseudorapidity related to the fourth jet
$QCMVA1$	CMVA related to the first jet
$QCMVA2$	CMVA related to the second jet
$QCMVA3$	CMVA related to the third jet
$QCMVA4$	CMVA related to the fourth jet
$QCent$	Centrality of the 4 jets

Variables of non selected jets

Name	Variable
AP_tmin	minimum p_t among non selected jets
AP_tmean	mean p_t among non selected jets
AP_tmax	maximum p_t among non selected jets
$AEtamin$	minimum η among non selected jets
$AEtamean$	mean η among non selected jets
$AEtamax$	maximum η among non selected jets
$ACMV Amin$	minimum CMVA among non selected jets
$ACMV Amean$	mean CMVA among non selected jets
$ACMV Amax$	maximum CMVA among non selected jets
$ACent$	centrality of non selected jets

Variables of pairs of jets, corresponding to each Higgs

Name	Variable
<i>DJ1mass</i>	mass of the first dijet
<i>DJ1P_t</i>	p_t of the first dijet
<i>DJ1Phi</i>	$\Delta\Phi$ of the first dijet
<i>DJ1Eta</i>	$\Delta\eta$ of the first dijet
<i>DJ1R</i>	ΔR of the first dijet
τ_1	twist of the first dijet
<i>DJ2mass</i>	mass of the second dijet
<i>DJ2P_t</i>	p_t of the second dijet
<i>DJ2Phi</i>	$\Delta\Phi$ of the second dijet
<i>DJ2Eta</i>	$\Delta\eta$ of the second dijet
<i>DJ2R</i>	ΔR of the second dijet
τ_2	twist of the second dijet

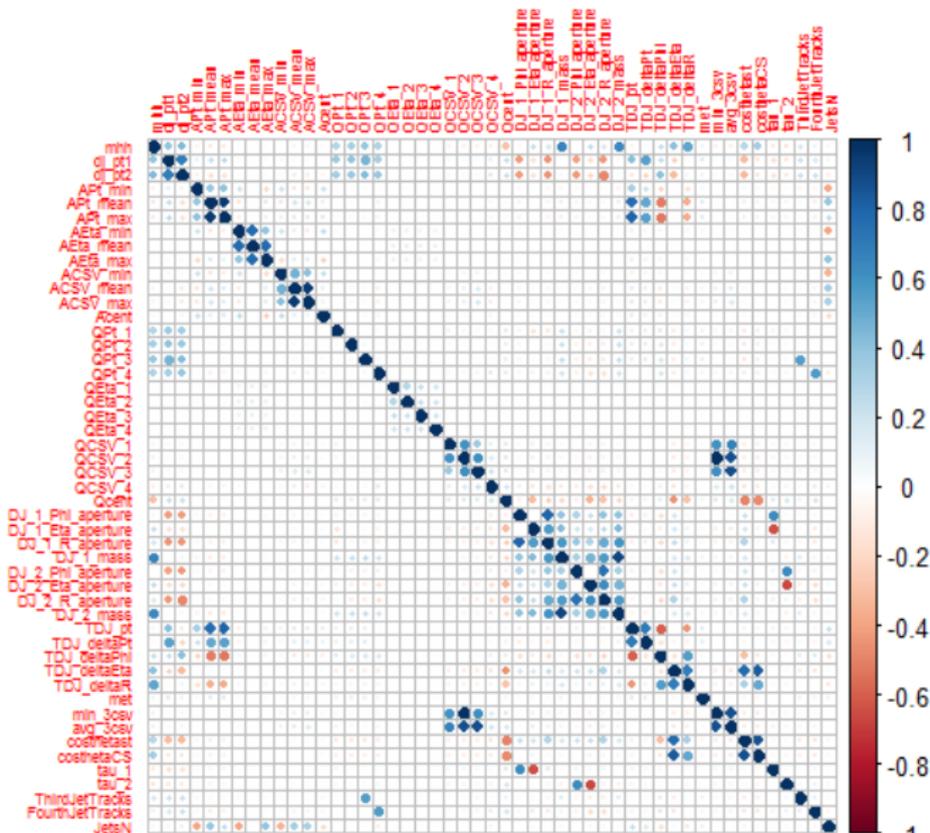
Other variables

Name	Variable
$TDJP_t$	vectorial sum of the P_t of the two dijet
$TDJ\Delta\Phi$	$\Delta\Phi$ between the two djets
$TDJ\Delta\eta$	$\Delta\eta$ between the two djets
$TDJ\Delta R$	ΔR between the two djets
HIM	invariant mass of the two djets
MET	Missing transverse energy
$min3cmva$	minimum CMVA among the first 3 jets
$avg3cmva$	mean CMVA among the first 3 jets
$cos\theta^*$	cosinus of θ on the c.o.m reference system of the two H
$cos\theta_{CS}$	cosinus of θ on the Collins Saper reference system
$JetsN$	number of jets in the event

Obtained variable

Name	Variable
$sumQP_t$	sum of the transverse momenta of the selected 4 jets

Available variables: Correlation Plot



Two approaches

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Estimation strategy

- Training set: 50 000 balanced events
- Test set: 16 000
- Validation set: 16 000

We consider a number of typical **statistical learning** models to best classify signal and background.

Linear and logistic regression, MARS, GAM, CART, Neural nets, Projection pursuit, etc. (e.g., Azzalini and Scarpa, 2012).

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Among the models with best performance on the test set:

- Random forests
- Gradient boosting
- Boosting decision tree

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- At each *tree split*, a random sample of m features (variables) is drawn, and only those m features are considered for splitting (Typically $m = \sqrt{p}$ or $\log_2 p$, where p is the number of features)

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Bias-variance trade off

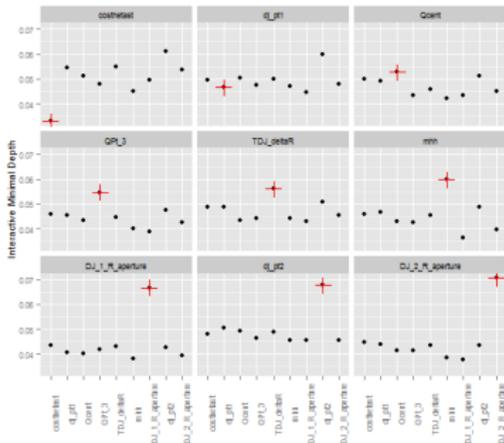
- *The small m , the lower the variance of the random forest ensemble*
- *Small m is also associated with higher bias, because important variables can be missed by the sampling*

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$DJ1P_t$	$DJ1R$	$QCent$	QP_t3
QP_t3	$TDJ\Delta R$	QP_t3	mhh
$TDJ\Delta R$	mhh	$TDJ\Delta R$	$DJ1R$
$DJ1R$	$DJ2P_t$	$DJ1R$	$DJ2R$
$QCent$	$DJ1R$	QP_t3	$DJ2R$
$QCent$	mhh	mhh	$DJ2R$
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- Final classifier is weighted average of classifiers

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- Boosting Decision Tree (BDT; Drucker, 1997): each error is normalized with the maximum of errors

Details

- Start with equal observation weights $p_i = 1/n$
- At iteration t , draw a bootstrap sample with the current probabilities p_1, p_2, \dots, p_n , compute the classifier and e_t , the error rate of the classifier on the original sample (for BDT normalized with the maximum error).

Let $\beta_t = e_t/(1 - e_t)$

- For those points that are classified correctly, decrease their probabilities by

$$p_i \leftarrow p_i \cdot \beta_t$$

and renormalise them

- Do this for many (say 1000) iterations.

At the end, take a weighted vote of the classifications, with weights $\log(1/\beta_t)$ (**more weight on classifiers with lower error**).

Boosting can improve bagging in many instances



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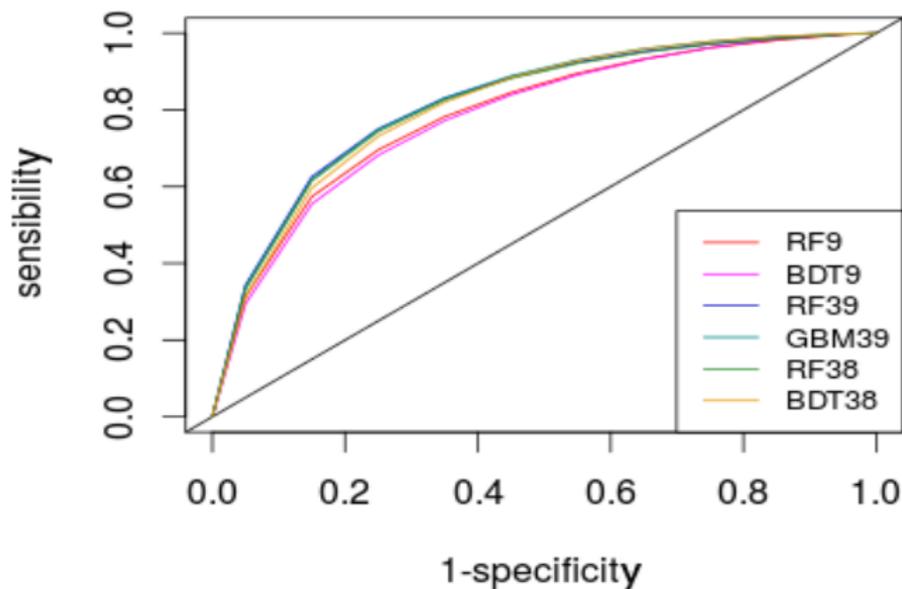
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- in Gradient Boosting, 'shortcomings' are identified by gradients (in Adaboost, 'shortcomings' are identified by high-weight data points).
- Both high-weight data points and gradients tell us how to improve the model.

ROC curve of the principal statistical learning classification models



Performances of the best statistical learning classification models



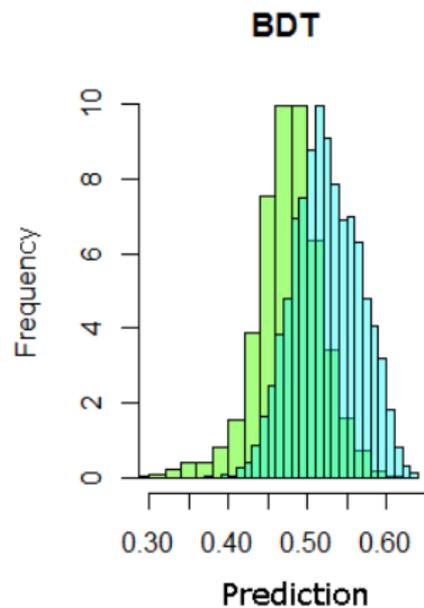
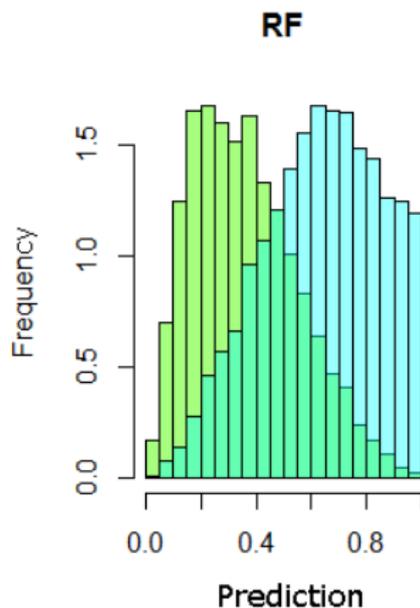
Model	Error test set	Error validation set	AUC test set	AUC validation set
BDT9	0.2821	0.2695	0.7969	0.7934
RF9	0.2756	0.2685	0.7982	0.8004
GBM9	0.2851	0.2858	0.7888	0.7911
BDT38	0.2596	0.2637	0.8204	0.8227
RF38	0.2540	0.2521	0.8224	0.8263
BDT39	0.2340	0.2464	0.8349	0.8400
RF39	0.2320	0.2424	0.8369	0.8424
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Comparison RF and BDT



Green: background
Blue: signal

- Let y_i be the binary variable encoding signal or background,
- the classical generalized mixed model formulation assumes

$$\begin{aligned}y_i | \pi_i &\sim \text{Bern}(\pi_i) \\ \text{logit}(\pi_i) &= \eta_i \\ \eta_i &= \mu_i + f(\mathbf{x}_i)\end{aligned}$$

where

- \mathbf{x}_i is the vector including all the explanatory variables for each event i
- β is a vector of parameters
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- Bayesian approach, assuming *priors* on parameters.
(*still “intrinsically” frequentist* - not a *subjective* approach)

- Dirichlet process (DP):

assume $\mu_i \sim P$ with $P \sim DP(\alpha P_0)$, $\alpha > 0$

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$$f(\mathbf{x}_i) = f_1(x_{i1}) + \dots + f_p(x_{ip})$$

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- **BART model (Bayesian Additive Regression Tree)**

$$f(\mathbf{x}_i) = \sum_{j=1}^m g(\mathbf{x}_i; T_j, M_j)$$

where $g(\mathbf{x}_i, T_j, M_j)$ denotes the predicting function assigning a value to \mathbf{x}_i given the Bayesian tree T_j and parameters M_j
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Priors distributions

- $\beta \sim \mathcal{N}_p(\mathbf{b}, \mathbf{B})$
- $\mathbf{b}_{p \times 1}, \mathbf{B}_{p \times p}$ are prior mean vector and covariance matrix
- $\mu_i \sim P$ with $P \sim DP(\alpha P_0)$, $\alpha > 0$, where DP indicates the **Dirichlet Process**.

The Dirichlet process $DP(\alpha P_0)$ represents a fully flexible prior with support on the set of distributions on the real line, allowing P to be unknown with

- P_0 indicating the best guess for such distribution and
- α quantifying the confidence in such guess.

In our case, we define P_0 as a normal distribution $\mathcal{N}(0, \sigma^2)$ where

- $\sigma^{-2} \sim \text{Gamma}(a, b)$ (i.e. prior for σ is inverse Gamma)
- $\alpha \sim \text{Gamma}(a_\alpha, b_\alpha)$ to favor learning of cluster effects in the data

We exploit the stick-breaking representation of the Dirichlet Process

Stick-breaking representation (Sethuraman, 1994)

Let

$$\begin{aligned} V_h &\stackrel{iid}{\sim} \text{Beta}(1, \alpha) & \theta_h &\stackrel{iid}{\sim} G_0 \\ \pi_h &\sim V_h \prod_{l < h} (1 - V_l) & G &= \sum_{h=1}^{\infty} \pi_h \delta_{\theta_h} \end{aligned}$$

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- Denoting with S_i the grouping variable, the stick-breaking representation shows clustering effects among events, providing $\mu_i = \theta_{S_i}$, with the number of clusters stochastically increasing with α
- This clustering property is particularly useful in our signal detection, favouring events with common kinematic features to share the same effect

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Let

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$$\pi_h \sim V_h \prod_{l < h} (1 - V_l) \quad G = \sum_{h=1}^{\infty} \pi_h \delta_{\theta_h}$$

therefore $G \sim DP(\alpha, G_0)$, where δ_{θ} indicates a mass point concentrated in θ .

- Key result: a realization of the Dirichlet process is **discrete** in nature
- It favours ties among random intercepts: events in the same cluster have equal random intercept values
- Denoting with S_i the grouping variable, the stick-breaking representation shows clustering effects among events, providing $\mu_i = \theta_{S_i}$, with the number of clusters stochastically increasing with α
- This clustering property is particularly useful in our signal detection, favouring events with common kinematic features to share the same effect
- Conditionally on the grouping indicator S_i , the Gaussian base measure P_0 is conjugate, favoring the implementation of a Gibbs sampler

For posterior computation we exploit

- *blocked Gibbs sampler* algorithm by Ishwaran and James (2001)

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- a recently proposed data-augmentation scheme based on Pólya-Gamma (PG) distribution

Pólya-Gamma data-augmentation

Assuming a Bayesian logistic regression setting where

$y_i \sim \text{Bern}(1/[1 + e^{-\phi_i}])$, $i = 1, \dots, n$, $\phi_i = \mathbf{x}_i^T \boldsymbol{\beta}$ and $\boldsymbol{\beta} \sim \mathcal{N}_p(\mathbf{b}, \mathbf{B})$, the resulting Gibbs sampler alternates between two full conditional conjugate steps

- $\omega_i \sim PG(1, \mathbf{x}_i^T \boldsymbol{\beta})$
- $\boldsymbol{\beta} | \mathbf{y}, \boldsymbol{\omega}, \mathbf{X} \sim \mathcal{N}_p(\boldsymbol{\mu}_\beta, \boldsymbol{\Sigma}_\beta)$

where $\boldsymbol{\Sigma}_\beta = (\mathbf{X}^T \boldsymbol{\Omega} \mathbf{X} + \mathbf{B}^{-1})$ and $\boldsymbol{\mu}_\beta = \boldsymbol{\Sigma}_\beta (\mathbf{X}^T \mathbf{z} + \mathbf{B}^{-1} \mathbf{b})$,
 $\mathbf{z} = [y_1 - 1/2, \dots, y_n - 1/2]$ and $\boldsymbol{\Omega} = \text{diag}(\omega_1, \dots, \omega_n)$

Results

	Test set classification error	False positives	False negatives	AUC
Logistic BNP	0.492	0.4121	0.5663	0.5118

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back to the generalized mixed model

$$\begin{aligned}y_i | \pi_i &\sim \text{Bern}(\pi_i) \\ \text{logit}(\pi_i) &\sim \eta_i \\ \eta_i &\sim \mu_i + f(\mathbf{x}_i)\end{aligned}$$



- Bayesian additive model

$$f(\mathbf{x}_i) = f_1(x_{i1}) + \cdots + f_p(x_{ip})$$

$f_1(\cdot), \dots, f_p(\cdot)$ estimated via Bayesian P -splines, and $\mu_i = 0$.

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$$f_j(x_j) = \sum_{r=1}^{M_j} \beta_{jr} B_{jr}(x_j)$$

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Additive model with P -splines

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$$f(\mathbf{x}_i) = \sum_{j=1}^m g(\mathbf{x}_i; T_j, M_j)$$

where $g(\mathbf{x}_i, T_j, M_j)$ denotes the predicting function assigning a value to \mathbf{x}_i given the Bayesian tree T_j and parameters M_j

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BART - Bayesian Additive Regression Tree

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	Test set	False	False	AUC
	classification error	positives	negatives	
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- Dirichlet process and P -splines

$$f(\mathbf{x}_i) = \mu_i + f_1(x_{i1}) + \cdots + f_p(x_{ip})$$

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- Dirichlet process with BART atoms

the atoms of the Dirichlet process depends on the \mathbf{x}_i and are described by a BART.

- Dirichlet model with BART atoms and P -splines

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- BART with P -SPLINES

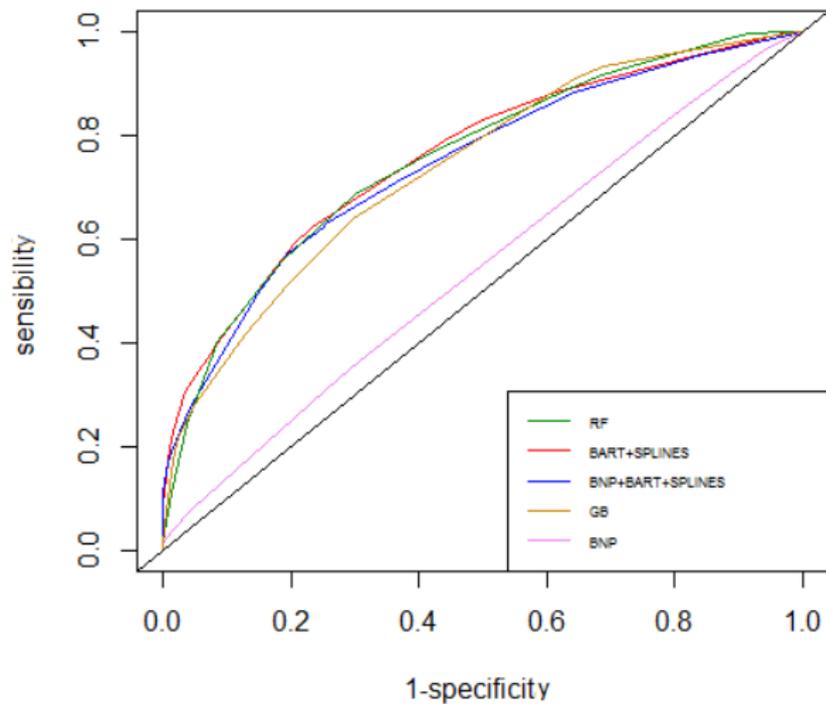
Si stima il modello

$$f(\mathbf{x}_i) = f_1(x_{i1}) + \cdots + f_p(x_{ip}) + \sum_{j=1}^m g(x; T_j, M_j)$$

where f_j are estimated via Bayesian P -splines, and sum of g is a BARTmodel.

Model	Classification Error	False Positives	False Negatives	AUC
RF	0.3200	0.2678	0.3678	0.7393
GB	0.3300	0.2971	0.3602	0.7234
BDT	0.3351	0.3461	0.3389	0.7198
DP	0.4920	0.4121	0.5663	0.5118
P-splines	0.3400	0.3975	0.2874	0.7102
BART	0.3480	0.3083	0.3846	0.7078
BART as atoms of DP	0.4620	0.4346	0.4867	0.5429
DP+P-splines	0.3300	0.3347	0.3295	0.7259
BART as atoms of DP + P-splines	0.3240	0.2970	0.3455	0.7321
BART+P-splines	0.3140	0.3138	0.3141	0.7417

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- The choice of the most appropriate model can be driven by knowledge of its strength. However, it is always better to fit different models and compare them on different data, in order to choose the best classifier.
- Combination of models may work better than single models if each model has a different strength compared to the others: committees can help when different learners have complementary strengths for a given task.
- Choice of 9 most predictive variables has been done by physicists by looking to marginal correlation and meaning of the variables. Most of models with 39 variables have some procedure of automatic selection of complexity and of variables, compromising between bias and variance. We used a test set to train and select complexity level of each model.