## Statistical errors, efficiency and acceptance

 corrections in cumulants of measured net-charge ( $\mathrm{N}^{+}-\mathrm{N}^{-}$) distributions, a theorem from Quantitative Finance and NBD fits to the PHENIX $\mathrm{N}^{+}$and $\mathrm{N}^{-}$distributions (New PHENIX results on T and $\boldsymbol{\mu}_{B}$ at freezeout) M. J. TannenbaumBrookhaven National Laboratory Upton, NY 11973 USA

Zimanyi School '15
Budapest Hungary
December 2015


BRDDKHRNEN
PH夷ENIX

NBD-p+p discoveryUA5 PLB 160, 193,199 (1985); 167, 476 (1986)
NBD in O+Cu central collisions at AGS vs $\Delta \eta$ ( Correlations due to to B-E don't vanish

E802 PRC 52, 2663 (1995)
E802 0+Cu Central Multiplicity data in eta bins


Poisson, no correlation

$$
\frac{\sigma^{2}}{\mu^{2}}-\frac{1}{\mu}=0
$$



$$
k(\delta \eta) \sim \frac{(\delta \eta / \xi)^{2}}{\left(\delta \eta / \xi-1+e^{-\delta \delta / \xi \xi}\right)}
$$

The rapidity correlation length $\xi=0.2$ for $\mathrm{O}+\mathrm{Cu}$ is from $\mathrm{B}-\mathrm{E}$.

E802, PRC56(1977) 1544
Zimanyi-Dec 2015
PH光 ENIX M.J. Tannenbaum 2

## BEAM Energy Scan Search for Critical Endpoint in Nuclear Matter Phase diagram Helped by Lattice QCD



## Hot off the presses-LBL Press release June 24,2011 Higher Moments of Net-Proton Distributions

- $1^{\text {st }}$ moment: mean $=\mu=<x>$
- $2^{\text {nd }}$ cumulant: variance $\kappa_{2}=\sigma^{2}=\left\langle(x-\mu)^{2}\right\rangle$ $3^{\text {rd }}$ cumulant: $\kappa_{3}=\mu_{3}=\left\langle(x-\mu)^{3}\right\rangle$
$3^{\text {rd }}$ standardized cumulant: skewness $=$ $\mathrm{S}=\kappa_{3} / \kappa_{2}{ }^{3 / 2}=<(\mathrm{x}-\mu)^{3}>/ \sigma^{3}$
- $4^{\text {th }}$ cumulant: $\kappa_{4}=\left\langle(x-\mu)^{4}\right\rangle-3 \kappa_{2}^{2}$
- $4^{\text {th }}$ standardized cumulant: kurtosis $=$ $\kappa=\kappa_{4} / \kappa_{2}{ }^{2}=\left\{\left\langle(x-\mu)^{4}>/ \sigma^{4}\right\}-3\right.$
- Calculate moments from the event-byevent net proton distribution.
$\checkmark$ Have similar plots for net-charge and netkaon distributions.


MJT-If you know the distribution, you know all the moments, but statistical mechanics and Lattice QCD use Taylor expansions, hence moments/cumulants

## Statistical Mechanics uses derivatives of

## the free energy to find susceptibilities

- Theoretical analyses tend to be made in terms of a Taylor expansion of the free energy $F=-T \ln Z$ around the critical temperature $\mathrm{T}_{\mathrm{c}}$ where Z is the partition function or sum over states, $\mathrm{Z} \approx \exp -\left[\left(\mathrm{E}-\Sigma_{i} \mu_{i} \mathrm{Q}_{\mathrm{i}}\right) / \mathrm{kT}\right]$ and $\mu_{\mathrm{i}}$ chemical potentials associated with conserved charges $\mathrm{Q}_{\mathrm{i}}$
- The terms of the Taylor expansion are called susceptibilities or $\chi_{(m)}$ which are proportional to the correlation length, e.g. $\chi_{(3)} \sim \xi^{6}, \chi_{(4)} \sim \xi^{8}$
- The only connection of this method to mathematical statistics is that the Cumulant generating function is also a Taylor expansion of the ln of an exponential:

$$
g_{x}(t)=\ln \left\langle e^{t x}\right\rangle=\sum_{n=1}^{\infty} \kappa_{n} \frac{t^{n}}{n!} \quad \kappa_{m}=\left.\frac{d^{m} g_{x}(t)}{d t^{m}}\right|_{t=0}
$$

## If you measure the distribution, then you know all the cumulants

Cumulants for Poisson, Binomial and Negative Binomial Distributions

| Cumulant | Poisson | Binomial | Negative Binomial |
| :--- | :--- | :--- | :--- |
| $\kappa_{1}=\mu$ | $\mu$ | $n p$ | $\mu$ |
| $\kappa_{2}=\mu_{2}=\sigma^{2}$ | $\mu$ | $\mu(1-p)$ | $\mu(1+\mu / k)$ |
| $\kappa_{3}=\mu_{3}$ | $\mu$ | $\sigma^{2}(1-2 p)$ | $\sigma^{2}(1+2 \mu / k)$ |
| $\kappa_{4}=\mu_{4}-3 \kappa_{2}^{2}$ | $\mu$ | $\sigma^{2}\left(1-6 p+6 p^{2}\right)$ | $\sigma^{2}\left(1+6 \mu / k+6 \mu^{2} / k^{2}\right)$ |
| $S \equiv \kappa_{3} / \sigma^{3}$ | $1 / \sqrt{\mu}$ | $(1-2 p) / \sigma$ | $(1+2 \mu / k) / \sigma$ |
| $\kappa \equiv \kappa_{4} / \kappa_{2}^{2}$ | $1 / \mu$ | $\left(1-6 p+6 p^{2}\right) / \sigma^{2}$ | $\left(1+6 \mu / k+6 \mu^{2} / k^{2}\right) / \sigma^{2}$ |
| $S \sigma=\kappa_{3} / \kappa_{2}$ | 1 | $(1-2 p)$ | $(1+2 \mu / k)$ |
| $\kappa \sigma^{2}=\kappa_{4} / \kappa_{2}$ | 1 | $\left(1-6 p+6 p^{2}\right)$ | $\left(1+6 \mu / k+6 \mu^{2} / k^{2}\right)$ |

Thanks to Gary Westfall of STAR in a paper presented at Erice-International School of Nuclear Physics 2012, I found out that the cumulants of the difference of samples from two such distributions $\mathrm{P}(\mathrm{n}-\mathrm{m})$ where $\mathrm{P}^{+}(\mathrm{n})$ and $\mathrm{P}(\mathrm{m})$ are both Poisson, Binomial or NBD with Cumulants $\kappa_{j}^{+}$and $\kappa_{j}^{-}$respectively is the same as if they were statistically independent, so long as they are not $100 \%$ correlated. I call this the NBD Cumulant Theorem

$$
\kappa_{j}=\kappa_{j}^{+}+(-1)^{j} \kappa_{j}^{-}
$$

## STAR publications 2014

## PRL 113(2014) 092301



So clearly favors NBD, not Poisson (!).
No non- monotonic behavior in S $\sigma$ or $\kappa \sigma^{2}$
but $\kappa \sigma^{2}=-1.5$ at $\sqrt{ } \mathrm{s}_{\mathrm{NN}}=20$ can't be ruled out

$\kappa \sigma^{2}=-1.5$ at $\sqrt{ } \mathrm{s}_{\mathrm{NN}}=20$ can be ruled out $\kappa \sigma^{2}$ changes for $\sqrt{\mathrm{s}_{\mathrm{NN}}} \leq 20 \mathrm{GeV}$ but antiprotons become negligible $<0.1 \mathrm{p}$

New PHENIX central cumulant ratios vs

## PHENIX arXiv:1506.07834



Note that the "data" • calculations from the $\Delta \mathrm{N}_{\mathrm{ch}}=\mathrm{N}^{+}-\mathrm{N}^{-}$distributions agree with the NBD fits to the $\mathrm{N}^{+}$and $\mathrm{N}^{-}$distribution and the NBD Cumulant Theorem.

## PHENIX and STAR comparison!!!




The key difference of the PHENIX and STAR results is that the error on all corrected cumulant ratios is $20-30 \%$ for PHENIX while for STAR the error on e.g. S $\sigma$ is $\sim 50 \%$, on $\kappa \sigma^{2}$ is $>100 \%$ but $<1 \%$ for $\sigma^{2} / \mu!!!$ (which turns out to be important)



BNL Lattice QCD group predictions for cumulant ratios $\kappa_{3} / \kappa_{1}=$ R31 and $\kappa_{1} / \kappa_{2}=$ R12 vs $\mathrm{T}_{\mathrm{f}}$ and $\mu_{\mathrm{B}}$ at freezout (when QGP hadronizes)
PoS(CPOD2014)005. PRL 109 (2012) 192302



STAR measurement of R31 $=\kappa_{3} / \kappa_{1}$ has such a huge error that the central $\approx$ could go anywhere in the dashed region, while R12 has such a small error that it is constrained to the region of the horizontal line by the assumption $140<\mathrm{T}_{\mathrm{f}}<150 \mathrm{MeV}$ PRL 113 (2014) 052301 PHENIX $\diamond$ measurement with comparable errors on R31 and R12 enables both $T_{f}$ and $\mu_{\mathrm{B}}$ to be determined from the Lattice QCD calculations: Pos(CPOD2014)005

## STAR's opinion of PHASE diagram 2014



$$
\frac{\bar{p}}{p}=\frac{e^{-\left(E+\mu_{B}\right) / T}}{e^{-\left(E-\mu_{B}\right) / T}}=e^{-\left(2 \mu_{B}\right) / T}
$$

## NEW: Experiment +Theory=Physical Quantity



Experimental result on net-charge cumulants + Lattice QCD calculation gives both freezeout $T_{f}+$ Baryon Chemical Potential $\mu_{\mathrm{B}}$ without particle identification!! I think this is a first and it also agrees with the best accepted calculations from baryon/anti-baryon ratios, PRC73(2006)034905

PHENIX $\mathrm{T}_{\mathrm{f}} \mu_{\mathrm{B}}$ from net-ch measurement cf. Calculation using STAR net-ch and net-p data
$\sqrt{ } \mathrm{s}_{\mathrm{NN}}$ (GeV)
$\mathrm{T}_{\mathrm{f}}$
(MeV)
(MeV)

## $\mu_{\mathrm{B}}$

$(\mathrm{MeV})^{* *}$

$$
\begin{array}{cccc}
27 & 164 \pm 6 & 181 \pm 21 & 136 \pm 13.8 \\
39 & 158 \pm 5 & 114 \pm 13 & 101 \pm 10 \\
62.4 & 163 \pm 5 & 71 \pm 8 & 66.6 \pm 7.9 \\
200 & 163 \pm 8 & 27 \pm 5 & 22.8 \pm 2.6
\end{array}
$$

**S. Borsanyi et al., Phys. Rev. Lett. 113, 052301 (2014) used $\mathrm{T}_{\mathrm{f}}=145 \pm \mathbf{5 ~ M e V}$ from STAR net-proton data averaged over above $4 \sqrt{ }{ }^{s_{\mathrm{NN}}}, \mu_{\mathrm{B}}$ from STAR net-charge $\mathrm{R}_{12}$

## PHENIX and STAR comparison!!!



The key difference of the PHENIX and STAR results is that the error on all corrected cumulant ratios is $20-30 \%$ for PHENIX while for STAR the error on e.g. So is $\sim 50 \%$, on $\kappa \sigma^{2}$ is $>100 \%$ but $<1 \%$ for $\sigma^{2} / \mu!!$ ! WHY?

## Why are STAR errors on R31 so large?

It must be that statistical errors and efficiency

## corrections are a BIG issue in these

measurements even though the correction is simply Binomial; and analytical for NBD $\mathrm{N}^{+}$ and $\mathrm{N}^{-}$distributions ( $k$ unchanged, $\mu_{\mathrm{t}}=\mu / \mathrm{p}$ where $p$ is the efficiency). So use the NBD "integer value Levy process" cumulant theorem: Tarnowsky, Westfall PLB 724 (2013) 51

Barndorff-Nielsen,Pollard,Shephard
http://www.economics.ox.ac.uk/materials/papers/4382/paper490.pdf

$$
\kappa_{j}=\kappa_{j}^{+}+(-1)^{j} \kappa_{j}^{-}
$$


$\Delta \mathrm{N}_{\mathrm{ch}}=\mathrm{N}^{+}-\mathrm{N}^{-}$distribution in $|\eta|<0.35, \delta \phi=\pi, 0.3<\mathrm{p}_{\mathrm{T}}<2.0 \mathrm{GeV} / \mathrm{c}$ Not corrected for detection efficiency $\varepsilon \approx 0.70$ in acceptance

The raw moments of the uncorrected distributions can be easily calculated

$$
\mu_{k}^{\prime} \equiv\left\langle x^{k}\right\rangle \equiv \sum_{i=1}^{n} x_{i}^{k} E\left(x_{i}\right) / \sum_{i=1}^{n} E\left(x_{i}\right)
$$

$\mu_{1}^{\prime} \equiv \mu=\langle x\rangle$ and $x_{i}$ is a bin in the $\Delta N_{\text {ch }}$ plot with $E\left(x_{i}\right)$ events.
$\underset{\text { BATIONAL LABORATORY }}{\text { BHENENE }} 2015$
PH少: ENIX M. J. Tannenbaum 18

Statistical errors--the complications begin

$$
\mu_{k}^{\prime} \equiv\left\langle x^{k}\right\rangle \equiv \sum_{i=1}^{n} x_{i}^{k} E\left(x_{i}\right) / \sum_{i=1}^{n} E\left(x_{i}\right)
$$

The statistical errors for every $\mu_{k}^{\prime}$ can be calculated from the statistical errors of each data point $E\left(x_{i}\right) \pm \sigma_{E\left(x_{i}\right)}$. Even though the $\sigma_{E\left(x_{i}\right)}$ on each point are independent, the errors on each $\mu_{k}^{\prime}$ are not independent because the same $\sigma_{E\left(x_{i}\right)}$ appears in all the moments.

Next one computes the cumulants $\kappa_{i}$ from the raw (aka) noncentral moments:

$$
\mu=\kappa_{1}=\mu_{1}^{\prime}
$$

$$
\begin{aligned}
\sigma^{2}=\mu_{2} \equiv\left\langle(x-\mu)^{2}\right\rangle & =\kappa_{2}
\end{aligned}=\mu_{2}^{\prime}-\mu_{1}^{\prime 2}, \mu^{\prime 2} .
$$

## Next correction---Efficiency

A certain random fraction of the tracks that fall on the acceptance are not detected because of inefficiency---a clearly random, thus binomial effect. This is further complicated if the $\mathrm{N}^{+}$and $\mathrm{N}^{-}$measurements have different efficiencies.

## Long Range Correlations: Binomial Split of NBD Carruthers and Shih PLB 165 (1985)209

If a population $n$ is distributed as $\operatorname{NBD}\left(\mu_{\mathrm{t}}, k\right)$ and then divided randomly into 2 subpopulations with probabilities $p$ and $q=1-p$, then the distribution on $p$ is NBD $\left(p \mu_{\mathrm{t}}, k\right)$ and on $q$ is NBD $\left(q \mu_{\mathrm{t}}, k\right)$ BUT the two sub-intervals are not statistically independent. Also $k$ does not change!!

## So if you measure $\mu=p \mu_{\mathrm{t}}$ with effieicncy $p$ the true value is $\mu_{\mathrm{t}}=\mu / p$

## Bzdak-Koch standard Binomial efficiency correction PRC 86 (2012) 044904

Efficiency corrected cumulants in terms of corrected double Factorial moments

$$
\begin{array}{rlrl}
\kappa_{1}= & \left\langle N_{+}\right\rangle-\left\langle N_{-}\right\rangle=\frac{\left\langle n_{+}\right\rangle}{\epsilon_{+}}-\frac{\left\langle n_{-}\right\rangle}{\epsilon_{-}}, & N=\left\langle N_{+}\right\rangle+\left\langle N_{-}\right\rangle \\
\kappa_{2}= & N-\kappa_{1}^{2}+F_{02}-2 F_{11}+F_{20}, & F_{i k}=\sum_{N_{1}=i}^{\infty} \sum_{N_{2}=k}^{\infty} P\left(N_{1}, N_{2}\right) \frac{N_{1}!}{\left(N_{1}-i\right)!} \frac{N_{2}!}{\left(N_{2}-k\right)!} \\
\kappa_{3}= & \kappa_{1}+2 \kappa_{1}^{3}-F_{03}-3 F_{02}+3 F_{12}+3 F_{20}-3 F_{21}+F_{30} \\
& -3 \kappa_{1}\left(N+F_{02}-2 F_{11}+F_{20}\right), \\
\kappa_{4}= & N-6 \kappa_{1}^{4}+F_{04}+6 F_{03}+7 F_{02}-2 F_{11}-6 F_{12}-4 F_{13} \\
& +7 F_{20}-6 F_{21}+6 F_{22}+6 F_{30}-4 F_{31}+F_{40} \\
& +12 \kappa_{1}^{2}\left(N+F_{02}-2 F_{11}+F_{20}\right)-3\left(N+F_{02}-2 F_{11}+F_{20}\right)^{2} \\
& -4 \kappa_{1}\left(\kappa_{1}-F_{03}-3 F_{02}+3 F_{12}+3 F_{20}-3 F_{21}+F_{30}\right)
\end{array}
$$

Here you can see the nice subtraction of the lower order moments; but new quantities, double Factorial Moments are introduced and very difficult to compute $\mathrm{P}\left(13^{+}, 11^{-}\right)=$? so you need to know both $\mathrm{N}_{+}$and $\mathrm{N}_{-}$distributions and their correlations. The $F_{i k}$ can be calculated from the data by making a 3 d Lego plot with base axes $\mathrm{N}_{+}$and $\mathrm{N}_{-}$and height $\mathrm{P}\left(\mathrm{N}_{+}, \mathrm{N}_{-}\right)$which costs statistical error but other methods, e.g Monte Carlo, are used.

If you measure the distribution, then you know all the corrected cumulants

Cumulants for Poisson, Negative Binomial Distributions Measured with efficiency $p$ corrected to true value, explicit in $\mu_{t}$ and $k$

| Measured Cumulant | Corrected Poisson | Corrected Negative Binomial Expanded |
| :--- | :--- | :--- |
| $\kappa_{1}=\mu$ | $\mu_{t} \equiv \mu / p$ | $\mu_{t} \equiv \mu / p$ |
| $\kappa_{2}=\mu_{2}=\sigma^{2}$ | $\mu_{t}$ | $\mu_{t}\left(1+\mu_{t} / k\right) \equiv \sigma_{t}^{2}$ |
| $\kappa_{3}=\mu_{3}$ | $\mu_{t}$ | $\mu_{t}\left(1+3 \mu_{t} / k+2 \mu_{t}^{2} / k^{2}\right)$ |
| $\kappa_{4}=\mu_{4}-3 \kappa_{2}^{2}$ | $\mu_{t}$ | $\mu_{t}\left(1+7 \mu_{t} / k+12 \mu_{t}^{2} / k^{2}+6 \mu_{t}^{3} / k^{3}\right)$ |
| $S \equiv \kappa_{3} / \sigma^{3}$ | $1 / \sqrt{\mu_{t}}$ | $\left(1+2 \mu_{t} / k\right) / \sqrt{\mu_{t}\left(1+\mu_{t} / k\right)}$ |
| $\kappa \equiv \kappa_{4} / \kappa_{2}^{2}$ | $1 / \mu_{t}$ | $\left(1+6 \mu_{t} / k+6 \mu_{t}^{2} / k^{2}\right) /\left[\mu_{t}\left(1+\mu_{t} / k\right)\right]$ |
| $S \sigma=\kappa_{3} / \kappa_{2}$ | 1 | $\left(1+2 \mu_{t} / k\right)$ |
| $\kappa \sigma^{2}=\kappa_{4} / \kappa_{2}$ | 1 | $\left(1+6 \mu_{t} / k+6 \mu_{t}^{2} / k^{2}\right)$ |
| $\mu / \sigma^{2}=\kappa_{1} / \kappa_{2}$ | 1 | $1 /\left(1+\mu_{t} / k\right)$ |
| $S \sigma^{3} / \mu=\kappa_{3} / \kappa_{1}$ | 1 | $\left(1+3 \mu_{t} / k+2 \mu_{t}^{2} / k^{2}\right)$ |

Use the NBD Cumulant Theorem allowing $\varepsilon=p$ to be different for $\mathrm{N}^{+}$and $\mathrm{N}^{-}$

$$
\kappa_{j}=\kappa_{j}^{+}+(-1)^{j} \kappa_{j}^{-}
$$

Zimanyi-Dec 2015

$$
\frac{\mu}{\sigma^{2}}=\frac{\kappa_{1}^{+}-\kappa_{1}^{-}}{\kappa_{2}^{+}+\kappa_{2}^{-}}=\frac{\mu_{t}^{+}-\mu_{t}^{-}}{\mu_{t}^{+}\left[1+\left(\frac{\mu_{t}^{+}}{k^{+}}\right)\right]+\mu_{t}^{-}\left[1+\left(\frac{\mu_{t}^{-}}{k^{-}}\right)\right]}
$$

$$
\mu_{t}=\frac{\mu}{\varepsilon}
$$

$$
\frac{S \sigma^{3}}{\mu}=\frac{\kappa_{3}^{+}-\kappa_{3}^{-}}{\kappa_{1}^{+}-\kappa_{1}^{-}}=\frac{\mu_{t}^{+}\left[1+3\left(\frac{\mu_{t}^{+}}{k^{+}}\right)+2\left(\frac{\mu_{t}^{+}}{k^{+}}\right)^{2}\right]-\mu_{t}^{-}\left[1+3\left(\frac{\mu_{t}^{-}}{k^{-}}\right)+2\left(\frac{\mu_{t}^{-}}{k^{-}}\right)^{2}\right]}{\mu_{t}^{+}-\mu_{t}^{-}}
$$

$$
S \sigma=\frac{\kappa_{3}^{+}-\kappa_{3}^{-}}{\kappa_{2}^{+}+\kappa_{2}^{-}}=\frac{\mu_{t}^{+}\left[1+3\left(\frac{\mu_{t}^{+}}{k^{+}}\right)+2\left(\frac{\mu_{t}^{+}}{k^{+}}\right)^{2}\right]-\mu_{t}^{-}\left[1+3\left(\frac{\mu_{t}^{-}}{k^{-}}\right)+2\left(\frac{\mu_{t}^{-}}{k^{-}}\right)^{2}\right]}{\mu_{t}^{+}\left[1+\frac{\mu_{t}^{+}}{k^{+}}\right]+\mu_{t}^{-}\left[1+\frac{\mu_{t}^{-}}{k^{-}}\right]}
$$

$$
\kappa \sigma^{2}=\frac{\kappa_{4}^{+}+\kappa_{4}^{-}}{\kappa_{2}^{+}+\kappa_{2}^{-}}=\frac{\mu_{t}^{+}\left[1+7\left(\frac{\mu_{t}^{+}}{k^{+}}\right)+12\left(\frac{\mu_{t}^{+}}{k^{+}}\right)^{2}+6\left(\frac{\mu_{t}^{+}}{k^{+}}\right)^{3}\right]+\mu_{t}^{-}\left[1+7\left(\frac{\mu_{t}^{-}}{k^{-}}\right)+12\left(\frac{\mu_{t}^{-}}{k^{-}}\right)^{2}+6\left(\frac{\mu_{t}^{-}}{k^{-}}\right)^{3}\right]}{\mu_{t}^{+}\left[1+\frac{\mu_{t}^{+}}{k^{+}}\right]+\mu_{t}^{-}\left[1+\frac{\mu_{t}^{-}}{k^{-}}\right]}
$$

The NBD only uses 4 quantities for this calculation: $\mu_{t}^{+}$and $\mu_{t}^{-}\left(\mu_{t} / k\right)^{+}$and $\left(\mu_{t} / k\right)^{-}$ The error on $\mu_{\mathrm{t}} \ll$ than the error on $\mu_{\mathrm{t}} / \mathrm{k}$ so is neglected. The errors are highly correlated for the sums of powers of $\mu_{\mathrm{t}} / \mathrm{k}$ in both the numerator and denominator. These correlations are handled by varying the $\left(\mu_{\mathrm{t}} / \mathrm{k}\right)^{+}$and $\left(\mu_{\mathrm{t}} / \mathrm{k}\right)^{-}$by $\pm 1 \sigma$ independently and adding the variations in quadrature

# The errors of the cumulants and ratios by 

## the direct method remain very complicated

A recent thorough treatment of both statistical errors and efficiency, with even more complicated formulas than Bzdak and Koch is given by Xiaofeng Luo, PRC 91 (2015) 034907 BUT to test the method:
"By deriving the covariance between factorial moments, one can obtain the general error formula for the efficiency corrected moments based on the error propagation derived from the Delta theorem. The Skellam-distribution-based Monto Carlo simulation is used to test the Delta theorem and bootstrap error estimation methods."

I note, of course, that Skellam is the difference between two Poissons so satisfies the integer Levy process theorem! I also note that Bzdak and Koch have not been idle PRC 91(2015) 027901

$$
\frac{S \sigma^{3}}{\mu}=\frac{\kappa_{3}^{+}-\kappa_{3}^{-}}{\kappa_{1}^{+}-\kappa_{1}^{-}}=\frac{\mu_{t}^{+}\left[1+3\left(\frac{\mu_{t}^{+}}{k^{+}}\right)+2\left(\frac{\mu_{t}^{+}}{k^{+}}\right)^{2}\right]-\mu_{t}^{-}\left[1+3\left(\frac{\mu_{t}^{-}}{k^{-}}\right)+2\left(\frac{\mu_{t}^{-}}{k^{-}}\right)^{2}\right]}{\mu_{t}^{+}-\mu_{t}^{-}}
$$

$$
R_{32}-R_{12}=S \sigma-\frac{\mu}{\sigma^{2}}=\frac{\mu_{t}^{+}\left[3\left(\frac{\mu_{t}^{+}}{k^{+}}\right)+2\left(\frac{\mu_{t}^{+}}{k^{+}}\right)^{2}\right]-\mu_{t}^{-}\left[3\left(\frac{\mu_{t}^{-}}{k^{-}}\right)+2\left(\frac{\mu_{t}^{-}}{k^{-}}\right)^{2}\right]}{\mu_{t}^{+}\left[1+\frac{\mu_{t}^{+}}{k^{+}}\right]+\mu_{t}^{-}\left[1+\frac{\mu_{t}^{-}}{k^{-}}\right]}
$$

Bzdak and Koch (and likely many others) have expressed concern about what is the "required acceptance" for an experimental result e.g. on the above quantities to compare with Lattice QCD calculations

The good news from the above equations and those on the previous pages is that if the ratios $\left(\mu_{\mathrm{t}} / \mathrm{k}\right)^{+}$and $\left(\mu_{\mathrm{t}} / \mathrm{k}\right)^{-}$don't change with the acceptance and if $\mu_{\mathrm{t}}^{+}$and $\mu_{\mathrm{t}}^{-}$scale by the same amount with the acceptance (e.g. dn/d $\eta$ constant in rapidity and azimuth) then the above formulas remain unchanged. What does nature say?

# Recall the NBD slide from E802 

E802 PRC52,2663(1995)
$\mathrm{k}(\delta \eta)$ vs $\mu(\delta \eta)$ from NBD fits


## PHENIX PRC76,0349033(2007)



The nice examples of short range correlation with $\xi$, indicated in the E802 plot, change dramatically in the newer PHENIX Au+Au ( 200 GeV ) measurement with the abrubt flattening of $k(\delta \eta)$ for $\mu(\delta \eta)>30,|\eta|>0.15$. This as far as I know is the only such measurement at RHIC or LHC. The E802 data has perfect centrality, all nucleons interact as measured in a ZDC, so the suggestion is that the flattening could be a long range correlation due to fluctuations in the number of participants in a centrality bin.

Cumulants are additive for independent processes -another NBD advantage

$$
\frac{1}{k^{\text {meas }}(\delta \eta)}=K_{2}^{\text {meas }}(\delta \eta)=K_{2}^{d y n}(\delta \eta)+K_{2}^{\text {bkg }}(\delta \eta)
$$

The two entries for E802 represent such a correction for background correlation from hits on adjacent wires.


In PRC78, PHENIX measured the effect of "geometry fluctuations" in $5 \%$ wide centrality bins and made a correction to $\mathrm{k}_{\mathrm{dyn}}=1 / \mathrm{K}_{2}{ }^{\text {dyn }}$ which is shown for the 1 overlapping bin in the PRC76 and PRC78 measurements. (This would appear to return to the trend $k / \mu \approx$ constant vs the $\delta \eta$ interval and if true at all $\delta \eta$ would preserve the cumulant ratios vs the $\delta \eta$ acceptance!)??

## Conclusions

-The NBD cumulant theorem brings a huge simplification to calculating the efficiency correction and statistical errors on net-charge cumulants.

- Acceptance corrections are much more difficult because of short range correlations in $\delta \eta$ and $\delta \phi$, but in certain cases discussed above the cumulant ratios will remain constant independent of acceptance, so would be one possible resolution to the question of the "required acceptance" to compare experiments with Lattice QCD calculations
- Fortunately, the two above issues can be further investigated by both experiment and theory. For instance if the STAR NBD data for net charge were available, I could calculate the corrected values and the errors for $\kappa \sigma^{2}$, etc. Similarly STAR could make cuts in acceptance in their measurements to determine the variation in the results and whether or where the "required acceptance" is satisfied.


## END

## Extras

- Theory OOPS
- 4 generating functions
- NBD fit plots
- $k(\delta \eta)$ PRC76,0349033(2007)
- other goodies

PH水ENIX M.J.Tannenbaum 30

## $\mathrm{T}_{\mathrm{f}}$ difference for STAR raw vs corrected

a

$\mathrm{T}_{\mathrm{f}}<155 \mathrm{MeV} 1$ std

$\mathrm{T}_{\mathrm{f}}<170 \mathrm{MeV} 1 \mathrm{std}$

## 4 Generating functions

Moment generating fn
Cumulant generating function
$M_{x}^{\prime}(t)=\left\langle e^{t x}\right\rangle \quad g_{x}(t)=\ln M_{x}^{\prime}(t)=\ln \left\langle e^{t x}\right\rangle$

Factorial moment gen fn.
$M_{x}(t)=\left\langle(1+t)^{x}\right\rangle$

Factorial cumulant gen fn.

$$
g_{x}(t)=\ln \left\langle(1+t)^{x}\right\rangle
$$

PHENIXAuAu Multiplicity $\mathrm{N}_{\mathrm{ch}}$ PRC 78, (2008) 044902

http://www.osti.gov/scitech/servlets/purl/10108142 U.S. DEPARTMENT OF ENERGY

BRDDKMANEN

It's not a Gaussian... it's a Gamma distribution!


Also: It's not Poisson, it's negative binomial


PHENIX: centrality 0-5\% $\sqrt{s}_{\mathrm{NN}}=7.7 \mathrm{GeV}$


## PHENIX k( $\delta \eta)$ PRC76,0349033(2007)



## New STAR net-p Preliminary




| Binomial p | So | ко^2 |
| :---: | :---: | :---: |
| p | 1-2p | $1-6 p+6 p^{\wedge} 2$ |
| 0.01 | 0.98 | 0.941 |
| 0.02 | 0.96 | 0.882 |
| 0.03 | 0.94 | 0.825 |
| 0.04 | 0.92 | 0.770 |
| 0.05 | 0.9 | 0.715 |
| 0.06 | 0.88 | 0.662 |
| 0.07 | 0.86 | 0.609 |
| 0.08 | 0.84 | 0.558 |
| 0.09 | 0.82 | 0.509 |
| 0.1 | 0.8 | 0.460 |
| 0.15 | 0.7 | 0.235 |
| 0.2 | 0.6 | 0.040 |
| 0.25 | 0.5 | -0.125 |
| 0.3 | 0.4 | -0.260 |
| 0.35 | 0.3 | -0.365 |
| 0.4 | 0.2 | -0.440 |
| 0.45 | 0.1 | -0.485 |
| 0.5 | 0 | -0.500 |


How can adding tracks $>0.8 \mathrm{GeV} / \mathrm{c}$ make such changes in $\kappa \sigma^{2}$ but not in $\mathrm{S} \sigma$

## Taylor expansion of the pressure

$$
\frac{p}{T^{4}}=\frac{1}{V T^{3}} \ln Z\left(V, T, \mu_{B}, \mu_{S}, \mu_{Q}\right)
$$

$$
=\sum_{i, j, k} \frac{1}{i!j!k!} \chi_{i j k}^{B Q S}\left(\frac{\mu_{B}}{T}\right)^{i}\left(\frac{\mu_{Q}}{T}\right)^{j}\left(\frac{\mu_{S}}{T}\right)^{k}
$$

generalized susceptibilities: $\quad \chi_{i j k}^{B Q S}=\left.\frac{\partial^{i+j+k} p / T^{4}}{\partial \hat{\mu}_{B}^{i} \partial \hat{\mu}_{Q}^{j} \partial \hat{\mu}_{S}^{k}}\right|_{\mu=0}$
conserved charge fluctuations: $\chi_{n}^{X}\left(T, \mu_{B}, \ldots\right)=\frac{\partial^{n} P / T^{4}}{\partial\left(\mu_{X} / T\right)^{n}}$

$$
X=B, Q, S
$$

$$
\frac{M_{X}}{\sigma_{X}^{2}}=\frac{\chi_{1}^{X}(T, \mu)}{\chi_{2}^{X}(T, \mu)}, S_{X} \sigma_{X}=\frac{\chi_{3}^{X}(T, \mu)}{\chi_{2}^{X}(T, \mu)}, \kappa_{X} \sigma_{X}^{2}=\frac{\chi_{4}^{X}(T, \mu)}{\chi_{2}^{X}(T, \mu)}
$$

## Hot off the presses-LBL Press release June 24, 2011

 Lattice and Experiment Compared-a first?
## Sourendu Gupta, et al., Science 332,1525 (2011)-LBL press release

(0) Eneray

A-Z INDEX | PHONE BOOK | JOBS | SEARCH

## When Matter Melts

By comparing theory with data from STAR, Berkeley Lab scientists and their colleagues map phase changes in the quark-gluon plasma

June 23, 2011
Theory:Lattice shows huge deviation of $\mathrm{T}^{2} \chi^{(4) /} \chi^{(2)}$ from 1 near 20 GeV , suggesting critical fluctuations. Expt $\kappa \sigma^{2}$ : maybe but with big errors.

I had to do lots of work to address this issue in my Erice lecture to understand whether this physics by press-release (not published in PRL) was also Baloney According to Karsch and later measurements, it was!!!

M. J. Tannenbaum 38

## Short range multiplicity correlations do not vanish in $\mathrm{A}+\mathrm{A}$ collisions!

- Short range multiplicity correlations in p-p collisons come largely from hadron decays such as $\rho \rightarrow \pi \pi, \Lambda \rightarrow \pi^{-} p$, etc., with correlation length $\xi \sim 1$ unit of rapidity
- In A+A collisions the chance of getting two particles from the same $\rho$ meson is reduced by $\sim 1 / \mathrm{N}_{\text {part }}$ so that the only remaining correlations are Bose-Einstein Correlations---when two identical Bosons, e.g. $\pi^{+} \pi^{+}$, occupy nearly the same coordinates in phase space so that constructive interference occurs due to the symmetry of the wave function from Bose statistics---a quantum mechanical effect, which remains at the same strength in A+A collisions:the amplitudes from the two different points add giving a large effect also called Hanbury-Brown Twiss (HBT).

$$
\begin{aligned}
& \text { See W.A.Zajc, et al, } \\
& \text { PRC } 29(1984) 2173
\end{aligned}
$$

## HBT effects in 2-particle Correlations

- The normalized two-particle short range rapidity correlation $R_{2}\left(y_{1}, y_{2}\right)$ is defined as

$$
\begin{equation*}
R_{2}\left(y_{1}, y_{2}\right) \equiv \frac{C_{2}\left(y_{1}, y_{2}\right)}{\rho_{1}\left(y_{1}\right) \rho_{1}\left(y_{2}\right)} \equiv \frac{\rho_{2}\left(y_{1}, y_{2}\right)}{\rho_{1}\left(y_{1}\right) \rho_{1}\left(y_{2}\right)}-1=R(0,0) e^{-\left|y_{1}-y_{2}\right| / \xi} \tag{8}
\end{equation*}
$$

where $\rho_{1}(y)$ and $\rho_{2}\left(y_{1}, y_{2}\right)$ are the inclusive densities for a single particle (at rapidity $y$ ) or 2 particles (at rapidities $y_{1}$ and $\left.y_{2}\right), C_{2}\left(y_{1}, y_{2}\right)=\rho_{2}\left(y_{1}, y_{2}\right)-\rho_{1}\left(y_{1}\right) \rho_{1}\left(y_{2}\right)$ is the Mueller correlation function for 2 particles (which is zero for the case of no correlation), and $\xi$ is the two-particle short-range rapidity correlation length $[3]$ for an exponential parameterization.

$$
K_{2}(\delta \eta)=2 R(0,0) \frac{\left(\delta \eta / \xi-1+e^{-\delta \eta / \xi}\right)}{(\delta \eta / \xi)^{2}} \quad \text { for NBD: } \mathbf{k}(\boldsymbol{\delta} \boldsymbol{\eta})=\mathbf{1} / \mathbf{K}_{\mathbf{2}}(\boldsymbol{\delta} \boldsymbol{\eta})
$$

The rapidity correlation length $\xi=0.2$ for $\mathrm{Si}+\mathrm{Au} \mathrm{E} 802$, PRC56(1977) 1544 is from HBT.

## if $\boldsymbol{\delta} \boldsymbol{\eta} \ll \xi, k \rightarrow 1 / R(\mathbf{0}, 0)=$ constant $\quad$ if $\boldsymbol{\delta} \boldsymbol{\eta} \gg \xi, k / \boldsymbol{\delta} \boldsymbol{\eta} \approx \mathrm{k} / \boldsymbol{\mu} \rightarrow$ constant

${ }^{\circ}$ For HBT analyses of two particles with $\mathbf{p}_{1}$ and $\mathbf{p}_{2}, \mathrm{C}^{\mathrm{HBT}}{ }_{2}(\mathbf{q})=R_{2}\left(\mathbf{p}_{1}-\mathbf{p}_{2}\right)+1$ and the random (un-correlated) distribution is taken from particles with $\mathbf{p}_{1}$ and $\mathbf{p}_{2}$ on different events. The HBT correlation function is taken as a Gaussian not an exponential as in (8) and is written:

$$
C_{2}^{H B T}=1+\lambda \exp -\left(R_{\text {side }}^{2} q_{\text {side }}^{2}+R_{\text {out }}^{2} q_{\text {out }}^{2}+R_{\text {long }}^{2} q_{\text {long }}^{2}\right)
$$

