

Markov Chain Monte Carlo technics applied to Parton Distribution Functions determination: proof of concept

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Introduction

- From minimization to Bayesian inference
 - Formulation of PDFs determination in terms of Bayesian inference
 - Markov chains
 - Metropolis algorithm
 - Hybrid (or Hamiltonian) Monte Carlo
- Markov chain analysis
 - Thermalisation, autocorrelation and all that....
- First results
 - PDFs Parameters
 - PDFs probability distribution functions

Conclusions



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Uncertainties in global PDFs analysis

Precise assessment of PDFs uncertainties has become crucial

PDFs subject both to theoretical (input parametrization, neglected higher order corrections, heavy flavors treatment,...) and experimental uncertainties.

To deal with experimental errors:

- Hessian method: based on linear error propagation
- Lagrange multiplier method
- . . .

Assumption on the permissible range of "acceptable" $\Delta\chi^2$ for the fit and choice of a tolerance parameter T^{-1} .

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¹This excepts Neural Networks technics

Traditional propagation of experimental uncertainties

- $\Delta\chi^2=1$ for 68% C.L. IF fitting consistent data sets with ideal gaussian errors
- In practice: inconsistencies between fitted data sets and unknown exp. and theoretical uncertainties, so not appropriate for global PDFs analysis \implies choice of tolerance criteria $\Delta \chi^2 \neq 1$.

Ex.: MRST[hep-ph/0211080]

"We estimate $\Delta \chi^2 = 50$ to be a conservative uncertainty (perhaps of the order of a 90% confidence level or a little less than 2σ) due to the observation that an increase of 50 in the global χ^2 [...]usually signifies that the fit to one or more data sets is becoming unacceptably poor. We find that an increase $\Delta \chi^2$ of 100 normally means that some data sets are very badly described by the theory."

What can MCMC methods tell us about PDFs uncertainties?

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Introduction

Very general flow diagram of PDFs extraction



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MCMC PDFs

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PDFs determination in terms of Bayesian inference

• \vec{q} : vector of PDFs parameters to be determined: $\vec{q} = (q^{(1)}, q^{(2)}, \dots, q^{(m)})^T$

xFitter functional form at $Q_0^2\sim 2~{\rm GeV^2}$:

 $xf_a(x) = A_a x^{B_a} (1-x)^{C_a} (1+D_a x + E_a x^2).$

where a labels a parton $(g, u_{val}, d_{val}, \dots)$.

• D: data

Bayesian inference:

Both model parameters and observables considered random quantities. Aims at determining a joint probability distribution $P(D, \vec{q})$ over all random quantities:

$$P(D,\vec{q}) = P(D|\vec{q})P(\vec{q})$$

- $P(\vec{q})$ prior distribution
- $P(D|\vec{q})$: likelihood of the data $\mathcal{L}(\vec{q})$

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Bayes theorem
Express
$$P(\vec{q}|D)$$
 in terms of the likelihood $P(D|\vec{q})$:

$$P(\vec{q}|D) = \frac{P(D|\vec{q})P(\vec{q})}{\int d\vec{q}P(D|\vec{q})P(\vec{q})}$$
(1)

- $P(\vec{q}|D)$: posterior probability density
- Can be sampled using a Monte Carlo algorithm

PDFs determination in terms of Bayesian inference

Likelihood

Let's note

- D_i and T_i : respectively the i^{th} experimental point and the corresponding theoretical calculation
- σ_i^2 the uncertainty associated with the measured data *i*.
- If $\frac{(D_i T_i)}{\sigma_i}$ independent and normally distributed:

$$\log \mathcal{L}(\hat{q}) = -\frac{1}{2} \sum_{i=1}^{n} \frac{(D_i - T_i)^2}{\sigma_i^2} = -\frac{1}{2} \chi^2$$
(2)

Possibility to construct more involved likelihood with correlated data

Prior distribution

We choose $P(\vec{q})$ uniform

Remark: If D_i distributed according to a Gaussian law, MLE and LSM equivalent a = b + a =

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In summary...



Metropolis [Metropolis et al., J. Chem Phys. 21 (1953)]

- one of the simplest Monte Carlo algorithm
- standard computational workhorse of MCMC methods
- in principle applicable to any system
- extremely straightforward to implement and to sample a target density $P(\vec{q}|D)$
- at each Monte Carlo time t, the next state \vec{q}_{t+1} is chosen by sampling a candidate point \vec{q}' from a proposal distribution^(*). The candidate point is then accepted with the probability

$$\alpha(\vec{q_t}, \vec{q'}) = \min\left(1, \frac{P(\vec{q'}|D)}{P(\vec{q_t}|D)}\right)$$

i.e. in our case

$$\alpha(\vec{q_t}, \vec{q'}) = \min\left(1, e^{-\frac{1}{2}\Delta\chi^2}\right)$$

(*) Proposal distribution assumed symmetric here

and why we do not use it.....



Acceptance test

- typically for 1 parameter: 30-50%
- decreases as ~ 0.5^m

 \rightarrow For m=10 parameters, unacceptable acceptance...

trial point far from the initial one \implies large change in the distribution to sample point close to the initial one \implies inefficient exploration of the parameter space

Metropolis algorithms (even flavored with multivariate Gaussian distributions or binary space partitioning) are NOT suited for realistic PDFs determination ($m \sim 25$).

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HMC

- Developed originally for lattice field theory [Duane et al, 1987]
- Combines molecular dynamics evolution with a Metropolis accept/reject step

Introduce for each set of parameters \vec{q} a set of conjugate momenta \vec{p} Associates an Hamiltonian $H(\vec{q},\vec{p}) = \vec{p}^T M^{-1} \vec{p}/2 + \mathcal{U}(q)$, where M is a mass matrix, and $\mathcal{U}(q)$ an arbitrary potential energy.

This allows to define a joint distribution as

$$P(q,p) = \frac{1}{Z}e^{-H(q,p)} = \frac{1}{Z}e^{-\mathcal{K}(p)}e^{-\mathcal{U}(q)} \qquad \qquad \text{Z normalizing constant}$$

We use for the potential energy $\mathcal{U}(q) = -\log[P(D|\vec{q})P(\vec{q})].$

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- $\bullet\,$ choose $\vec{p_0}$ normally distributed
- let the system evolve deterministically
- candidate point $\vec{q_1}$ accepted with probability $min(1, e^{-\Delta H})$ i.e. 100% !!!

In practice: acceptance degraded because of numerical resolution of Hamilton equations, but still very high (typically $\sim 80-90\%$, independently of the dimension of the chain).



Sampling of a 100D gaussian. Values of the variable with largest standard deviation.

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[R. M. Neal, arXiv:1206.1901]

Implementation of HMC in xFitter package

- parametrized HERA PDFs: xu_v , xd_v , xg, $x\overline{U} = x\overline{u}$, $x\overline{D} = x\overline{d} + x\overline{s}$
- functional form at $Q_0^2 \sim 2 \ {\rm GeV^2}$:

$$xf_a(x) = A_a x^{B_a} (1-x)^{C_a} (1+D_a x + E_a x^2)$$
(3)

where a labels a parton (g, u_{val} , d_{val} , ...).

- we consider 10 free parameters: B_g , C_g , $B_{u_{val}}$, $C_{u_{val}}$, $E_{u_{val}}$, $C_{d_{val}}$, $C_{\overline{U}}$, $A_{\overline{D}}$, $B_{\overline{D}}$ and $C_{\overline{D}}$
- validation study, with combined inclusive $e^{\pm}p$ scattering cross-sections from H1 and ZEUS (\sim 600 data points), and ZMVFN scheme.
- comparison with HERAPDF1.0 with ZMVFNS.

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Analysis (1/3)

Extracting observables and their statistical errors from Monte Carlo simulations: subtle task, requiring careful treatment of the Markov chain.

• The thermalization time (or burn-in length b) : number of states $\{\hat{q}_t\}_{t=1,\dots,b}$ to be discarded from the beginning so that the chain forgets its starting point.



We have taken $P(\hat{q}_b|D) > P_{1/2}$

Analysis (2/3)

• Autocorrelation: inherent correlations from one point to the next. Usual estimate of root-mean-square deviation of an observable O:

$$\sigma_{naive}^2 = \frac{N}{N-1} \left(\langle O^2 \rangle - \langle O \rangle^2 \right)$$

- Relies on the assumption that measurements performed on the Markov chain are NOT correlated
- Several methods to account for the correlations: jackkniffe binning (pre-averaging over blocks of data), Γ-method (explicit determination of autocorrelation functions and times),...
- in practice for this work: $\tau_{corr} \sim 2$.

Analysis (3/3)

• Convergence: start from different points, and check we reach the same distribution





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Ready to compare standard minimization and Monte Carlo method....

Comparison of results obtained for 2 parameters of the gluon PDF using Markov Chain procedure or χ^2 (MINUIT) minimisation:

parameter	values	MCMC	MINUIT minimization
	mean	-0.0537 ± 0.0002	
$\mathbf{B}_{\mathbf{g}}$	most probable	-0.0632 ± 0.0168	-0.0559
	standard deviation	0.0299 ± 0.0001	0.0288
	mean	5.9483 ± 0.0025	
C_{g}	most probable	5.8952 ± 0.0615	5.9274
	standard deviation	0.5037 ± 0.0019	0.5078

Much more information with MCMC

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PDF Parameters: Marginal probability distribution and correlations



PDF Parameters: Marginal probability distribution and correlations



PDF Parameters: χ^2 distribution

 χ^2 distribution for a 10D MCMC. The solid line is an adjustment with a χ^2 distribution law.



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Probability distribution functions of PDFs (PDFs of PDFs...)

For each MC parameter set $\vec{q_t}$, compute the corresponding PDFs



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Gluon PDF probability distribution function for $x \approx 10^{-4}$ (l.h.s.) and $x \approx 0.83$ at fixed $Q^2 = 10$ GeV². The 68% confidence interval is obtained considering the region of the distribution containing 68% of the data remaining on each side of the most probable value.

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MCMC PDFs



The parton distribution functions obtained using MCMC (right) compared to HERAPDF1.0 (ZMVFN scheme) from xFitter output (left) for xu_{val} and xg, at $Q^2 = 10 \text{ GeV}^2$. The bands show the 68% confidence interval around the most probable value for the MCMC PDFs, and the standard $\Delta\chi^2 = 1$ deviation for HERAPDF.

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What we have done....

- MCMC, well-suited to multi-parameter determination, applicable to PDFs extraction
- Overcome the technical difficulties by a nice recycling of a lattice algorithm (HMC)
- Obtained probability densities of PDFs and as by-product, confidence intervals

What we would like to do....

- Consider more complex χ^2 functions including correlations and study prior influence
- Extend this work to a competitive PDF ensemble with more parameters and data
- Extract $lpha_s$ and possibly m_c, m_b from data using MCMC

What we have NOT done....

• Tackled the problem of potentially incompatible data, but....

Bayesian approach applied to global analyses can lead to a deeper insight into PDFs uncertainties determination

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