

Search and Discovery Statistics in HEP

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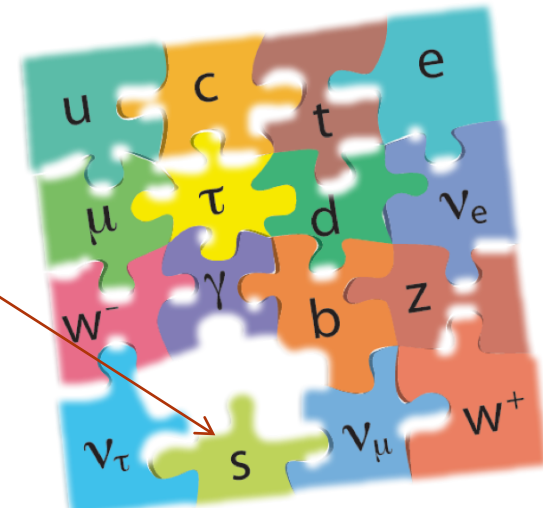
This presentation would have not been possible without the tremendous help of the following people throughout many years

Louis Lyons, Alex Read, Glen Cowan, Kyle Cranmer
Ofer Vitells & Bob Cousins

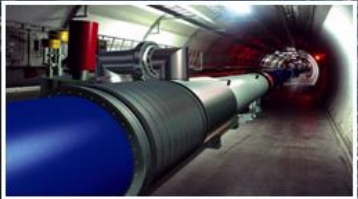
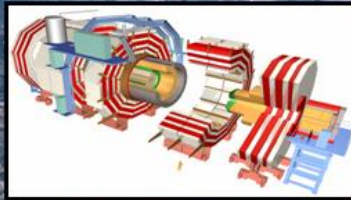


What is the statistical challenge in HEP?

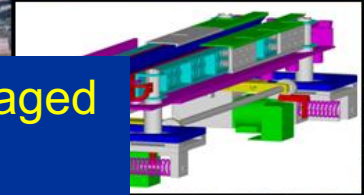
- High Energy Physicists (**HEP**) have an hypothesis:
The Standard Model.
- This model relies on the existence of the 2012 discovery of
the Higgs Boson
- The minimal content of the Standard Model includes the Higgs Boson , but extensions of the Model include other particles which are yet to be discovered
- The challenge of HEP is to generate tons of data and to develop powerful analyses to tell if the data indeed contains evidence for the new particle, and confirm if it is the expected Higgs Boson (Mass, Spin, CP) or a member of a family of Scalar Bosons



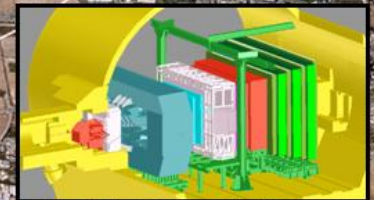
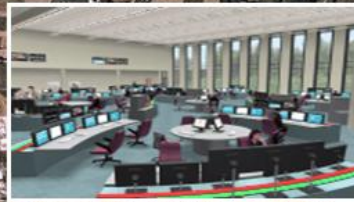
The Large Hadron Collider (LHC)



The LHC is a very powerful accelerator which managed to hunt a Higgs with a 10^{-12} production probability




This is statistics of rare events!



Higgs Hunter's Independence Day

July 4th 2012

From Wikipedia: On 4 July 2012, the discovery of a new particle with a mass between 125 and 127 GeV/c² was announced; physicists suspected that it was the Higgs boson. Since then, the particle has been shown to behave, interact, and decay in many of the ways predicted by the Standard Model.



I was looking
for you for
over 20 years

Now you
have found
me

The Charge of the Lectures



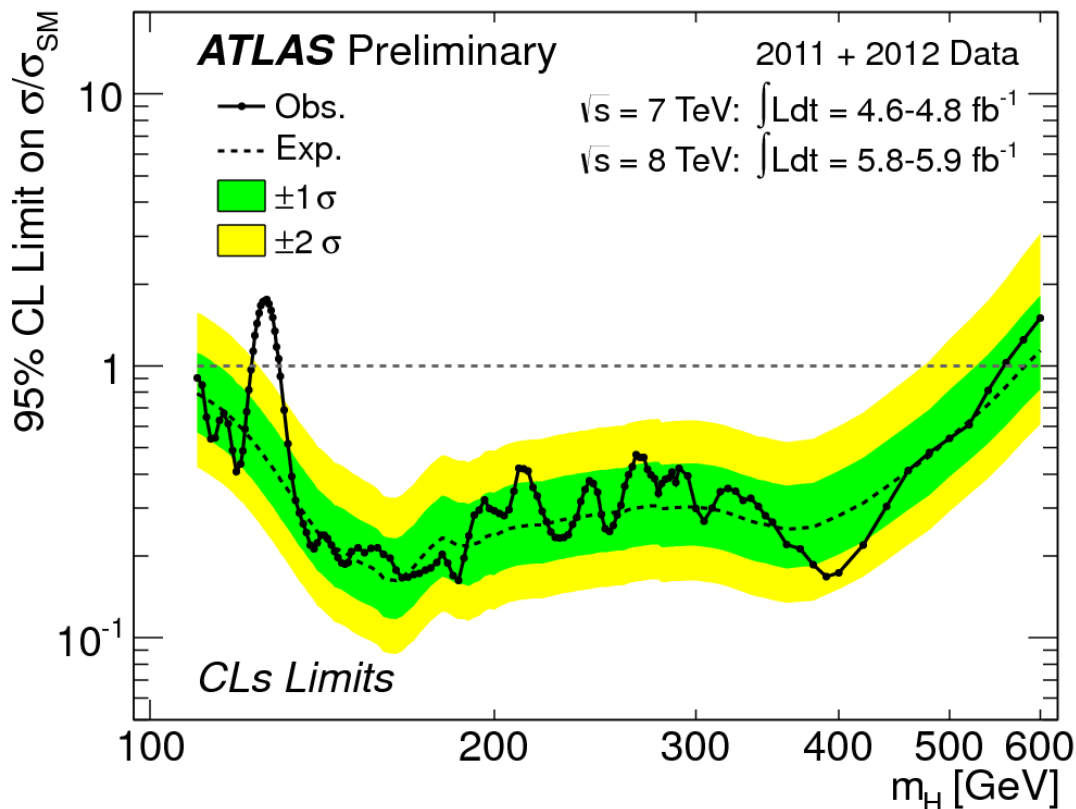
The Brazil Plot, what does it mean?

How do we exclude theory parameters?

Observed Limit

Bands

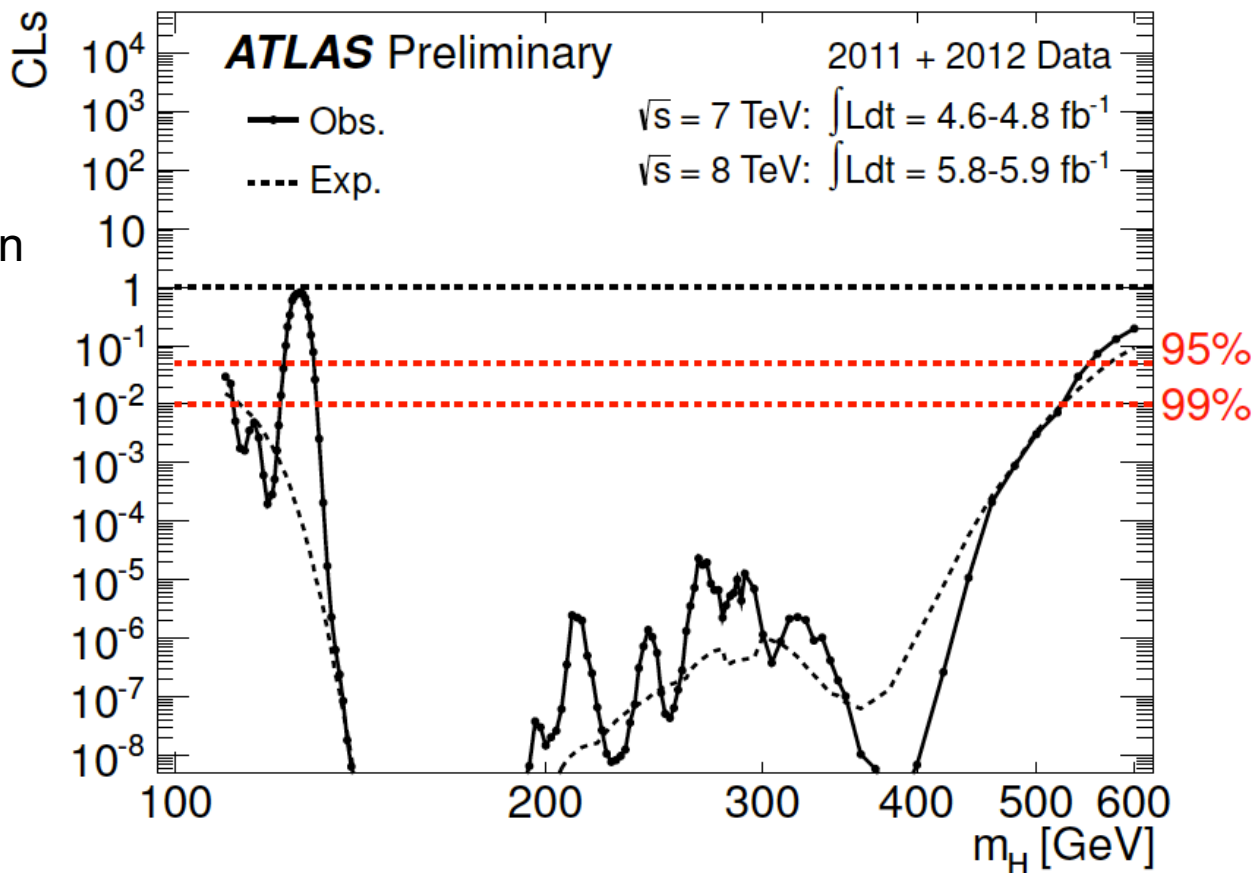
Expected Limit



What the --- CLs?

What is exclusion
at the 95% CL?

99% CL?



How do we make discoveries?

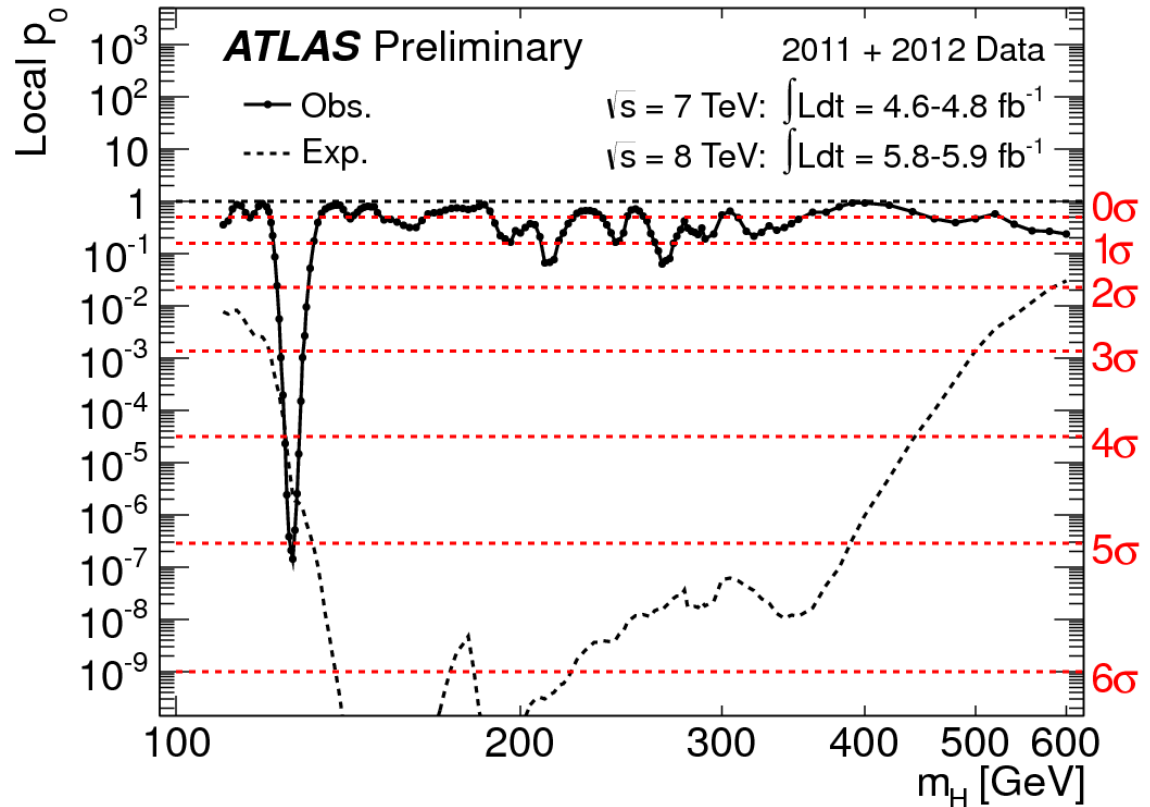
The p_0 discovery plot, how to read it?

p-value

Local p_0

Expected p_0

Observed local p_0



Global p_0 and the Look Elsewhere Effect

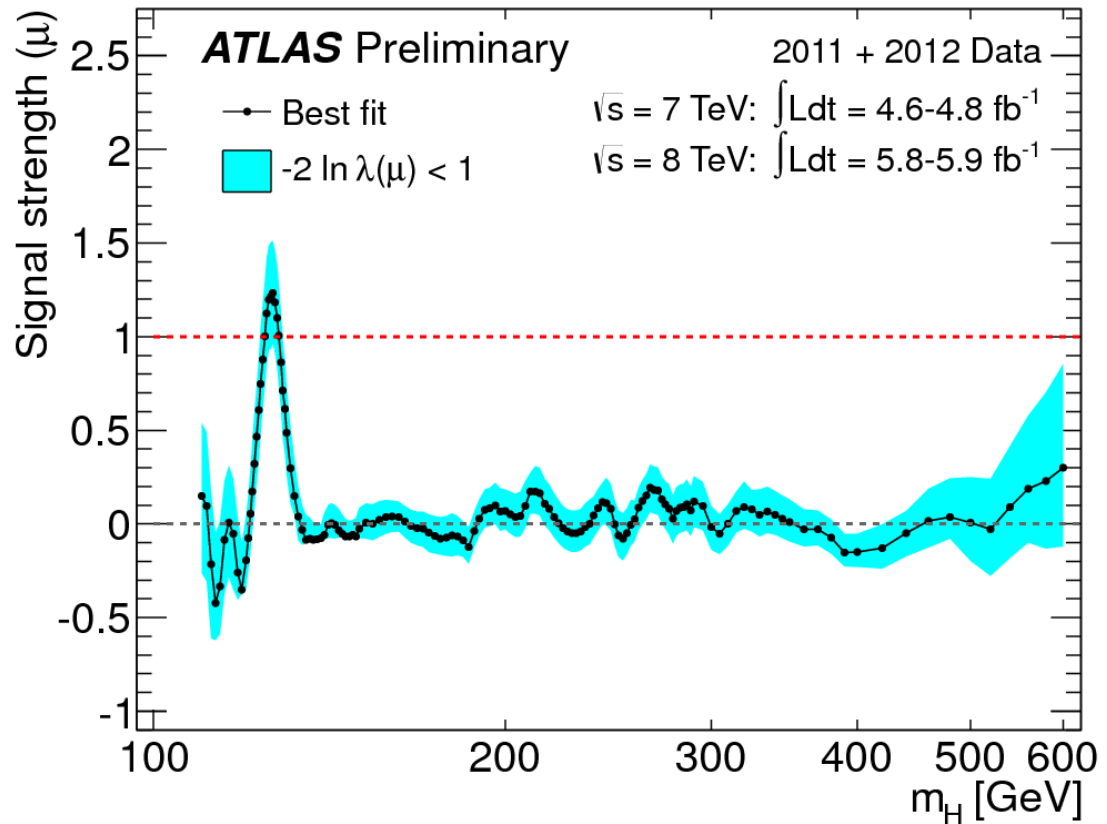


How do we make discoveries?

The cyan band plot, what is it?

What is mu hat?

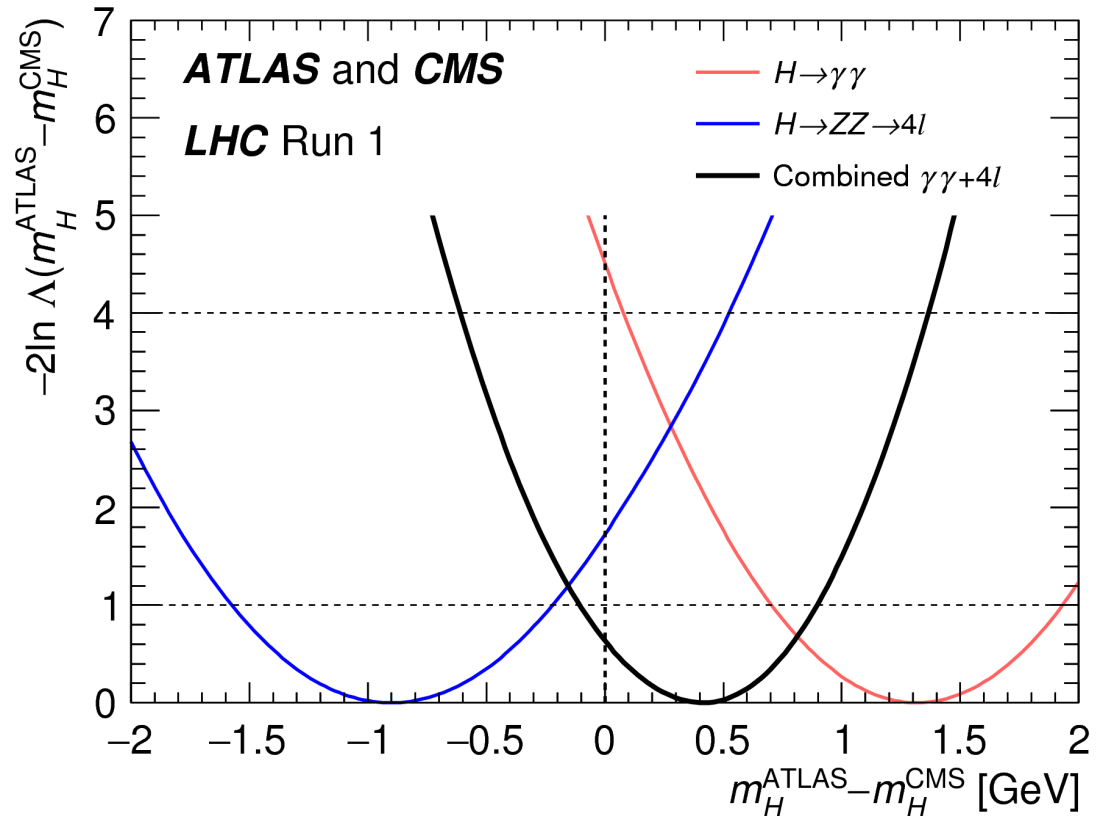
$\hat{\mu}$



How do we estimate parameters?

Towards a measurement

Likelihoods Scans



How do we take uncertainties into account?

Towards a measurement

Measurements &
Systematics vs
Stat errors

ATLAS $H \rightarrow \gamma\gamma$

CMS $H \rightarrow \gamma\gamma$

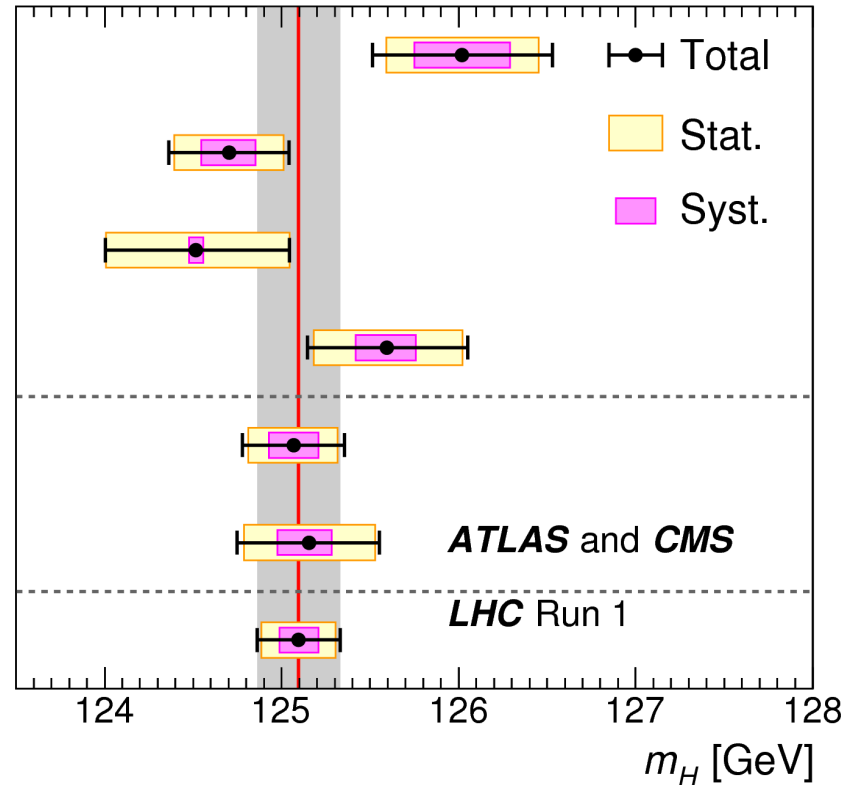
ATLAS $H \rightarrow ZZ \rightarrow 4l$

CMS $H \rightarrow ZZ \rightarrow 4l$

ATLAS+CMS $\gamma\gamma$

ATLAS+CMS $4l$

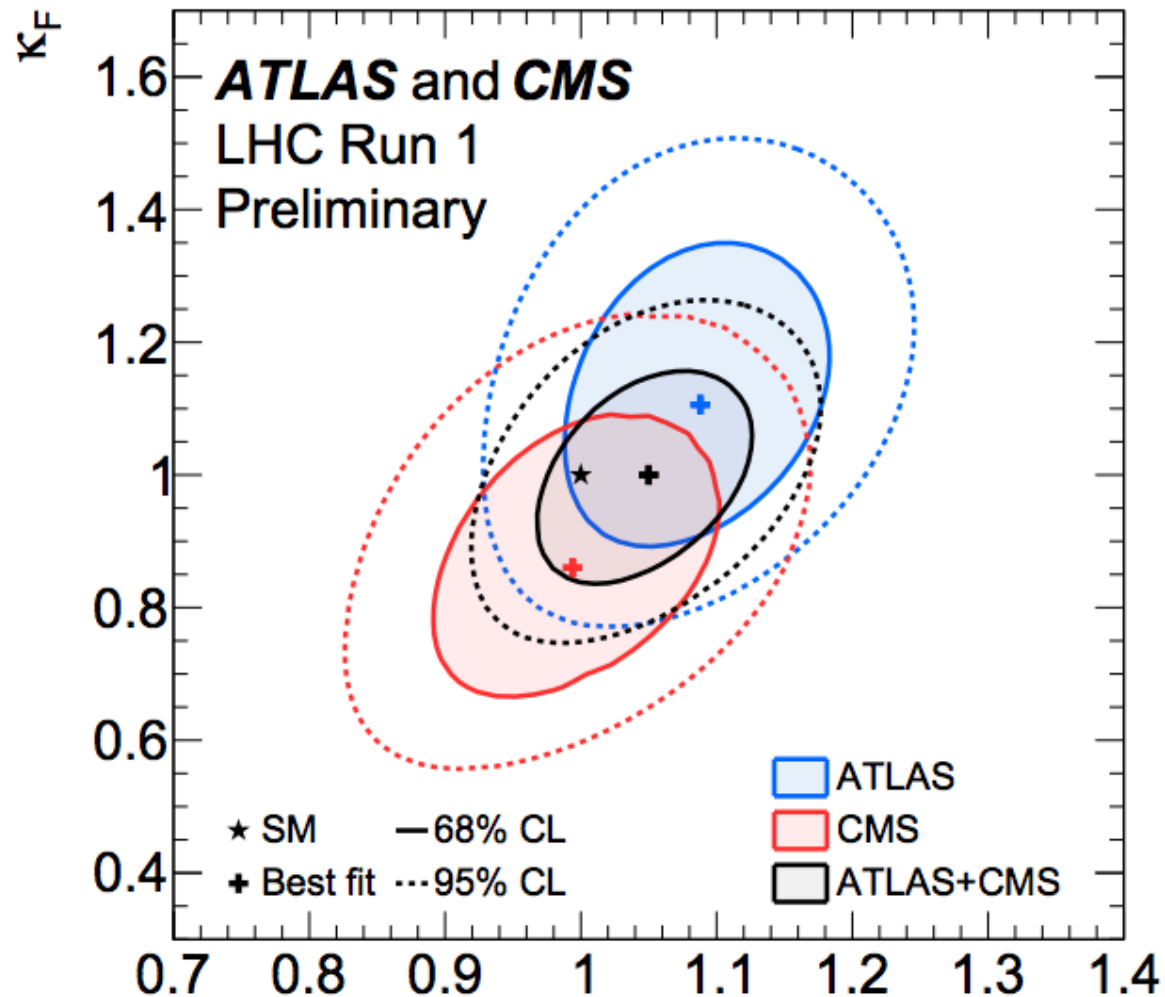
ATLAS+CMS $\gamma\gamma+4l$



How do we estimate parameters?

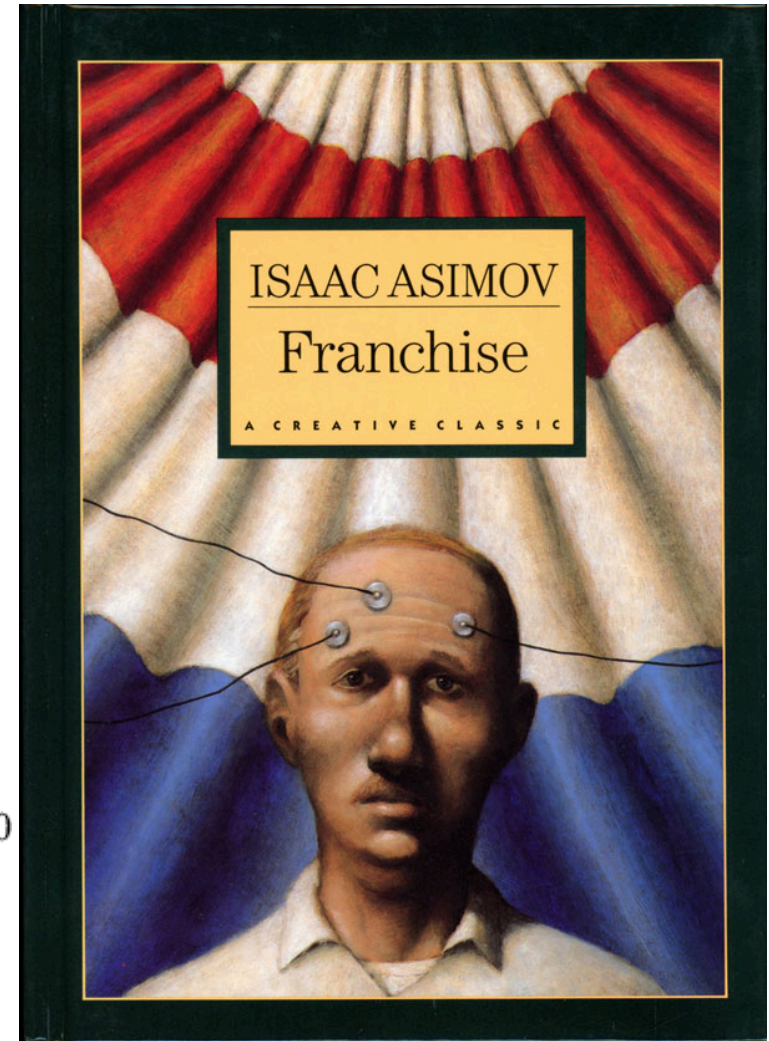
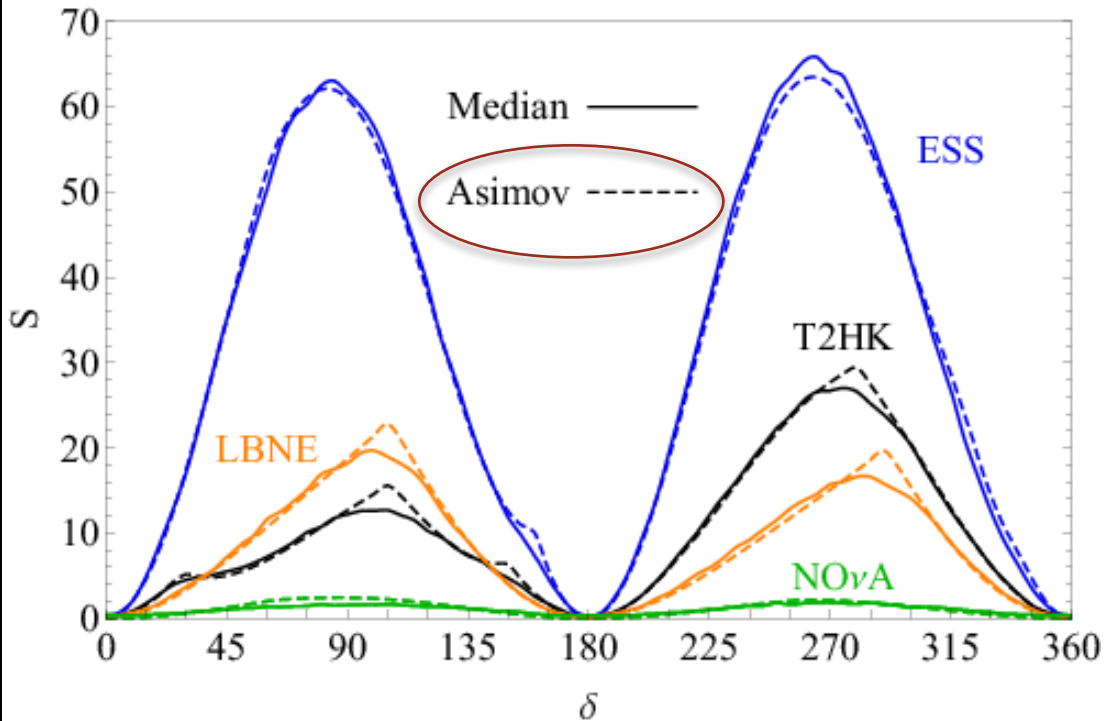
Towards a measurement

2-D Likelihoods



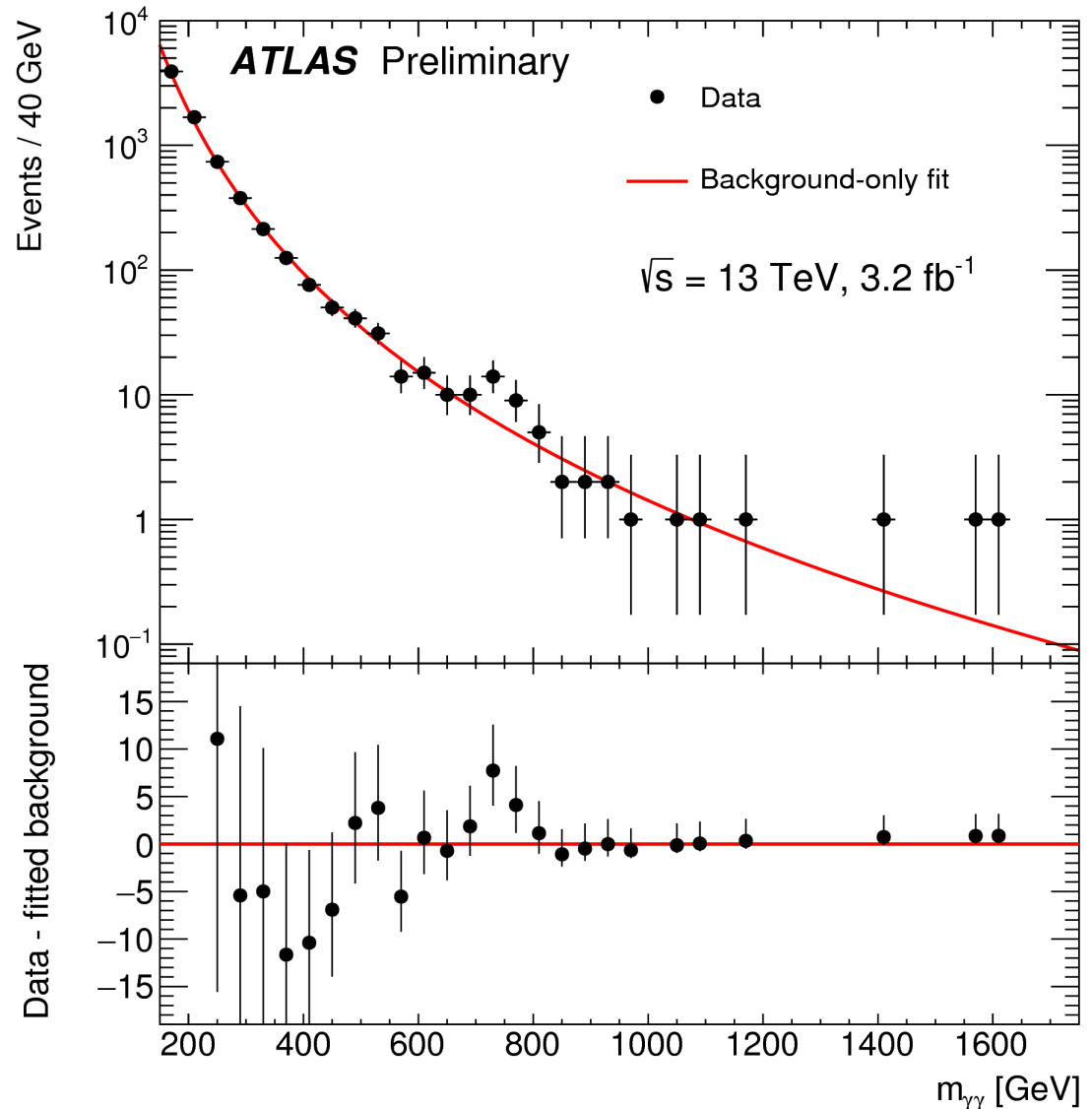
How do we estimate the expected sensitivity?

The Asimov Data Set



The Look Elsewhere Effect in 1 and 2D

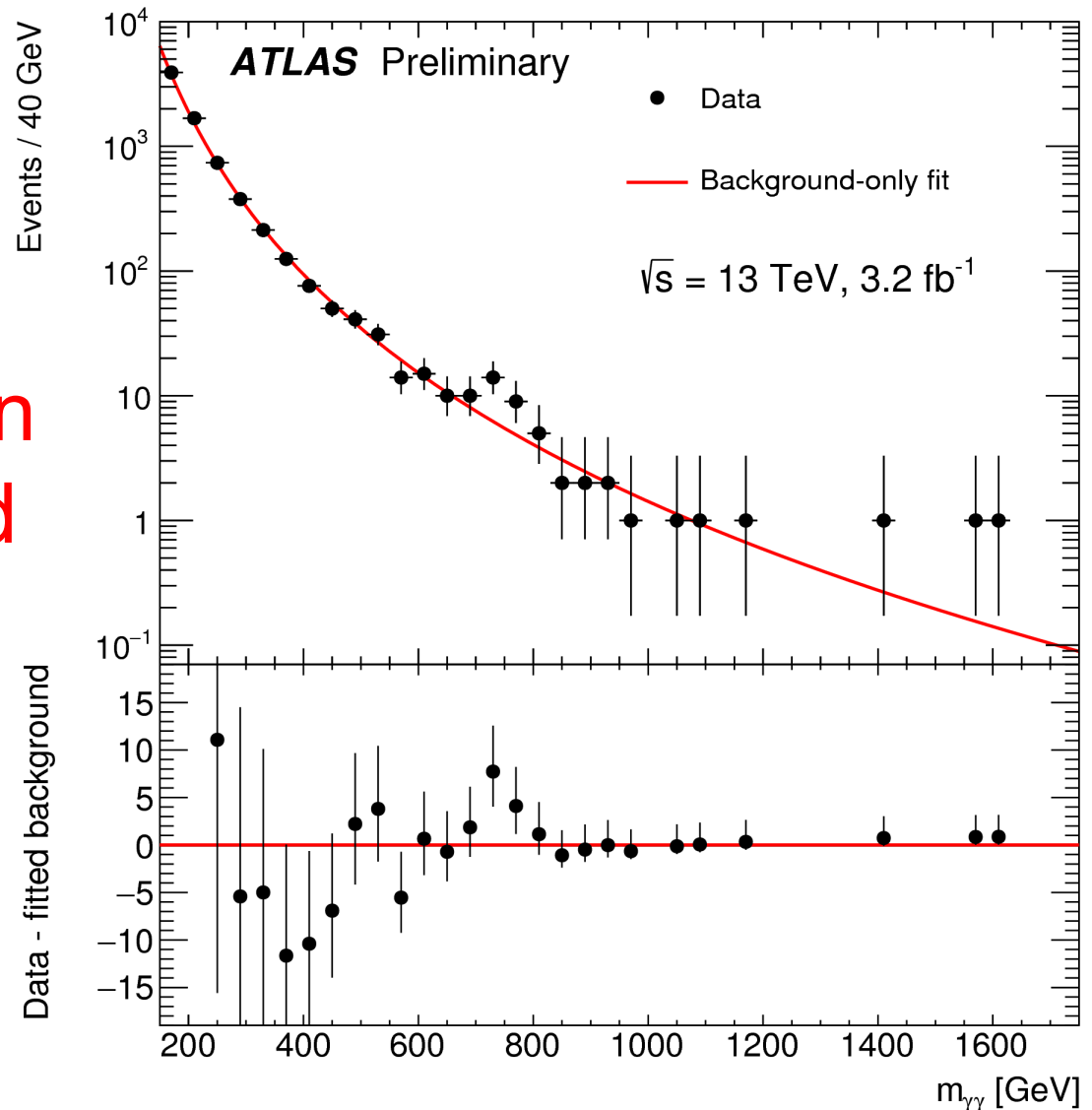
What is the probability for a random fluctuation of the background at a mass of 750 GeV with a width of 45 GeV ?



The Look Elsewhere Effect in 1 and 2D

LEE

What is the probability for a random fluctuation of the background at **SOME** mass with **SOME** width anywhere in the search range?



References in the Discovery Papers

ATLAS

- PL**[26] G. Cowan, K. Cranmer, E. Gross and O. Vitells, *Asymptotic formulae for likelihood-based tests of new physics*, *Eur. Phys. J.* **C71** (2011) 1554. **CCGV**
- CLs**[27] A. L. Read, *Presentation of search results: The CL(s) technique*, *J. Phys.* **G28** (2002) 2693–2704.
- LEE**[28] E. Gross and O. Vitells, *Trial factors for the look elsewhere effect in high energy physics*, *Eur. Phys. J.* **C70** (2010) 525–530.

CMS

- PL**[90] G. Cowan et al., “Asymptotic formulae for likelihood-based tests of new physics”, *Eur. Phys. J. C* **71** (2011) 1–19, doi:10.1140/epjc/s10052-011-1554-0, arXiv:1007.1727. **CCGV**
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- CLs**[93] A. L. Read, “Presentation of search results: the CLs technique”, *J. Phys. G: Nucl. Part. Phys.* **28** (2002) 2693, doi:10.1088/0954-3899/28/10/313.
- L**[94] Gross, E. and Vitells, O., “Trial factors for the look elsewhere effect in high energy physics”, *Eur. Phys. J. C* **70** (2010) 525–530, doi:10.1140/epjc/s10052-010-1470-8, arXiv:1005.1891.

More Refs (taken from CMS legacy Run 1 Paper)

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Wald Approximation

- [187] A. Wald, “Tests of statistical hypotheses concerning several parameters when the number of observations is large”, *Trans. Amer. Math. Soc.* **54** (1943) 426, doi:10.1090/S0002-9947-1943-0012401-3.

Wald Approximation

- [188] R. F. Engle, “Chapter 13 Wald, likelihood ratio, and Lagrange multiplier tests in econometrics”, in *Handbook of Econometrics*, Z. Griliches and M. D. Intriligator, eds., volume 2, p. 775. Elsevier, 1984. doi:10.1016/S1573-4412(84)02005-5.

- [189] G. J. Feldman and R. D. Cousins, “Unified approach to the classical statistical analysis of small signals”, *Phys. Rev. D* **57** (1998) 3873, doi:10.1103/PhysRevD.57.3873, arXiv:physics/9711021.

Feldman-Cousins

From ATLAS di-photon 750 GeV “resonance”

- [41] O. Vitells and E. Gross, *Estimating the significance of a signal in a multi-dimensional search*, *Astropart. Phys.* **35** (2011) 230, arXiv: 1105.4355 [astro-ph.IM].

LEE 2D



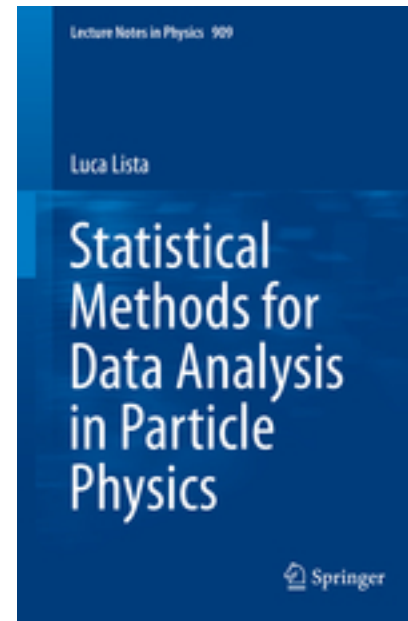
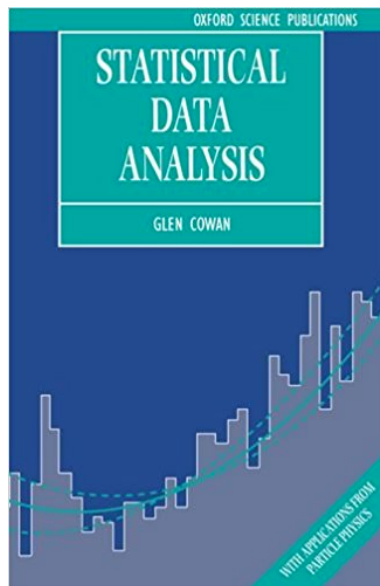
Books

G. Cowan, *Statistical Data Analysis*, Clarendon Press, Oxford, 1998.

R.J.Barlow, *A Guide to the Use of Statistical Methods in the Physical Sciences*, John Wiley, 1989;

L.Lyons, *Statistics for Nuclear and Particle Physics*, CUP, 1986

L. Lista *Statistical methods for Data Analysis*, Springer, 2015



Preliminaries

X is a random variable

Probability Distribution Function

PDF

$$P(x \in [x, x + dx]) = f(x)dx$$

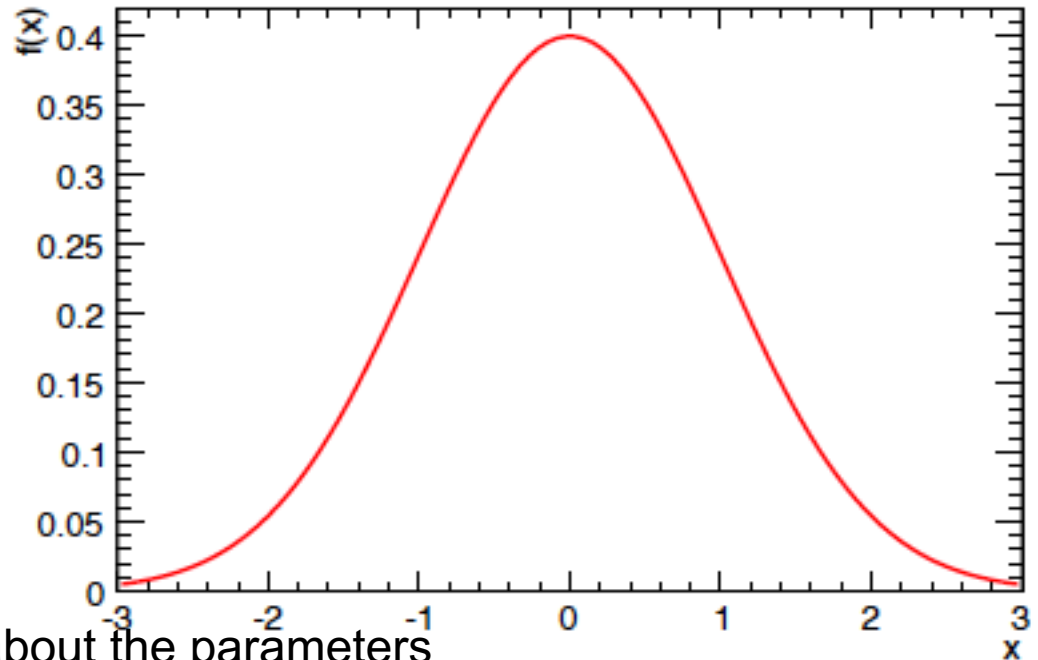
$$\int_{-\infty}^{\infty} f(x)dx = 1$$

f(x) is not a probability
f(x)dx is a probability

$$G(x|\mu, \sigma)$$

Is a pdf parametrized by (μ, σ)

We would like to make inference about the parameters



Likelihood is NOT a PDF

A Poisson distribution describes a discrete event count n for a real valued Mean μ .

$$Pois(n|\mu) = \mu^n \frac{e^{-\mu}}{n!}$$

Say, we observe n_o events

What is the likelihood of μ ?

The likelihood of μ is given by

$$L(\mu) = Pois(n_o | \mu)$$

It is a continuous function of μ but it is NOT a PDF

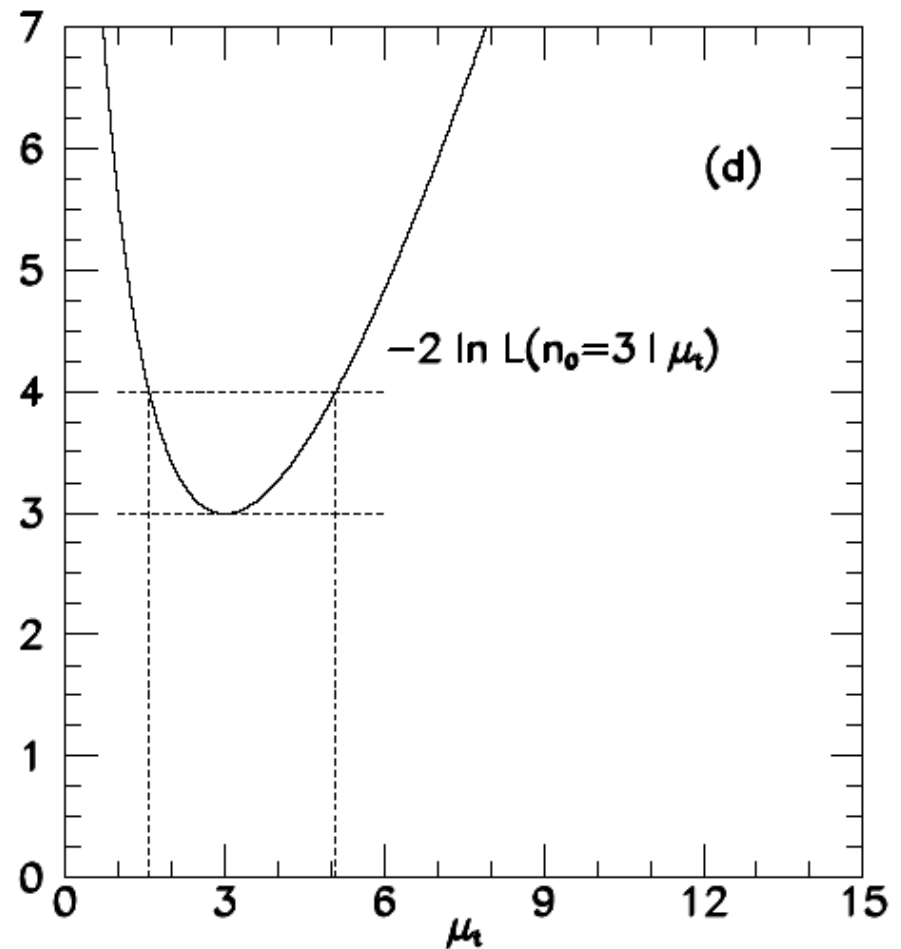


Figure from R. Cousins,
Am. J. Phys. 63 398 (1995)



Use of Likelihood Ratio

$$x \rightarrow y(x)$$

$$p(y(x)|\theta) = p(x|\theta) / |dy/dx| \quad (\text{prove})$$

probabilities are invariant under change of variable

$$P(y(x_1) < y < y(x_2)) = P(x_1 < x < x_2)$$

$$p(y(x)|\theta) \neq p(x|\theta)$$

But Likelihood ratio is invariant under change of variable x
(cancellation of Jacobian)

Likelihood is not a PDF in θ

$$L(\theta) = L(u(\theta))$$



Uniform PDF

$$x \in (a, b)$$

$$y(x) = \int_a^x p(x') dx'$$

then

$y \in (0, 1)$ and $p(y)$ is uniform for all y

$$p\text{-value} \equiv \int_{q_{obs}}^{\infty} f(q | H_{null}) dq$$

In statistics, when a p-value is used to test a simple null hypothesis, and the distribution of the test statistic, q , is continuous, then the p-value is uniformly distributed between 0 and 1 if the null hypothesis is true.



Bayesian vs Frequentist

Is there a Higgs Boson?

Frequentist

$Prob(Data | Theory)$

Run simulations of the theory and calculate the probability on the limit of long term frequency

Bayesia will try to answer

$Prob(Theory | Data)$

i.e. $Prob(Higgs | Data)$



Bayesian vs Frequentist

- Is there a Higgs Boson? What do you mean?
Given the data, is there a Higgs Boson?
- Can you really answer that without any a priori knowledge of the Higgs Boson?
Change your question: What is your degree of belief in the Higgs Boson given the data... Need a prior degree of belief regarding the Higgs Boson itself...

$$P(\text{Higgs} | \text{Data}) = \frac{P(\text{Data} | \text{Higgs})P(\text{Higgs})}{P(\text{Data})} = \frac{L(\text{Data} | \text{Higgs})\pi(\text{Higgs})}{\int L(\text{Data} | \text{Higgs})\pi(\text{Higgs})d(\text{Higgs})}$$

- If not, make sure that when you quote your answer you also quote your prior assumption!
- The most refined question is:
 - Assuming there is a Higgs Boson with some mass m_H , how well the data agrees with that?
 $P(\text{Data} | \text{Higgs})$
 - But even then the answer relies on the way you measured the data (i.e. measurement uncertainties), and that might include some pre-assumptions, priors!



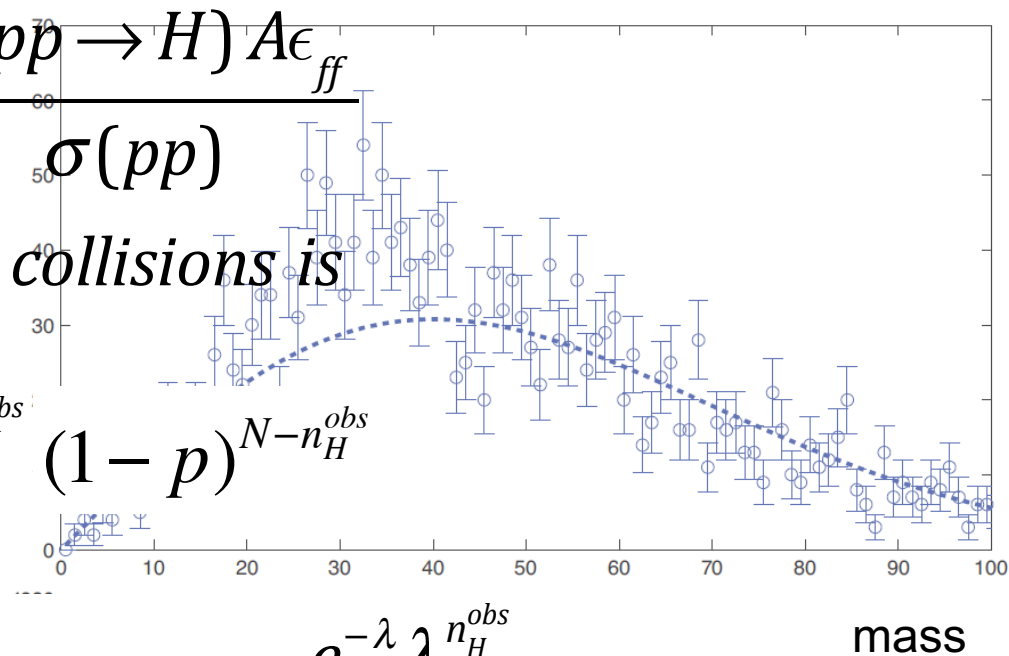
Histograms

N collisions

$$p(\text{Higgs event}) = \frac{\sigma(pp \rightarrow H) A \epsilon_{ff}}{\sigma(pp)}$$

Prob to see n_H^{obs} in N collisions is

$$P(n_H^{obs}) = \binom{N}{n_H^{obs}} p^{n_H^{obs}} (1-p)^{N-n_H^{obs}}$$



$$\lim_{N \rightarrow \infty} P(n_H^{obs}) = \text{Poiss}(n_H^{obs}, \lambda) = \frac{e^{-\lambda} \lambda^{n_H^{obs}}}{n_H^{obs}!}$$

$$\lambda = Np = n_H^{exp}$$



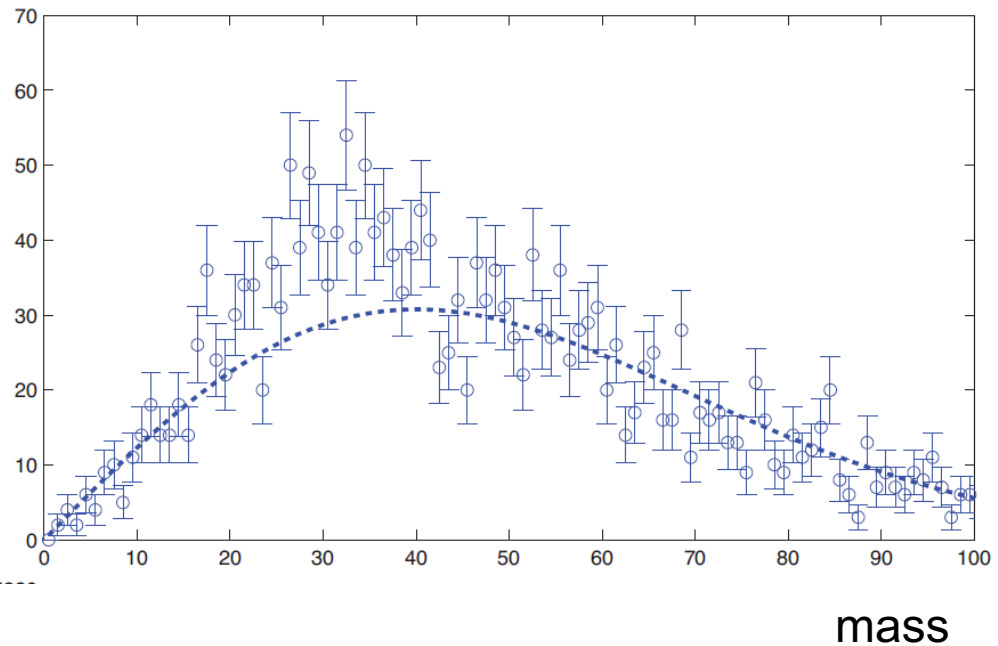
The Statistical Challenge of HEP

The DATA: Billions of Proton-Proton collisions

which could be visualized with histograms

The searched particle mass is unknown (for the sake of this lecture)

In this TOY example, we ask if the expected background (e.g. the Standard Model WITHOUT the Higgs Boson) contains a Higgs Boson, which would manifest itself as a peak in the distribution



WE NEED TO KNOW WHAT WE SEARCH FOR.....

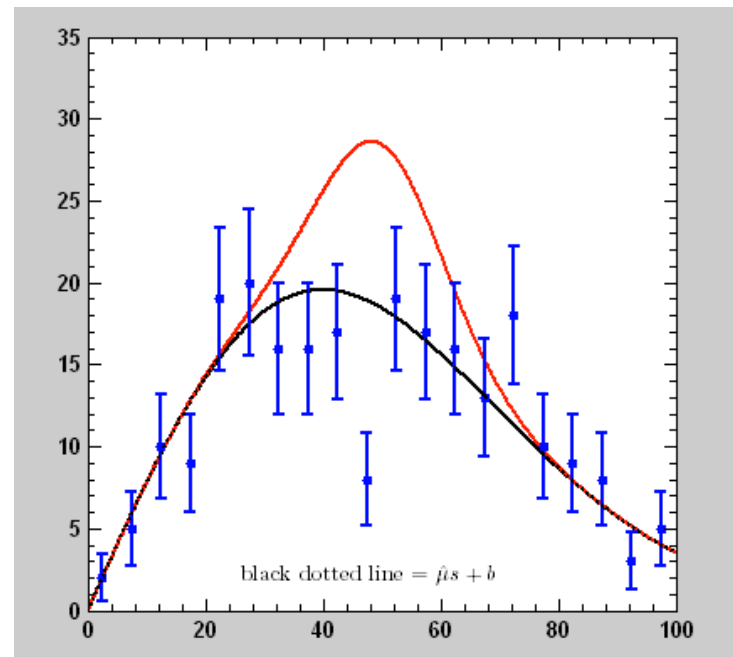
We need to have a model

We need to have two hypotheses if we want a powerful test



What is the statistical challenge?

- The black line represents the Standard Model (SM) expectation (Background only),
- How compatible is **the data (blue)** with the **SM expectation (black)**?
- Is there a signal hidden in this data?
- What is its statistical significance?
- What is the most powerful test statistic that can tell the SM (black) from an **hypothesized signal (red)**?



The Model

- The Higgs hypothesis is that of signal $s(m_H)$

$$s(m_H) = L \cdot \sigma_{SM}(m_H) \cdot A \cdot eff$$

For simplicity unless otherwise noted $s(m_H) = L \cdot \sigma_{SM}(m_H)$

- In a counting experiment

$$n = \mu \cdot s(m_H) + b$$

$$\mu = \frac{L \cdot \sigma(m_H)}{L \cdot \sigma_{SM}(m_H)} = \frac{\sigma(m_H)}{\sigma_{SM}(m_H)}$$

- μ is the strength of the signal (with respect to the expected Standard Model one)
- The hypotheses are therefore denoted by H_μ
- H_1 is the SM with a Higgs, H_0 is the background only model



A Frequentist Tale of Two Hypotheses

NULL

ALTERNATE

- Test the Null hypothesis and try to reject it
- Fail to reject it OR reject it in favor of the Alternate hypothesis



The Null Hypothesis

- The Standard Model without the Higgs is an hypothesis, (BG only hypothesis) many times referred to as **the null hypothesis** and is denoted by H_0
(remember that it is the null hypothesis **ONLY** if we aim at a discovery)
- In the absence of an alternate hypothesis, one would like to test the compatibility of the data with H_0
- This is actually a **goodness of fit test**,
NOT an hypothesis test



A Tale of Two Hypotheses

NULL

H_0 - SM w/o Higgs

ALTERNATE

H_1 - SM with Higgs

- Test the Null hypothesis and try to reject it
- Fail to reject it OR reject it in favor of the Alternate hypothesis



The Alternate Hypothesis?

- Let's zoom on

H_1 - SM with Higgs

- Higgs with a specific mass m_H
OR
- Higgs anywhere in a specific mass-range
→ • The look elsewhere effect

A Tale of Two Hypotheses

NULL

H_0 - SM w/o Higgs

ALTERNATE

H_1 - SM with Higgs

- Reject H_0 in favor of H_1 – A DISCOVERY

Swapping Hypotheses \rightarrow exclusion

NULL

H_0 - SM w/o Higgs

ALTERNATE

H_1 - SM with Higgs

- Reject H_1 in favor of H_0

Excluding H_1 (m_H) \rightarrow Excluding the Higgs with a mass m_H

Testing an Hypothesis (wikipedia...)

- The first step in any hypothesis testing is to state the relevant **null, H_0** and **alternative hypotheses**, say, H_1
- The next step is to define a test statistic, q , under the null hypothesis
- Compute from the observations the observed value q_{obs} of the test statistic q .
- Decide (based on q_{obs}) to **either fail to reject the null hypothesis or reject it in favor of an alternative hypothesis**
- **next: How to construct a test statistic, how to decide?**



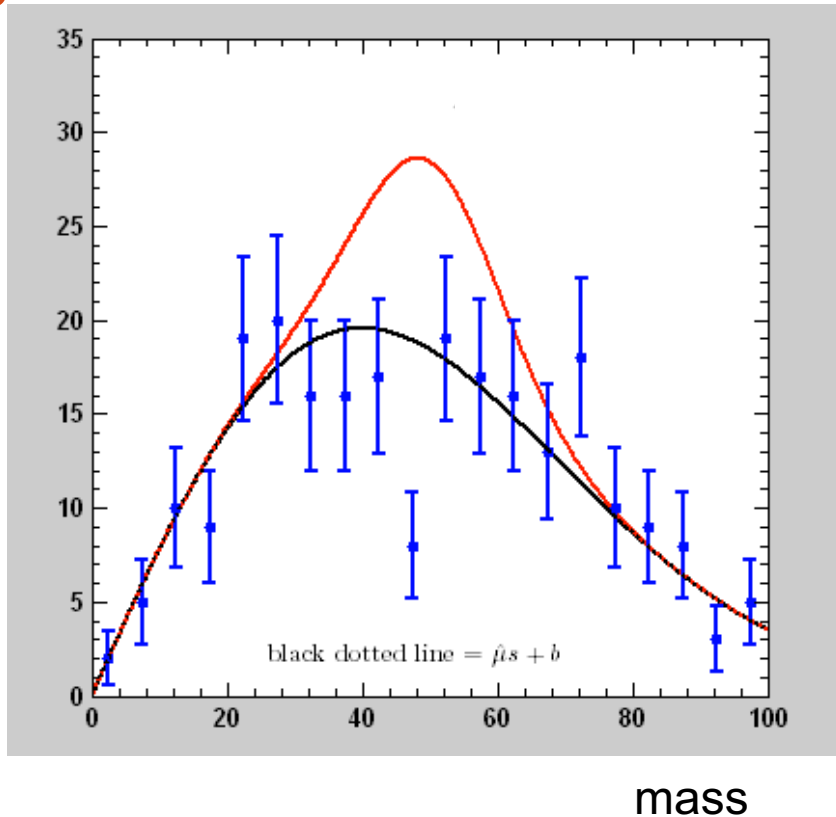
Test statistic and p-value



The Physics Model

- SM without Higgs Background

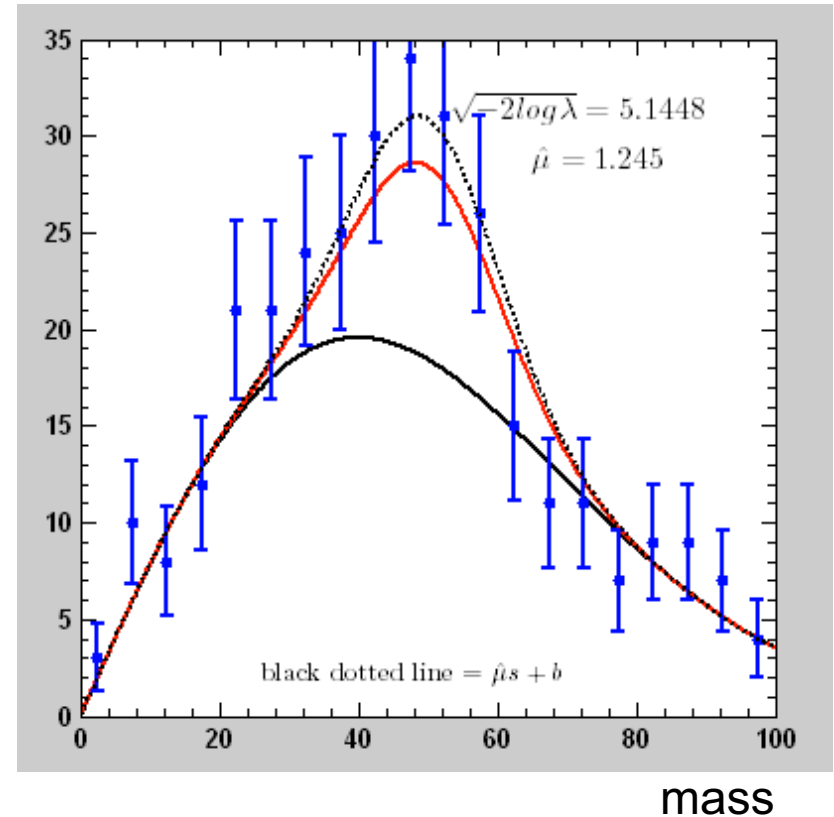
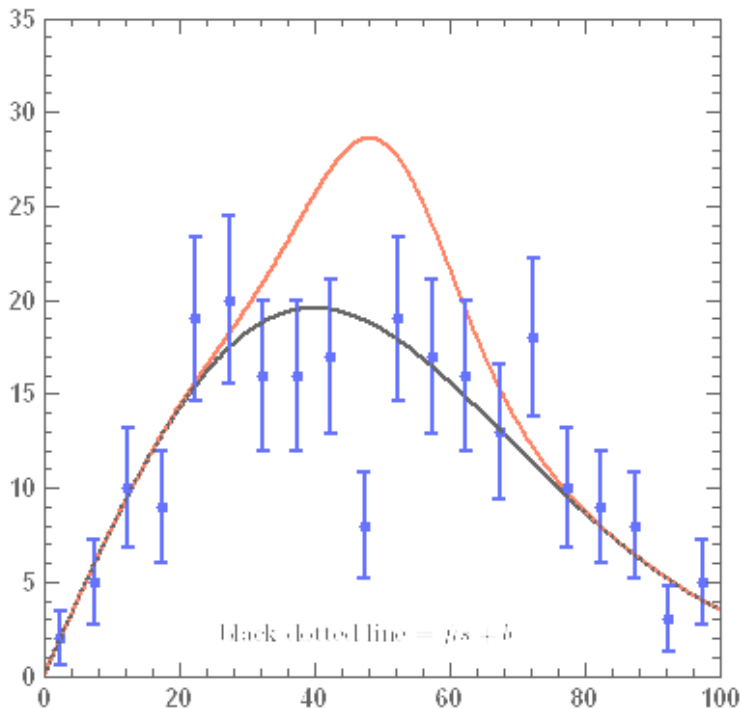
No signal $\langle n \rangle = b$



The Physics Model

- SM without Higgs Background Only $\langle n \rangle = b$

- SM with a Higgs Boson with a mass m_H $\langle n \rangle = s(m_H) + b$



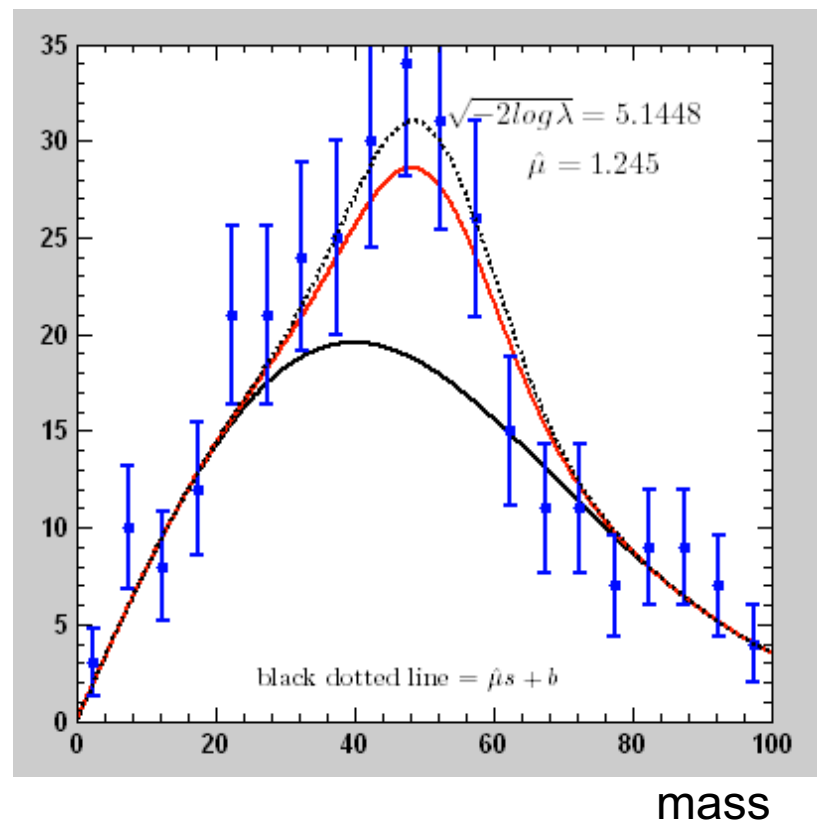
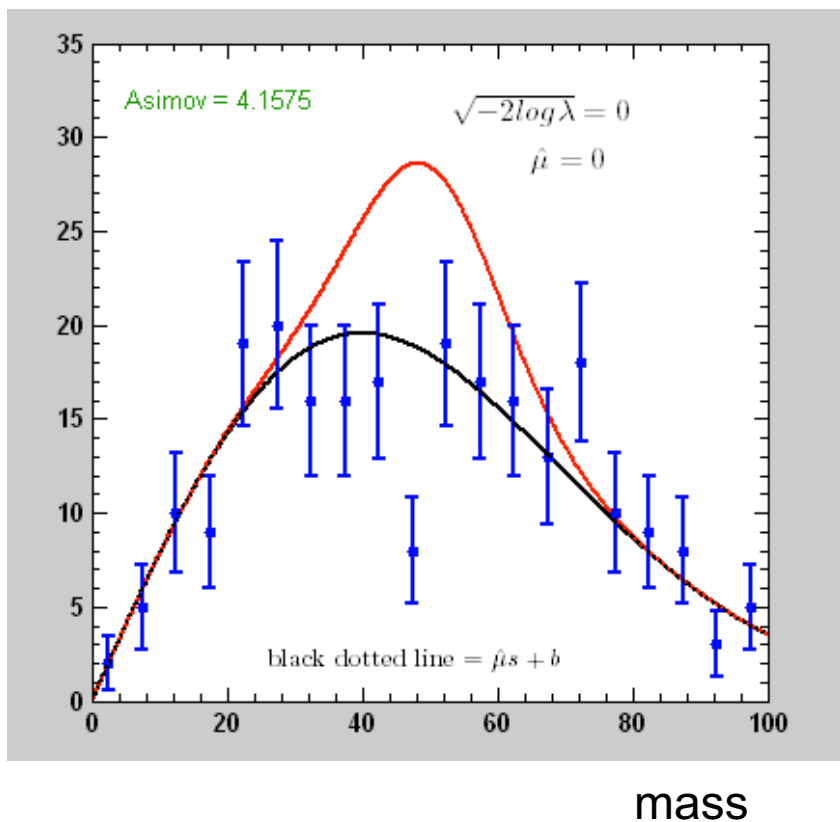
The Physics Model

$$n = \mu s + b$$

$$\langle \hat{\mu} \rangle = 0 \text{ under } H_0$$

$$MLE \quad \hat{\mu}$$

$$\langle \hat{\mu} \rangle = 1 \text{ under } H_1$$



The Neyman-Pearson Test Statistic

- NP test statistic

$$Q = -2 \ln \frac{L(H_0)}{L(H_1)}$$

$$n = \mu s + b$$

$$H_0; \hat{\mu} = 0, \langle n_{obs} \rangle = b$$

$$H_1; \hat{\mu} = 1, \langle n_{obs} \rangle = s + b$$

$$L(H_0) = \text{prob}(x | H_0) = \text{prob}(x | b)$$

$$L(H_1) = \text{prob}(x | H_1) = \text{prob}(x | s + b)$$

$$L = f(x), L(H_\mu) = \text{prob}(x | H_\mu) = f(\hat{\mu})$$

Likelihood

- Likelihood is a function of the data

$$L(H) = L(H | x) = f(x)$$

$$L(H | x) = P(x | H)$$

$$q_{NP} = q_{NP}(x) = -2 \ln \frac{L(H_1 | x)}{L(H_0 | x)}$$

Bayes Theorem

- Likelihood is not the probability of the hypothesis given the data

$$P(H | x) = \frac{P(x | H) \cdot P(H)}{\sum_H P(x | H) P(H)}$$

$$P(H | x) \approx P(x | H) \cdot \mathbf{P(H)}$$

Prior



Frequentist vs Bayesian

- The Bayesian infers from the data using **priors**

posterior $P(H | x) \approx P(x | H) \cdot P(H)$

- Priors is a science on its own.

Are they objective? Are they subjective?

- The Frequentist calculates the probability of an hypothesis to be inferred from the data based on a large set of hypothetical experiments

Ideally, the frequentist does not need priors, or any degree of belief while the Bayesian posterior based inference **is** a “Degree of Belief”.

- However, NPs inject a Bayesian flavour to any Frequentist analysis



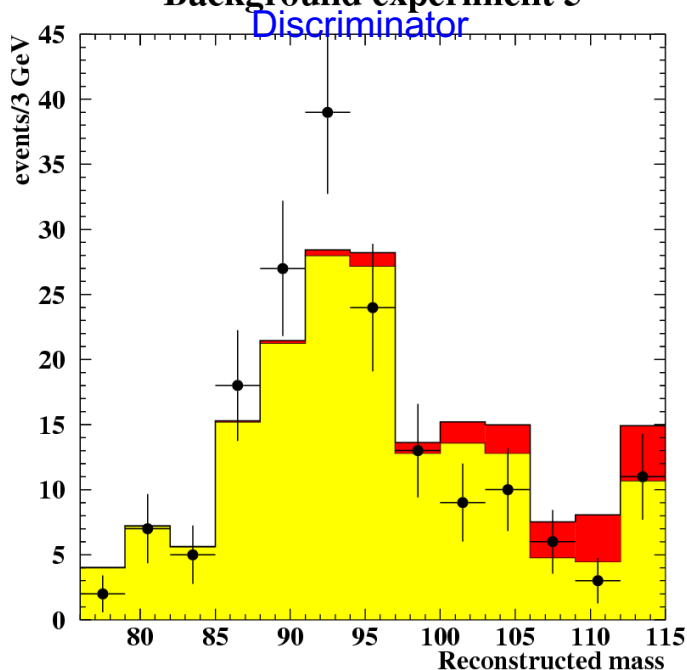
Example: Simulating BG Only Experiments

$$q_{NP}(m) = \frac{L(H_1)}{L(H_0)} = \frac{L(s(m)+b)}{L(b)}$$

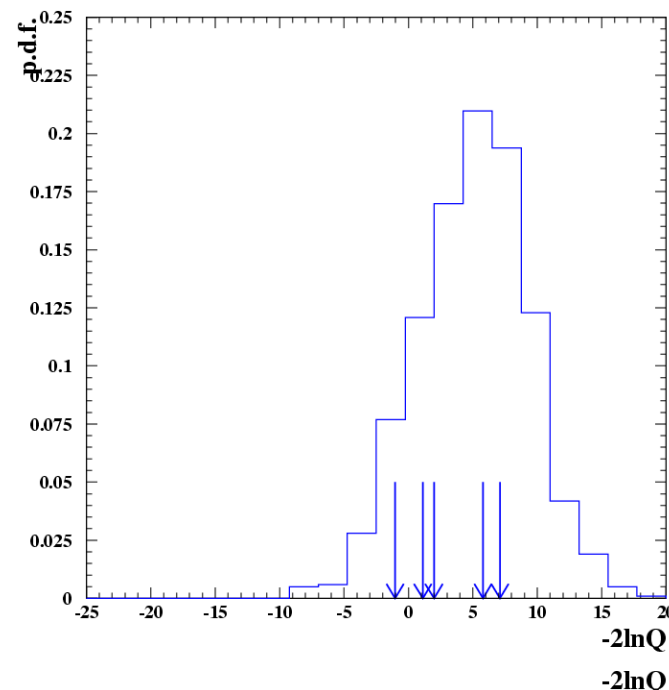
- The likelihood ratio, $-2\ln q(m_H)$ tells us how much the outcome of an experiment is signal-like
- **NOTE**, here the $s+b$ pdf is plotted to the left (it's the null hypothesis)!
- Test the $s+b$ hypothesis

$$f(q^{NP} | b)$$

Background experiment 5



Test mass = 115 GeV



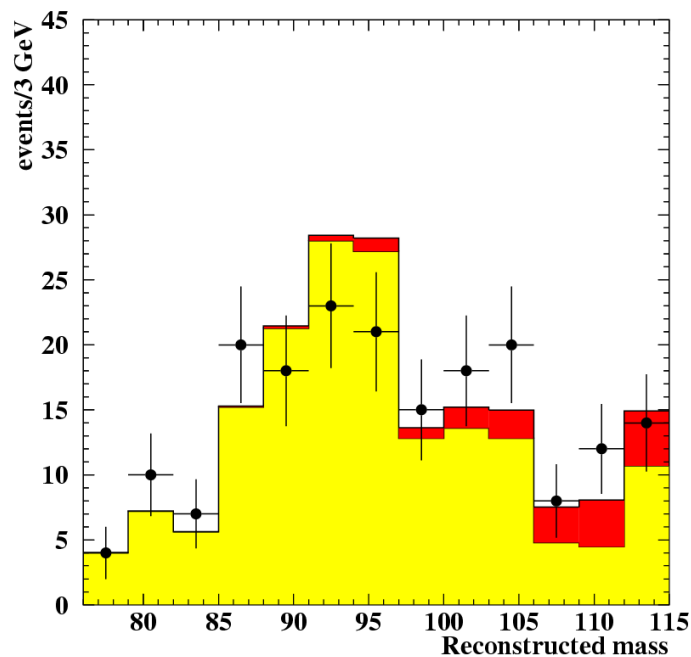
Example:

Simulating $S(m_H)+b$ Experiments

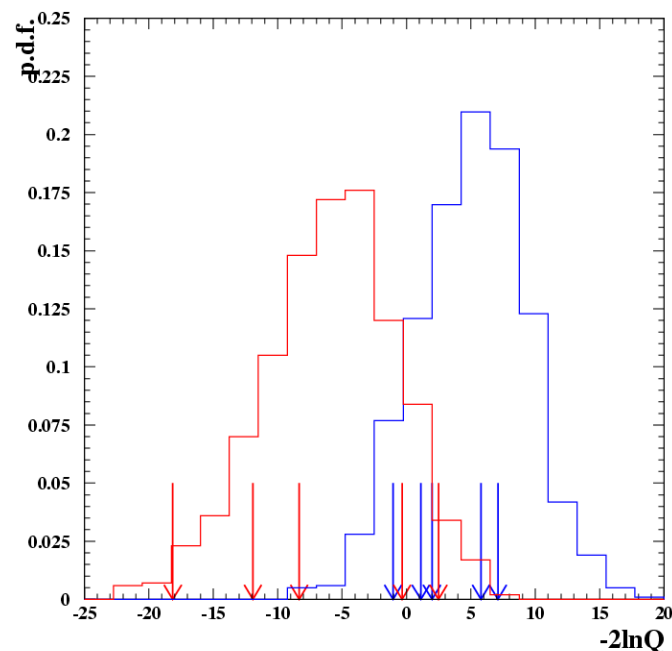
$$f(q^{NP} | s(m_H) + b)$$

$$f(q^{NP} | b)$$

Signal+bkg. Experiment 5



Test mass = 115 GeV



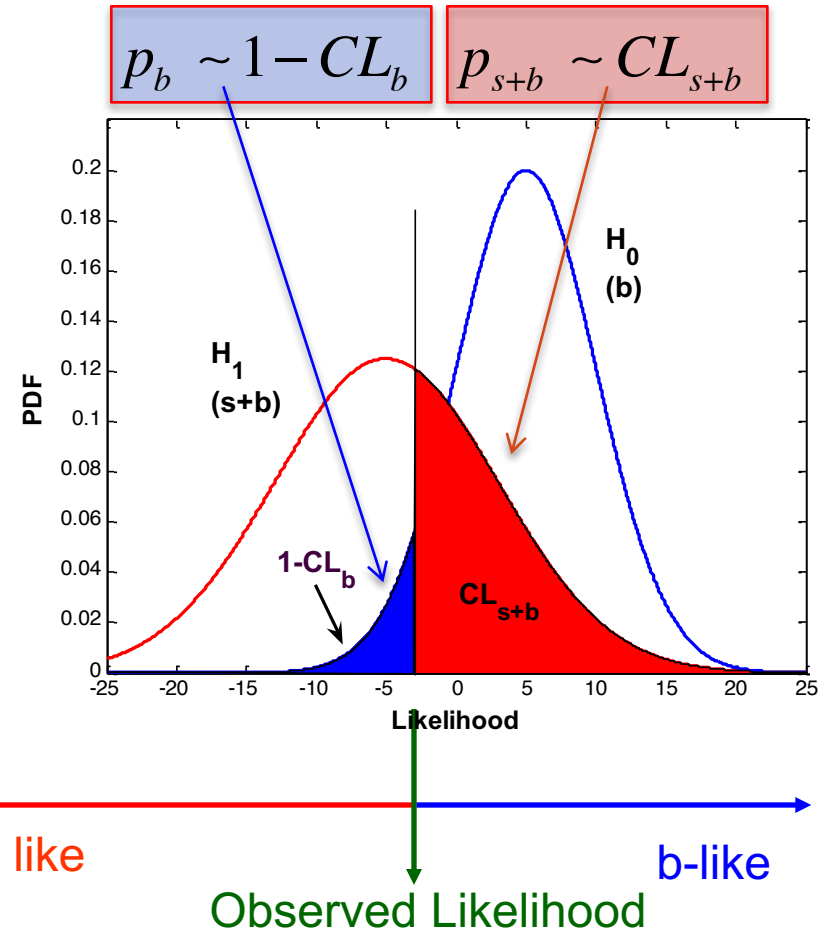
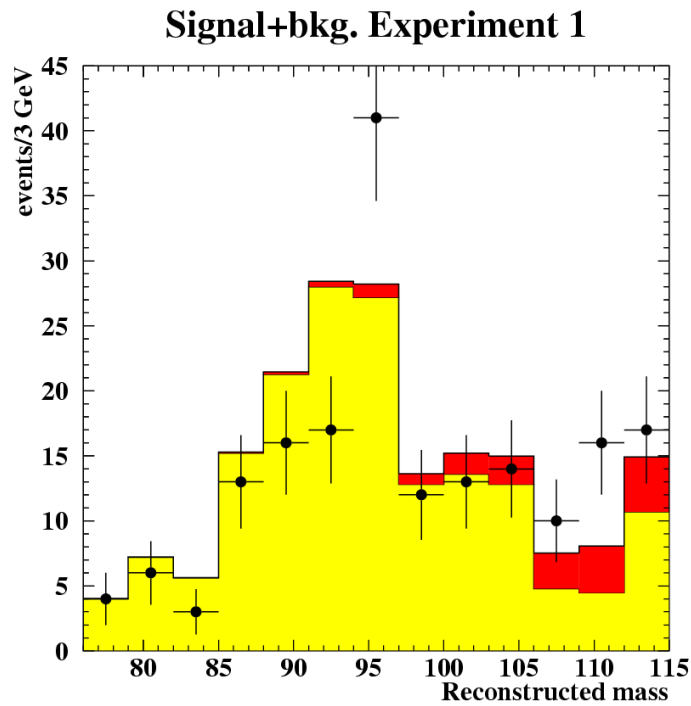
$s+b$ like

b -like



Example:

Simulating $S(m_H)+b$ Experiments

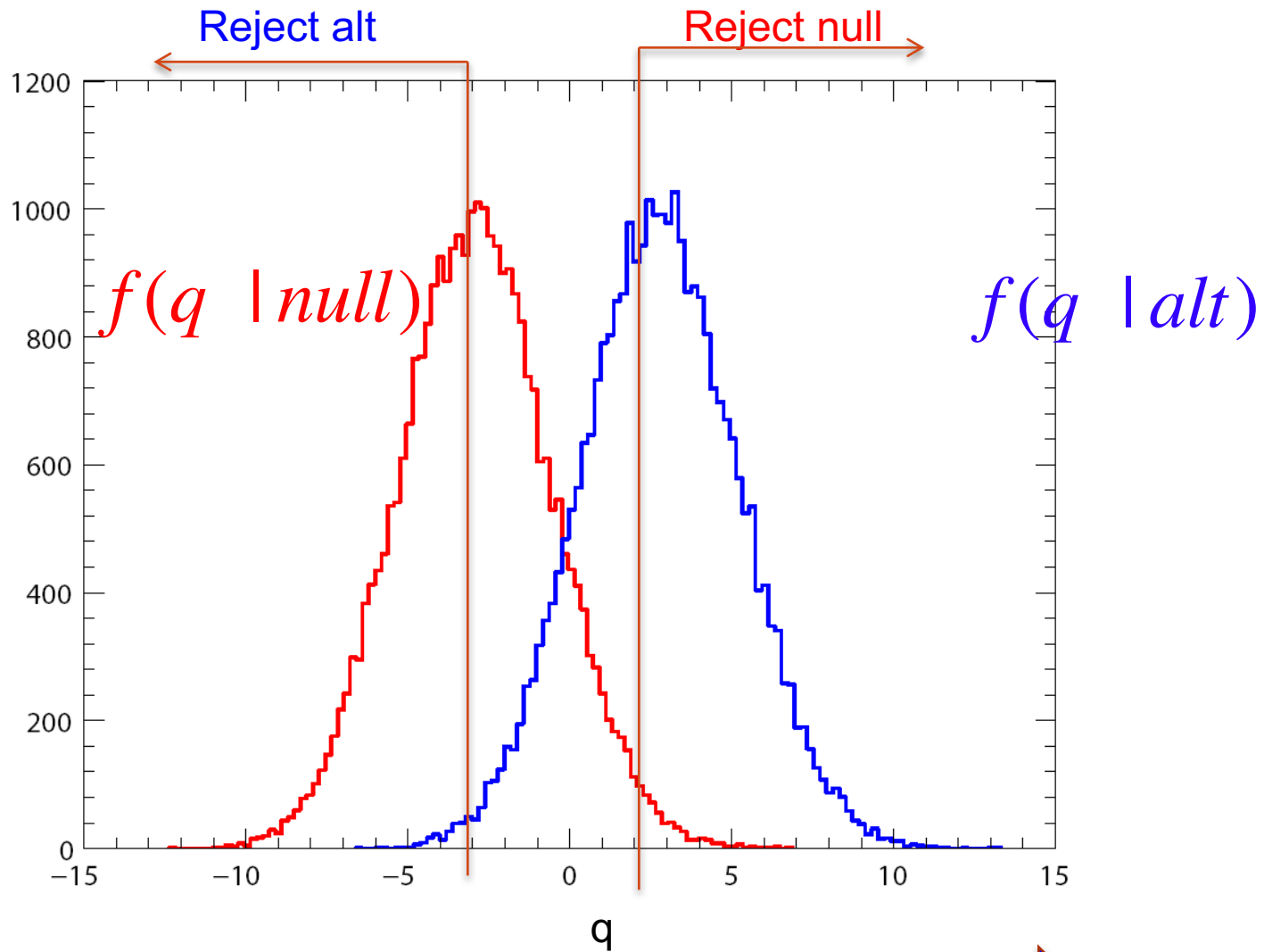


p-Value

- The observed *p-value* is a measure of the incompatibility of the data with the tested hypothesis.
- It is the probability, under assumption of the null hypothesis H_{null} , of finding data of equal or greater incompatibility with the predictions of H_{null}
- Wikipedia: *An important property of a test statistic is that its sampling distribution under the null hypothesis must be calculable, either exactly or approximately, which allows p-values to be calculated.*



PDF of a test statistic



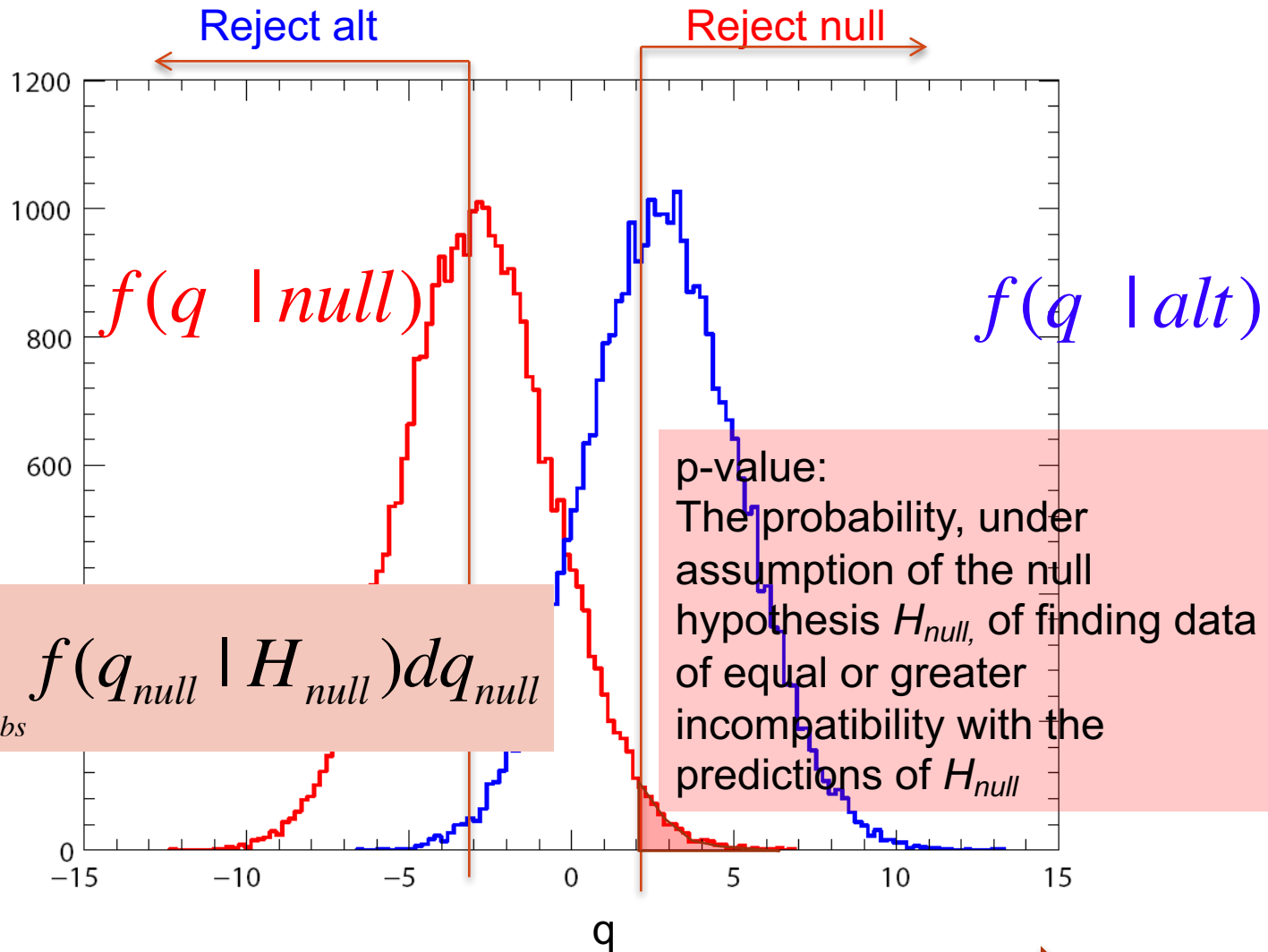
Null like



alt like



PDF of a test statistic



Basic Definitions: type I-II errors

- By defining α you determine your tolerance towards mistakes... (accepted mistakes frequency)
- **type-I error:** the probability to reject the tested (null) hypothesis (H_0) when it is true

$$\alpha = \Pr ob(reject H_0 | H_0)$$

$$\alpha = \textit{type I error}$$

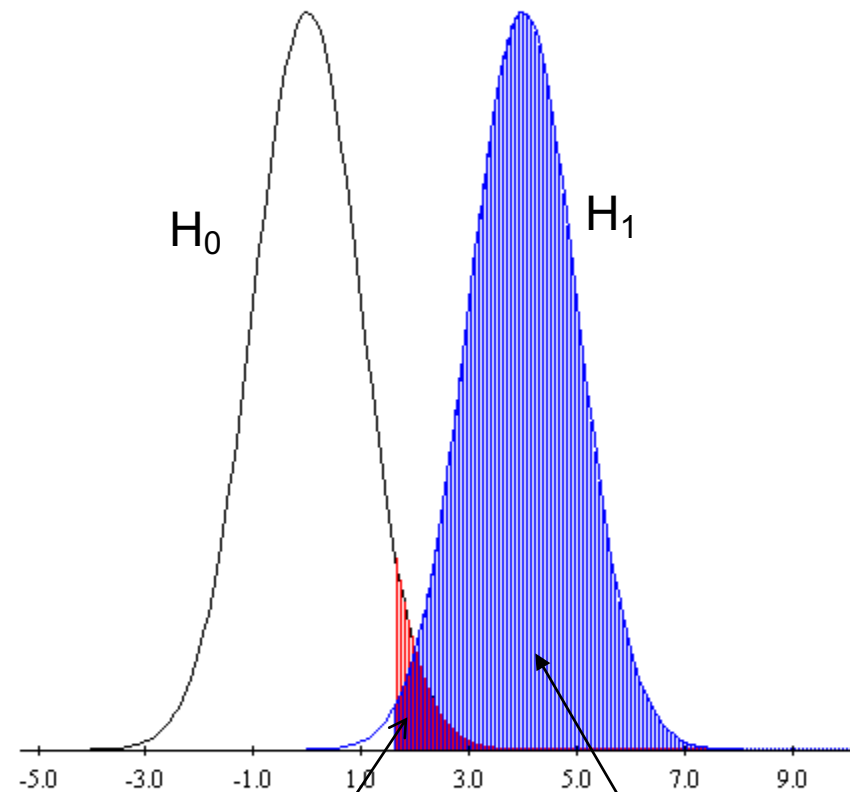
- **Type II:** The probability to accept the null hypothesis when it is wrong

$$\beta = \Pr ob(accept H_0 | \bar{H}_0)$$

$$= \Pr ob(reject H_1 | H_1)$$

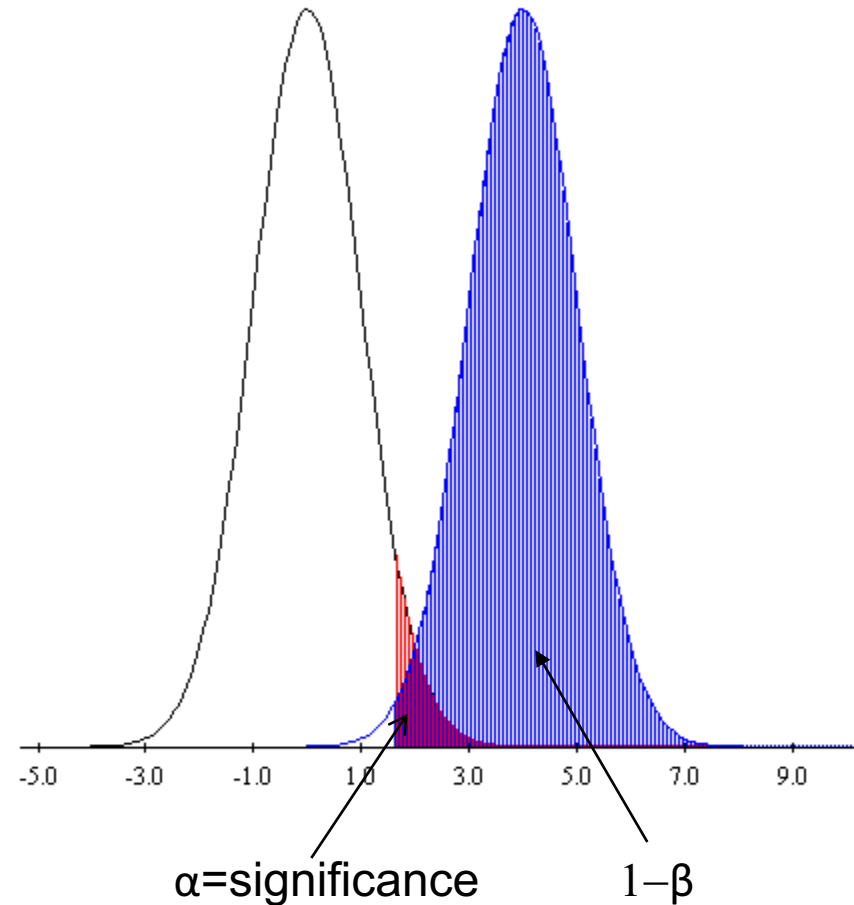
$$\beta = \textit{type II error}$$

- The pdf of q



Basic Definitions: POWER

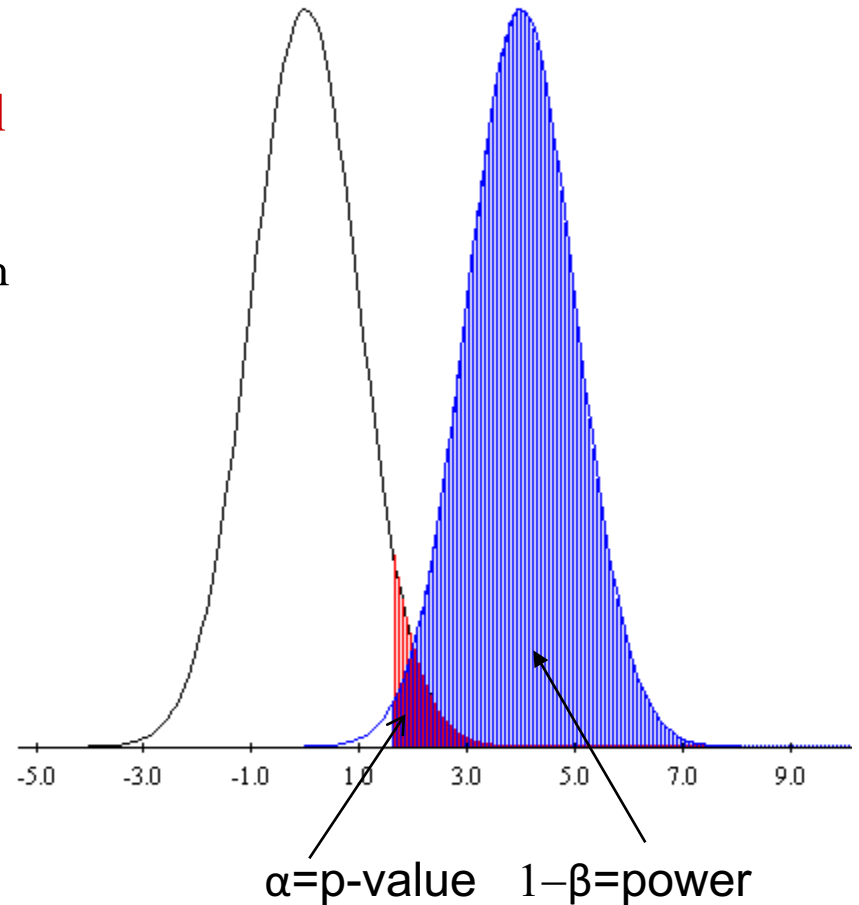
- $\alpha = \text{Pr ob}(\text{reject } H_0 \mid H_0)$
- The POWER of an hypothesis test is the probability to reject the null hypothesis when the alternate analysis is true!
- $\text{POWER} = \text{Prob}(\text{reject } H_0 \mid H_1)$
 $\beta = \text{Pr ob}(\text{reject } H_1 \mid H_1) \Rightarrow$
 $1 - \beta = \text{Pr ob}(\text{accept } H_1 \mid H_1) =$
 $1 - \beta = \text{Pr ob}(\text{reject } H_0 \mid H_1) \Rightarrow$
 $\text{POWER} = 1 - \beta$
- The power of a test increases as the rate of type II error decreases



Which Analysis is Better

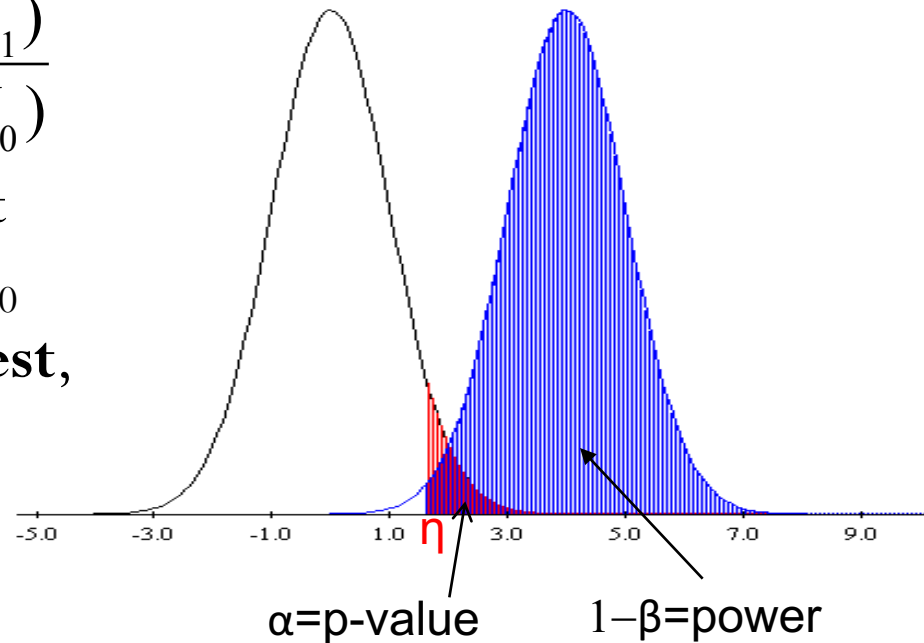
- To find out which of two methods is better plot the p-value vs the power for each analysis method
- Given the p-value, the one with the higher power is better
- p-value \sim significance

Different Test Statistics
Differ by
POWER
and
Asymptotic Properties



The Neyman-Pearson Lemma

- Define a **test statistic** $\lambda = \frac{L(H_1)}{L(H_0)}$
- When performing a hypothesis test between two simple hypotheses, H_0 and H_1 , **the Likelihood Ratio test**, which rejects H_0 in favor of H_1 , **is the most powerful test** of size α for a threshold η
- **Note:** Likelihoods are functions of the data, even though we often not specify it explicitly



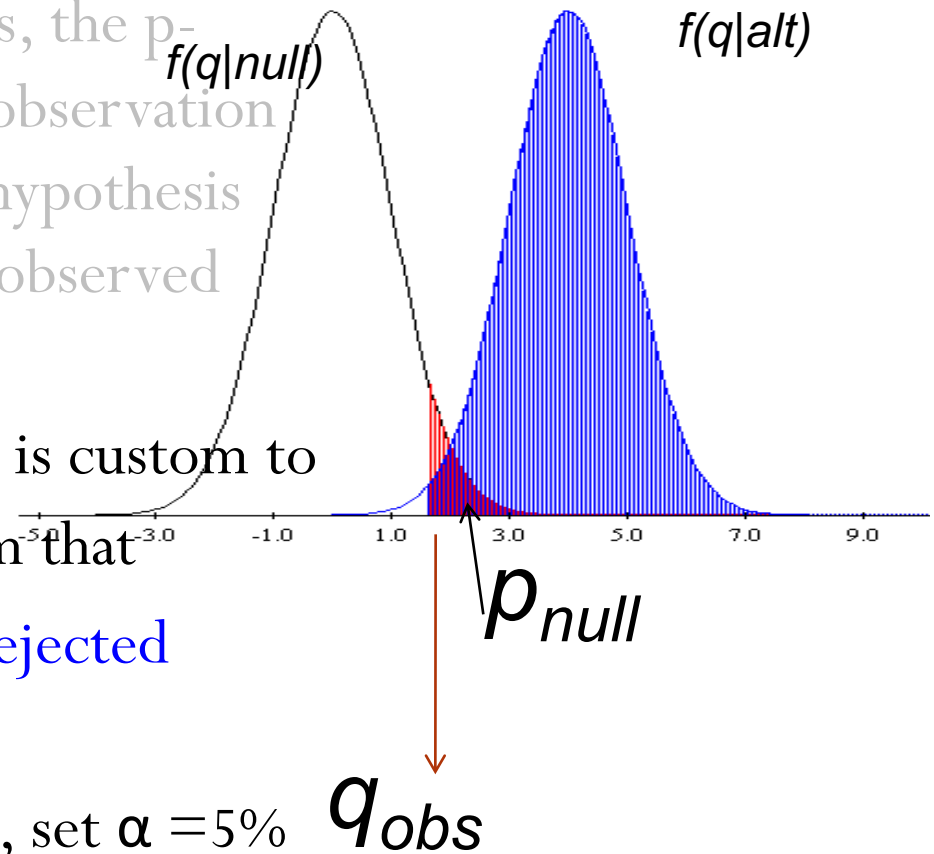
p-value – testing the null hypothesis

- When testing the null hypothesis, the p-value is the probability that the observation is less compatible with the null hypothesis (more alternative like) than the observed one
- When testing the b hypothesis, it is custom to set $\alpha = 2.9 \cdot 10^{-7}$ and it is custom that

if $p_b < 2.9 \cdot 10^{-7}$ the b hypothesis is rejected
→ discovery

- When testing the s+b hypothesis, set $\alpha = 5\%$

if $p_{s+b} < 5\%$ the signal hypothesis is rejected at the 95% Confidence Level (CL) → Exclusion

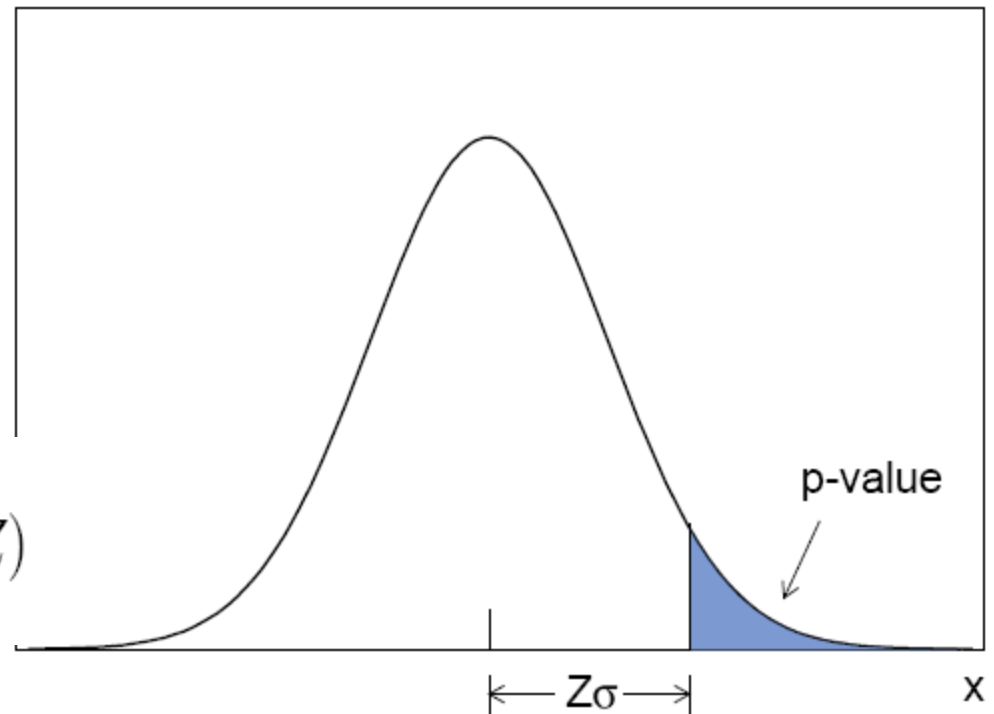


From p-values to Gaussian significance

It is a custom to express the p-value as the significance associated to it, had the pdf were Gaussians

$$p = \int_Z^{\infty} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx = 1 - \Phi(Z)$$

$$Z = \Phi^{-1}(1 - p)$$

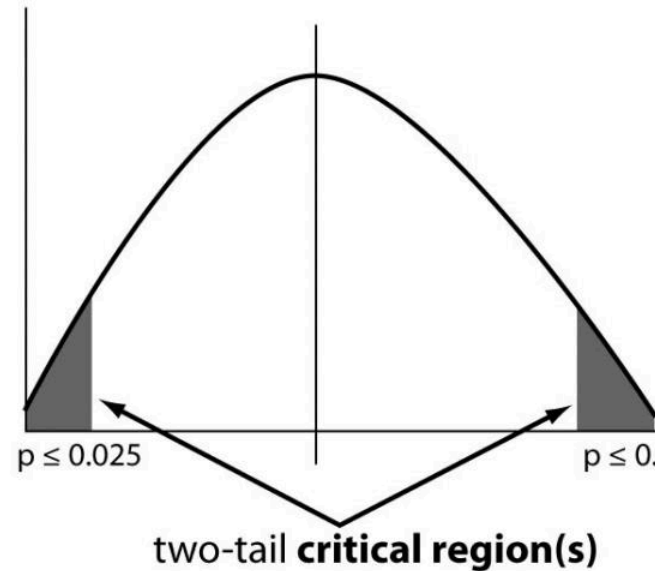


A significance of $Z = 5$ corresponds to $p = 2.87 \times 10^{-7}$

Beware of 1 vs 2-sided definitions!

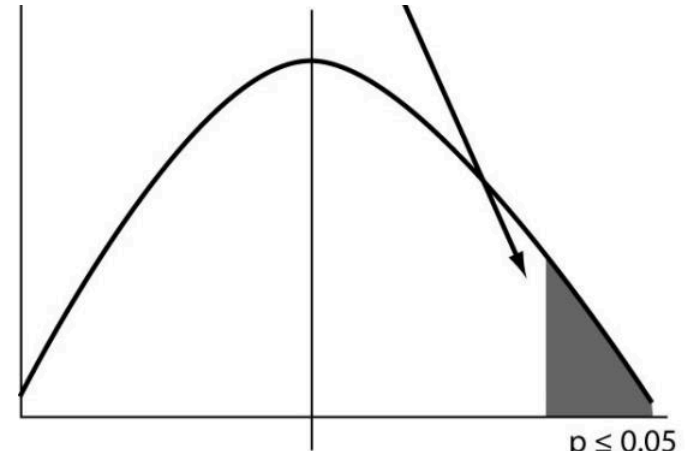
2-Sided p-value

- When performing a measurement, any deviation above or below the mean is drawing our attention and might serve an indication of some anomaly or new physics.
- Here we use a 2-sided p-value



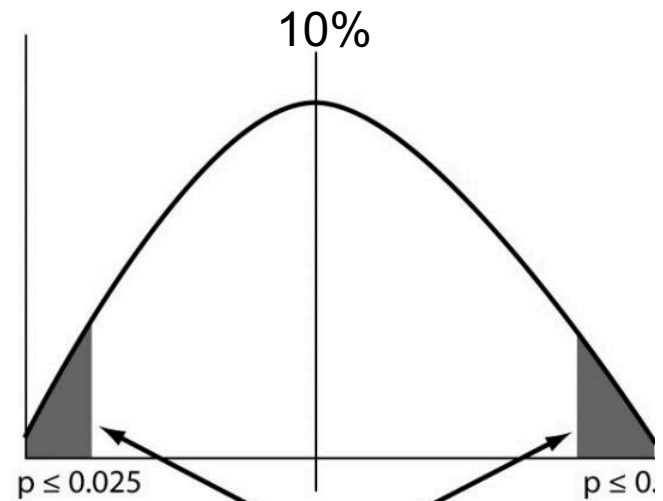
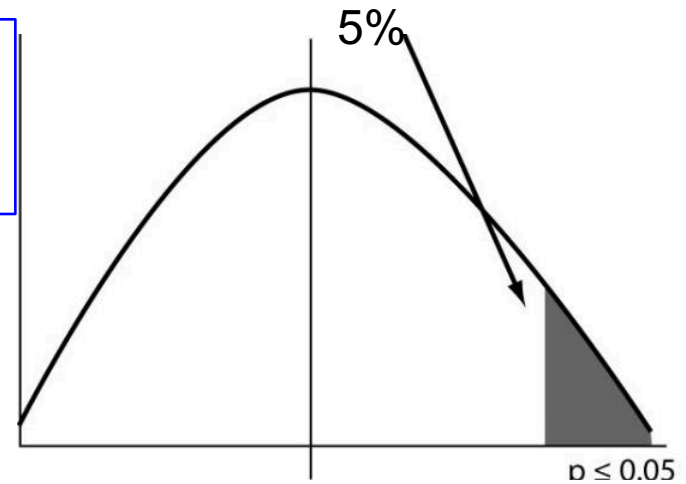
1-Sided p-value

- When trying to reject an hypothesis while performing searches, one usually considers only one-sided tail probabilities.
- Downward fluctuations of the background will not serve as an evidence against the background
- Upward fluctuations of the signal will not be considered as an evidence against the signal



1-sided 2-sided

- To determine a 1 sided 95% CL, we sometimes need to set the critical region to 10% 2 sided



two-tail **critical region(s)**



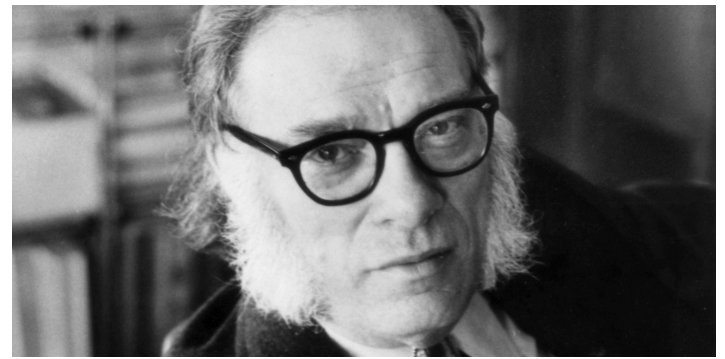
Estimating the Sensitivity of an Experiment

- Estimate the expected significance one could achieve (for discovering the Higgs Boson) with a given analysis, a given Luminosity and CM energy..
- Option 1:
 - Toss, say, 1000000 BG only events (null) and derive the BG-only pdf of q , $f(q_{\text{null}} | \text{BG})$.
Toss 1000000 S+BG (alt) events and find the significance for each one of them Find the median significance....
- This may take ages..., is there a shortcut?

←→ Option2: Asimov Data Set

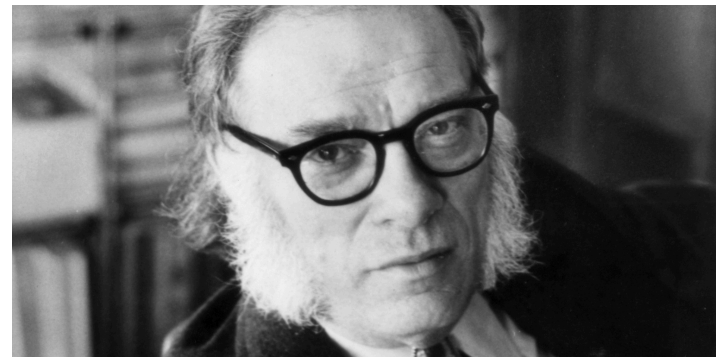


The Asimov Data Set



In the future, the United States has converted to an "electronic democracy" where the computer Multivac selects a single person to answer a number of questions. Multivac will then use the answers and other data to determine what the results of an election would be, avoiding the need for an actual election to be held.

The Asimov Data Set



- *The use of a single representative individual to stand in for the entire population can help in evaluating the sensitivity of a statistical method.*
- *The "Asimov data set":
an ensemble of simulated experiments can be replaced by a single representative one.*

Estimating the Sensitivity of an Experiment

- Option 2: the Asimov Data Set
 - one can replace each ensemble of the alternate-hypothesis experiments with one data set that represents the typical experiment. This “Asimov” data set delivers the desired median sensitivity. Hence, one is exempted from the need to perform an ensemble of experiments for each set of parameters.
- The Asimov data set is constructed such that when one uses it to evaluate the estimators for all parameters, one obtains the true parameter values.
- the Asimov data set can trivially be constructed from the true parameters values. For example, a set corresponding to the H_1 hypothesis is $n_A = s + b$. and the one correspond to the H_0 hypothesis is $n_A = b$.
- As strange as it reads, the Asimov data set is not necessarily an integer.



The Profile Likelihood

The choice of the LHC for hypothesis inference in Higgs search



The Profile Likelihood (“PL”)

The best signal $\hat{\mu} = 0.3 \rightarrow 1.27\sigma$

For discovery we test the H_0 null hypothesis

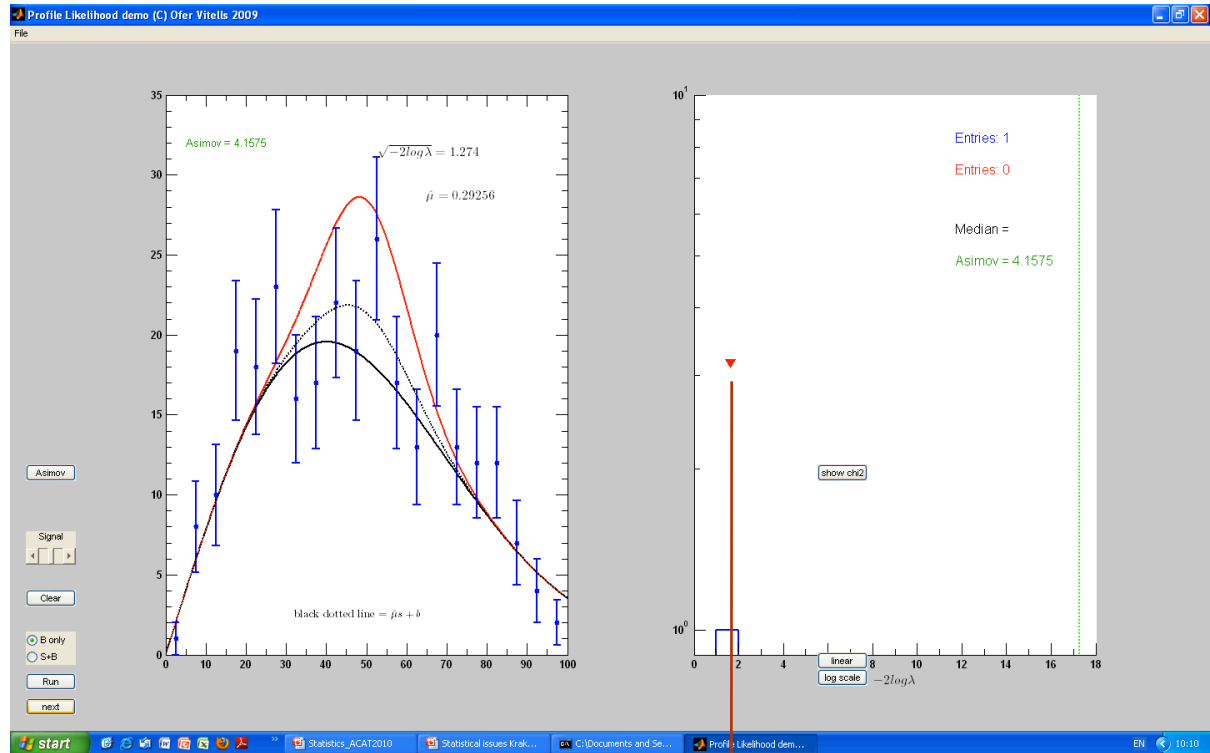
For
$$q_0 = -2 \ln \frac{L(b)}{L(\hat{\mu}s + b)}$$

$\hat{\mu} \sim 0$, q_0 small

$\hat{\mu} \sim 1$, q_0 large

In general: testing the H_μ hypothesis i.e., a SM with a signal of strength μ ,

$$q_\mu = -2 \ln \frac{L(\mu)}{L(\hat{\mu})}$$



$$Z_{obs} = \sqrt{q_{0,obs}}$$

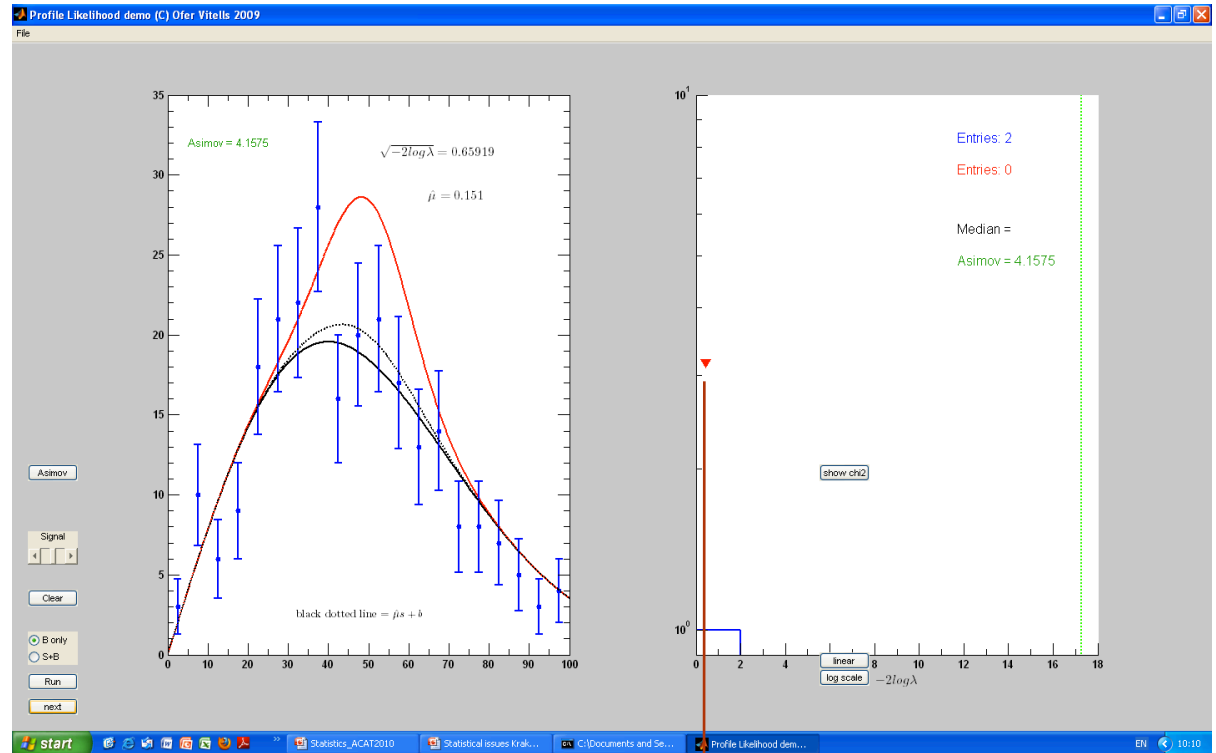
$$q_0 = 1.6 \rightarrow Z = \sqrt{1.6} = 1.27$$



PL: test q_0 under BG only ; $f(q_0 | H_0)$

$\hat{\mu} = 0.15 \rightarrow 0.6\sigma$

$$q_0 = -2 \ln \frac{L(b)}{L(\hat{\mu}s + b)}$$



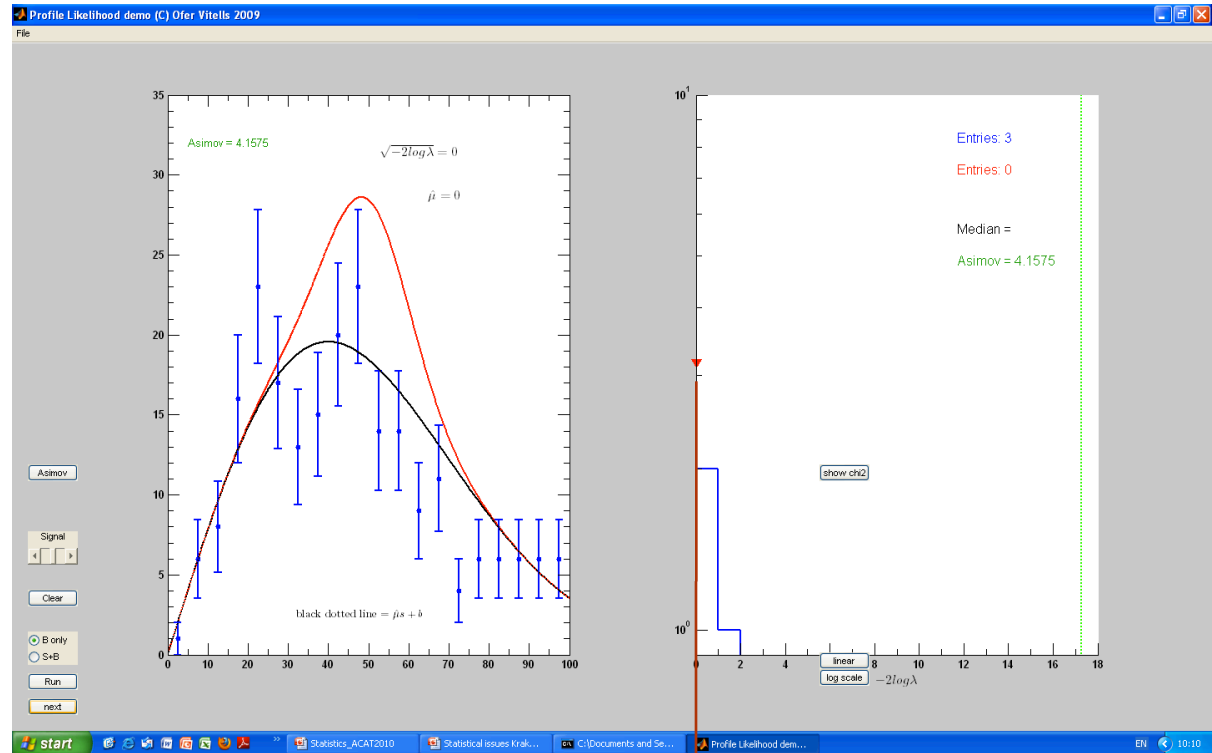
$$Z_{obs} = \sqrt{q_{0,obs}} \quad q_0 = 0.43 \rightarrow Z = 0.66\sigma$$



PL: test q_0 under BG only ; $f(q_0 | H_0)$

$\hat{\mu} = 0$

$$q_0 = -2 \ln \frac{L(b)}{L(\hat{\mu}s + b)}$$

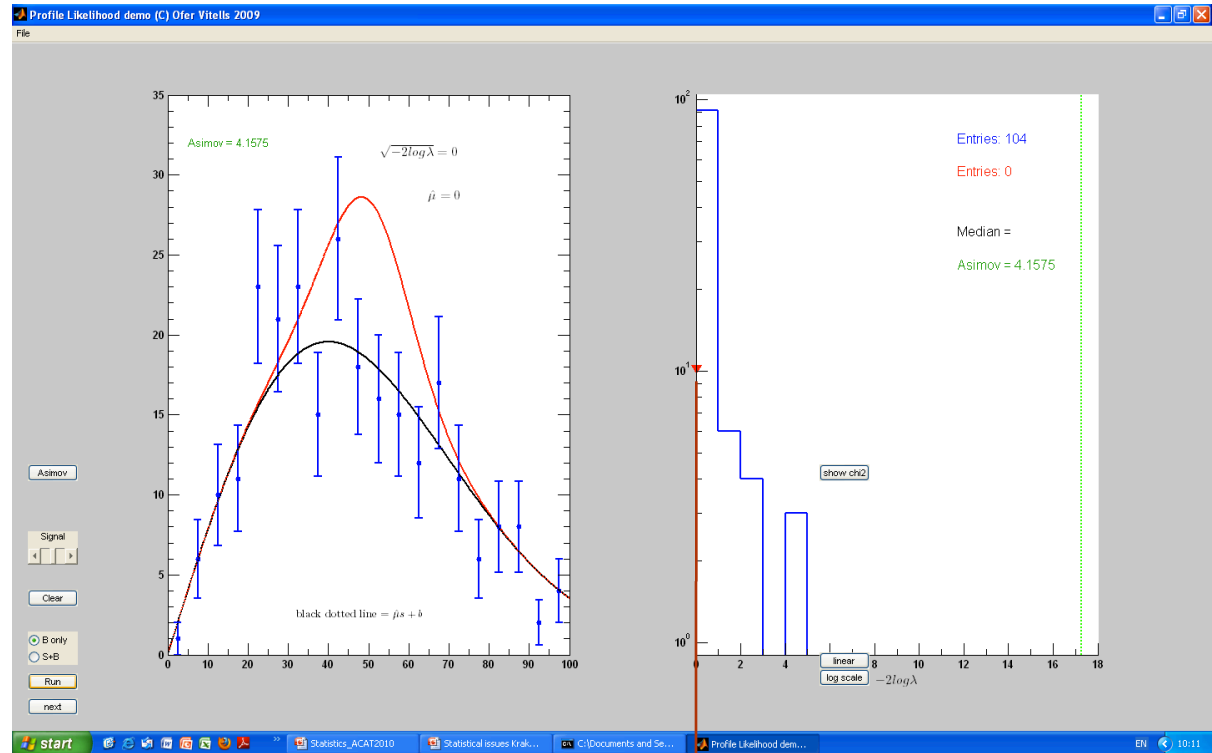


$$Z_{obs} = \sqrt{q_{0,obs}} \quad q_0 = 0$$



PL: test q_0 under BG only ; $f(q_0 | H_0)$ $\hat{\mu} = 0$

$$q_0 = -2 \ln \frac{L(b)}{L(\hat{\mu}s + b)}$$



$$Z_{obs} = \sqrt{q_{0,obs}}$$

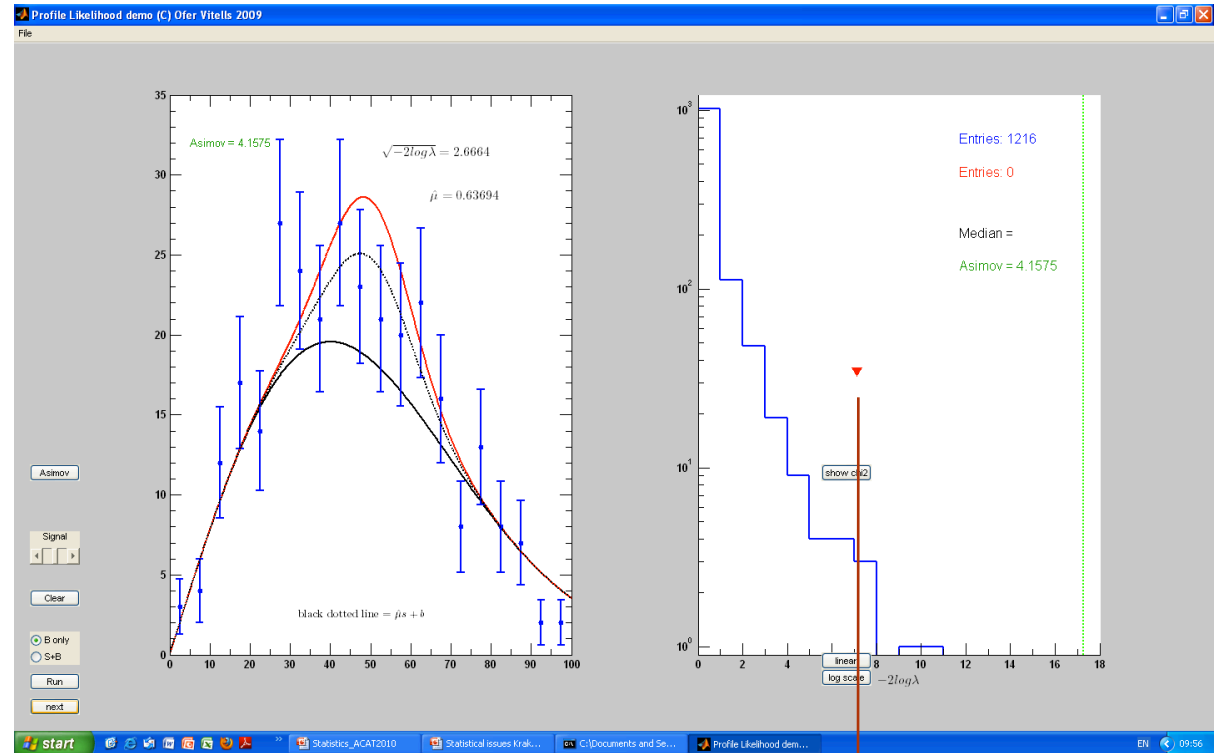
$$q_0 = 0$$



PL: test q_0 under BG only ; $f(q_0 | H_0)$

$\hat{\mu} = 0.6 \rightarrow 2.6\sigma$

$$q_0 = -2 \ln \frac{L(b)}{L(\hat{\mu}s + b)}$$



$$Z_{obs} = \sqrt{q_{0,obs}}$$

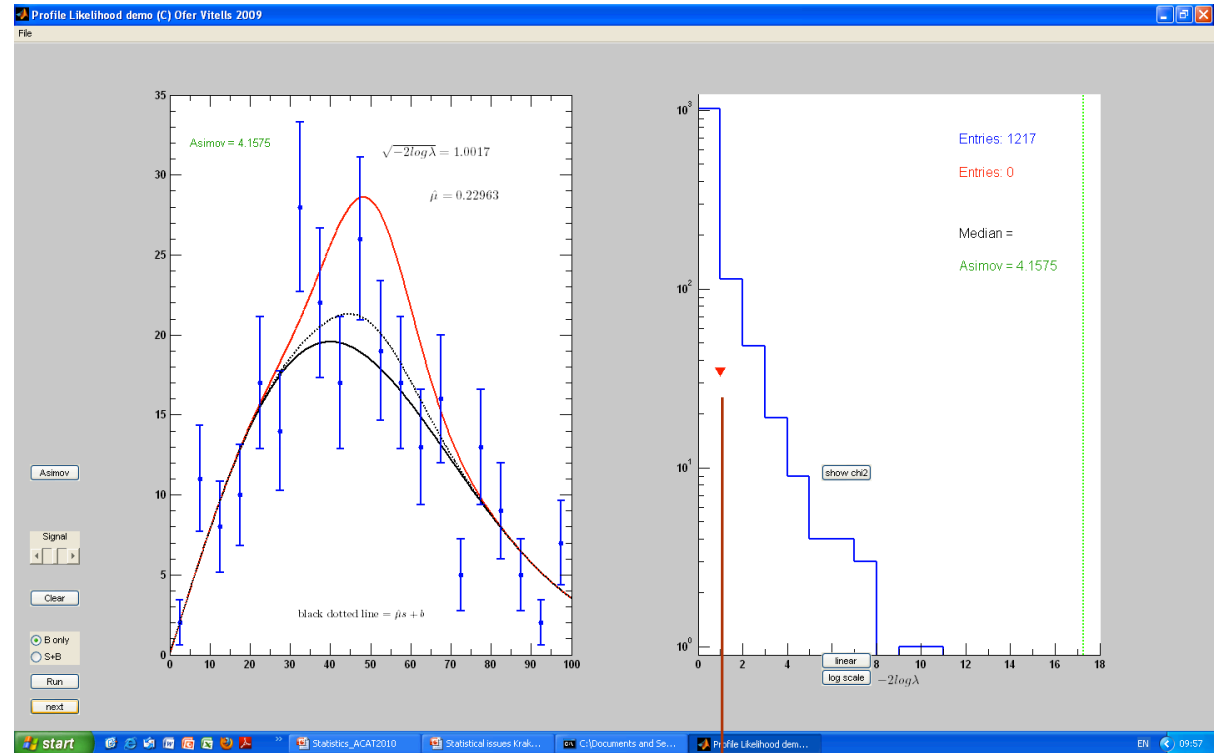
$$q_0 = 6.76 \rightarrow Z = 2.6\sigma$$



PL: test q_0 under BG only ; $f(q_0 | H_0)$

$\hat{\mu} = 0.22 \rightarrow 1.1\sigma$

$$q_0 = -2 \ln \frac{L(b)}{L(\hat{\mu}s + b)}$$



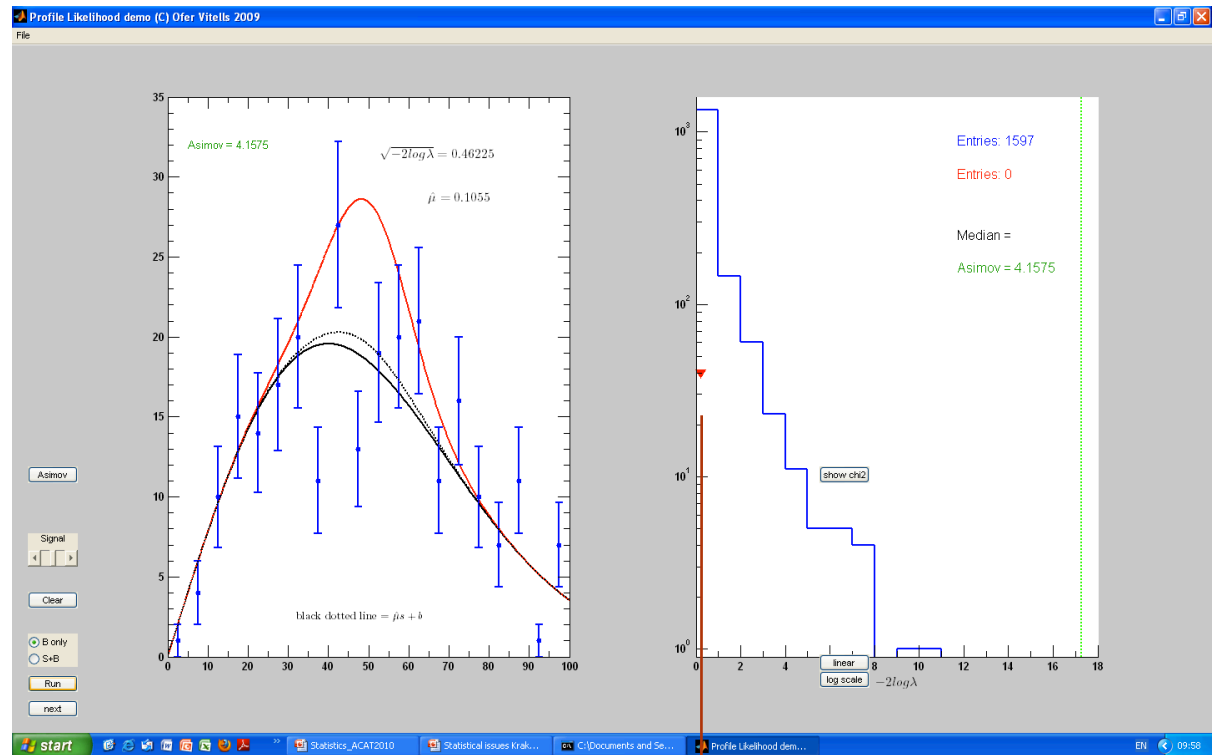
$$Z_{obs} = \sqrt{q_{0,obs}} \quad q_0 = 1.2 \rightarrow Z = 1.1\sigma$$



PL: test q_0 under BG only ; $f(q_0 | H_0)$

$\hat{\mu} = 0.11 \rightarrow 0.4\sigma$

$$q_0 = -2 \ln \frac{L(b)}{L(\hat{\mu}s + b)}$$



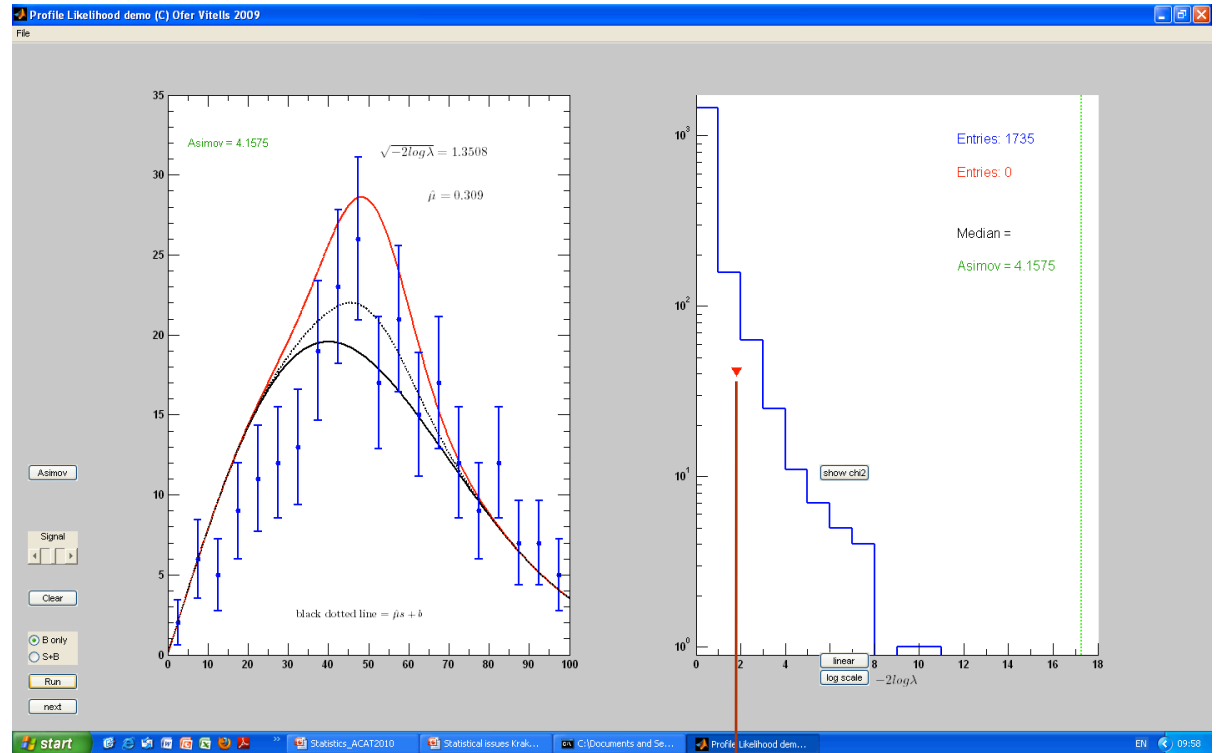
$$Z_{obs} = \sqrt{q_{0,obs}} \quad q_0 = 0.16 \rightarrow Z = 0.4\sigma$$



PL: test q_0 under BG only ; $f(q_0 | H_0)$

$\hat{\mu} = 0.31 \rightarrow 1.35\sigma$

$$q_0 = -2 \ln \frac{L(b)}{L(\hat{\mu}s + b)}$$



$$Z_{obs} = \sqrt{q_{0,obs}}$$

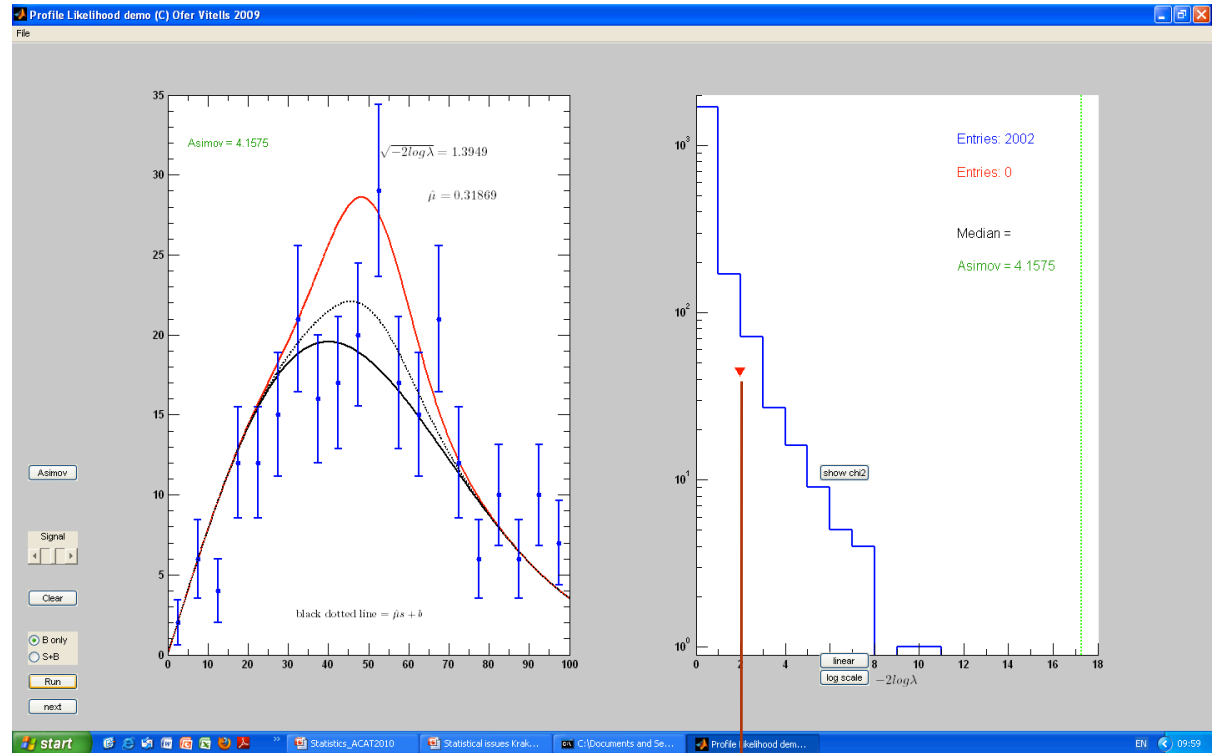
$$q_0 = 1.8 \rightarrow Z = 1.35\sigma$$



PL: test q_0 under BG only ; $f(q_0 | H_0)$

$\hat{\mu} = 0.32 \rightarrow 1.39\sigma$

$$q_0 = -2 \ln \frac{L(b)}{L(\hat{\mu}s + b)}$$



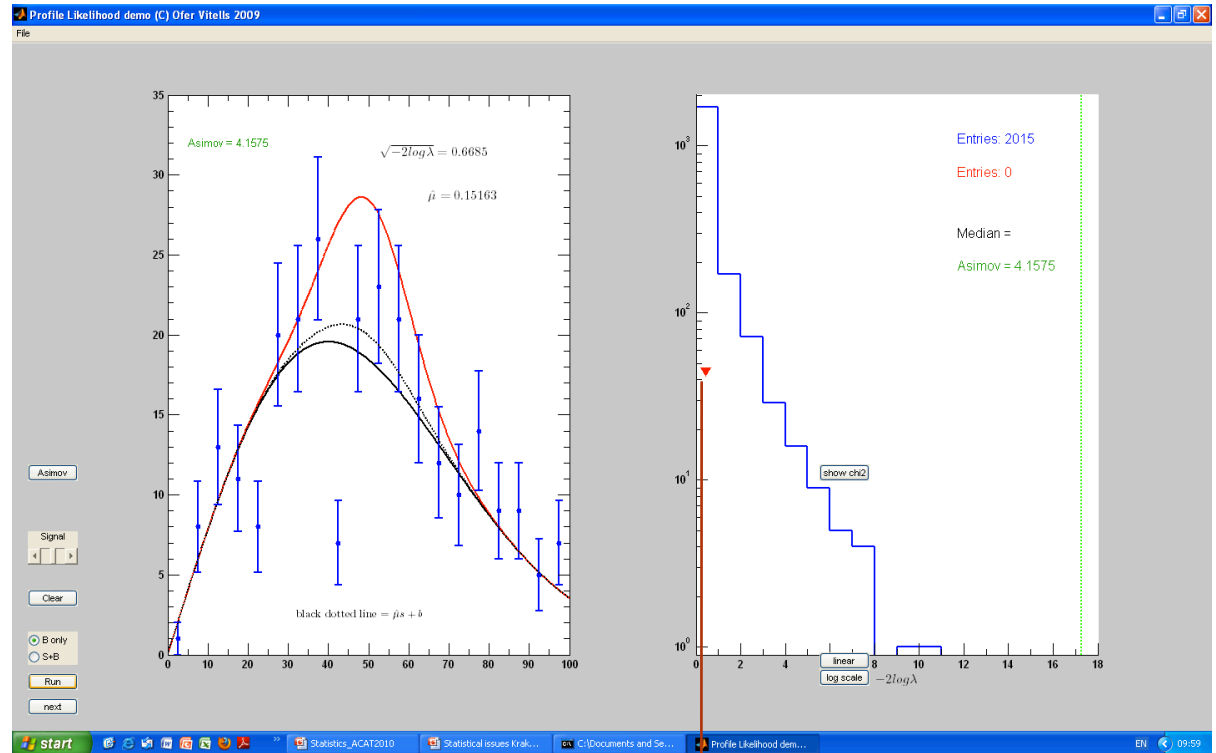
$$Z_{obs} = \sqrt{q_{0,obs}} \quad q_0 = 1.9 \rightarrow Z = 1.39\sigma$$



PL: test q_0 under BG only ; $f(q_0 | H_0)$

$\hat{\mu} = 0.15 \rightarrow 0.66\sigma$

$$q_0 = -2 \ln \frac{L(b)}{L(\hat{\mu}s + b)}$$



$$Z_{obs} = \sqrt{q_{0,obs}} \quad q_0 = 0.43 \rightarrow Z = 0.66\sigma$$



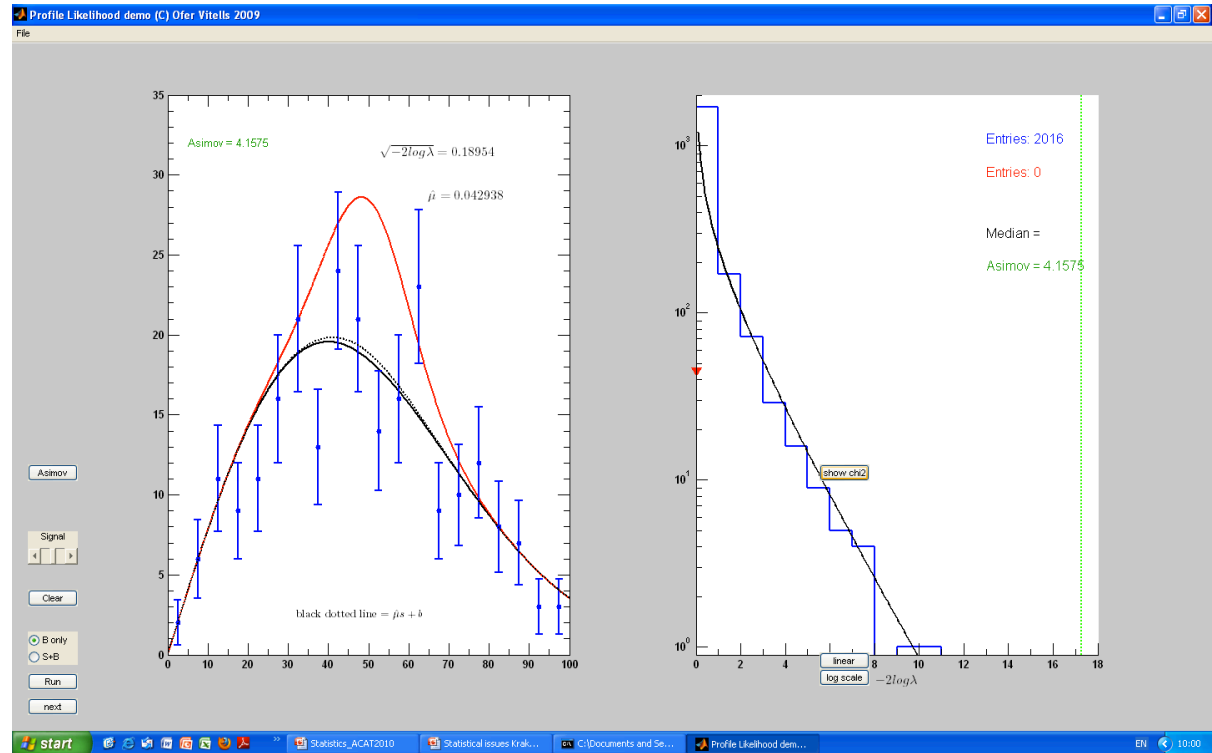
Confirm Wilks Theorem

$$q_0 = -2 \ln \frac{L(b)}{L(\hat{\mu}s + b)}$$

For the test statistic

$$q_0 = -2 \ln \frac{L(b)}{L(\hat{\mu}s + b)}$$

$$f(q_0 | H_0) = \chi_1^2$$



χ^2 Distribution

Chi Squared Distribution

In probability theory and statistics,

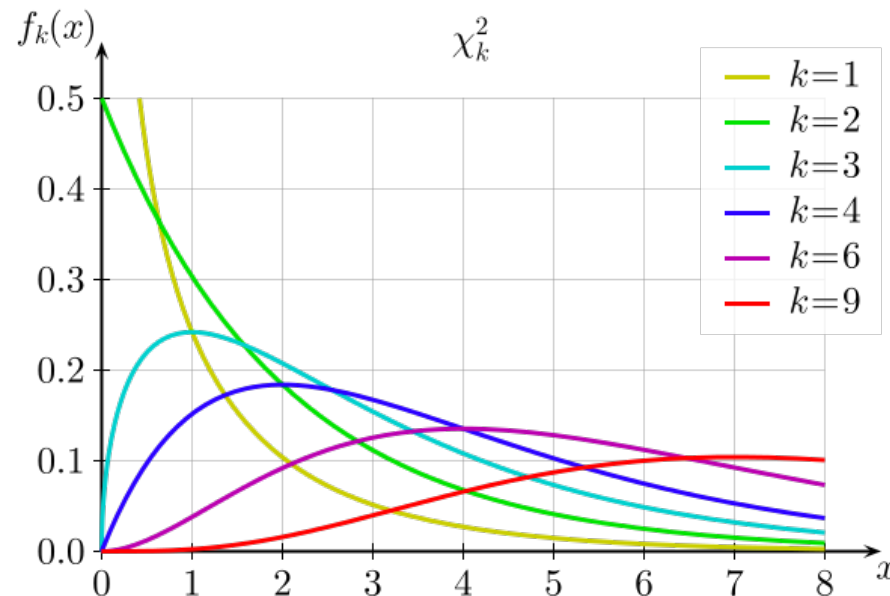
the **chi – squared distribution** (also **chi – square** or **χ^2 – distribution**) with k degrees of freedom is the distribution of a sum of the squares of k independent standard normal random variables.

$$\chi_k^2 = \sum_i^k \frac{(x_i - \mu_i)^2}{\sigma_i^2}$$

q_0 distributes like a χ^2

$$\sqrt{q_{0,obs}} = Z$$

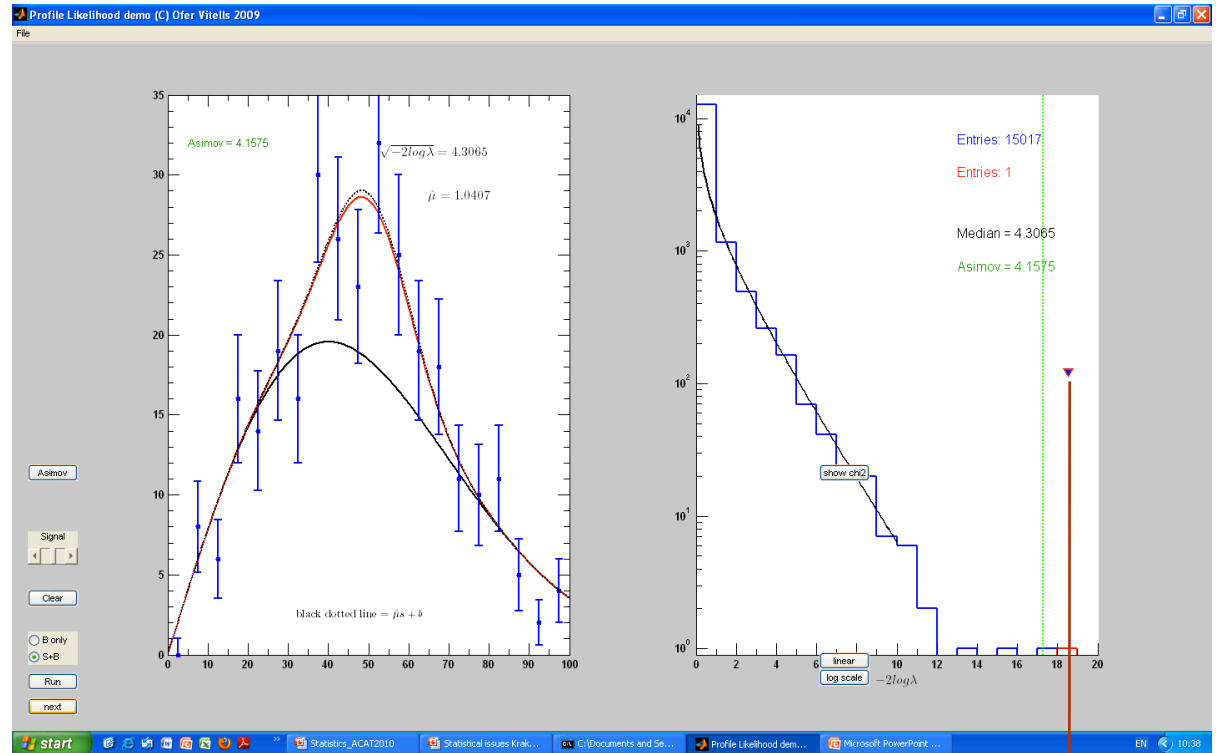
$$\chi^2 = \frac{(x - \mu)^2}{\sigma^2} = Z^2$$



The PDF of q_0 under s+b experiments (H_1)

$$q_0 = -2 \ln \frac{L(b)}{L(\hat{\mu}s + b)} = -2 \ln \frac{L(b | H_1)}{L(\hat{\mu}s + b | H_1)}$$

$$\hat{\mu} = 1.04 \rightarrow 4.3\sigma$$



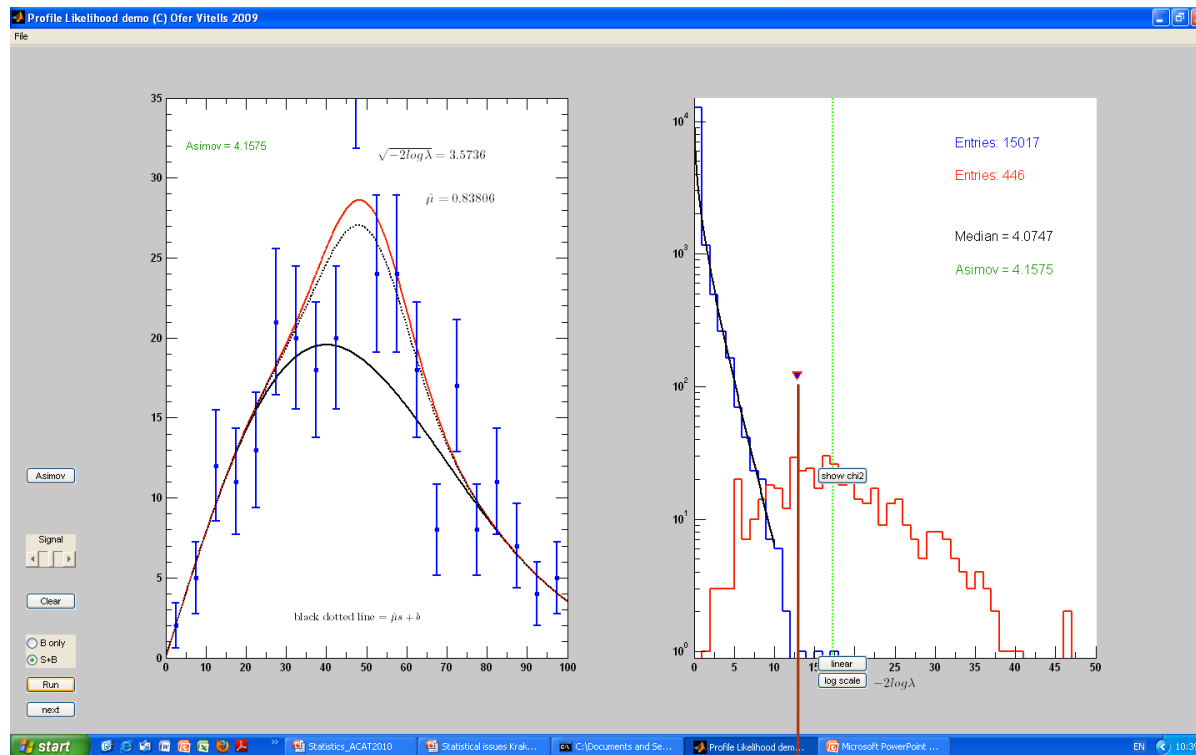
$$q_0 = 18.5 \rightarrow Z = 4.3\sigma$$



The PDF of q_0 under s+b experiments (H_1)

$$q_0 = -2 \ln \frac{L(b)}{L(\hat{\mu}s + b)} = -2 \ln \frac{L(b | H_1)}{L(\hat{\mu}s + b | H_1)}$$

$$\hat{\mu} = 0.83 \rightarrow 3.6\sigma$$



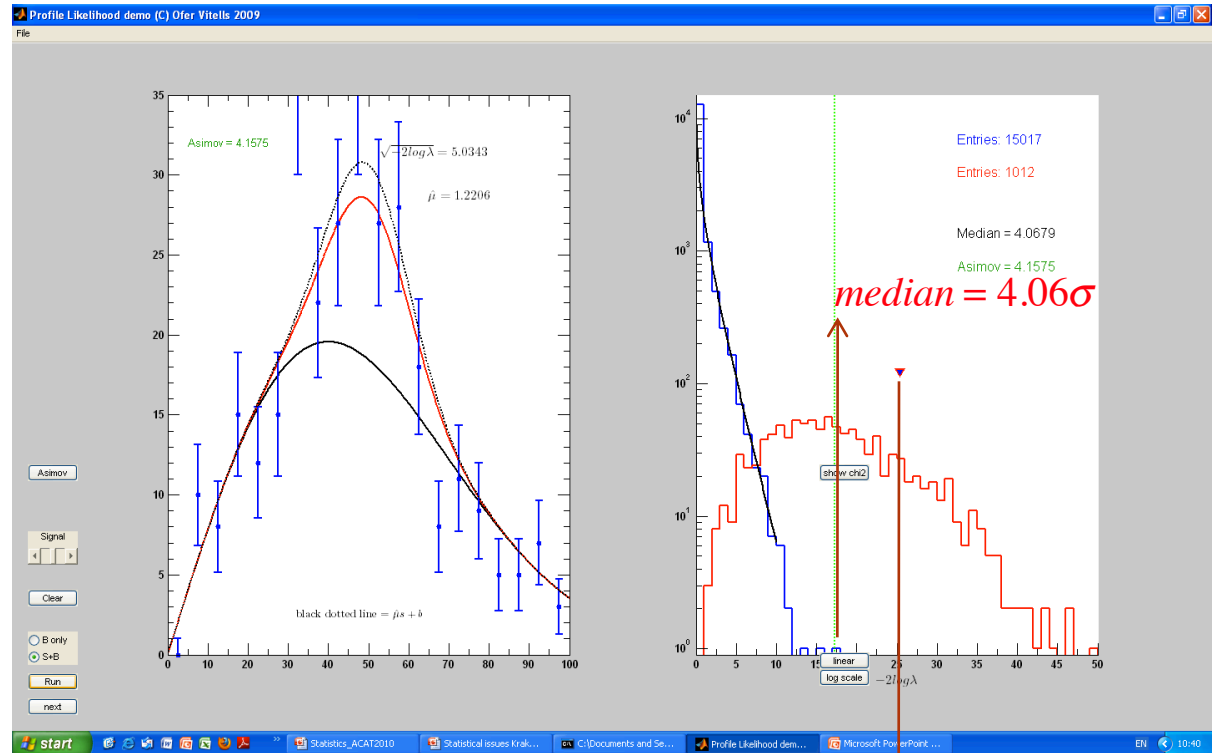
$$q_0 = 12.9 \rightarrow Z = 3.6\sigma$$



The PDF of q_0 under s+b experiments (H_1)

$$q_0 = -2 \ln \frac{L(b)}{L(\hat{\mu}s + b)} = -2 \ln \frac{L(b | H_1)}{L(\hat{\mu}s + b | H_1)}$$

$$\hat{\mu} = 1.22 \rightarrow 5.0\sigma$$



$$q_0 = 25 \rightarrow Z = 5.0\sigma$$



Median sensitivity in a Click (Asimov)

Franchise (short story)

From Wikipedia, the free encyclopedia



This article **needs additional citations for verification**. Please help improve this article by adding citations to reliable sources. Unsourced material may be **challenged** and **removed**. (December 2009)

Franchise is a **science fiction short story** by **Isaac Asimov**. It first appeared in the August 1955 issue of the magazine *Amazing Stories* and was reprinted in the collections *Earth Is Room Enough* (1957) and *Robot Dreams* (1986). It is one of a loosely connected series of fictional **computer** called **Multivac**. It is the first story in which Asimov dealt with computers as *computers* and not as *robots*.

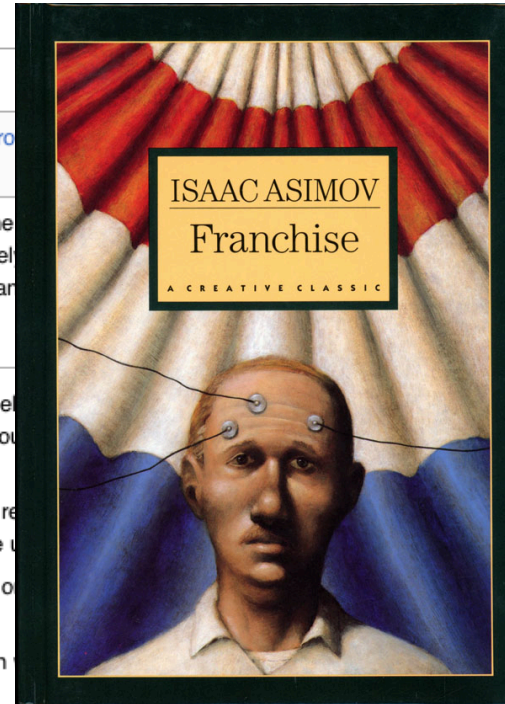
Plot summary

In the future, the **United States** has converted to an "electronic **democracy**" where the computer Multivac selects the candidates for office and asks them questions. Multivac will then use the answers and other data to determine what the results of an **election** would be if an election were to be held.

The story centers around Norman Muller, the man chosen as "Voter of the Year" in **2008**. Although the law requires that he represent the entire **electorate**, he is not sure that he wants the responsibility of representing the entire **electorate**, worrying that the result will be too close to call.

However, after 'voting', he is very proud that the citizens of the United States had, through him, "exercised their right to vote" and makes a statement that is somewhat ironic as the citizens didn't actually get to vote.

The idea of a computer predicting whom the electorate would vote for instead of actually holding an election is a common theme in science fiction. The correct prediction of the result of the **1952 election**.



S. Unsourced	
"Franchise"	
Author	Isaac Asimov
Country	United States
Language	English
Series	Multivac
Genre(s)	science fiction
Published in	<i>If</i>
Publisher	Quinn Publications
Media type	Magazine
Publication date	August 1955
Preceded by	"Question"
Followed by	"The Dead End"

Influence

The use of a single representative individual to stand in for the entire population can help in evaluating the sensitivity of a statistical method. *Franchise* was cited as the inspiration of the "data set", where an ensemble of simulated experiments can be replaced by a single representative one. ^[1]

References

- [↑] G. Cowan, K. Cranmer, E. Gross, and O. Vitells (2011). "Asymptotic formulae for likelihood-based tests of new physics". *Eur.Phys.J.* **C71**: 1554. DOI:10.1140/epjc/s10052-011-1554-0 [↗](#).

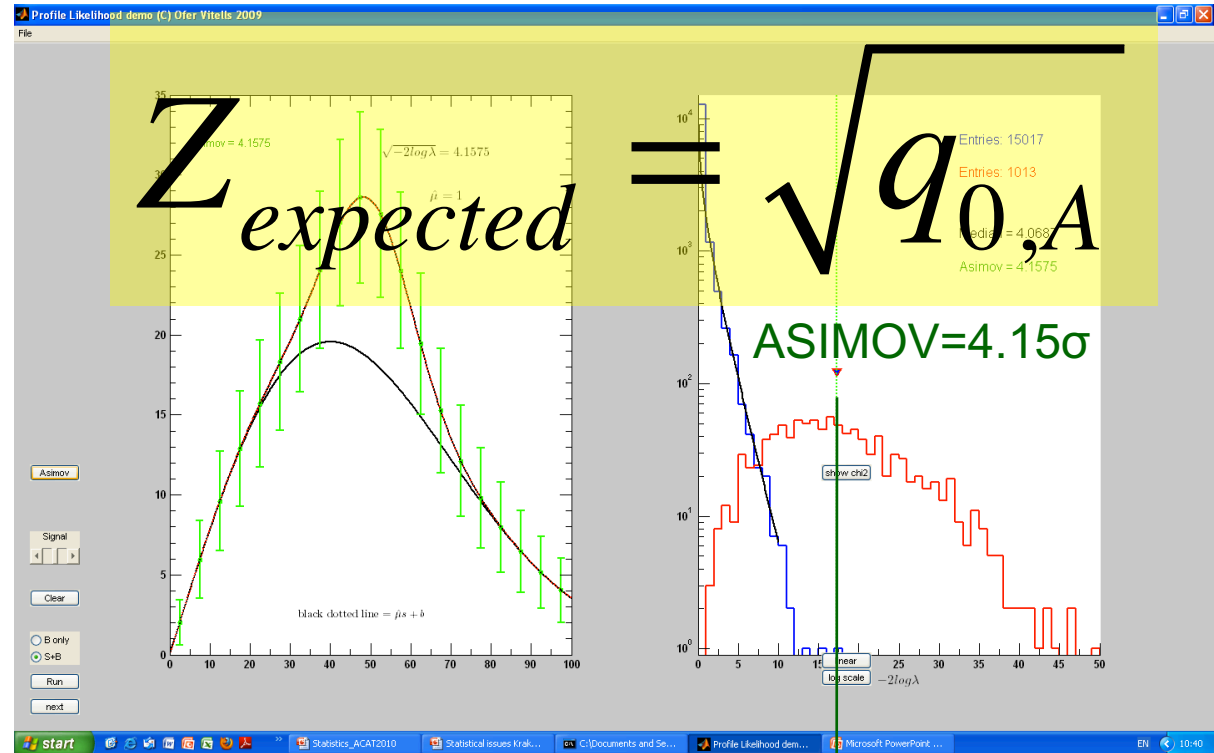
CCGV ref



The Median Sensitivity (via ASIMOV)

To estimate the median sensitivity of an experiment (before looking at the data), one can either perform lots of $s+b$ experiments and estimate the median $t_{0,med}$ or evaluate t_0 with respect to a representative data set, the ASIMOV data set with $\hat{\mu}=1$, i.e. $x=s+b$

$$\hat{\mu} = 1.00 \rightarrow 4.15\sigma$$



$$= \sqrt{q_{0,A}}$$

ASIMOV=4.15σ

$$q_A = 17.22 \rightarrow Z_A = 4.15$$

$$q_{o,med} \approx q_0(\hat{\mu} = \mu_A = 1) = -2 \ln \frac{L(b | x = x_A = s + b)}{L(\hat{\mu}s + b | x = x_A = s + b)} = -2 \ln \frac{L(b)}{L(1 \cdot s + b)}$$



Asymptotic Distributions – a Teaser

Tossing Monte Carlos to get the test statistic distribution functions (PDF) is sometimes beyond the experiment technical capability.

Knowing both PDF

$$f(q_{null} | H_{null})$$

$$f(q_{null} | H_{alternate})$$

enables calculating both the observed and expected significance (or exclusion) without a single toy....

