

# Multi moment cancellation of participant fluctuations - MMCP method

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- The only experimentally controllable way to probe the **QCD phase diagram** is by studying interactions of **different system size** nuclei at various **energies**.
- **Participant fluctuations** is one of the **main background** effects in such study.
- It is the number of nucleons,  **$N_p$** , that interacted inelastically and produced other particles during the collision.
- These fluctuations may hide the fluctuations from other sources.

There are several popular ways of reducing participant fluctuations:

- (i) the selection of as **narrow centrality bins** as possible
- (ii) the Centrality Bin Width Correction procedure (**CBWC**) (STAR, Luo (2011)),
- (iii) the use of **strongly intensive** quantities (Gazdzicki, Mrowczynski (1992), Gorenstein, Gazdzicki (2011)),

**We propose** a different approach - to **cancel participant fluctuations** in a **combination** of several **high fluctuation moments** (v.B., Mackowiak-Pawlowska 1705.01110).

A **multiplicity distribution**,  $P(N)$ , can be characterized by **central moments**,  $m_n$ ,

$$m_n = \sum_N (N - \langle N \rangle)^n P(N), \quad \text{where } \langle N^n \rangle = \sum_N N^n P(N).$$

They are related to **cumulants**,  $\kappa_n$ ,

$$\kappa_2 = m_2, \quad \kappa_3 = m_3, \quad \kappa_4 = m_4 - 3m_2^2, \quad \dots,$$

and **susceptibilities**,  $\chi_n$ ,

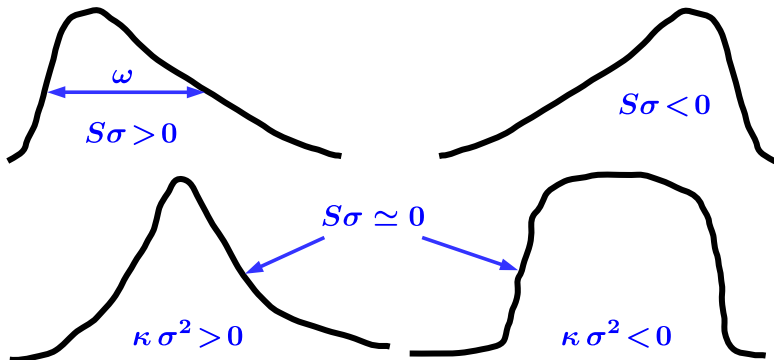
$$\chi_n = \frac{\partial^n (\mathcal{P}/T^4)}{\partial (\mu/T)^n} = \frac{\kappa_n}{V T^3}, \quad \chi_{n,k} = \frac{\chi_n}{\chi_k} = \frac{\kappa_n}{\kappa_k},$$

where  $\mathcal{P}$  is pressure,  $T$  - temperature,  $\mu$  - chemical potential, and  $V$  - volume.

Frequently used cumulant ratios - **scaled variance**, normalized **skewness** and normalized **kurtosis** - are:

$$\omega = \frac{\kappa_2}{\langle N \rangle} = \frac{\sigma^2}{\langle N \rangle}, \quad S \sigma = \frac{\kappa_3}{\kappa_2}, \quad \kappa \sigma^2 = \frac{\kappa_4}{\kappa_2},$$

where  $\sigma = \sqrt{\langle N^2 \rangle - \langle N \rangle^2} = \sqrt{\kappa_2}$  is **standard deviation**.



**Poisson** distribution:  $\omega = S\sigma = \kappa\sigma^2 = 1$

**Normal (Gauss)** distribution:  $\omega$  - free parameter,  $S\sigma = \kappa\sigma^2 = 0$

**Log-normal** distribution:  $\omega \sim \langle N \rangle$ ,  $S\sigma \sim \langle N \rangle$ ,  $\kappa\sigma^2 \sim \langle N \rangle^2$

- The approach '**just take negative binomial** (Poisson, Gauss...)' is **not working**, because it imposes a certain relation between moments, which might not exist.

# Independent participant model

The only assumptions are that participants are **identical** and **independent**. Then mean multiplicity  $\mathbf{N}$  is the sum of contributions from  $\mathbf{N}_P$  participants,

$$\mathbf{N} = n_1 + n_2 + \dots + n_{N_P}, \quad \text{and} \quad \langle \mathbf{N} \rangle = \langle \mathbf{N}_P \rangle \langle n_A \rangle,$$

where **identical** and **independent** means that,  $\langle n_i \rangle = \langle n_j \rangle = \langle n_1 \rangle = \langle n_A \rangle$ , and  $\langle n_i n_j \dots n_k \rangle = \langle n_A \rangle^k$ . Using multinomial theorem

$$N^k = (n_1 + n_2 + \dots + n_{N_P})^k = \sum_{k_1, k_2, \dots, k_{N_P}} \frac{k!}{k_1! k_2! \dots k_{N_P}!} n_1^{k_1} n_2^{k_2} \dots n_{N_P}^{k_{N_P}} \delta \left( k - \sum_{i=1}^{N_P} k_i \right),$$

where  $\delta$  is the Kronecker delta, one can obtain arbitrarily high moments, e.g.

$$\omega = \omega_A + \langle n_A \rangle \omega_P,$$

$$S \sigma = \frac{\omega_A S_A \sigma_A + \langle n_A \rangle \omega_P [3 \omega_A + \langle n_A \rangle S_P \sigma_P]}{\omega_A + \langle n_A \rangle \omega_P},$$

$$\kappa \sigma^2 = \frac{\omega_A \kappa_A \sigma_A^2 + \langle n_A \rangle \omega_P \left[ \langle n_A \rangle^2 \kappa_P \sigma_P^2 + \omega_A (3 \omega_A + 4 S_A \sigma_A + 6 \langle n_A \rangle S_P \sigma_P) \right]}{\omega_A + \langle n_A \rangle \omega_P},$$

**red** - what we would like to measure, **black** - what we measure, **blue** - participant fluctuations (V.B. 1606.05358, Skokov, Friman, Redlich (2013), Braun-Munzinger, Rustamov, Stachel (2017))

# The problem

- A moment of a rank  $n$  is a function of all lower moments for both participants,  $\langle N_{\mathbf{P}}^n \rangle$ , and a source,  $\langle n_{\mathbf{A}}^n \rangle$ ,

$$\langle N^n \rangle = \mathcal{F} \left( \langle n_{\mathbf{A}}^1 \rangle, \langle n_{\mathbf{A}}^2 \rangle, \dots, \langle n_{\mathbf{A}}^n \rangle, \langle N_{\mathbf{P}}^1 \rangle, \langle N_{\mathbf{P}}^2 \rangle, \dots, \langle N_{\mathbf{P}}^n \rangle \right) .$$

Therefore, one has only  $n$  **measures**, but  $2n$  **unknowns** for their description.

- **Strongly intensive** measures **require** two types of values, e.g. pions and kaons,

$$\langle N_{\mathbf{A}}^n \rangle = \mathcal{F} \left( \langle n_{\mathbf{A}}^1 \rangle, \langle n_{\mathbf{A}}^2 \rangle, \dots, \langle n_{\mathbf{A}}^n \rangle, \langle N_{\mathbf{P}_A}^1 \rangle, \langle N_{\mathbf{P}_A}^2 \rangle, \dots, \langle N_{\mathbf{P}_A}^n \rangle \right) ,$$

$$\langle N_{\mathbf{B}}^n \rangle = \mathcal{F} \left( \langle n_{\mathbf{B}}^1 \rangle, \langle n_{\mathbf{B}}^2 \rangle, \dots, \langle n_{\mathbf{B}}^n \rangle, \langle N_{\mathbf{P}_B}^1 \rangle, \langle N_{\mathbf{P}_B}^2 \rangle, \dots, \langle N_{\mathbf{P}_B}^n \rangle \right) .$$

and the **assumption** that all corresponding participant fluctuations moments are the same  $\langle N_{\mathbf{P}_A}^n \rangle = \langle N_{\mathbf{P}_B}^n \rangle = \langle N_{\mathbf{P}}^n \rangle$ , which gives  $3n$  **unknowns** for  $2n$  **measured** values.

- **Wounded nucleon** model gives all  $\langle N_{\mathbf{P}}^n \rangle$ , but it is **not working** - **participants are not protons**. The new SPS data of the **NA49** and **NA61/SHINE** show that

$$\omega_{\mathbf{p}+\mathbf{p}} > \omega_{\mathbf{Ar}+\mathbf{Sc}} , \omega_{\mathbf{Pb}+\mathbf{Pb}} \quad \text{at SPS (Rybczynski (2013), Aduszkiewicz (2015), Seryakov (2017))}$$

i.e.  $\omega_{\mathbf{p}}$  can be negative, which is forbidden by definition. At higher energies the wounded nucleon model clearly contradicts the data, because

$$\omega_{\mathbf{p}+\mathbf{p}} \gg \omega_{\mathbf{Pb}+\mathbf{Pb}} \quad \text{at LHC (v.B. 1606.05358)}$$

- **Suppose** that **current** experimental **methods** are **effective enough** to make scaled variance for the fluctuations from a source close to the measured fluctuations,

$$\omega \simeq \omega_A, \quad \text{then } \alpha = \frac{\omega - \omega_A}{\omega_A} = \langle n_A \rangle \frac{\sigma_P}{\omega_A} \ll 1 \text{ is a small parameter}$$

- $\omega_A$  competes with  $\langle n_A \rangle S_P \sigma_P$  and  $\langle n_A \rangle^2 \kappa_P \sigma_P^2$ , **assume** that their ratio is also small,

$$|\beta| = \langle n_A \rangle \frac{|S_P \sigma_P|}{\omega_A} \ll 1, \quad |\gamma| = \langle n_A \rangle^2 \frac{|\kappa_P \sigma_P^2|}{\omega_A^2} \ll 1,$$

The  $\alpha, |\beta|, |\gamma| \ll 1$  is the **mathematical meaning** of the 'small participant fluctuations'. Then

$$\begin{aligned} \omega_A &\simeq \omega - \alpha \omega, \\ S_A \sigma_A &\simeq S \sigma + \alpha (S \sigma - 3 \omega_A), \\ \kappa_A \sigma_A^2 &\simeq \kappa \sigma^2 + \alpha (\kappa \sigma^2 - 3 \omega_A^2 - 4 \omega_A S_A \sigma_A), \quad \alpha, |\beta|, |\gamma| \ll 1. \end{aligned}$$

The skewness  $S_A \sigma_A$  has a contribution  $\omega_A$ , while kurtosis  $\kappa_A \sigma_A^2$  includes  $\omega_A^2$ . Therefore, it is reasonable to introduce two more notations:

$$\delta = \frac{S_A \sigma_A}{\omega_A} \quad \text{and} \quad \varepsilon = \frac{\kappa_A \sigma_A^2}{\omega_A^2}$$

Then, without any approximation

$$\begin{aligned}\omega &= \omega_A(1 + \alpha), \\ S\sigma &= \omega_A \frac{3\alpha + \delta + \alpha\beta}{1 + \alpha}, \\ \kappa\sigma^2 &= \omega_A^2 \frac{3\alpha + \varepsilon + \alpha [6\beta + \gamma + 4\delta]}{1 + \alpha}.\end{aligned}$$

One can make  $\alpha$ ,  $|\beta|$ ,  $|\gamma| \ll 1$  small by **decreasing bin width**, or by decreasing  $\langle n_A \rangle$ , choosing **rare particles**, or **net charges** for analysis (Braun-Munzinger, Rustamov, Stachel (2017)).

**Close to critical point**  $\kappa_n \sim \xi^{\frac{5(n-1)-1}{2}}$ ,  $\xi \rightarrow \infty$  (Stephanov (2008), Mukherjeea, Venugopalan, Yin (QM2017)), then

$$\alpha \sim \langle n_A \rangle^2 \frac{\omega_P}{\xi^2} \rightarrow 0, \quad |\beta| \sim \langle n_A \rangle^2 \frac{|S_P \sigma_P|}{\xi^2} \rightarrow 0, \quad |\gamma| \sim \langle n_A \rangle^4 \frac{|\kappa_P \sigma_P^2|}{\xi^4} \rightarrow 0,$$

while  $\delta \sim \langle n_A \rangle \xi^{0.5} \rightarrow \infty$ , and  $\varepsilon \sim \langle n_A \rangle^2 \xi \rightarrow \infty$ .

**Away from critical point**, e.g. in hadron gas, (Karsch, Redlich (2011))

$$\delta = \varepsilon = \frac{1}{\omega_A^2} = \tanh^2(\mu_B/T) \rightarrow 0, \quad \mu_B/T \rightarrow 0$$

$\delta = \varepsilon = 0$  for Gauss, and  $\delta = \varepsilon = 1$  for Poisson distribution.



For  $\alpha$ ,  $|\beta|$ ,  $|\gamma| \ll 1$ , and neglecting **either**  $\delta$  or  $\varepsilon$ , one can solve the system

$$\begin{aligned}\omega &= \omega_A(1 + \alpha), \\ S\sigma &= \omega_A \frac{3\alpha + \delta + \alpha\beta}{1 + \alpha}, \\ \kappa\sigma^2 &= \omega_A^2 \frac{3\alpha + \varepsilon + \alpha[6\beta + \gamma + 4\delta]}{1 + \alpha}.\end{aligned}$$

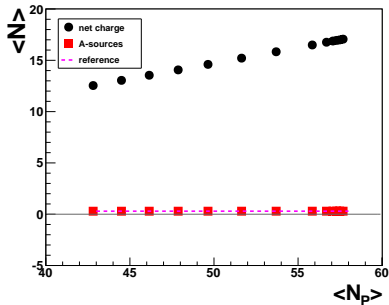
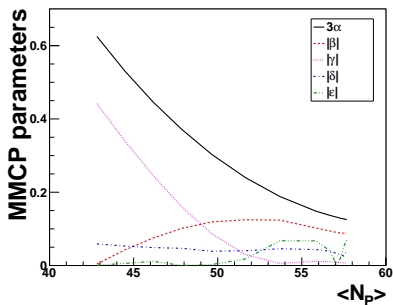
In case of  $\delta \ll \varepsilon$  there is a simple **analytic solution**:

$$\begin{aligned}\alpha &\simeq \frac{S\sigma}{3\omega - S\sigma}, \\ \omega_A &\simeq \omega - \frac{S\sigma}{3}, \\ \kappa_A\sigma_A^2 &\simeq \kappa\sigma^2 - \omega_A S\sigma, \quad \alpha, |\beta|, |\gamma| \ll 1, \quad \delta \ll \varepsilon.\end{aligned}$$

These approximate equations **remove** fluctuations of **participants** and **obtain** the fluctuation of **sources** through **measured** values. This is the meaning of the **MMCP method**.

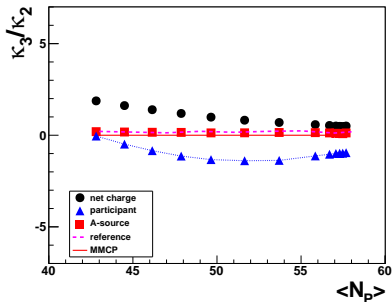
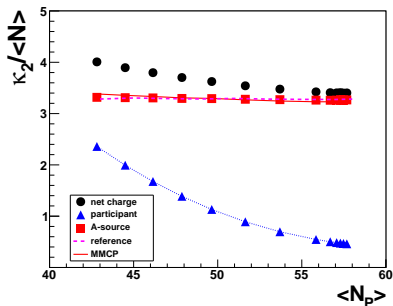
# Test of the MMCP in EPOS – parameters and net electric charge

Net electric charge in  $^{40}_{18}\text{Ar} + ^{45}_{21}\text{Sc}$  at  $p_{lab} = 150 \text{ GeV}/c$  and with  $y^{CMS} > 0$  in centrality, left to right: 20%, 17.5%, 15%, 12.5%, 10%, 7.5%, 5%, 2.5%, 1.5%, 1%, 0.75%, 0.5% and 0.2%.



- The condition  $\alpha, |\beta|, |\gamma| \ll 1$  is valid for the system that we study
- The  $|\delta| \ll |\epsilon|$  is not always satisfied, however,  $|\delta|, |\epsilon| < 3\alpha$ .
- $\langle N_{\text{net charge}} \rangle / \langle N_p^{\text{max}} \rangle = (18 + 21) / (40 + 45) \approx 0.5$ , condition  $y^{CMS} > 0$  corresponds to 1/2 of created system, therefore,  $\langle n_A \rangle \approx 0.5 * 0.5 = 0.25 \approx 0.3$ .
- $\langle n_A \rangle$  is independent of centrality. (V.B., Mackowiak-Pawlowska 1705.01110)

# Test of the MMCP in EPOS – scaled variance and normalized skewness

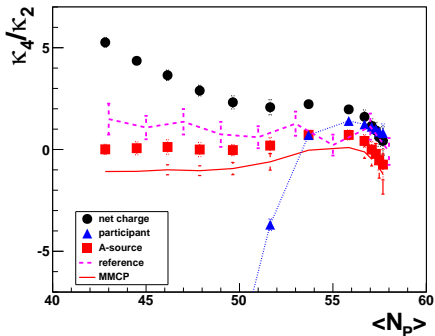


**red** - what we would like to measure, **black** - what we measure, **blue** - participants, **magenta** - the 'reference' is obtained selecting  $N_p = \text{const}$ .

- $\omega_A$  is independent of centrality window (V.B., Mackowiak-Pawlowska 1705.01110)
- $\langle n_A \rangle$  is small, but its fluctuations give the main contribution  $\omega \sim \omega_A \approx 3.3 \gg \omega_P \gtrsim 0.5$
- $\omega_P$  is small, but it exists, even if bin width goes to zero
- The skewness of a source,  $S_A \sigma_A$ , is also independent on centrality and is close to zero.
- The large values of the net charge  $S \sigma$  are due to the second moment fluctuations of participants,  $S_A \sigma_A \approx 0.1 \ll S \sigma \approx 3 \langle n_A \rangle \omega_P > 0$ , while  $S_P \sigma_P \leq 0$ , for all  $\delta c$

# Test of the MMCP in EPOS – normalized kurtosis

Left to right: 20%, 17.5%, 15%, 12.5%, 10%, 7.5%, 5%, 2.5%, 1.5%, 1%, 0.75%, 0.5% and 0.2%.



**red** - what we would like to measure, **black** - what we measure, **blue** - participants, **magenta** - the 'reference' is obtained selecting  $N_p = \text{const}$ .

- The **higher** the **order**, the **stronger** is the **dependence** on the **bin width**  $\delta c$
- In very central collision in a **fixed target** experiment  $N_p \approx \text{const}$ , while in **collider mode** and in **peripheral collisions**  $N_p \neq \text{const}$ . Therefore, **STAR**, **NA61** and **STAR fixed target** data may be **hard to compare** (v.B., Mackowiak-Pawlowska 1705.01110).

# Test of the CBWC method in EPOS

The **CBWC** procedure used by **STAR** means that a value  $X$  is measured in  $r$  sub-samples, and then summed up with the relative weights  $w_r$  of the sub-samples  $r$ ,

$$X = \sum_r w_r X_r, \quad w_r = n_r / \sum_r n_r,$$

where  $n_r$  is the number of events in the bin  $r$ .

bin width	$n_r$	$w_r$	$\langle N \rangle$	$\sigma^2 / \langle N \rangle$	$S\sigma$	$\kappa\sigma^2$
0-1%	624827	$\approx 0.2$	16.88(1)	3.405(3)	0.503(4)	1.2(3)
1-2%	626043	$\approx 0.2$	16.36(1)	3.417(3)	0.603(4)	2.4(3)
<b>2-3%</b>	611242	$\approx 0.2$	<b>15.83(1)</b>	<b>3.447(3)</b>	<b>0.609(4)</b>	<b>2.2(3)</b>
3-4%	665988	$\approx 0.2$	15.32(1)	3.467(3)	0.660(4)	2.9(3)
4-5%	623110	$\approx 0.2$	14.81(1)	3.486(3)	0.741(3)	2.3(3)
<b>0-5%</b>	3151210	1.0	<b>15.834(4)</b>	<b>3.442(1)</b>	<b>0.625(2)</b>	<b>2.2(1)</b>

The sub-bin values for the collected number of events  $n_r$ , the weight of the sub-bin  $w_r$ , the average net charge  $\langle N \rangle$ , scaled variance  $\omega = \kappa_2 / \langle N \rangle = \sigma^2 / \langle N \rangle$ , normalized skewness  $S\sigma = \kappa_3 / \kappa_2$ , and normalized kurtosis  $\kappa\sigma^2 = \kappa_4 / \kappa_2$  (v.B., Mackowiak-Pawlowska 1705.01110).

- The **CBWC reduces** statistical **uncertainty**.
- The **CBWC gives average measured** fluctuations over the selected sub-bins.

# Comparison of different methods for 0%–5% and 0%–20% centrality

0 – 5%	CBWC	net charge	reference	A-source	MMCP
$\kappa_2/\langle N \rangle$	3.4	3.5	3.3	3.3	3.3
$\kappa_3/\kappa_2$	0.6	0.7	0.2	0.2	0
$\kappa_4/\kappa_2$	2.2(1)	2.2(2)	1.3(5)	0.7(2)	0.0(2)

0 – 20%	net charge	A-source	MMCP
$\kappa_2/\langle N \rangle$	4.0	3.3	3.4
$\kappa_3/\kappa_2$	1.9	0.2(1)	0
$\kappa_4/\kappa_2$	5.3(2)	0.0(2)	-1.1(2)

- The **MMCP coincides** within the uncertainty with the **A-source** for  $\omega = \kappa_2/\langle N \rangle$  in the **0 – 5%** centrality, and deviates only for **2%** from A-source in the **0 – 20%** centrality.
- The **CBWC overestimates**  $S\sigma = \kappa_3/\kappa_2$  and  $\kappa\sigma^2 = \kappa_4/\kappa_2$  **three times**.
- The  $\kappa_3/\kappa_2$  of sources is zero by definition in the **MMCP**. It **agrees** within three standard deviations with the **A-source** generated by EPOS.
- The  $\kappa_4/\kappa_2$ , in MMCP for 0-5% bin underestimates the A-source. This is the result of neglecting skewness of the participants  $S_P \sigma_P$  (v.B., Mackowiak-Pawlowska 1705.01110).

- The average number of particles produced by a source,  $\langle n_A \rangle$ , and its fluctuations of the second,  $\omega_A$ , and the third order,  $S_A \sigma_A$ , **do not depend on centrality**.
- However, the fourth order fluctuations of a source,  $\kappa_A \sigma_A^2$ , **change non-monotonously** for the bin width smaller than **5%** in the range from  $-1$  to  $+1$ . This effect **should be studied**, especially **for higher moments**.
- $S \sigma$  and  $\kappa \sigma^2$  depend on the lower order fluctuations, which give the largest contribution to their values. The fluctuations from a source that one would like to access,  $S_A \sigma_A$  and  $\kappa_A \sigma_A^2$ , **are almost zero** in the considered example:

$$S \sigma \simeq 3 \langle n_A \rangle \omega_P > 0, \quad \kappa \sigma^2 \simeq 3 \langle n_A \rangle \omega_A \omega_P > 0, \quad S_A \sigma_A \simeq 0, \quad \kappa_A \sigma_A^2 \simeq 0.$$

- The **CBWC** is **unable to remove participant fluctuations**. It gives the average of the 'total', i.e. non-processed fluctuations, which are dominated by participant fluctuations. Its results **depend** on the **width**, **weight**, and the **position** of the sub-bins, i.e. may give arbitrary result.
- **The same is true if one** does not mix the bins, but **selects a particular centrality bin** and increases statistics.

- A **fluctuation of participants** and a contribution from lower moments may **persist** in higher moments, even if the bin width approaches zero.
- Summary of the methods reducing participant fluctuations:
  - (i) **Narrowing centrality bin** gives  $\omega = \omega_A + X$ , while  $\omega_A$  is needed,
  - (ii) **CBWC** gives  $\langle \omega \rangle = \langle \omega_A \rangle + \langle X \rangle$ , while  $\omega_A$  is needed ,
  - (iii) **Strongly Intensive** quantities give  $\langle n_B \rangle \omega_A \pm \langle n_A \rangle \omega_B$ , while  $\omega_A$  or  $\omega_B$  is needed,
  - (iv) **MMCP gives  $\omega_A$  (!)**, if  $X/\omega_A$  is a small parameter.
- We introduce the way how to **quantify smallness** and **largeness** of fluctuations:

$$\alpha = \langle n_A \rangle \frac{\omega_P}{\omega_A}, \quad \beta = \langle n_A \rangle \frac{S_P \sigma_P}{\omega_A}, \quad \gamma = \langle n_A \rangle^2 \frac{\kappa_P \sigma_P^2}{\omega_A^2}, \quad \delta = \frac{S_A \sigma_A}{\omega_A}, \quad \varepsilon = \frac{\kappa_A \sigma_A^2}{\omega_A^2}.$$

- The **MMCP works well** for the **scaled variance**. Is it possible to obtain  $S_A \sigma_A$  and  $\kappa_A \sigma_A^2$  measuring  $\kappa_5$  and  $\kappa_6$ , and **further developing MMCP?**



# Thank you!

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