



NAMIBIA UNIVERSITY
OF SCIENCE AND TECHNOLOGY



UNAM
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African School of Fundamental
Physics and Applications

Fifth African School of Fundamental Physics and its Applications

Search and Discovery Statistics in HEP Lecture 4




Eilam Gross, Weizmann Institute of Science

This presentation would have not been possible without the tremendous
help of
the following people throughout many years


Louis Lyons, Alex Read, Bob Cousins, Glen Cowan, Kyle Cranmer,
Ofer Vitells & Jonathan Shlomi



What can you expect from the Lectures

-  Lecture 1-2: Basic Concepts
Histograms, PDF, Testing Hypotheses,
LR as a Test Statistics, p-value, POWER, CLs
Measurements
-  Lecture 3: Wald Theorem, Asymptotic Formalism, Asimov Data Set, Feldman-Cousins, PL & CLs
-  Lecture 4: Asimov Significance
Look Elsewhere Effect
1D LEE the non-intuitive thumb rule
(upcrossings, trial #~Z)
2D LEE (Euler Characteristic)

Look Elsewhere Effect

A group of people are sitting on a wooden pier, looking out at the ocean during a sunset. The sky is a mix of orange, yellow, and pink, and the water is dark blue. The pier is made of wooden planks and has some graffiti on the concrete base. The people are silhouetted against the bright sky.

E.G., O. Vitells “Trial factors for the look elsewhere effect in high energy physics”,

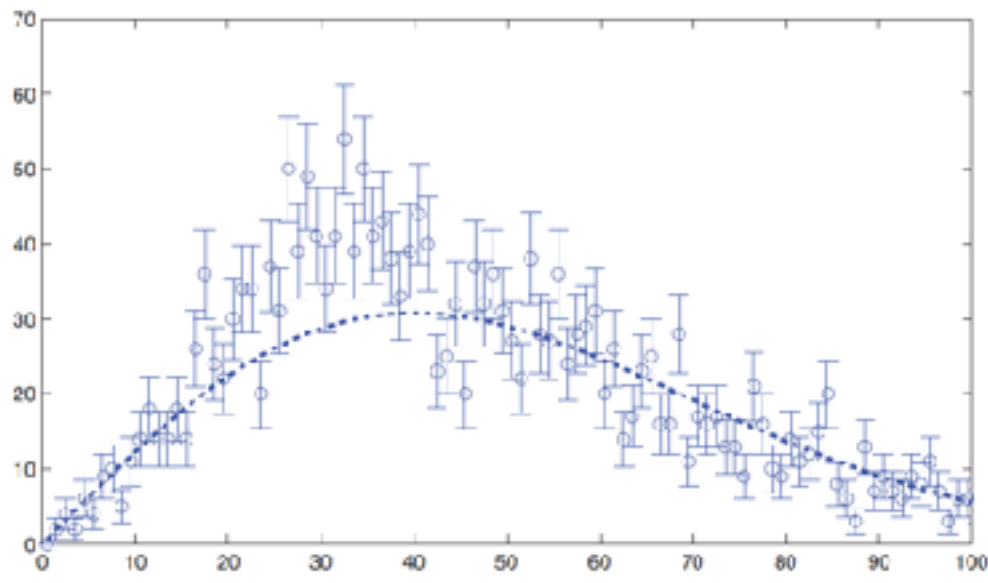
Eur. Phys. J. C 70 (2010) 525

O. Vitells and E. G., Estimating the significance of a signal in a multi-dimensional search,

1669 Astropart. Phys. 35 (2011) 230, arXiv:1105.4355

Look Elsewhere Effect

- Is there a signal here?

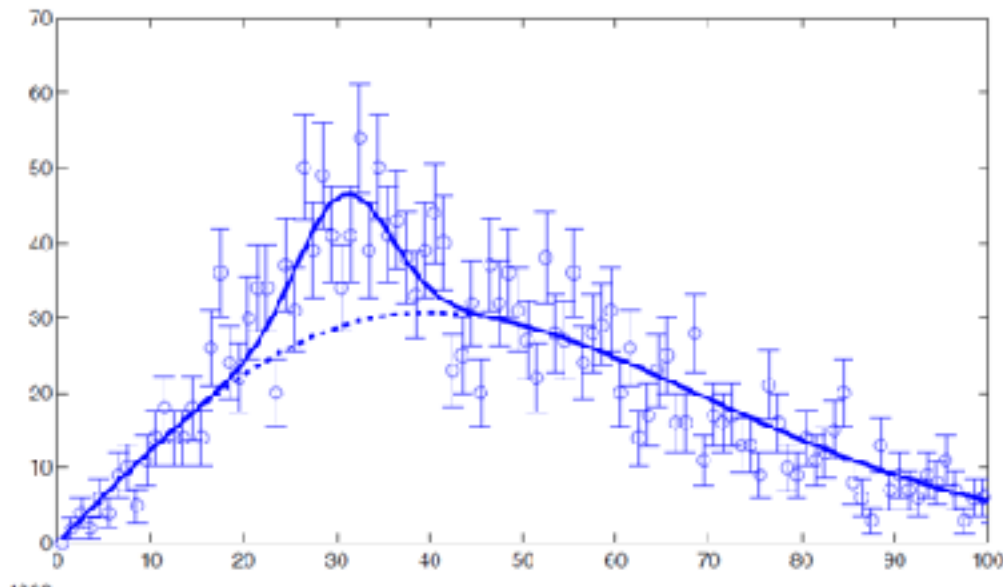


Look Elsewhere Effect

- Looks like a signal at $m=30$
- What is its significance?

Test the BG hypothesis
At $m=30$

$$q_0(\theta) = \begin{cases} -2 \log \frac{L(\mu = 0)}{L(\hat{\mu}, \theta)} \\ 0 \end{cases}$$



$$q_{fix,obs} = -2 \ln \frac{L(b)}{L(\hat{\mu}_s(m=30) + b)}$$

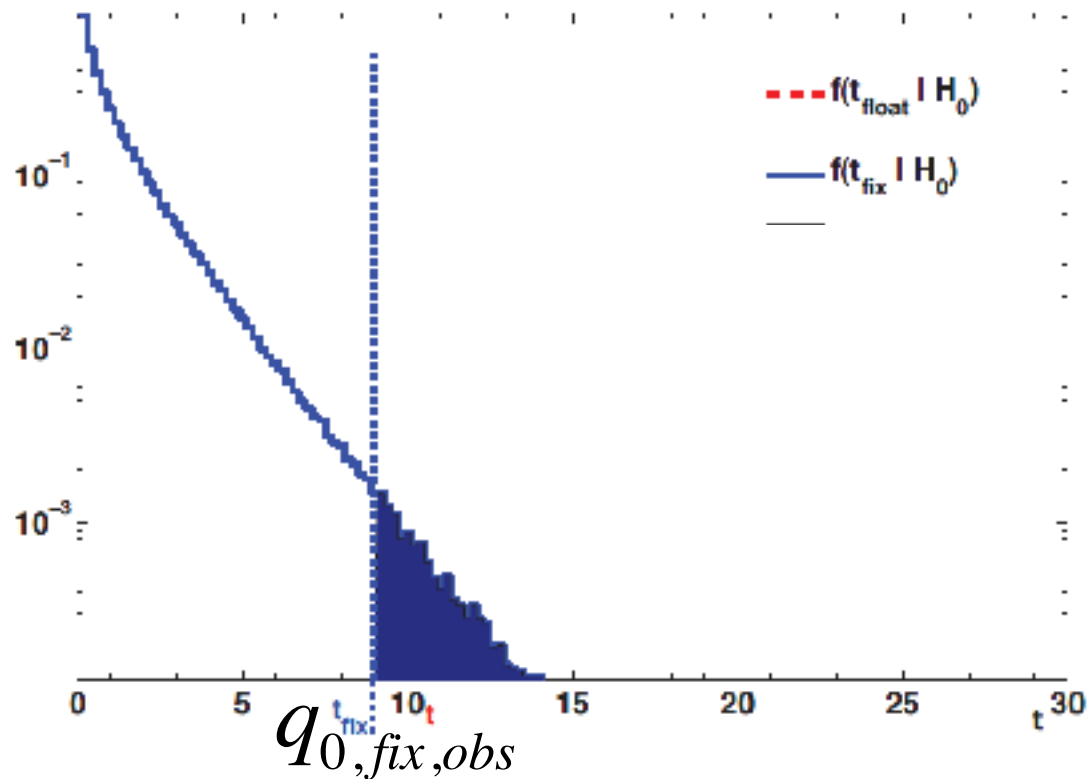
$$Z = \sqrt{q_{0,fix,obs}}$$

Look Elsewhere Effect

$$q_{0,fix} = -2 \ln \frac{L(\mu = 0)}{L(\hat{\mu}_s(30) + b)}$$

$$f(q_{0,fix} | H_0) \sim \chi^2$$

$$p_{fix} = \int_{q_{fix,obs}}^{\infty} f(q_0 | H_0) dq_0$$

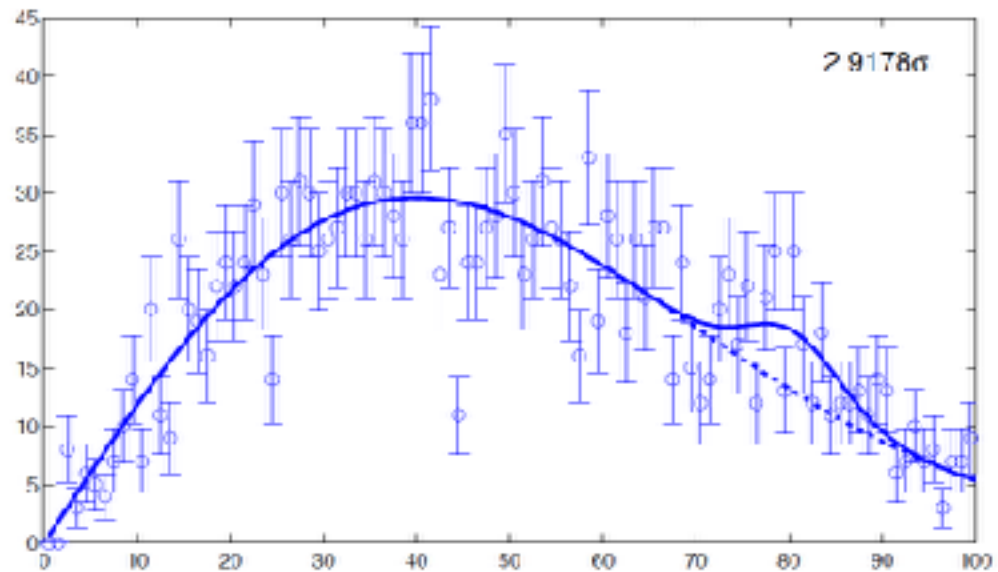


p_{fix} answers the question :

What is the probability to have a fluctuation as or bigger than the observed one?

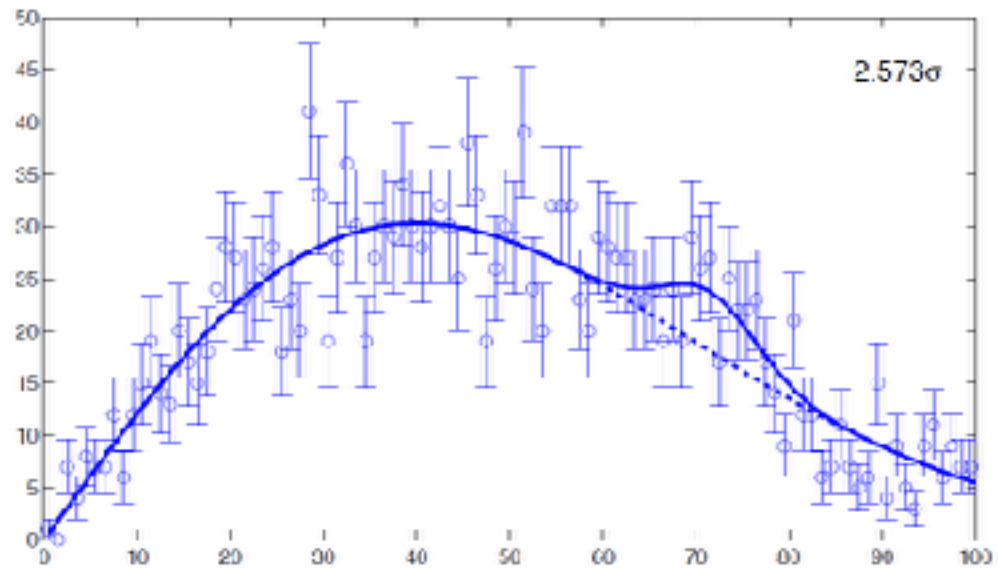
Look Elsewhere Effect

- Would you ignore this signal, had you seen it?



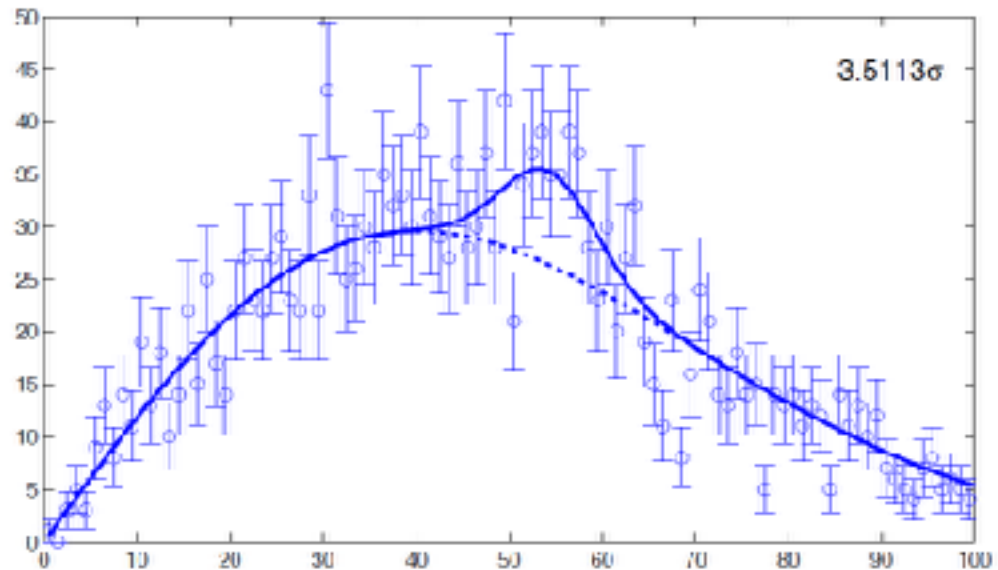
Look Elsewhere Effect

- Or this?



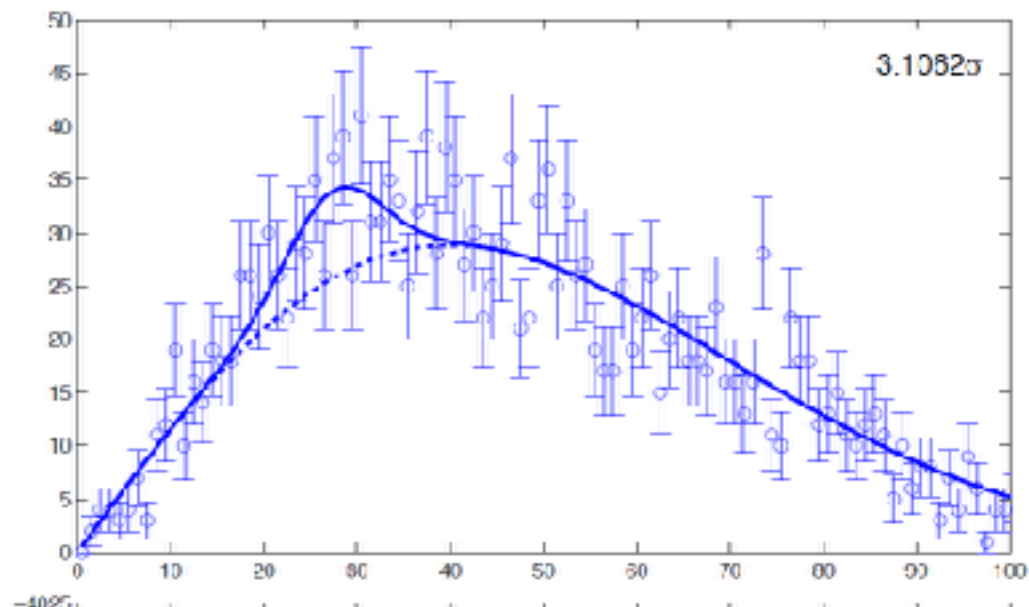
Look Elsewhere Effect

- Or this?



Look Elsewhere Effect

- Or this?
- Obviously NOT!
- ALL THESE "SIGNALS" ARE BG FLUCTUATIONS



The right question :

What is the probability to have a fluctuation as or bigger than the observed one

***ANYWHERE** in the mass search range?*

Look Elsewhere Effect

- Having no idea where the signal might be there are two equivalent options

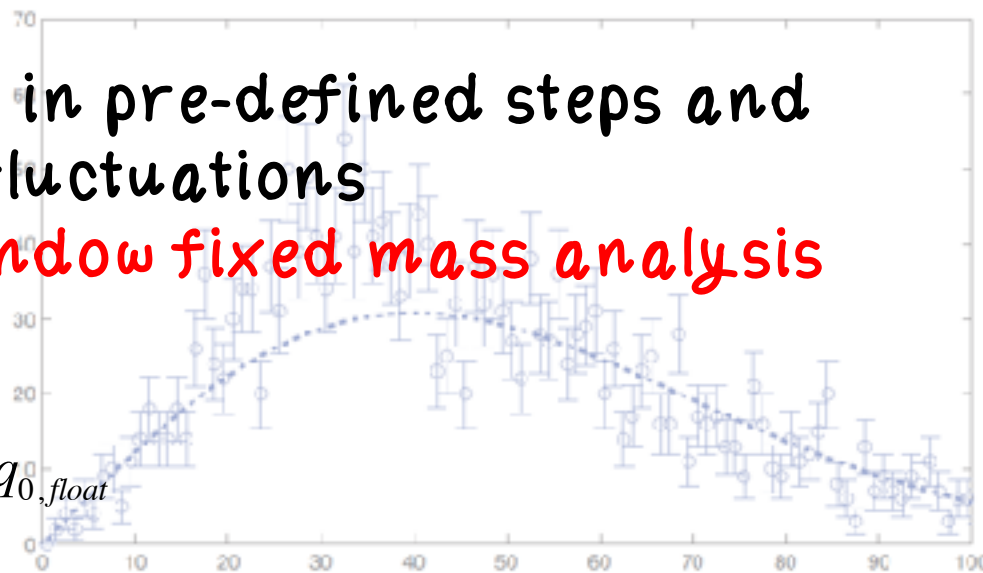
- **OPTION I:**

scan the mass range in pre-defined steps and test any disturbing fluctuations

Perform a sliding window fixed mass analysis

$$q_{0, \text{float}} = \max_m (q_0(m))$$

$$P_{\text{float}} = \int_{q_{\text{float, obs}}}^{\infty} f(q_{0, \text{float}} | H_0) dq_{0, \text{float}}$$



- **OPTION II:**

Perform a floating mass analysis

$$q_{0, \text{float}} = q_0(\hat{m}) = -2 \ln \frac{L(b)}{L(\hat{\mu}_s(\hat{m}) + b)}$$

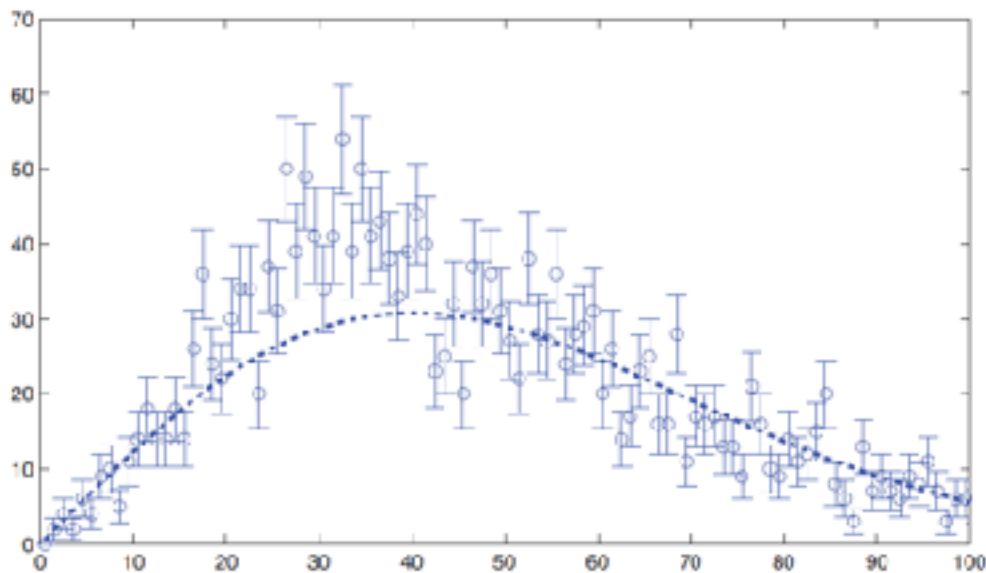
$$P_{\text{float}} = \int_{q_{\text{float, obs}}}^{\infty} f(q_{0, \text{float}} | H_0) dq_{0, \text{float}}$$



Sliding Window

- Scan and perform a fixed mass analysis at each point

$$q_0 = -2 \ln \frac{L(\mu = 0)}{L(\hat{\mu}_s(m) + b)}$$



- The scan resolution must be less than the signal mass resolution

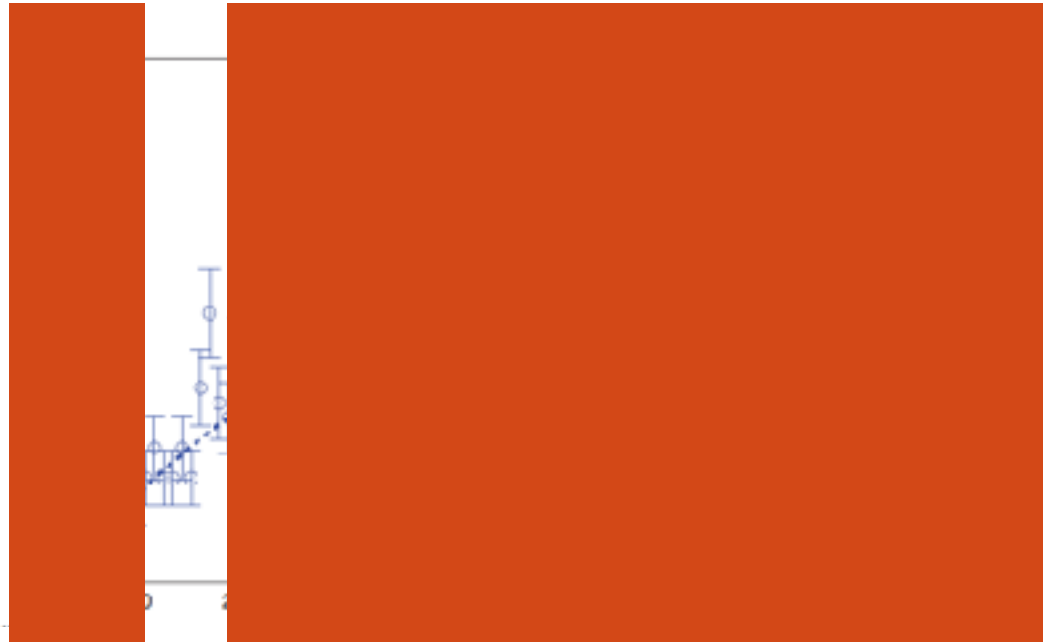
Sliding Window

$$q_0 = -2\ln \frac{L(\mu = 0)}{L(\hat{\mu}_s(m) + b)}$$



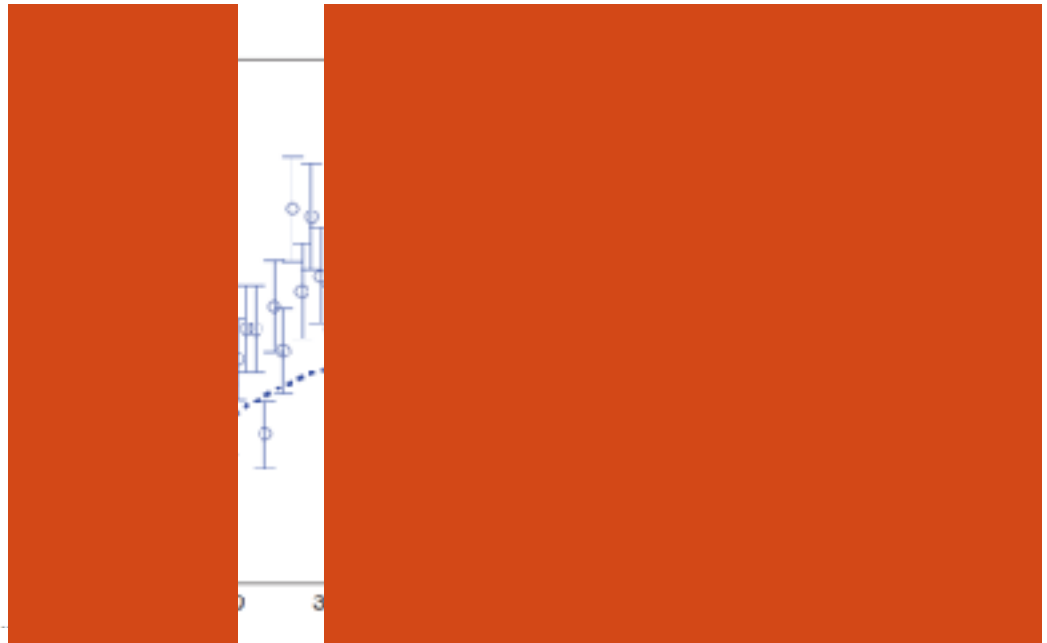
Sliding Window

$$q_0 = -2\ln \frac{L(\mu = 0)}{L(\hat{\mu}_s(m) + b)}$$



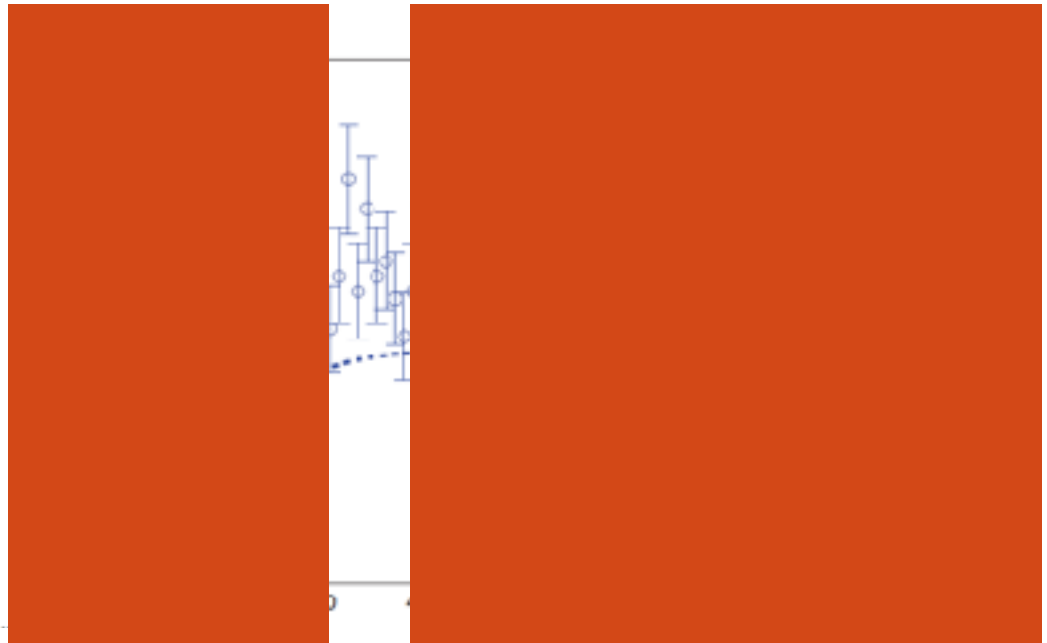
Sliding Window

$$q_0 = -2\ln \frac{L(\mu = 0)}{L(\hat{\mu}_s(m) + b)}$$



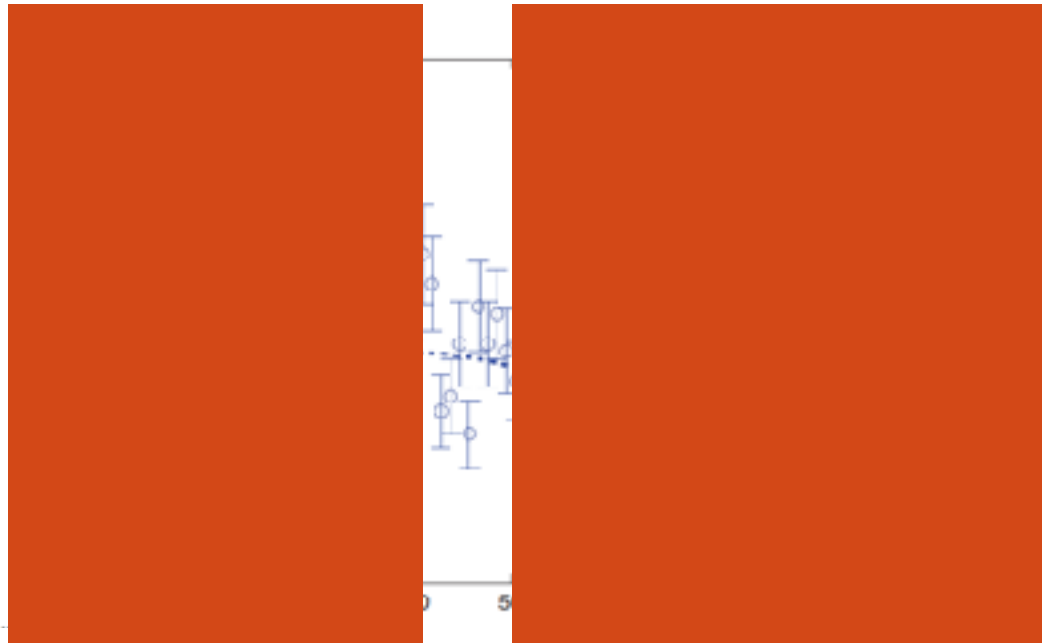
Sliding Window

$$q_0 = -2\ln \frac{L(\mu = 0)}{L(\hat{\mu}_s(m) + b)}$$



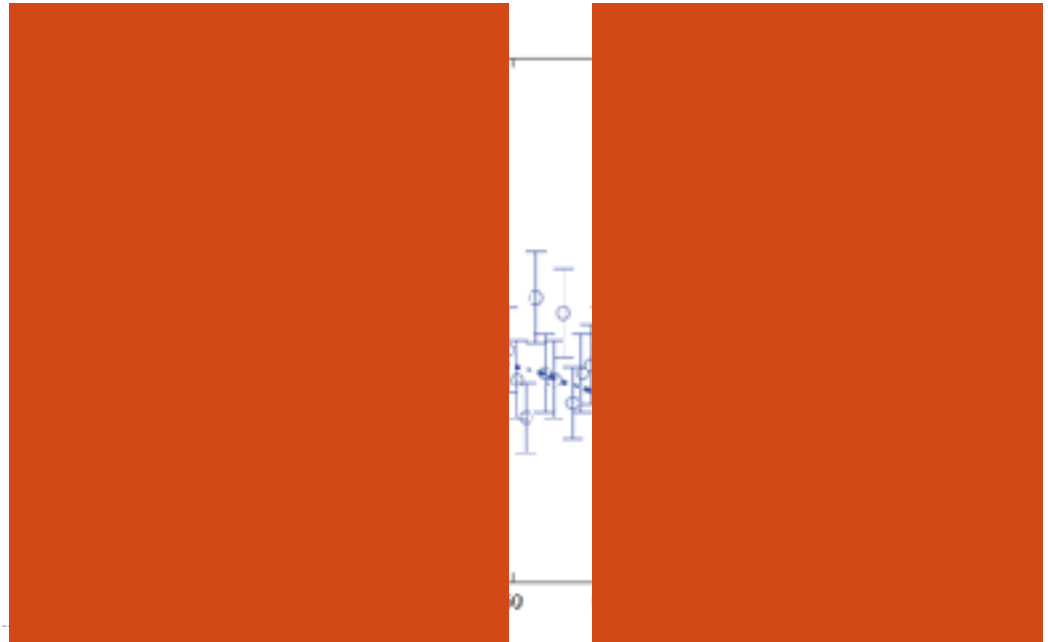
Sliding Window

$$q_0 = -2\ln \frac{L(\mu = 0)}{L(\hat{\mu}_s(m) + b)}$$



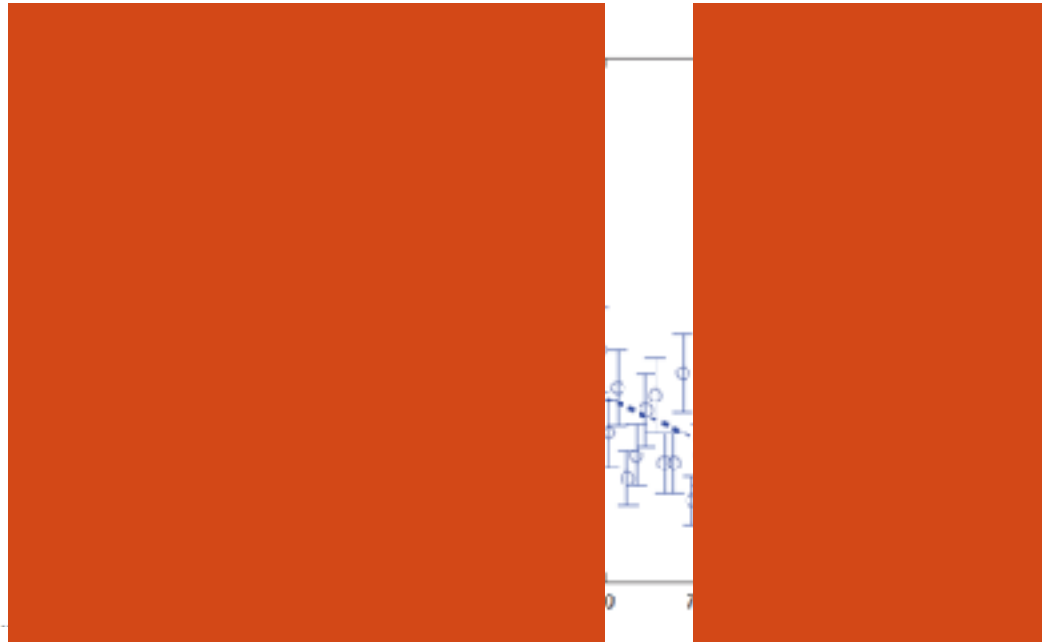
Sliding Window

$$q_0 = -2\ln \frac{L(\mu = 0)}{L(\hat{\mu}_s(m) + b)}$$



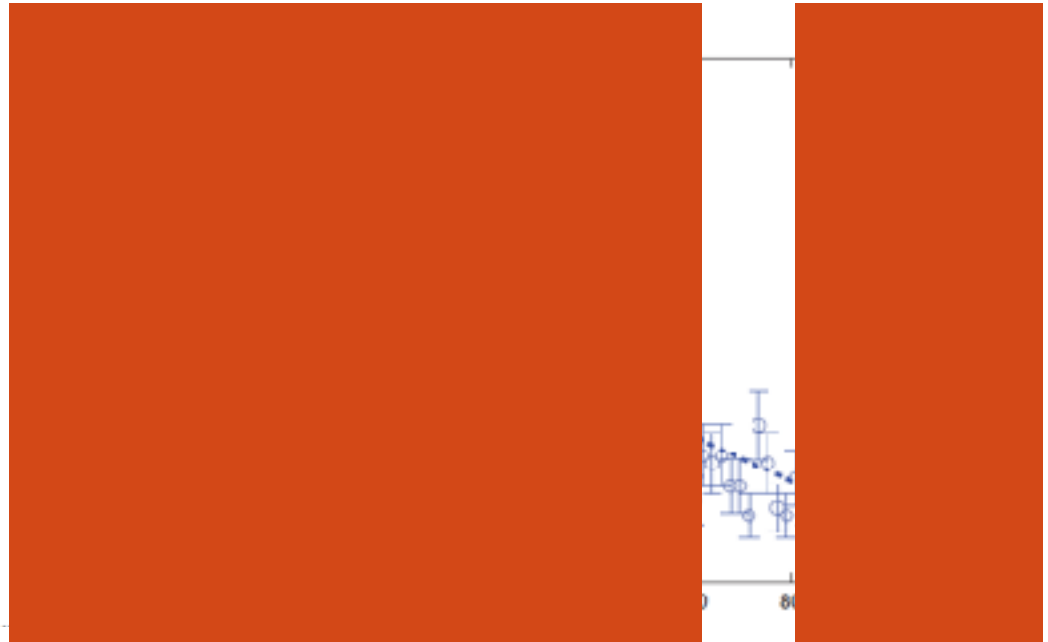
Sliding Window

$$q_0 = -2 \ln \frac{L(\mu = 0)}{L(\hat{\mu}_s(m) + b)}$$



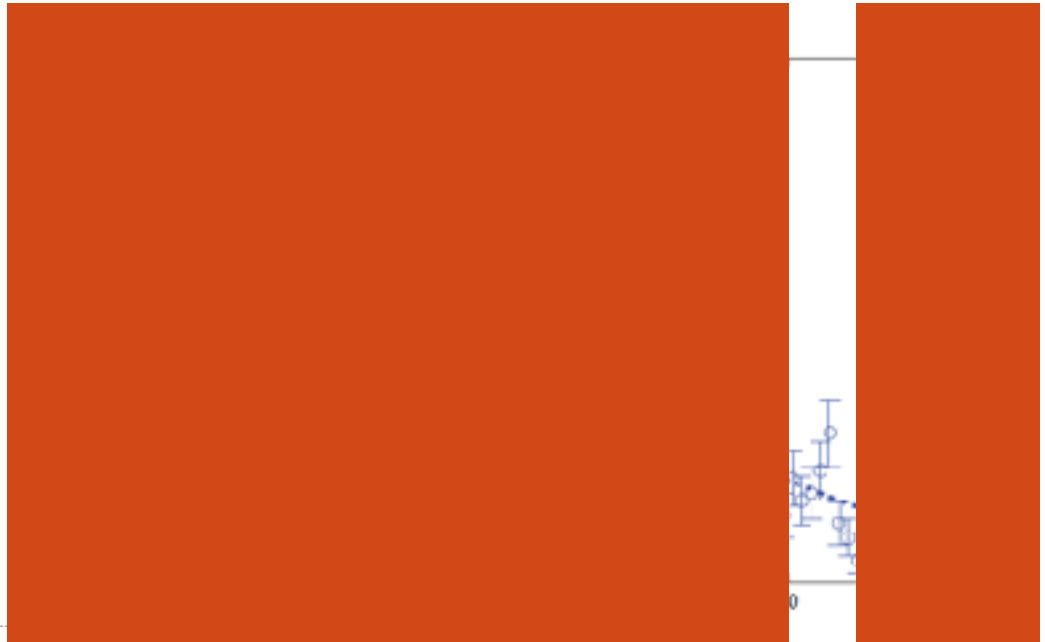
Sliding Window

$$q_0 = -2\ln \frac{L(\mu = 0)}{L(\hat{\mu}_s(m) + b)}$$



Sliding Window

$$q_0 = -2\ln \frac{L(\mu = 0)}{L(\hat{\mu}_s(m) + b)}$$



Sliding Window

- Assuming the signal can be only at one place
- pick the one with the **MAXIMUM SIGNIFICANCE**



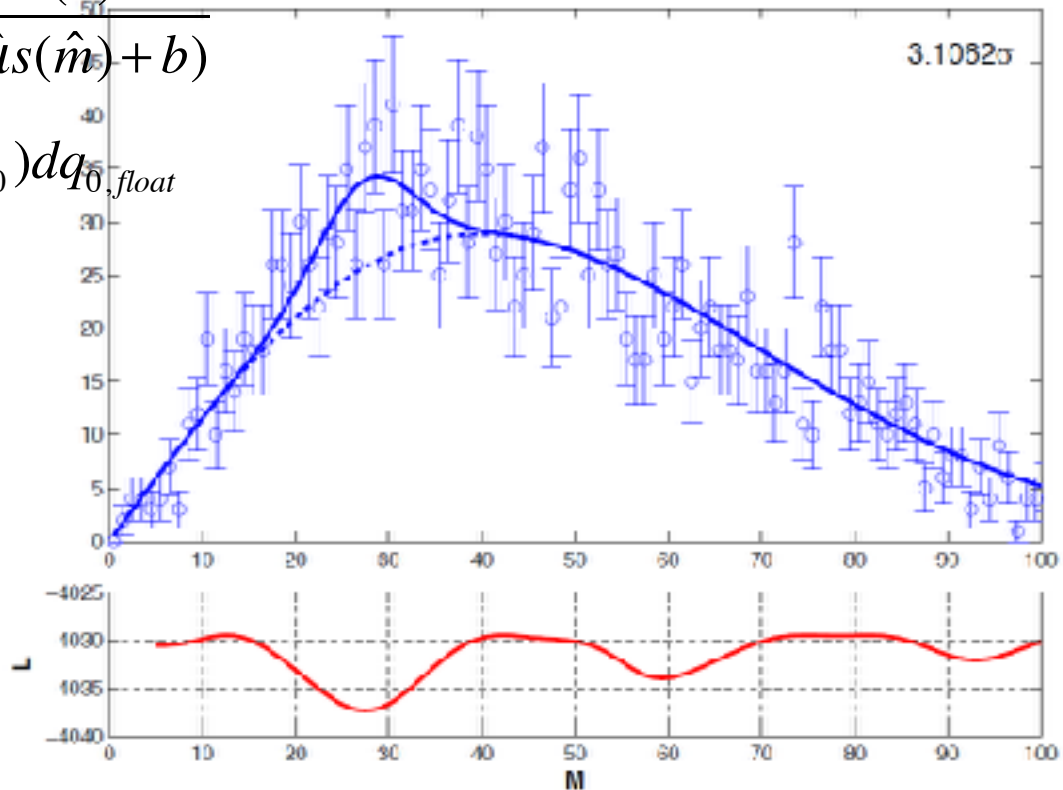
$$q_{0, float} = \max_m (q_0(m))$$

Look Elsewhere Effect: Floating Mass

OPTION II

$$q_{0, \text{float}} = q_0(\hat{m}) = -2 \ln \frac{L(b)}{L(\hat{\mu}_s(\hat{m}) + b)}$$

$$P_{\text{float}} = \int_{q_{\text{float, obs}}}^{\infty} f(q_{0, \text{float}} | H_0) dq_{0, \text{float}}$$

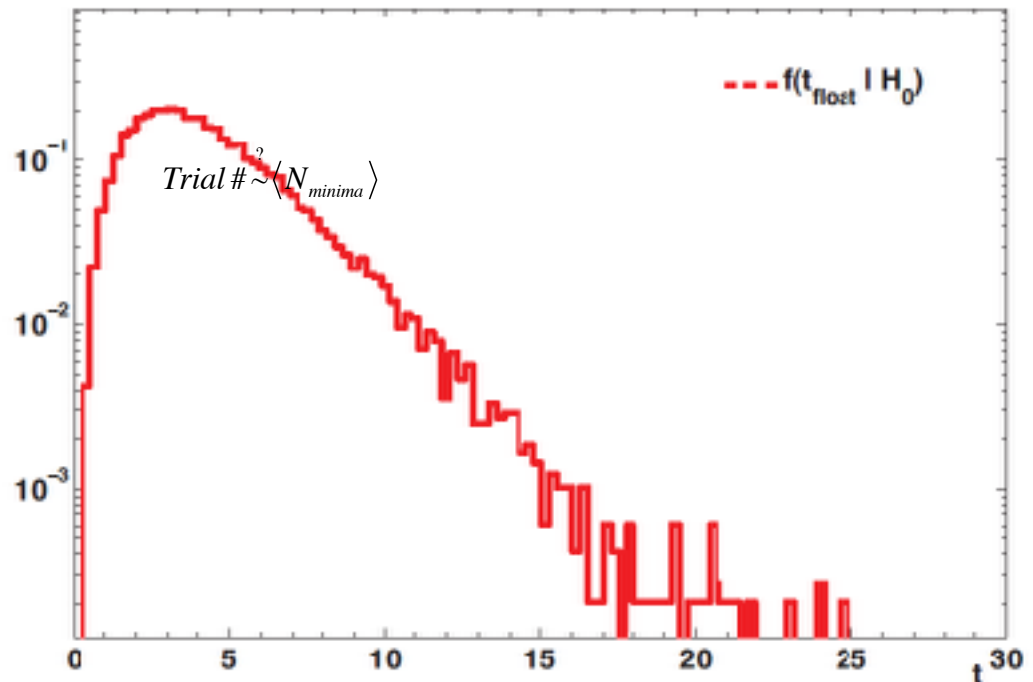


Look Elsewhere Effect

- The distribution $f(q_{\text{float}}|H_0)$ does not follow a chi-squared with 2 dof because the mass parameter is not defined under the null hypothesis

$$\exists m_{\text{fix}} \quad q_0(\hat{m}) \geq q_0(m_{\text{fix}})$$

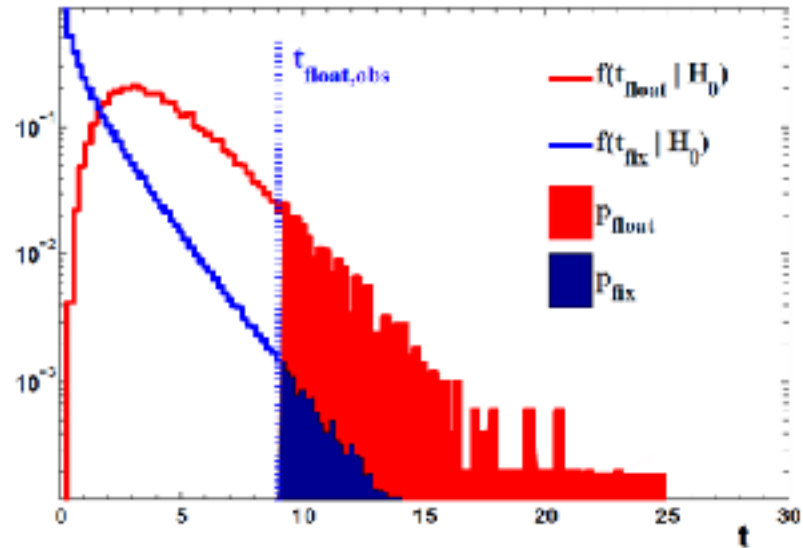
The χ^2_1 distribution is pushed to the right



trial#

- Assume a maximal local fluctuation at mass $\hat{m} = 30$
- The observed q_0 is given by

$$q_{0,obs} = -2 \ln \frac{L(\mu = 0)}{L(\hat{\mu}_s(m) + b)}$$



$$p_{fix} = \int_{q_{0,obs}}^{\infty} f(q_{0,fix} | H_0) dq_{0,fix}$$

$$p_{float} = \int_{q_{0,obs}}^{\infty} f(q_{0,float} | H_0) dq_{0,float}$$

$$trial \# = \frac{p_{float}}{p_{fix}}$$

Can we calculate analytically the floating mass p-value

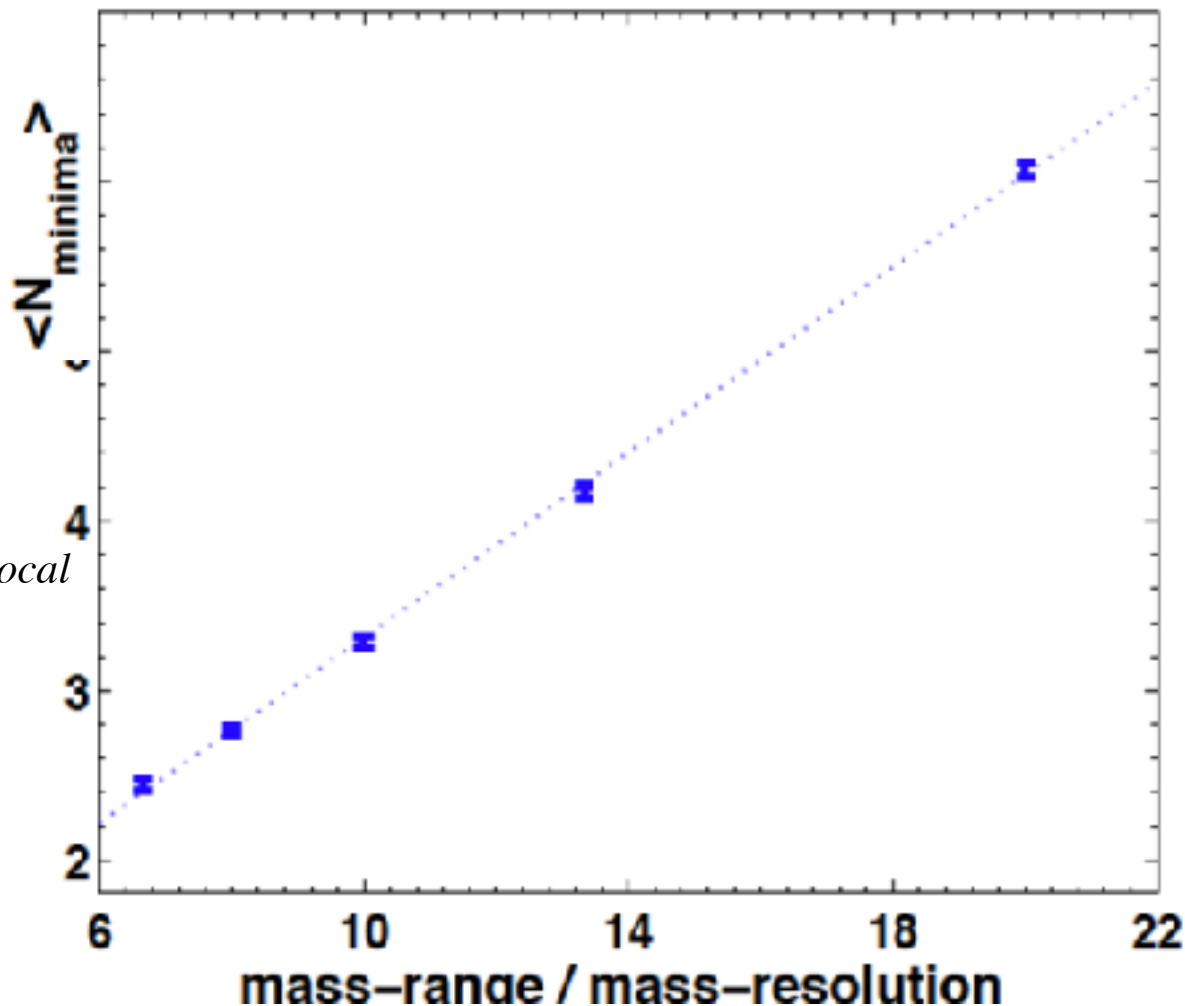
(Wrong) Thumb Rule

$$\langle N_{\text{minima}} \rangle \sim \frac{\text{Mass Range}}{\text{Mass Resolution}}$$

$$\text{Trial \#} \sim \langle N_{\text{minima}} \rangle$$

$$\text{Trial \#} \stackrel{?}{=} \langle N_{\text{minima}} \rangle P_{\text{local}}$$

The answer is NO



The right question :

*What is the probability to have a fluctuation
as or bigger than the observed one*

***ANYWHERE** in the mass search range?*

*Let θ be a nuisance parameter
undefined under the null hypothesis.*

Define $q(\hat{\theta}) = \max_{\theta} (q(\theta))$

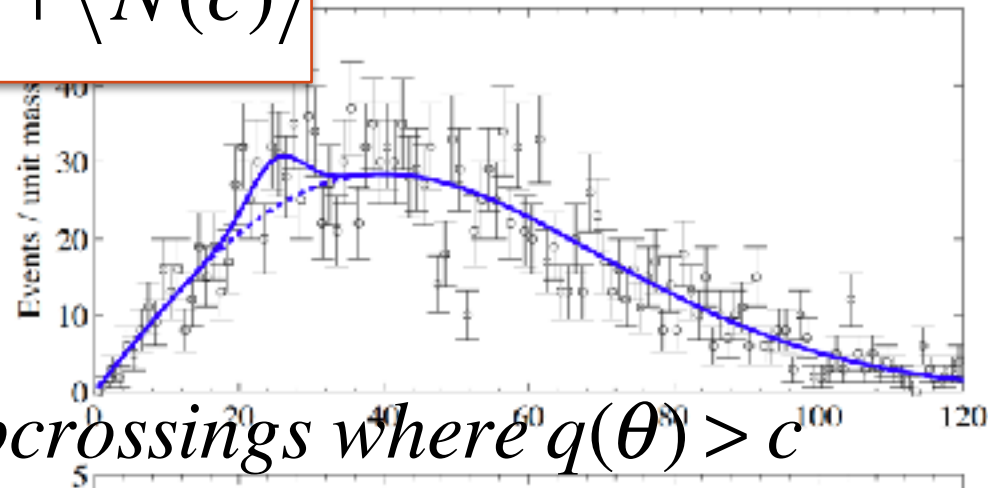
Davies (1987) finds, for $c \gg 1$

$P(q(\hat{\theta}) > c) \sim P(\chi_1^2 > c) + \langle N(c) \rangle$

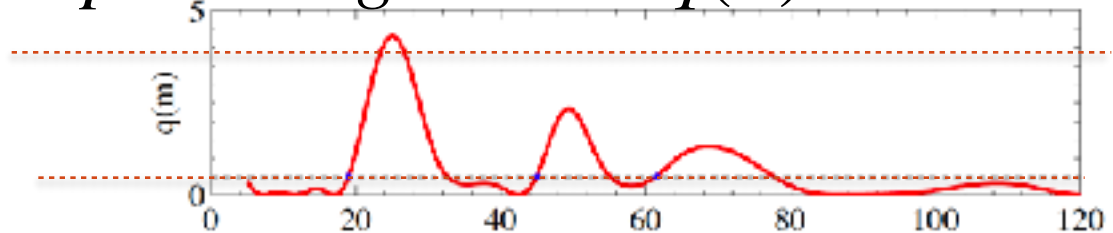
$\langle N(c) \rangle =$ Number of
upcrossings $q(\theta) > c$

Davies Formula

$$P(q(\hat{\theta}) > c) \sim P(\chi_1^2 > c) + \langle N(c) \rangle$$

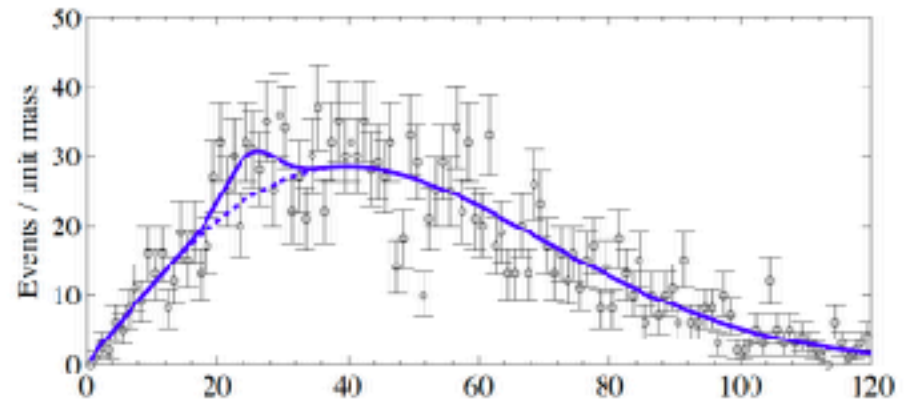


$\langle N(c) \rangle =$ Number of upcrossings where $q(\theta) > c$



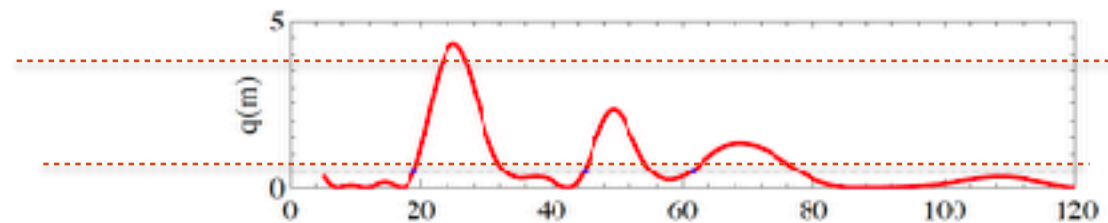
or $c \gg 1 \rightarrow \langle N(c) \rangle \ll 1$

Making Davies Formula Accessible



$$\langle N(c) \rangle \ll 1$$

$$\langle N(c) \rangle \sim e^{-c/2}$$



$$P(q(\hat{\theta}) > c) \sim P(\chi_1^2 > c) + \langle N(c_0) \rangle \frac{\langle N(c) \rangle}{\langle N(c_0) \rangle}$$

$$P(q(\hat{\theta}) > c) \sim P(\chi_1^2 > c) + \langle N(c_0) \rangle e^{-(c-c_0)/2}$$

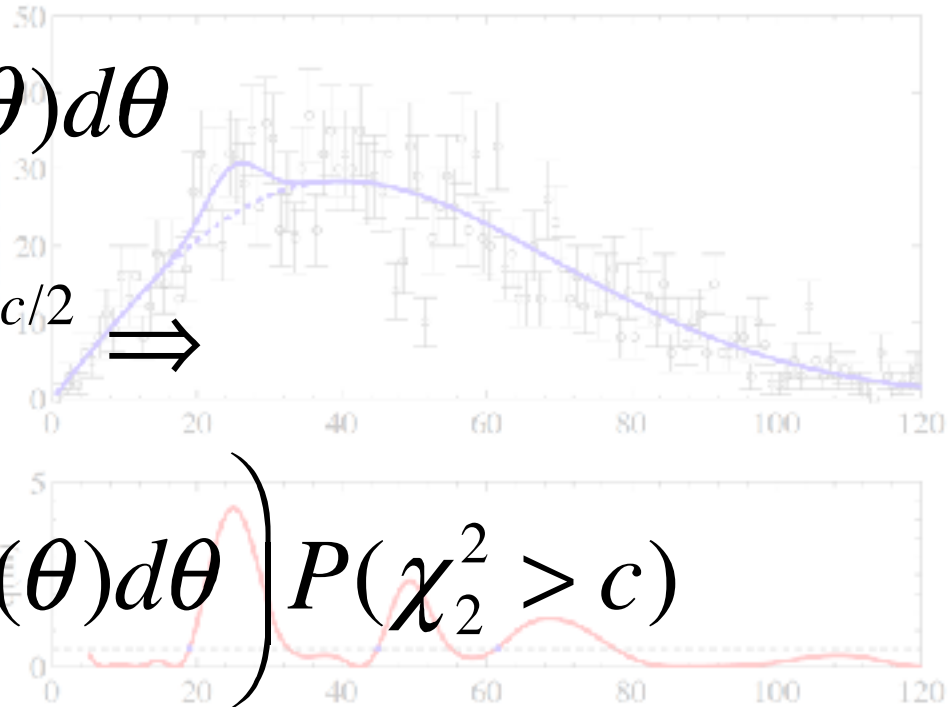
Davies Formula

$$\langle N(c) \rangle \approx \frac{e^{-c/2}}{\sqrt{2\pi}} \int_{\theta} C(\theta) d\theta$$

$$P(\chi_2^2 > c) \xrightarrow{c \rightarrow \infty} e^{-c/2} \Rightarrow$$

$$\langle N(c) \rangle \approx \left(\frac{1}{\sqrt{2\pi}} \int_{\theta} C(\theta) d\theta \right) P(\chi_2^2 > c)$$

$$\langle N(c) \rangle = \mathcal{N} P(\chi_2^2 > c)$$



$$P(q(\hat{\theta}) > c) \sim P(\chi_1^2 > c) + \mathcal{N} P(\chi_2^2 > c)$$

Trial

$$P(\chi_1^2 > c) \xrightarrow{c \gg 1} \sqrt{\frac{2}{c}} \frac{e^{-c/2}}{\Gamma\left(\frac{1}{2}\right)}$$

$$P(\chi_2^2 > c) \xrightarrow{c \gg 1} e^{-c/2}$$

$$\text{trial \#} = \frac{P(q(\hat{\theta}) > c)}{P(q(\theta) > c)} \approx$$

$$\approx 1 + \mathcal{N} \frac{P(\chi_2^2 > c)}{P(\chi_1^2 > c)} \Rightarrow$$

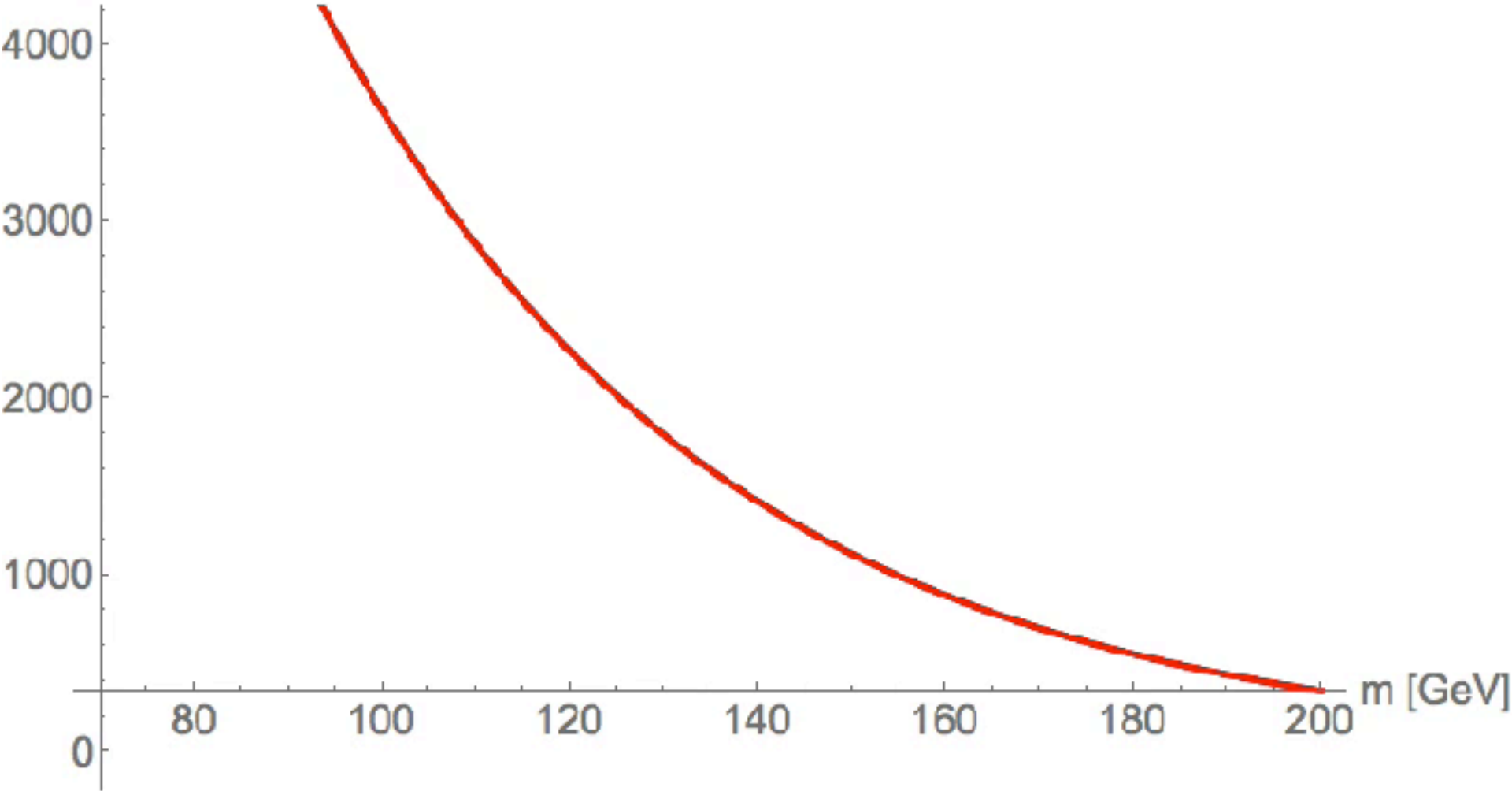
$$\text{trial \#} \approx 1 + \mathcal{N} \sqrt{\frac{c}{2}} \Gamma(1/2) \Rightarrow$$

$$\text{trial \#} \approx 1 + \sqrt{\frac{\pi}{2}} \mathcal{N} Z_{fix}$$

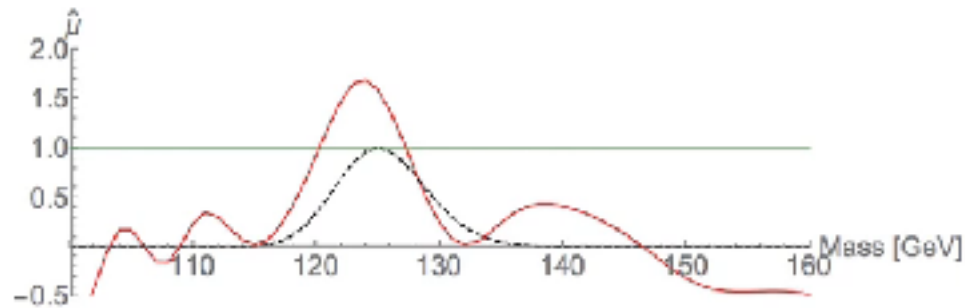
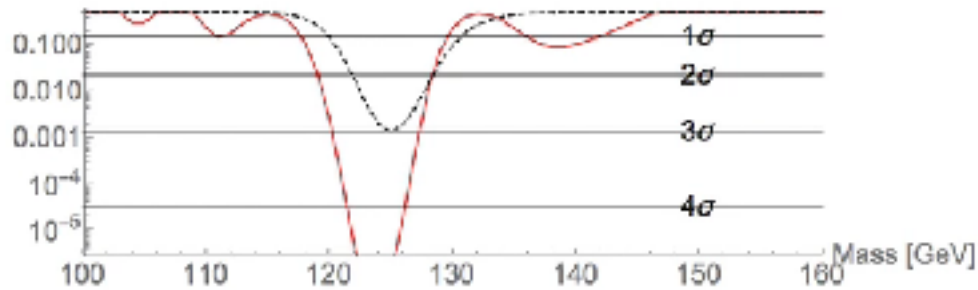
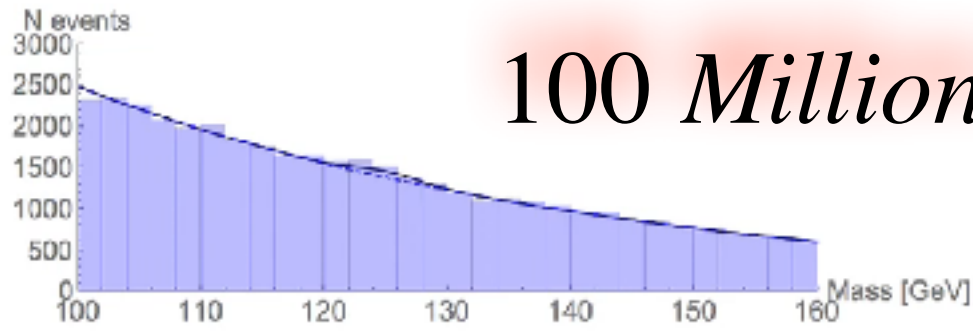
What is Going On?



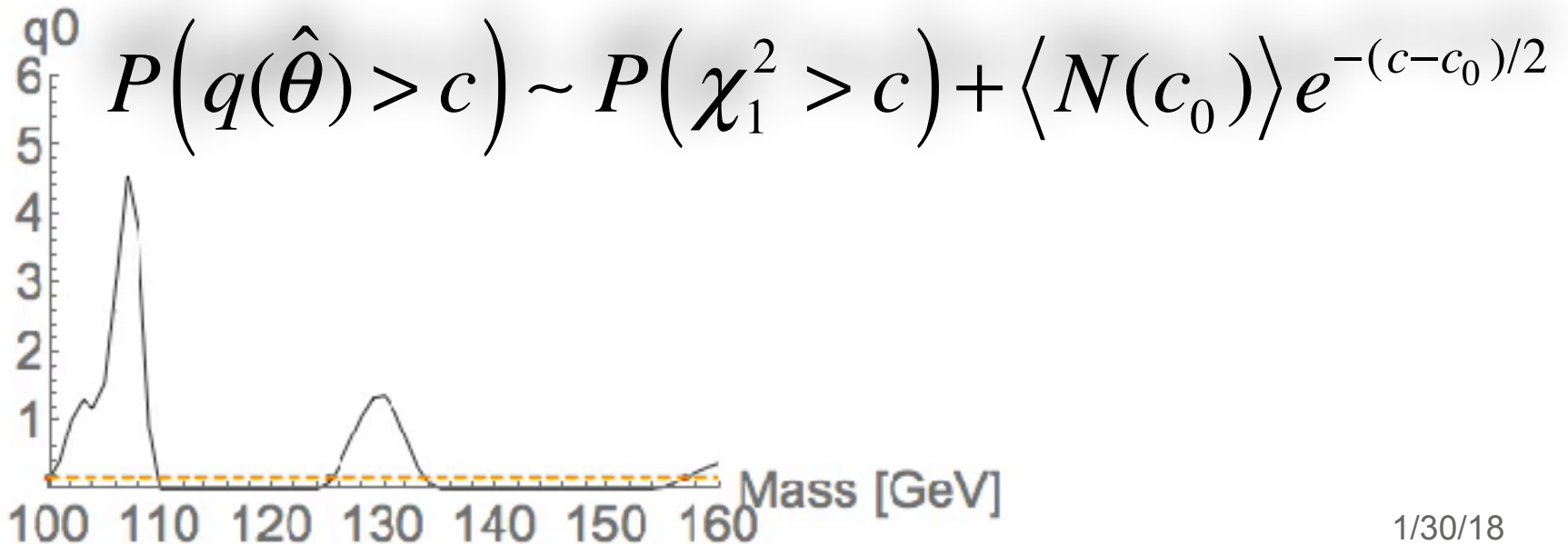
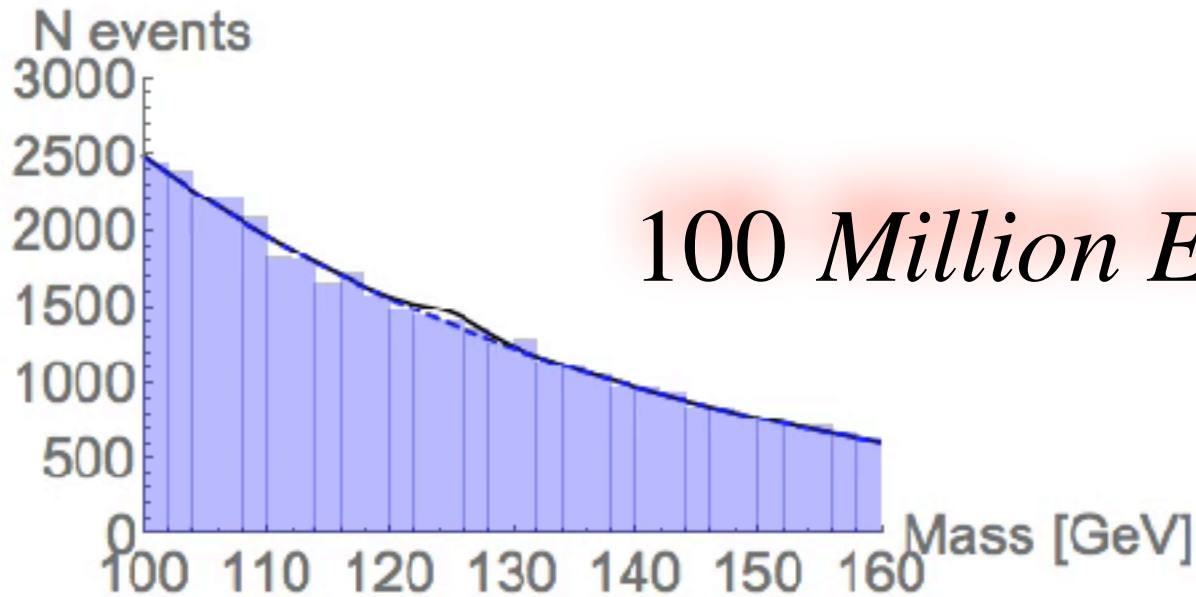
N events

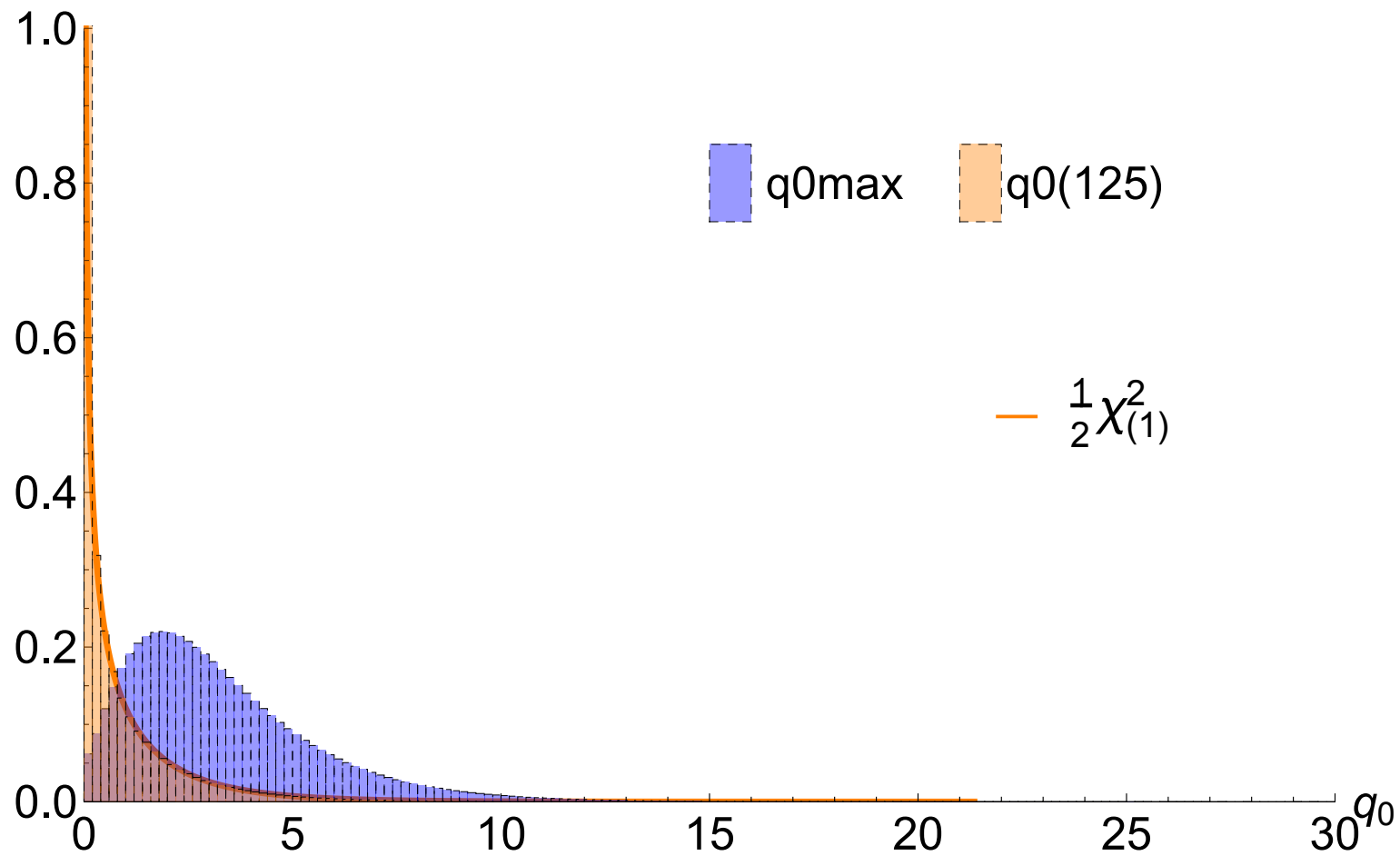


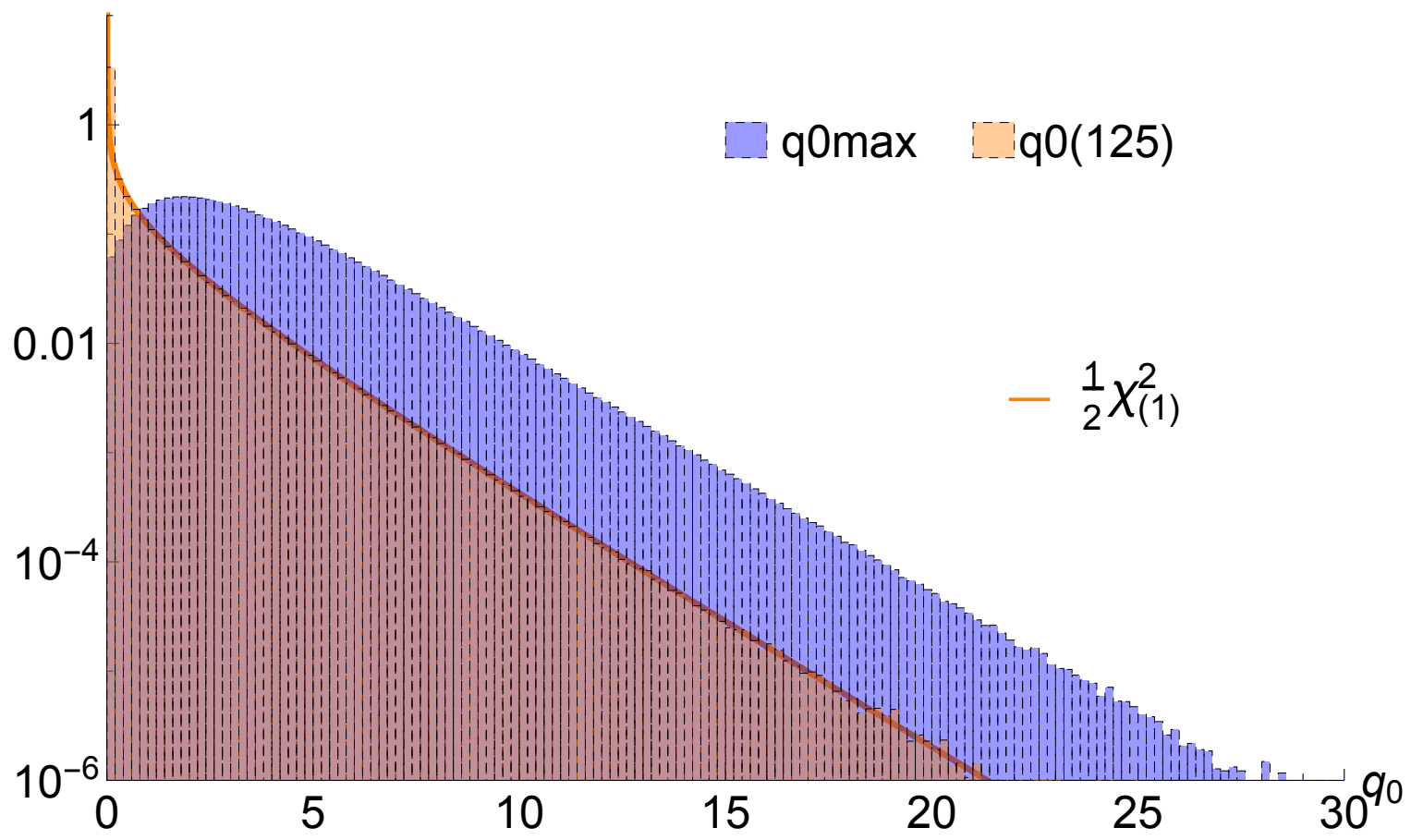
100 Million Experiments

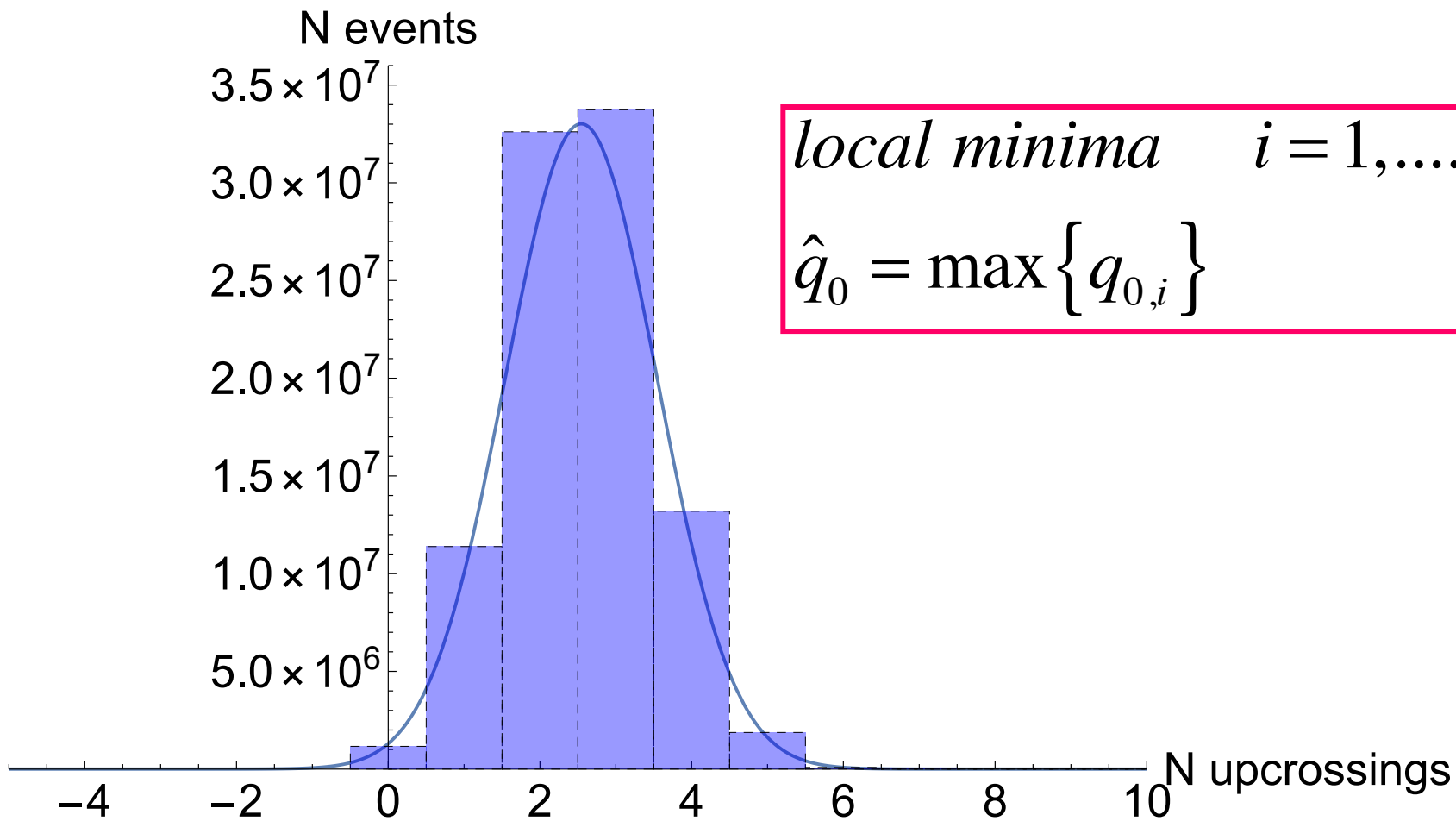


100 Million Experiments



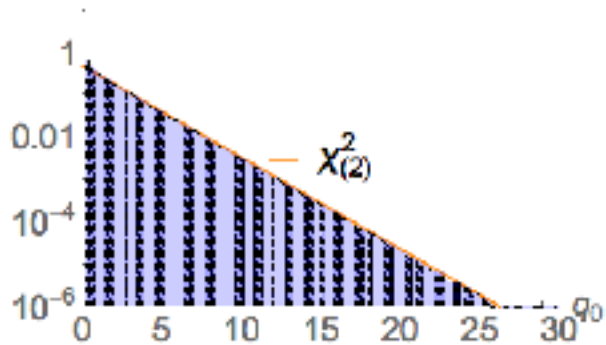




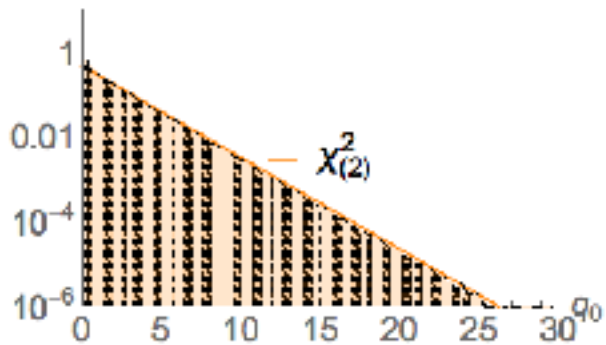


$$\forall i \quad q_{0,i} \sim \chi_2^2$$

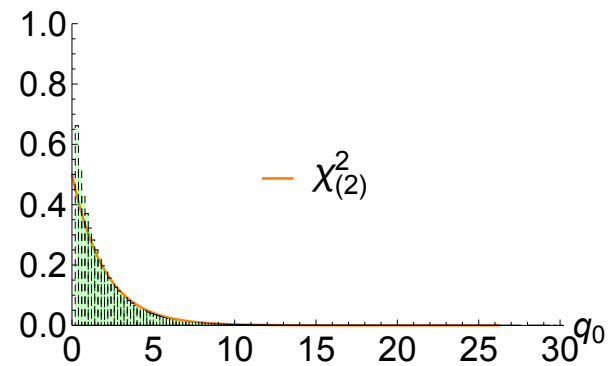
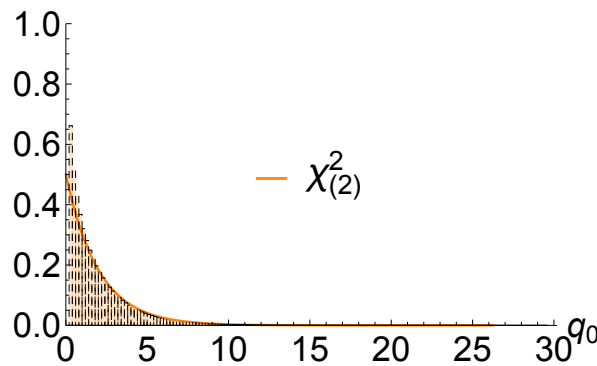
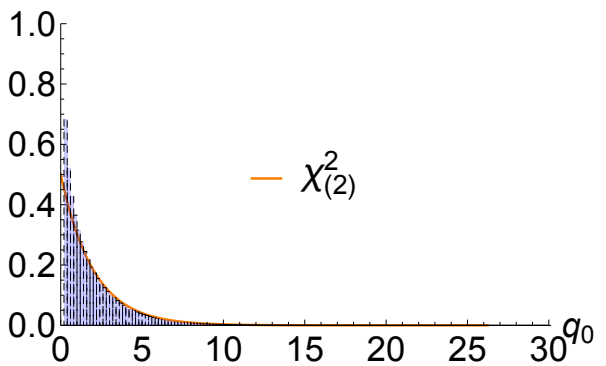
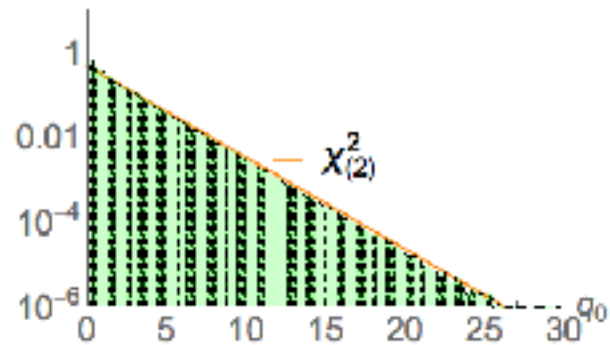
$q_{0,1}$

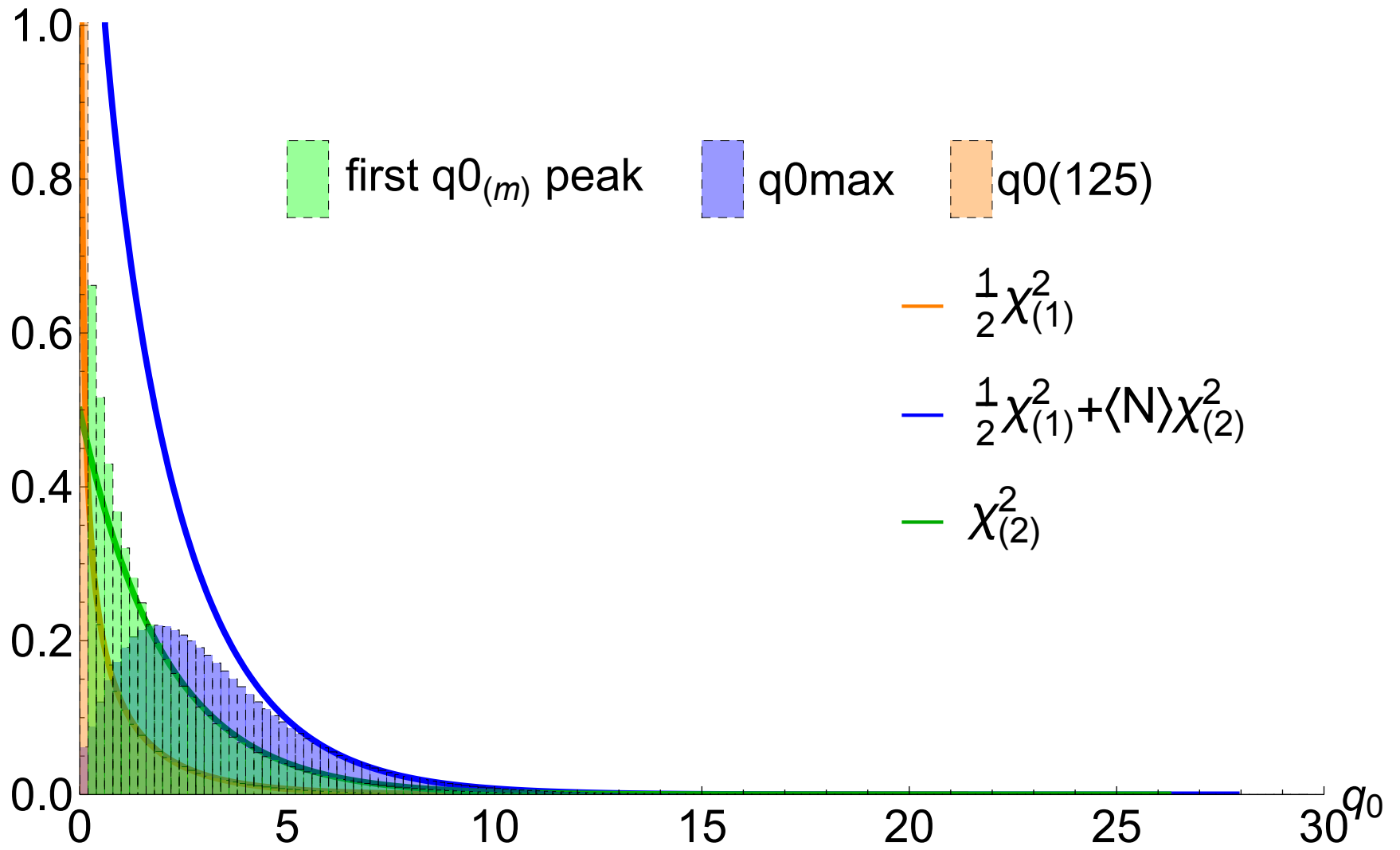


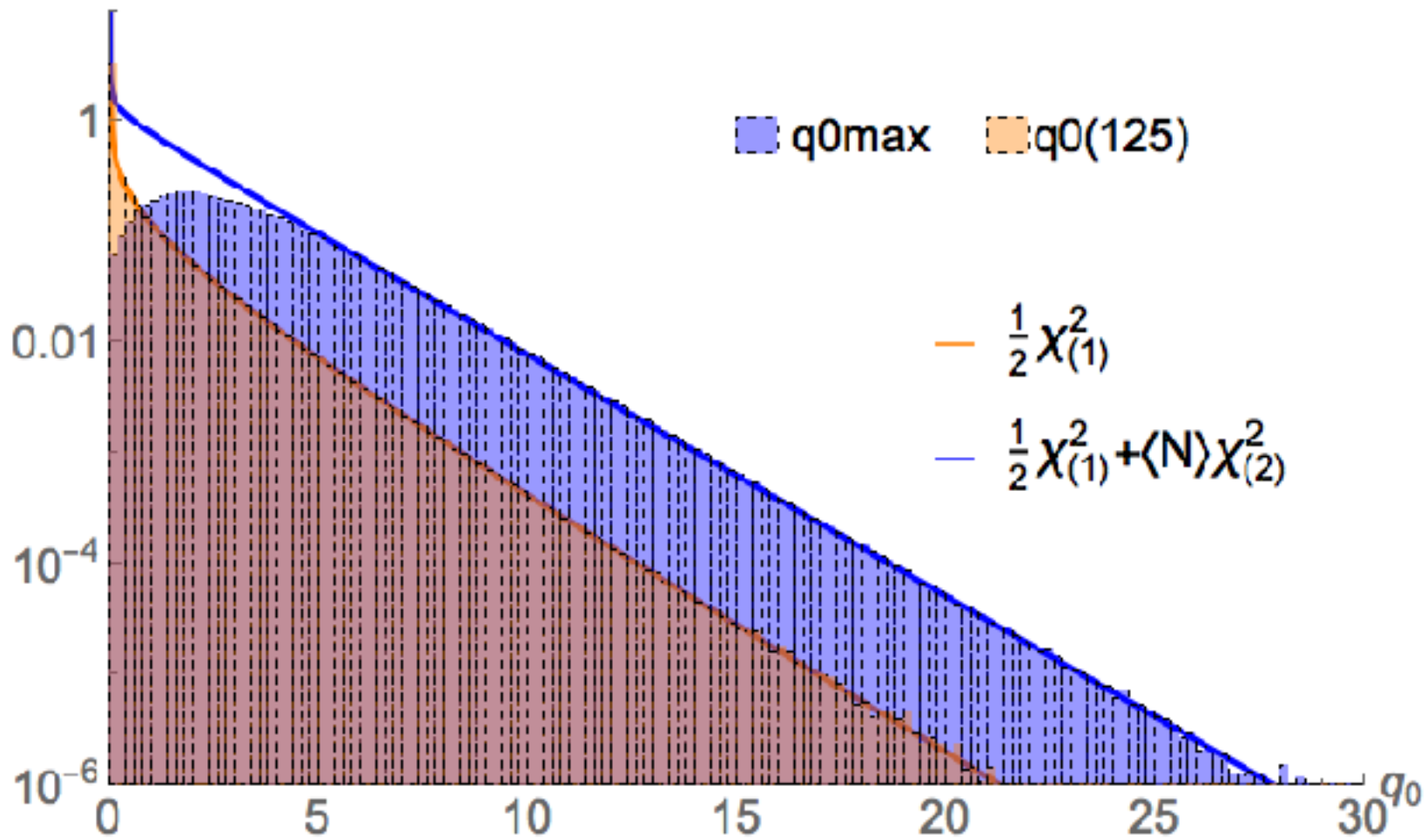
$q_{0,2}$



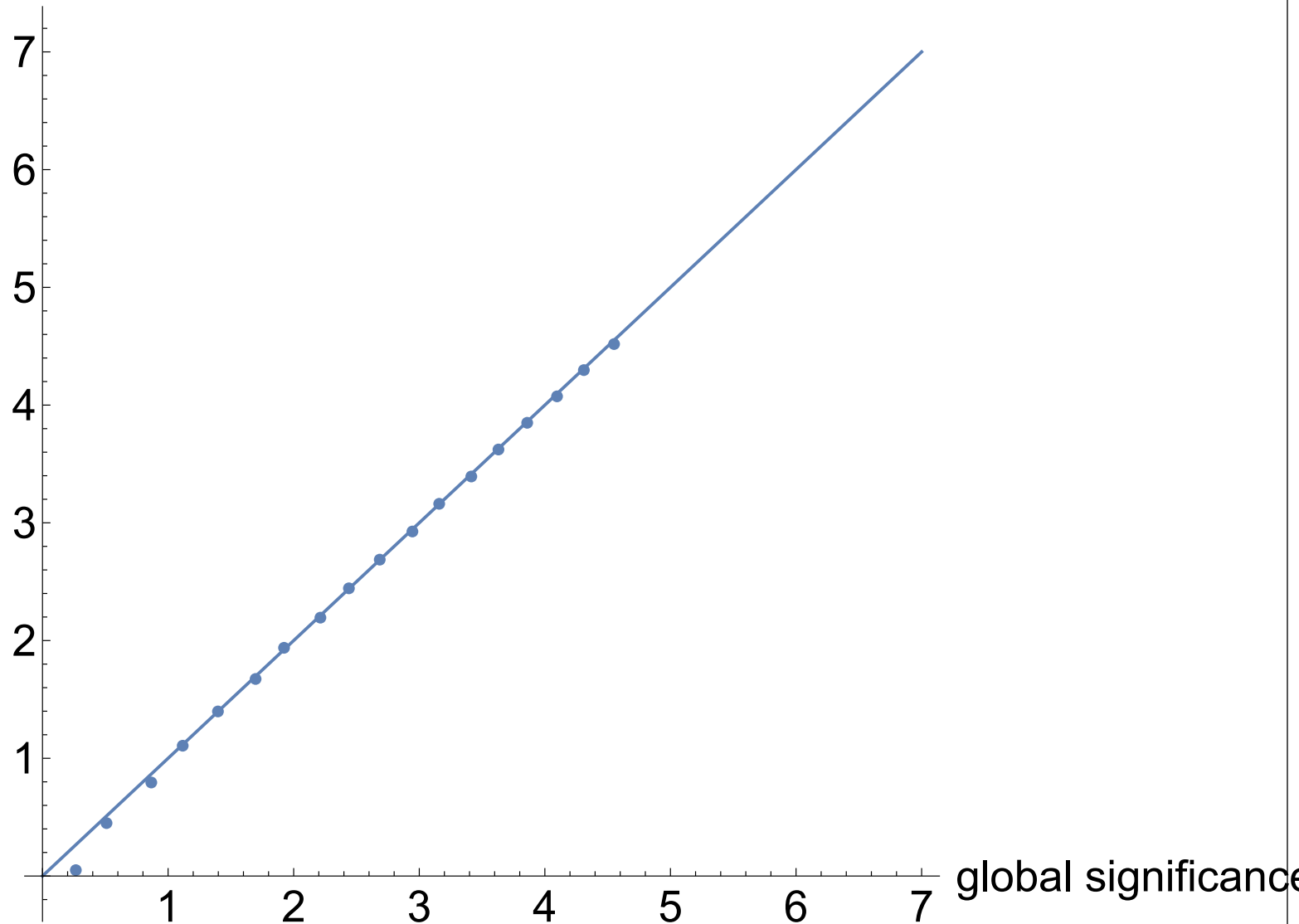
$q_{0,3}$







global significance formula



Trial Factor

40

30

20

10

0

— Formula
• Toys

0

1

2

3

4

5

6

7

Local Significance

$$trial\ # \sim \sqrt{\frac{\pi}{2}} \mathcal{N} Z_{fix}$$



Why $\text{Trial\#} \sim Z_{\text{fix}}$?



Solution of the LEE problem

$$P(q(\hat{\theta}) > c) \sim P(\chi_1^2 > c) + \langle N(c_0) \rangle e^{-(c-c_0)/2}$$

$$\text{trial \#} \sim 1 + \sqrt{\frac{\pi}{2}} \mathcal{N} Z_{\text{fix}}$$

$$\langle N(c_0) \rangle = \mathcal{N} P(\chi_2^2 > c_0)$$

Where does the Z dependence come from?

View the results as if there are \mathcal{N} independent search regions

In each one there is a χ_2^2 distribution of $q_0(\mu, m)$

The mass is a dof even though it is undefined under the null

$$\sigma_{\hat{m}} \sim \frac{1}{Z} \Rightarrow \frac{\Delta m}{\sigma_{\hat{m}}} \sim Z$$

Why trial# ~ Z

$$\text{Var}(m) = \left[-E \left(\frac{\partial^2 \log \mathcal{L}}{\partial m^2} \right) \right]^{-1}$$

$$n \sim \text{Pois}(\mu s(m) + b) \approx e^{-(\mu s(m) + b)} (\mu s(m) + b)^n$$

$$\log \mathcal{L} = -\mu s(m) - b + n \log(\mu s(m) + b)$$

$$\frac{\partial \log \mathcal{L}}{\partial m} = -\mu \frac{\partial s(m)}{\partial m} + n \frac{\mu}{\mu s(m) + b} \frac{\partial s(m)}{\partial m}$$

$$\frac{\partial^2 \log \mathcal{L}}{\partial m^2} = -\mu \frac{\partial^2 s(m)}{\partial m^2} + n \frac{\mu}{\mu s(m) + b} \frac{\partial^2 s(m)}{\partial m^2} - n \frac{\mu^2}{(\mu s(m) + b)^2} \left(\frac{\partial s(m)}{\partial m} \right)^2$$

$$E[n] = \mu s(m) + b$$

$$E \left[\frac{\partial^2 \log \mathcal{L}}{\partial m^2} \right] = -\frac{\mu^2}{\mu s(m) + b} \left(\frac{\partial s(m)}{\partial m} \right)^2$$

$$\text{Var}[m] \sim \frac{1}{\mu} \sim \frac{1}{Z} \Rightarrow \sigma_{\hat{m}} \sim Z \Rightarrow \text{trial \#} \sim \frac{\text{range}}{\sigma_{\hat{m}}} \sim Z$$



A real life example

$$P(q_0 > u) \leq E[N_u] + P(q_0(0) > u)$$

$$E[N_u] = N_1 e^{-u/2}$$

$$N_1 \cong \langle N_{u_0} \rangle e^{u_0/2}$$

$$P(q_0 > u) = N_1 e^{-u/2} + \frac{1}{2} P(\chi_1^2 > u)$$

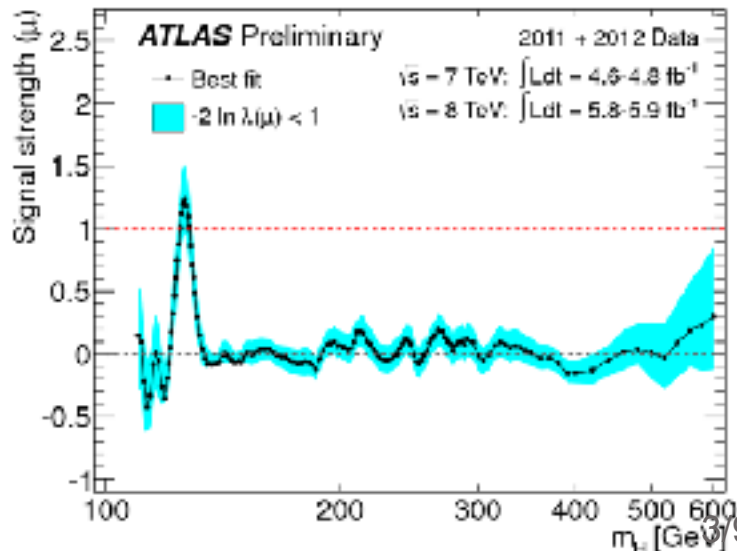
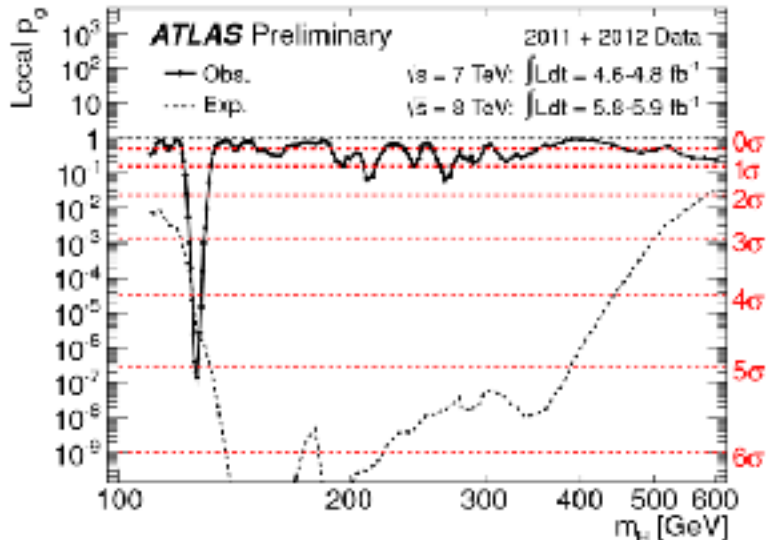
$$p_{global} = N_1 e^{-u/2} + p_{local}$$

$$p_{global} = \langle N_{u_0} \rangle e^{\frac{u_0 - u}{2}} + p_{local}$$

$$N_{u_0=0} = 9 \pm 3$$

$$p_{global} = 9 \cdot e^{-25/2} + O(10^{-7}) = 3.3 \cdot 10^{-5}$$

$$5\sigma \rightarrow 4\sigma \text{ trial}\# \sim 100$$



Example: The 750 GeV Resonance

Spin 0 2015

Largest significance

$m_x \sim 750 \text{ GeV}, \Gamma_x \sim 45 \text{ GeV}$ (6)

Local $Z = 3.9\sigma$

Any peak with $Z > 3.8\sigma$
with $m = 500 - 2000$ will draw our attention

$$P_{global}(u) \approx p_{local}(u) + E(n_{u_0}) e^{-\frac{u_0 - u}{2}}$$

$$p_{local} = 5 \cdot 10^{-5}$$

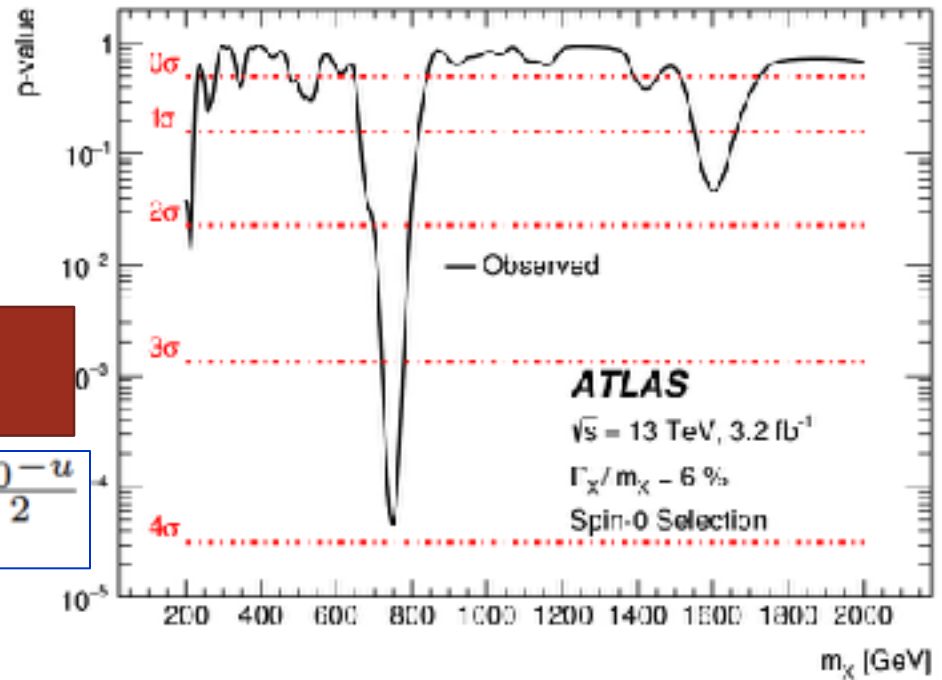
$$u_0 = 0$$

$$n_{u_0} = 7 \pm 2.6$$

$$u = Z^2 = 3.9^2 = 15.2$$

$$p_{global} = 5 \cdot 10^{-5} + (7 \pm 2.6) e^{-15.2/2} = (2.2 - 4.8) 10^{-3}$$

$$Z_{global} \sim 2.7 \pm 0.1\sigma$$



The LEE is even stronger when you consider another dimension
(the width range (0-10%) m should also be taken into account)

End of Lectures
Thank You

Eilam Gross

