

## ROC curves, AUC's and alternatives in HEP event selection and in other domains

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*Disclaimer: I last did physics analyses more than 15 years ago  
(mainly statistically-limited precision measurements and combinations – e.g. no searches)*

# Why and when I got interested in this topic

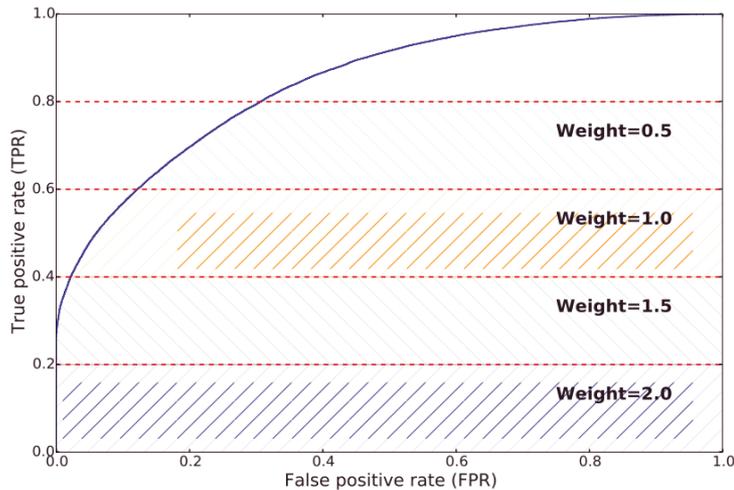


Figure 3: Weights assigned to the different segments of the ROC curve for the purpose of submission evaluation. The  $x$  axis is the False Positive Rate (FPR), while the  $y$  axis is True Positive Rate (TPR).

T. Blake et al., *Flavours of Physics: the machine learning challenge for the search of  $\tau \rightarrow \mu\mu\mu$  decays at LHCb* (2015, unpublished). <https://kaggle2.blob.core.windows.net/competitions/kaggle/4488/media/lhcb.description.official.pdf> (accessed 15 January 2018)

## The 2015 LHCb Kaggle ML Challenge

- Event selection in search for  $\tau \rightarrow \mu\mu\mu$
- **Classifier wins if it maximises a weighted ROC AUC**
- Simplified for Kaggle – real analysis uses CLs

- First time I saw an **Area Under the Roc Curve (AUC)**
- My reaction: what is this? is this relevant in HEP?
  - try to understand why the AUC was introduced in other scientific domains
  - review *common knowledge* for optimizing several types of HEP analyses

*Questions for you – How extensively are AUC's used in HEP, particularly in event selection?  
Are there specific HEP problems where it can be shown that AUC's are relevant?*

# Spoiler! – What I will argue in this talk

- **Different disciplines / problems → different challenges → different metrics**
  - Tools from other domains → assess their relevance before using them in HEP
- **Most relevant metrics in HEP event selection: purity  $\rho$  and signal efficiency  $\epsilon_s$** 
  - “Precision and Recall” – HEP closer to Information Retrieval than to Medicine
  - “True Negatives”, ROCs and AUCs irrelevant in HEP event selection\*
    - **AUCs → Higher not always better. Numerically, no relevant interpretation.**
- **HEP specificity: fits of differential distributions → binning / partitioning of data**
  - local efficiency and purity in each bin → more relevant than global averages of  $\rho, \epsilon_s$
  - scoring classifiers → more useful for partitioning data than for imposing cuts
    - optimize statistical errors on parameter estimates → metrics based on local  $\rho_i^* \epsilon_{s,i}$
    - optimal partitioning: split into bins of uniform purity  $\rho_i$  and sensitivity  $\frac{1}{s_i} \frac{\partial s_i}{\partial \theta}$

\* ROCs are relevant in particle-ID – but this is largely beyond the scope of this talk

# Outline

- Introduction to binary classifiers: the confusion matrix, ROCs, AUCs, PRCs
- Binary classifier evaluation: domain-specific challenges and solutions
  - Overview of Diagnostic Medicine and Information Retrieval
  - A systematic analysis and summary of optimizations in HEP event selection
- Statistical error optimization in HEP parameter estimation problems
  - Information metrics and the effect of local efficiency and purity in binned fits
  - Optimal binning and the relevance of local purity
- Conclusions

# Binary classifiers: the “confusion matrix”

- Data sample containing instances of two classes:  $N_{tot} = S_{tot} + B_{tot}$ 
  - HEP: signal  $S_{tot} = S_{sel} + S_{rej}$
  - HEP: background  $B_{tot} = B_{sel} + B_{rej}$
- Discrete binary classifiers assign each instance to one of the two classes
  - HEP: classified as signal and selected  $N_{sel} = S_{sel} + B_{sel}$
  - HEP: classified as background and rejected  $N_{rej} = B_{rej} + S_{rej}$

	<u>true class</u> : Positives + (HEP: signal)	<u>true class</u> : Negatives - (HEP: background)
<u>classified as</u> : positives (HEP: selected)	<b>True Positives (TP)</b> (HEP: selected signal <b>Ssel</b> )	<b>False Positives (FP)</b> (HEP: selected bkg <b>Bsel</b> )
<u>classified as</u> : negatives (HEP: rejected)	<b>False Negatives (FN)</b> (HEP: rejected signal <b>Srej</b> )	<b>True Negatives (TN)</b> (HEP: rejected bkg <b>Brej</b> )

T. Fawcett, *Introduction to ROC analysis*, Pattern Recognition Letters 27 (2006) 861. doi:10.1016/j.patrec.2005.10.010

*I will not discuss multi-class classifiers (useful in HEP particle-ID)*

# The confusion matrix about the confusion matrix...

Different domains → focus on different concepts → different terminologies

TP ( $S_{sel}$ )	FP ( $B_{sel}$ )
FN ( $S_{rej}$ )	TN ( $B_{rej}$ )

TP ( $S_{sel}$ )	FP ( $B_{sel}$ )
FN ( $S_{rej}$ )	TN ( $B_{rej}$ )

TP ( $S_{sel}$ )	FP ( $B_{sel}$ )
FN ( $S_{rej}$ )	TN ( $B_{rej}$ )

$$TPR = \frac{TP}{TP + FN}$$

$$PPV = \frac{TP}{TP + FP}$$

$$TNR = \frac{TN}{TN + FP} = 1 - FPR$$

HEP: “efficiency”

$$\epsilon_s = \frac{S_{sel}}{S_{tot}}$$

HEP: “purity”

$$\rho = \frac{S_{sel}}{S_{sel} + B_{sel}}$$

HEP: “background rejection”

$$1 - \epsilon_b = 1 - \frac{B_{sel}}{B_{tot}}$$

IR: “recall”

IR: “precision”

—

MED: “sensitivity”

—

MED: “specificity”

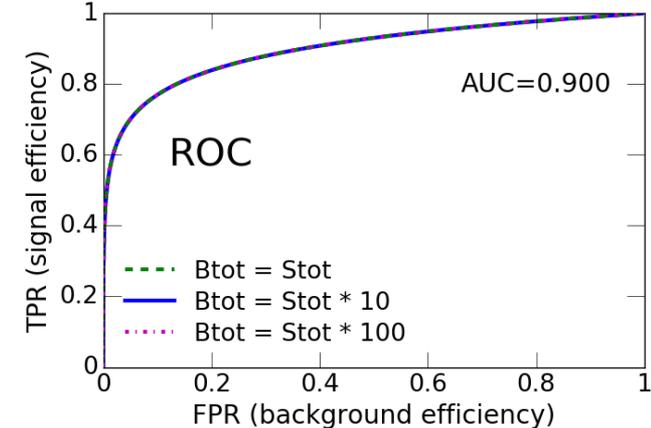
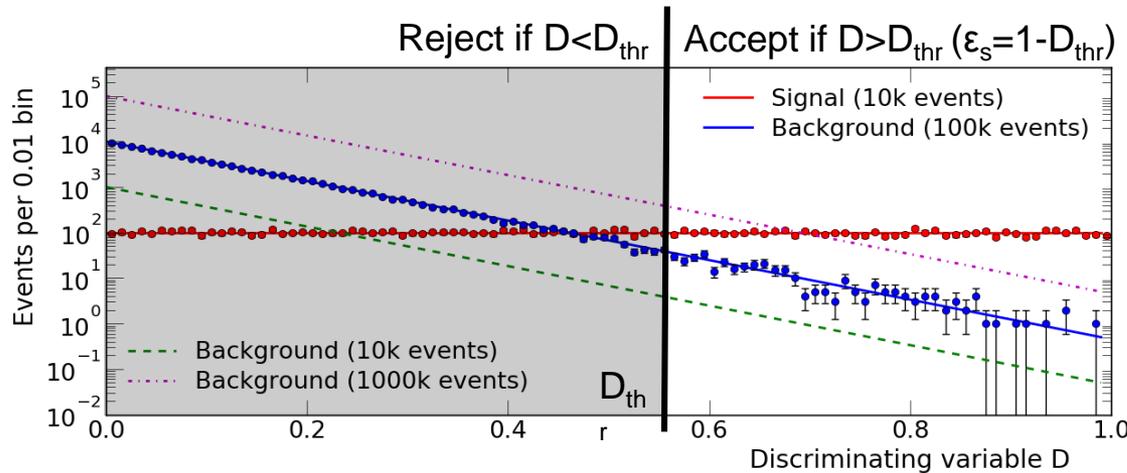
I will cover three domains:

- **Medical Diagnostics (MED)**  
*does Mr. A. have cancer?*
- **Information Retrieval (IR)**  
*Google documents about “ROC”*
- **HEP event selection (HEP)**  
*select Higgs event candidates*

MED: prevalence

$$\pi_s = \frac{S_{tot}}{S_{tot} + B_{tot}}$$

# Discrete vs. Scoring classifiers – ROC curves



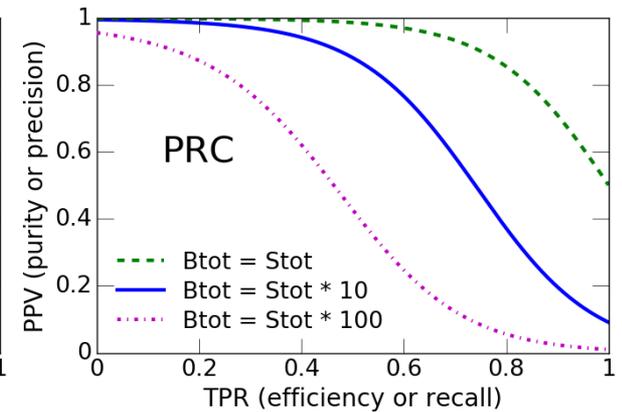
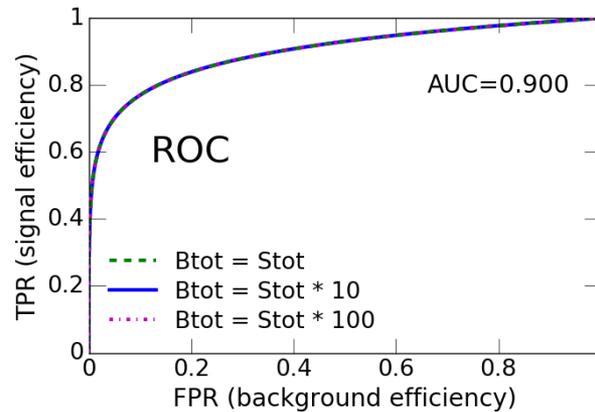
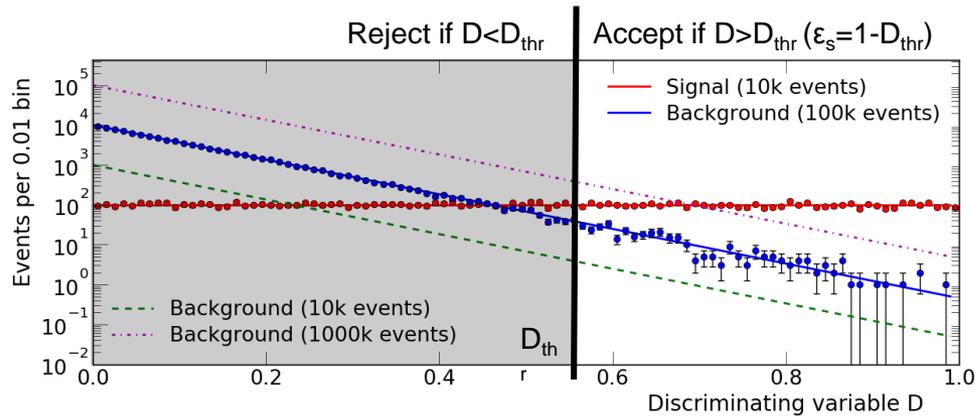
- Discrete classifiers → either select or reject → confusion matrix
- Scoring classifiers → assign score  $D$  to each event (e.g. BDT)
  - ideally related to likelihood that event is signal or background (Neyman-Pearson)
  - from scoring to discrete: choose a threshold → classify as signal if  $D > D_{thr}$
- ROC curves describe how  $FPR(\epsilon_b)$  and  $TPR(\epsilon_s)$  are related when varying  $D_{thr}$ 
  - used initially in radar signal detection and psychophysics (1940-50's)

W. W. Peterson, T. G. Birdsall, W. C. Fox, *The theory of signal detectability*, Transactions of the IRE Professional Group on Information Theory 4 (1954) 171. doi:10.1109/TIT.1954.1057460  
 W. P. Tanner, J. A. Swets, *A decision-making theory of visual detection*, Psychological Review 61 (1954), 401. doi:10.1037/h0058700

J. A. Swets, *Is There a Sensory Threshold?*, Science 134 (1961) 168. doi:10.1126/science.134.3473.168  
 J. A. Swets, W. P. Tanner, T. G. Birdsall, *Decision processes in perception*, Psychological Review 68 (1961) 301. doi:10.1037/h0040547

# ROC and PRC (precision-recall) curves

- Different choice of ratios in the confusion matrix:  $\varepsilon_s, \varepsilon_b$  (ROC) or  $\rho, \varepsilon_s$  (PRC)
- When  $B_{tot}/S_{tot}$  (“prevalence”) varies  $\rightarrow$  PRC changes, ROC does not



# Understanding domain-specific challenges

- Many domain-specific details → but also general cross-domain questions:
  - **1. Qualitative imbalance?**
    - Are the two classes equally relevant?
  - **2. Quantitative imbalance?**
    - Is the prevalence of one class much higher?
  - **3. Prevalence known? Time invariance?**
    - Is relative prevalence known in advance? Does it vary over time?
  - **4. Dimensionality? Scale invariance?**
    - Are all 4 elements of the confusion matrix needed?
    - Is the problem invariant under changes of some of these elements?
  - **5. Ranking? Binning?**
    - Are all selected instances equally useful? Are they partitioned into subgroups?
- Point out properties of MED and IR, attempt a systematic analysis of HEP

M. Sokolova, G. Lapalme, *A Systematic Analysis of Performance Measures for Classification Tasks*, Information Processing and Management 45 (2009) 427.  
[doi:10.1016/j.ipm.2009.03.002](https://doi.org/10.1016/j.ipm.2009.03.002)

# Medical diagnostics (1)

## and ML research

H. Sox, S. Stern, D. Owens, H. L. Abrams, *Assessment of Diagnostic Technology in Health Care: Rationale, Methods, Problems, and Directions*, The National Academies Press (1989). doi:10.17226/1432

X. H. Zhou, D. K. McClish, N. A. Obuchowski, *Statistical Methods in Diagnostic Medicine* (Wiley, 2002). doi:10.1002/9780470317082

- Medical Diagnostics (MED)  
does Mr. A. have cancer?

- Binary classifier optimisation goal: maximise “diagnostic accuracy”
  - patient / physician / society have different goals → many possible definitions

- Most popular metric: “accuracy”, or “probability of correct test result”:

$$\text{ACC} = \frac{\text{TP} + \text{TN}}{\text{TP} + \text{TN} + \text{FP} + \text{FN}} = \pi_s \times \text{TPR} + (1 - \pi_s) \times \text{TNR}$$

TP (correctly diagnosed as ill)	FP (truly healthy, but diagnosed as ill)
FN (truly ill, but diagnosed as healthy)	TN (correctly diagnosed as healthy)

- Symmetric → all patients important, both truly ill (TP) and truly healthy (TN)
- Also “by far the most commonly used metric” in ML research in the 1990s

F. J. Provost, T. Fawcett, *Analysis and Visualization of Classifier Performance: Comparison Under Imprecise Class and Cost Distributions*, Proc. 3rd Int. Conf. on Knowledge Discovery and Data Mining (KDD-97), Newport Beach, USA (1997). <https://aaai.org/Library/KDD/1997/kdd97-007.php>

L. B. Lusted, *Signal Detectability and Medical Decision-Making*, Science 171 (1971) 1217 doi:10.1126/science.171.3977.1217

J. A. Swets, *Measuring the accuracy of diagnostic systems*, Science 240 (1988) 1285. doi:10.1126/science.3287615

- Since the ‘90s → shift from ACC to ROC in the MED and ML fields

- TPR (sensitivity) and TNR (specificity) studied separately
  - solves ACC limitations (imbalanced or unknown prevalence – rare diseases, epidemics)
- Evaluation often AUC-based → two perceived advantages for MED and ML fields
  - **AUC interpretation: “probability that test result of randomly chosen sick subject indicates greater suspicion than that of randomly chosen healthy subject”**
  - ROC comparison without prior  $D_{\text{thr}}$  choice (prevalence-dependent  $D_{\text{thr}}$  choice)

F. J. Provost, T. Fawcett, R. Kohavi, *The Case against Accuracy Estimation for Comparing Induction Algorithms*, Proc. 15th Int. Conf. on Machine Learning (ICML '98), Madison, USA (1998). <https://www.researchgate.net/publication/2373067>

A. P. Bradley, *The use of the area under the ROC curve in the evaluation of machine learning algorithms*, Pattern Recognition 30 (1997) 1145. doi:10.1016/S0031-3203(96)00142-2

J. A. Hanley, B. J. McNeil, *The meaning and use of the area under a receiver operating characteristic (ROC) curve*, Radiology 143 (1982) 29. doi:10.1148/radiology.143.1.7063747



# Medical diagnostics (2)

## and ML research

- ROC and AUC metrics → currently widely used in the MED and ML fields
  - Remember: moved because *ROC better than ACC with imbalanced data sets*
- Limitation: evidence that *ROC not so good for highly imbalanced data sets*
  - may provide an overly optimistic view of performance
  - PRC may provide a more informative assessment of performance in this case
    - PRC-based reanalysis of some data sets in life sciences has been performed
- Very active area of research → other options proposed (CROC, cost models)
  - Take-away message: *ROC and AUC not always the appropriate solutions*

J. Davis, M. Goadrich, *The relationship between Precision-Recall and ROC curves*, Proc. 23rd Int. Conf. on Machine Learning (ICML '06), Pittsburgh, USA (2006). doi:10.1145/1143844.1143874

C. Drummond, R. C. Holte, *Explicitly representing expected cost: an alternative to ROC representation*, Proc. 6th Int. Conf. on Knowledge Discovery and Data Mining (KDD-00), Boston, USA (2000). doi:10.1145/347090.347126

D. J. Hand, *Measuring classifier performance: a coherent alternative to the area under the ROC curve*, Mach Learn (2009) 77: 103. doi:10.1007/s10994-009-5119-5

S. J. Swamidass, C.-A. Azencott, K. Daily, P. Baldi, *A CROC stronger than ROC: measuring, visualizing and optimizing early retrieval*, Bioinformatics 26 (2010) 1348. doi:10.1093/bioinformatics/btq140

D. Berrar, P. Flach, *Caveats and pitfalls of ROC analysis in clinical microarray research (and how to avoid them)*, Briefings in Bioinformatics 13 (2012) 83. doi:10.1093/bib/bbr008

H. He, E. A. Garcia, *Learning from Imbalanced Data*, IEEE Trans. Knowl. Data Eng. 21 (2009) 1263. doi:10.1109/TKDE.2008.239

T. Saito, M. Rehmsmeier, *The Precision-Recall Plot Is More Informative than the ROC Plot When Evaluating Binary Classifiers on Imbalanced Datasets*, PLoS One 10 (2015) e0118432. doi:10.1371/journal.pone.0118432

# Information Retrieval

- Information Retrieval (IR)  
Google documents about "ROC"

- Qualitative distinction between “relevant” and “non-relevant” documents
  - also a very large quantitative imbalance
- Binary classifier optimisation goal: make users happy in web searches
  - minimise # relevant documents not retrieved → maximise “recall” i.e. efficiency
  - minimise # of irrelevant documents retrieved → maximise “precision” i.e. purity
  - retrieve the more relevant documents first → ranking very important
  - maximise speed of retrieval
- IR-specific metrics to evaluate classifiers based on the PRC (i.e. on  $\epsilon_s$ ,  $\rho$ )
  - unranked evaluation → e.g. F-measures  $F_\alpha = \frac{1}{\alpha/\epsilon_s + (1-\alpha)/\rho}$ 
    - $\alpha \in [0,1]$  *tradeoff between recall and precision* → equal weight gives  $F1 = \frac{2\epsilon_s\rho}{\epsilon_s + \rho}$
  - ranked evaluation → precision at k documents, mean average precision (MAP), ...
    - MAP approximated by the Area Under the PRC curve (AUCPR)

C. D. Manning, P. Raghavan, H. Schütze, *Introduction to Information Retrieval* (Cambridge University Press, 2008).  
<https://nlp.stanford.edu/IR-book>

NB: Many different of meanings of “Information”!  
IR (web documents), HEP (Fisher), Information Theory (Shannon)...



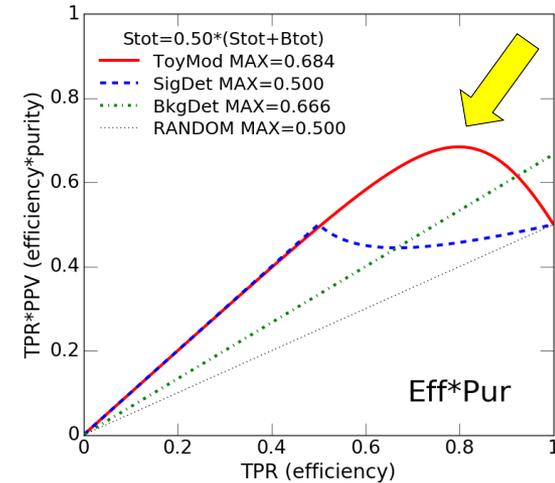
# First (simplest) HEP example

- HEP event selection (HEP)  
select Higgs event candidates

- Measurement of a total cross-section  $\sigma_s$  in a counting experiment
- To minimize statistical errors: **maximise  $\epsilon_s * \rho$**  (well-known since decades)
  - global efficiency  $\epsilon_s = S_{sel}/S_{tot}$  and global purity  $\rho = S_{sel}/(S_{sel} + B_{sel})$  – “1 single bin”

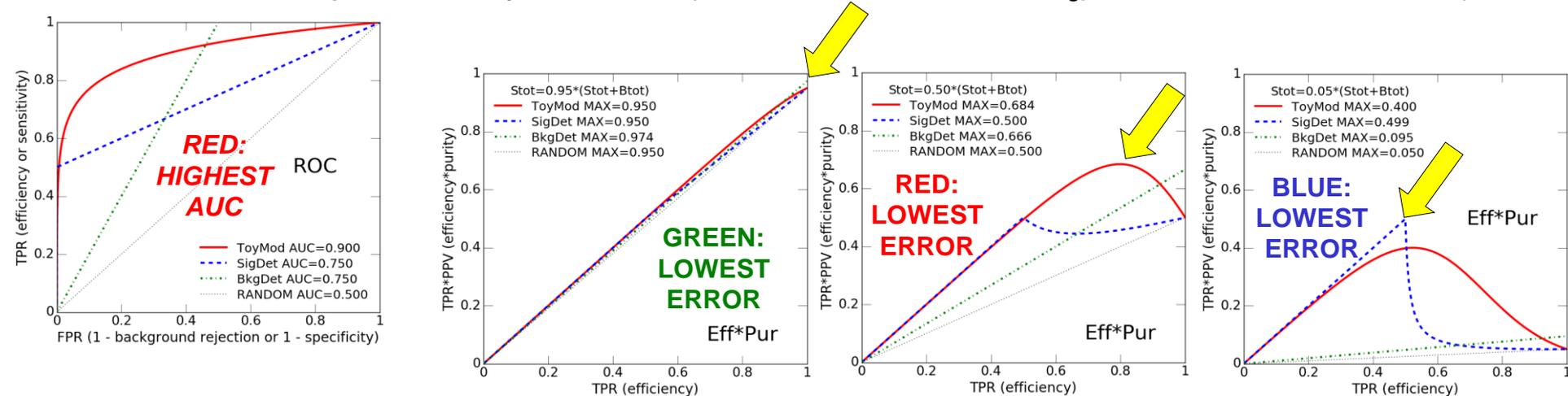
$$\frac{1}{(\Delta\sigma_s)^2} = \frac{1}{\sigma_s} \mathcal{L} \epsilon_s \rho = \frac{1}{\sigma_s^2} S_{tot} \epsilon_s \rho$$

- To compare classifiers (red, green, blue, black):
  - in each classifier  $\rightarrow$  vary Dthr cut  $\rightarrow$  vary  $\epsilon_s$  and  $\rho$
  - $\rightarrow$  find maximum of  $\epsilon_s * \rho$  (choose “operating point”)
  - chose classifier with maximum of  $\epsilon_s * \rho$  out of the four
- $\epsilon_s * \rho$ : metric between 0 and 1
  - qualitatively relevant: the higher, the better
  - numerically: fraction of Fisher information ( $1/\text{error}^2$ ) available after selecting
  - **correct metric only for  $\sigma_s$  by counting!**  $\rightarrow$  table with more cases on a next slide



# Examples of issues with AUCs – crossing ROCs

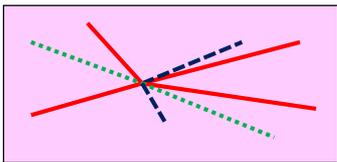
- Choice of classifier easy if one ROC “dominates” another (higher TPR  $\forall$  FPR)
  - PRC “dominates” too, then – and of course AUC is higher, too
- Choice is less obvious if ROCs cross!
- Example: cross-section by counting
  - maximise product  $\epsilon_s \rho \rightarrow$  i.e. minimise the statistical error  $\Delta\sigma^2$
  - depending on  $S_{\text{tot}}/B_{\text{tot}}$ , a different classifier (green, red, blue) should be chosen
  - in two out of three scenarios, **the classifier with the highest AUC is not the best**
    - AUC is qualitatively irrelevant (higher is not always better)
    - AUC is quantitatively irrelevant (0.75, 0.90, so what? –  $\epsilon_s \rho$  instead means  $1/\Delta\sigma^2$ ...)



# Binary classifiers in HEP

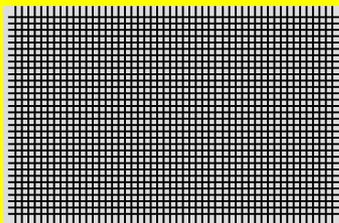
- HEP event selection (HEP)  
select Higgs event candidates

Binary classifier optimisation goal: maximise physics reach at a given budget



**Tracking and particle-ID (event reconstruction)** – e.g. fake track rejection  
→ maximise identification of particles (*all particles within each event are important*)

Instances: tracks within one event, created by earlier reconstruction stage.  
→ P = real tracks, N = fake tracks (ghosts) → goal: keep real tracks, reject ghosts  
→ TN = fake tracks identified as such and rejected: **TN are relevant** (IIUC...)  
[Optimisation: should translate tracking metrics into measurement errors in physics analyses]



**Trigger** → maximise signal event throughput, within the computing budget – e.g. HLT

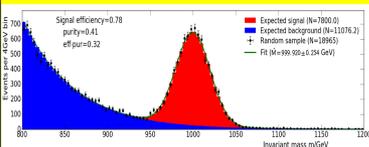
Instances: events, from the earlier trigger stage (e.g. L0 hardware trigger)  
→ P = signal events, N = background events [per unit time: trigger rates]  
→ goal: **maximise retained signal efficiency** TP/(TP+FN) at a given trigger rate FP (as TP << FP)  
→ TN = background events identified as such and rejected: **TN are irrelevant**  
→ constraint: max HLT rate (from HLT throughput), whatever the input L0 rate is: **TN are ill-defined**

## EVENT SELECTION – I WILL FOCUS ON THIS IN THIS TALK

**Physics analyses** → maximise the physics reach, given the available data sets

Instances: events, from pre-selected data sets  
→ P = signal events, N = background events  
→ goal: **minimise measurement errors** or maximise significance in searches  
→ TN = background events identified as such and rejected: **TN are irrelevant**  
→ physics results independent of pre-selection or MC cuts: **TN are ill-defined**

TP = S <sub>sel</sub>	FP = B <sub>sel</sub>
FN = S <sub>rej</sub>	<del>TN = B<sub>rej</sub></del>



Domain Property	Medical diagnostics	Information retrieval	HEP event selection
Qualitative class imbalance	<b>NO.</b> Healthy and ill people have “equal rights”. <i>TN are relevant.</i>	<b>YES.</b> “Non-relevant” documents are a nuisance. <i>TN are irrelevant.</i>	<b>YES.</b> Background events are a nuisance. <i>TN are irrelevant.</i>
Quantitative class imbalance	<b>From small to extreme.</b> From common flu to very rare disease.	<b>Generally very high.</b> Only very few documents in a repository are relevant.	<b>Generally extreme.</b> Signal events are swamped in background events.
Varying or unknown prevalence $\pi$	<b>Varying and unknown.</b> Epidemics may spread.	<b>Varying and unknown</b> in general (e.g. WWW).	<b>Constant in time</b> (quantum cross-sections). <b>Unknown</b> for searches. <b>Known</b> for precision measurements.
Dimensionality and invariances <small>M. Sokolova, G. Lapalme, A Systematic Analysis of Performance Measures for Classification Tasks, Information Processing and Management 45 (2009) 427. doi:10.1016/j.ipm.2009.03.002</small>	<b>3 ratios <math>\epsilon_s, \epsilon_b, \pi</math> + scale.</b> New metrics under study because ROC ignores $\pi$ . Costs scale with $N_{tot}$ .	<b>2 ratios <math>\epsilon_s, \rho</math> + scale.</b> $\epsilon_s, \rho$ enough in many cases. Costs and speed scale with $N_{tot}$ . Show only $N_{sel}$ docs in one page. <i>TN are irrelevant.</i>	<b>2 ratios <math>\epsilon_s, \rho</math> + scale.</b> $\epsilon_s, \rho$ enough in many cases. Lumi is needed for: trigger, syst. vs stat., searches. <i>TN are irrelevant.</i>
Different use of selected instances	<b>Binning – NO.</b> <b>Ranking – YES?</b> Treat with higher priority patients who are more likely to be ill?	<b>Binning – NO.</b> <b>Ranking – YES.</b> Precision at k, R-precision, MAP all involve <i>global</i> precision-recall (“top $N_{sel}$ documents retrieved”)	<b>Binning – YES.</b> Fits to distributions: <i>local <math>\epsilon_{st}, \rho</math> in each bin</i> rather than global $\epsilon_s, \rho$ .

# Different HEP problems → Different metrics

## Binary classifiers for HEP event selection (signal-background discrimination)

Statistical error minimization (or statistical significance maximization)	Cross-section (1-bin counting)	Only 2 or 3 global/local variables – TN, AUC irrelevant	2 variables: global $\epsilon_s, \rho$ (given $S_{tot}$ )	Maximise $S_{tot} * \epsilon_s * \rho$ (at any $S_{tot}$ )
	Searches (1-bin counting)		Simple and CCGV – 2 variables: global $S_{sel}, B_{sel}$ (or equivalently $\epsilon_s, \rho$ )	Maximise $\frac{S_{sel}}{\sqrt{S_{sel} + B_{sel}}}$ (i.e. $\sqrt{S_{tot} * \epsilon_s * \rho}$ )
			HiggsML – 2 variables: global $S_{sel}, B_{sel}$	Maximise $\sqrt{2((S_{sel} + B_{sel}) \log(1 + \frac{S_{sel}}{B_{sel}}) - S_{sel})}$
			Punzi – 2 variables: global $\epsilon_s, B_{sel}$	Maximise $\frac{\epsilon_s}{A/2 + \sqrt{B_{sel}}}$
Cross-section (binned fits)	Parameter estimation (binned fits)		2 variables: local $\epsilon_{s,i}$ and $\rho_i$ in each bin (given $s_{tot,i}$ in each bin)	Maximise $\sum_i s_{tot,i} * \epsilon_{s,i} * \rho_i$ Partition in bins of equal $\rho_i$
				Maximise $\sum_i s_{tot,i} * \epsilon_{s,i} * \rho_i * (\frac{1}{S_{tot,i}} \frac{\partial S_{tot,i}}{\partial \theta})^2$ Partition in bins of equal $\rho_i * (\frac{1}{S_{tot,i}} \frac{\partial S_{tot,i}}{\partial \theta})$
Searches (binned fits)			3 variables: local $s_{sel}, S_{tot}, s_{sel}$ in each bin (2 counts or ratios enough?)	Maximise a sum? *
Statistical + Systematic error minimization			3 variables: $\epsilon_s, \rho, \text{lumi}$ (lumi: tradeoff stat. vs. syst.)	No universal recipe * (may use local $S_{sel}, B_{sel}$ in side band bins)
Trigger optimization			2 variables: global $B_{sel}/\text{time}, \text{global } \epsilon_s$	Maximise $\epsilon_s$ at given trigger rate

## Binary classifiers for HEP problems other than event selection

Tracking and Particle-ID optimizations	All 4 variables? * (NB: TN is relevant)	ROC relevant – is AUC relevant? *
Other? *	? *	? *

\* Many open questions for further research



# Predict and optimize statistical errors in binned fits

- Fit  $\theta$  from a binned multi-dimensional distribution
  - expected counts  $y_i = \int f(x_i, \theta) dx = \epsilon_i * s_i(\theta) + b_i \rightarrow$  depend on parameter  $\theta$  to fit
- Statistical error related to Fisher information  $\boxed{(\Delta\hat{\theta})^2 = \text{var}(\hat{\theta}) \geq \frac{1}{\mathcal{I}_\theta}}$  (Cramer-Rao)
  - binned fit  $\rightarrow$  combine measurements in each bin, weighed by information

- Easy to show (backup slides) that Fisher information in the fit is:

$$\mathcal{I}_\theta^{(\text{real classifier})} = \sum_{i=1}^m \epsilon_i \rho_i \times \frac{1}{S_i} \left( \frac{\partial S_i}{\partial \theta} \right)^2$$

$$\mathcal{I}_\theta^{(\text{ideal classifier})} = \sum_{i=1}^m \frac{1}{S_i} \left( \frac{\partial S_i}{\partial \theta} \right)^2$$

–  $\epsilon_i$  and  $\rho_i \rightarrow$  local signal efficiency and purity in the  $i^{\text{th}}$  bin

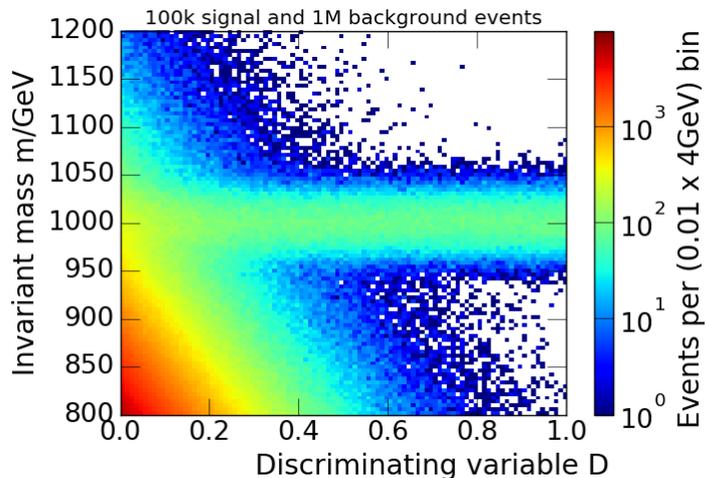
- Define a binary classifier metric as information fraction to ideal classifier:
  - in  $[0, 1] \rightarrow 1$  if keep all signal and reject all backgrounds
  - higher is better  $\rightarrow$  maximise IF
  - interpretation:  $\boxed{(\Delta\hat{\theta}^{(\text{real classifier})})^2 \geq \frac{1}{\text{IF}} (\Delta\hat{\theta}^{(\text{ideal classifier})})^2}$

$$\text{IF} = \frac{\mathcal{I}_\theta^{(\text{real classifier})}}{\mathcal{I}_\theta^{(\text{ideal classifier})}} = \frac{\sum_{i=1}^m \epsilon_i \rho_i \times \frac{1}{S_i} \left( \frac{\partial S_i}{\partial \theta} \right)^2}{\sum_{i=1}^m \frac{1}{S_i} \left( \frac{\partial S_i}{\partial \theta} \right)^2}$$

**NB:** global  $\epsilon * \rho$  is the IF for measuring  $\theta = \sigma_s$  in a 1-bin fit (counting experiment)!

# Numerical tests with a toy model

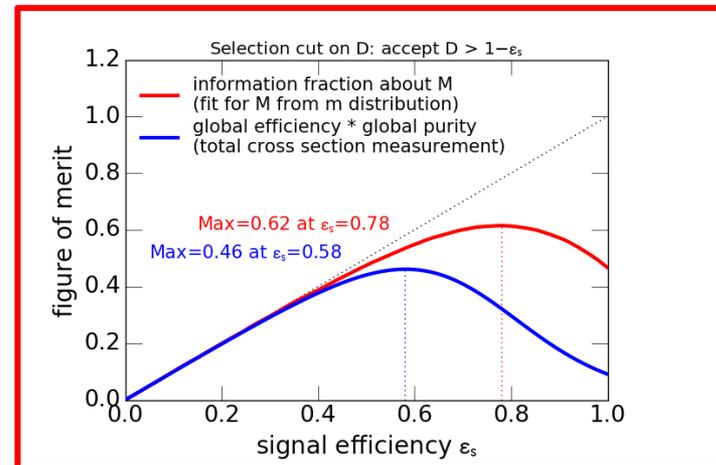
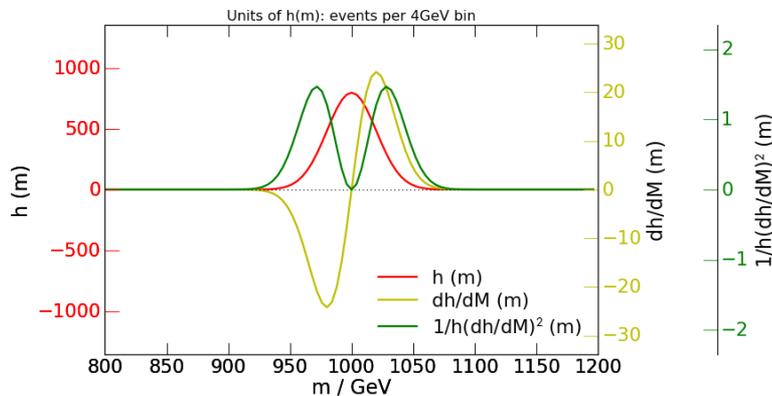
- I used a simple toy model to make some numerical tests
  - Verify that my formulas are correct – and also illustrate them graphically
  - Two-dimensional distribution (m,D) → signal Gaussian, background exponential
- Two measurements:
  - total cross-section measurement by counting and 1-D or 2-D fit
  - mass measurement by 1-D or 2-D fits
- Details in the backup slides



*Using scipy / matplotlib / numpy  
and iminuit in Python from SWAN*

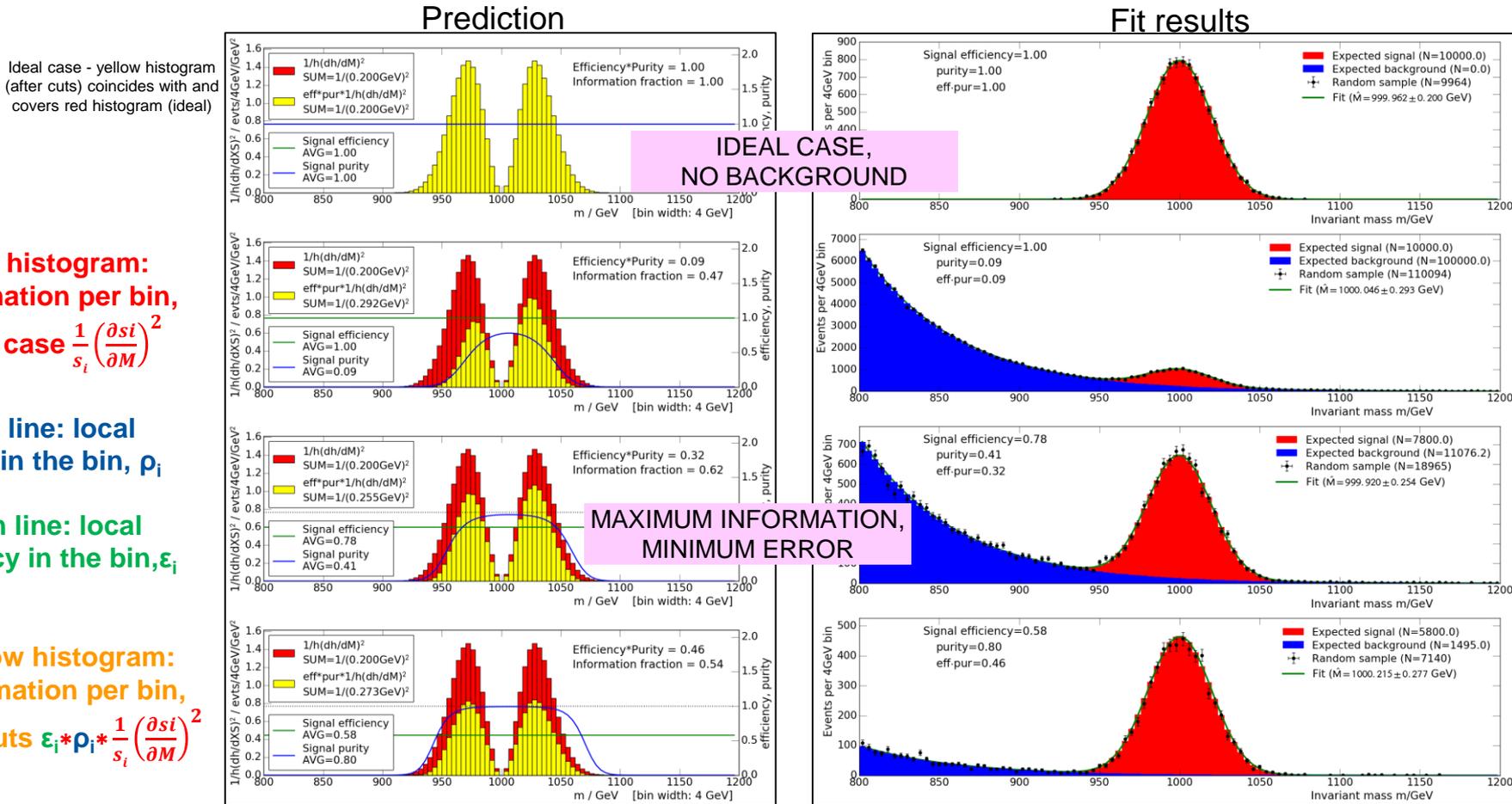
# M by 1D fit to m – optimizing the classifier

- Choose operating point  $D_{\text{thr}}$  optimizing information fraction for  $\theta=M$  in m-fit
  - NB: different to operating point maximising  $\varepsilon^*\rho$  (IF for  $\theta=\sigma_s$  in a 1-bin fit)
- To compute IF as sum over bins  $\rightarrow$  need average  $\frac{1}{s} \frac{\partial s}{\partial \theta}$  in each bin
  - proof-of-concept  $\rightarrow$  integrate by toy MC with *event-by-event weight derivatives*
    - in a real MC, could save  $\frac{1}{|\mathcal{M}|^2} \frac{\partial |\mathcal{M}|^2}{\partial \theta}$  for the matrix element squared  $|\mathcal{M}|^2$



# M by 1D fit to m – visual interpretation

- Information after cuts:  $\sum_i \frac{1}{s_i} \left( \frac{\partial s_i}{\partial M} \right)^2 * \epsilon_i * \rho_i \rightarrow$  show the 3 terms in each bin  $i$
- fit = combine N different measurements in N bins  $\rightarrow$  local  $\epsilon_i, \rho_i$  relevant!



# Optimal partitioning – information inflow

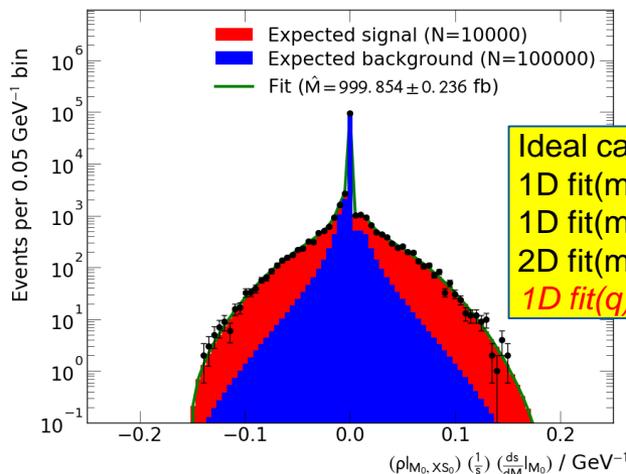
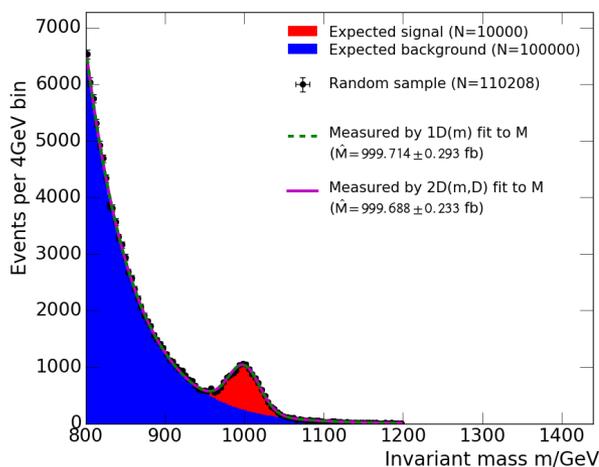
- Information about  $\theta$  in a binned fit  $\rightarrow \mathcal{I}_\theta = \sum_{i=1}^m \frac{1}{y_i} \left( \frac{\partial y_i}{\partial \theta} \right)^2$
- Do I gain anything by splitting bin  $y_i$  into two separate bins?  $y_i = w_i + z_i$ 
  - i.e. is the “information inflow”<sup>\*</sup> positive?
 
$$\frac{1}{w_i} \left( \frac{\partial w_i}{\partial \theta} \right)^2 + \frac{1}{z_i} \left( \frac{\partial z_i}{\partial \theta} \right)^2 - \frac{1}{w_i + z_i} \left( \frac{\partial (w_i + z_i)}{\partial \theta} \right)^2 = \frac{(w_i \frac{\partial z_i}{\partial \theta} - z_i \frac{\partial w_i}{\partial \theta})^2}{w_i z_i (w_i + z_i)} \geq 0$$
  - information increases (errors on parameters decrease) if  $\frac{1}{w_i} \frac{\partial w_i}{\partial \theta} \neq \frac{1}{z_i} \frac{\partial z_i}{\partial \theta}$
  - effect of the classifier  $\rightarrow$  **information increases if  $\rho_w \frac{1}{s_w} \frac{\partial s_w}{\partial \theta} \neq \rho_z \frac{1}{s_z} \frac{\partial s_z}{\partial \theta}$**
- In summary: **try to partition the data into bins of equal  $\rho_i \frac{1}{s_i} \frac{\partial s_i}{\partial \theta}$** 
  - for cross-section measurements (and searches?): split into bins of equal  $\rho_i$ 
    - “use the scoring classifier D to partition the data, not to reject events”

\*A. van den Bos, *Parameter Estimation for Scientists and Engineers* (Wiley, 2007).

# Optimal partitioning – optimal variables

- The previous slide implies that  $q = \rho \frac{1}{s} \frac{\partial s}{\partial \theta}$  is an optimal variable to fit for  $\theta$ 
  - proof of concept → 1-D fit of  $q$  has the same precision on  $M$  as 2-D fit of  $(m, D)$
  - closely related to the “optimal observables” technique

M. Davier, L. Duflot, F. LeDiberder, A. Roug , *The optimal method for the measurement of tau polarization*, Phys. Lett. B 306 (1993) 411. doi:10.1016/0370-2693(93)90101-M  
M. Diel, O. Nachtmann, *Optimal observables for the measurement of three-gauge-boson couplings in  $e^+e^- \rightarrow W^+W^-$* , Z. Phys. C 62 (1994) 397. doi:10.1007/BF01555899  
O. Nachtmann, F. Nagel, *Optimal observables and phase-space ambiguities*, Eur. Phys. J. C40 (2005) 497. doi:10.1140/epjc/s2005-02153-9



Ideal case:	$\pm 0.200$
1D fit(m), no cut(D):	$\pm 0.292$
1D fit(m), optimal cut(D):	$\pm 0.254$
2D fit(m,D), no cuts:	$\pm 0.233$
1D fit(q):	$\pm 0.236$

- In practice: train one ML variable to reproduce  $\frac{1}{s} \frac{\partial s}{\partial \theta}$ 
  - not needed for cross-sections or searches (this is constant)

# Conclusion and outlook

- *Different disciplines / problems → different challenges → different metrics*
  - there is no universal magic solution – and the AUC definitely is not one
  - I proposed a systematic analysis of many problems in HEP event selection only
- True Negatives, ROCs & AUCs are irrelevant in HEP event selection
  - PRC approach (like IR, unlike MED) more appropriate → purity  $\rho$ , efficiency  $\varepsilon_s$
- Binning in HEP analyses → global averages of  $\rho$ ,  $\varepsilon_s$  irrelevant in that case
  - FOM integrals that are relevant to HEP use local  $\rho$ ,  $\varepsilon_s$  in each bin
  - AUC is an integral of global  $\rho$ ,  $\varepsilon_s$  → one more reason why it is irrelevant
  - optimal partitioning exists to minimise statistical errors on fits
- What am I proposing about ROCs and AUCs, essentially?
  - **stop using AUCs and ROCs in HEP event selection**
    - ROCs confusing → they make you think in terms of the wrong metrics
  - **identify the metrics most appropriate to your specific problem**
    - I summarized many metrics that exist for some problems in event selection
    - *more research needed* in other problems (e.g. pID, systematics in event selection...)

*I am preparing a paper on this – thank you for your feedback in this meeting!*

# BACKUP SLIDES

# Statistical error in binned fits

- Observed data: event counts  $n_i$  in  $m$  bins of a (multi-D) distribution  $f(x)$ 
  - the expected counts  $y_i = f(x_i, \theta) dx$  depend on a parameter  $\theta$  that we want to fit
  - [NB here  $f$  is a differential cross section, it is not normalized to 1 like a pdf]
- Fitting  $\theta$  is like combining the independent measurements in the  $m$  bins
  - expected error on  $n_i$  in bin  $x_i$  is  $\Delta n_i = \sqrt{y_i} = \sqrt{f(x_i, \theta) dx}$
  - expected error on  $f(x_i, \theta)$  in bin  $x_i$  is  $\Delta f = f * \Delta n_i / n_i = \sqrt{f / dx}$
  - expected error on estimated  $\hat{\theta}_i$  in bin  $x_i$  is  $\frac{1}{(\Delta \hat{\theta})_{(\text{bin } dx)}^2} = \left(\frac{\partial f}{\partial \theta}\right)^2 \frac{1}{(\Delta f)^2} = \left(\frac{\partial f}{\partial \theta}\right)^2 \left(\frac{\sqrt{dx}}{\sqrt{f}}\right)^2 = \left(\frac{\partial f}{\partial \theta}\right)^2 \frac{dx}{f}$
  - expected error on estimated  $\hat{\theta}$  by combining the  $m$  bins is  $\left(\frac{1}{\Delta \hat{\theta}}\right)^2 = \int \frac{1}{f} \left(\frac{\partial f}{\partial \theta}\right)^2 dx$
- A bit more formally, joint probability for observing the  $n_i$  is  $P(\mathbf{n}; \theta) = \prod_{i=1}^m \frac{e^{-y_i} y_i^{n_i}}{n_i!}$ 
  - Fisher information on  $\theta$  from the data available is then
 
$$\mathcal{I}_\theta = E \left[ \frac{\partial \log P(\mathbf{n}; \theta)}{\partial \theta} \right]^2 \quad \text{i.e.} \quad \mathcal{I}_\theta = \sum_{i=1}^m \frac{1}{y_i} \left( \frac{\partial y_i}{\partial \theta} \right)^2 = \int \frac{1}{f} \left( \frac{\partial f}{\partial \theta} \right)^2 dx$$
  - The minimum variance achievable (Cramer-Rao lower bound) is  $(\Delta \hat{\theta})^2 = \text{var}(\hat{\theta}) \geq \frac{1}{\mathcal{I}_\theta}$

# Effect of realistic classifiers on fits

- Previous slide: variance on estimated  $\hat{\theta}$  is  $(\Delta\hat{\theta})^2 = \text{var}(\hat{\theta}) \geq \frac{1}{\mathcal{I}_\theta}$  where  $\mathcal{I}_\theta = \sum_{i=1}^m \frac{1}{y_i} \left(\frac{\partial y_i}{\partial \theta}\right)^2$
- With an *ideal classifier*, all signal events and only signal events are selected, i.e.  $y_i = S_i$ , hence:  $\mathcal{I}_\theta^{(\text{ideal classifier})} = \sum_{i=1}^m \frac{1}{S_i} \left(\frac{\partial S_i}{\partial \theta}\right)^2$
- With a realistic classifier, only a fraction of all available signal events are selected, as well as some background events:  $y_i(\theta) = \epsilon_i S_i(\theta) + b_i$ 
  - here  $\epsilon_i$  is the local signal efficiency in bin  $x_i$
  - note that  $\frac{1}{y_i} = \rho_i \frac{1}{\epsilon_i S_i}$  where the local signal purity is defined as  $\rho_i = \frac{s_i}{s_i + b_i}$
  - the available information is therefore reduced to  $\mathcal{I}_\theta^{(\text{real classifier})} = \sum_{i=1}^m \epsilon_i \rho_i \times \frac{1}{S_i} \left(\frac{\partial S_i}{\partial \theta}\right)^2$
- In summary, with respect to an ideal classifier, a realistic classifier leads to a higher error on the fitted parameter,  $(\Delta\hat{\theta}^{(\text{real classifier})})^2 \geq \frac{1}{\text{IF}} (\Delta\hat{\theta}^{(\text{ideal classifier})})^2$
- “IF” is the “information fraction” available after cuts:

$$\text{IF} = \frac{\mathcal{I}_\theta^{(\text{real classifier})}}{\mathcal{I}_\theta^{(\text{ideal classifier})}} = \frac{\sum_{i=1}^m \epsilon_i \rho_i \times \frac{1}{S_i} \left(\frac{\partial S_i}{\partial \theta}\right)^2}{\sum_{i=1}^m \frac{1}{S_i} \left(\frac{\partial S_i}{\partial \theta}\right)^2}$$

# Information fraction vs. AUC

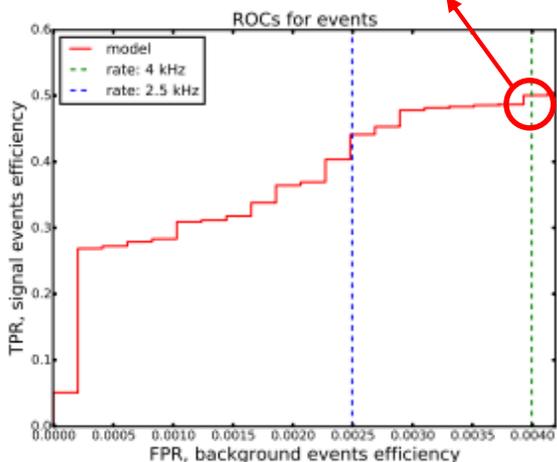
- “IF” is a figure of merit between 0 and 1 (like the AUC...) 
$$\text{IF} = \frac{\mathcal{I}_\theta^{(\text{real classifier})}}{\mathcal{I}_\theta^{(\text{ideal classifier})}} = \frac{\sum_{i=1}^m \epsilon_i \rho_i \times \frac{1}{S_i} \left( \frac{\partial S_i}{\partial \theta} \right)^2}{\sum_{i=1}^m \frac{1}{S_i} \left( \frac{\partial S_i}{\partial \theta} \right)^2}$$
  - it depends on efficiency and purity (PRC rather than ROC)
    - True Negatives are irrelevant...
  - it depends on local efficiencies and purities
    - but also applies to counting experiments (1 single “bin”) – see examples
  - it depends on the choice of a point on the PRC/ROC (a threshold on D)
    - but one can also use it in a fit to the full distribution of D – see examples
  - it is qualitatively (higher is better) and quantitatively ( $\Delta \hat{\theta} \sim 1/\text{IF}$ ) relevant
- *A different figure of merit is needed for every different problem!*
  - I derived this for statistical errors in parameter fits (precision measurements)
  - A similar f.o.m. can certainly be derived for optimizing searches
    - “combining” the different bins of the distribution is done slightly differently...
  - Systematic errors need to be handled differently...

# Systematic errors

- Statistical errors  $\propto \frac{1}{\sqrt{N}}$   $\rightarrow$  systematics become more relevant as N grows
  - Minimise statistical errors at low N  $\rightarrow$  only depends on  $\epsilon_s, \rho$
  - Minimise stat+syst errors at high N  $\rightarrow$  also depends on luminosity scale ( $S_{\text{tot}}$ )
    - i.e. need all three numbers TP, FP, FN  $\rightarrow$  but TN remains irrelevant
- Simple example  $\rightarrow$  measure  $\sigma_s$  by counting, 1% relative uncertainty in  $\sigma_b$ 
  - systematic error is lower than statistical error if  $\left(\frac{1-\rho}{\sqrt{\rho}}\right) \leq \frac{1}{\sqrt{\epsilon_s S_{\text{tot}}}} \times \frac{1}{\Delta\sigma_b/\sigma_b}$
  - optimizing total systematic + statistical error is a tradeoff involving  $\epsilon_s, \rho, S_{\text{tot}}$
- Complex problem, no universal recipe  $\rightarrow$  interesting problem to work on!
  - more in-depth discussion is *beyond the scope of this talk*

# Trigger

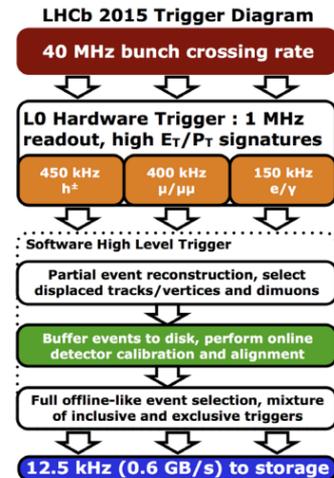
Maximise  $\epsilon_s$  at 4 kHz



T. Likhomanenko et al., *LHCb Topological Trigger Reoptimization*, Proc. CHEP 2015, J. Phys. Conf. Series 664 (2015) 082025. doi:10.1088/1742-6596/664/8/082025

**Figure 2.** Trigger events ROC curve. An output rate of 2.5 kHz corresponds to an FPR of 0.25%, 4 kHz — 0.4%. Thus to find the signal efficiency for a 2.5 kHz output rate, we take 0.25% background efficiency and find the point on the ROC curve that corresponds to this FPR.

IIUC, 4kHz is  $\epsilon_b$  (FPR) = 0.4% of 1 MHz L0 hw rate



F. Dordei, *LHCb detector and trigger performance in Run II*, Proc. 5th Int. Conf. on New Frontiers in Physics (IC-NFP 2016), EPJ Web of Conferences 164, 01016 (2017). doi:10.1051/epjconf/201716401016

- Different meaning of absolute numbers in the confusion matrix
  - Trigger → events per unit time i.e. trigger rates
  - (Physics analyses → total event sample sizes i.e. total integrated luminosities)
- Binary classifier optimisation goal: maximise  $\epsilon_s$  for a given  $B_{sel}$  per unit time
  - i.e. maximise  $TP/(TP+FN)$  for a given  $FP \rightarrow TN$  irrelevant
- Relevant plot →  $\epsilon_s$  vs.  $B_{sel}$  per unit time (i.e.  $TPR$  vs  $FP$ )
  - ROC curve ( $TPR$  vs.  $FPR$ ) confusing and irrelevant
  - e.g. maximise  $\epsilon_s$  for 4 kHz trigger rate, whether L0 rate is 1 MHz or 2MHz

# Event selection in HEP searches

- Statistical error in searches by counting experiment → “significance”
  - several metrics → but optimization always involves  $\epsilon_s$ ,  $\rho$  alone → TN irrelevant

$$Z_0 = \frac{S_{\text{sel}}}{\sqrt{S_{\text{sel}} + B_{\text{sel}}}} \implies (Z_0)^2 = S_{\text{tot}} \epsilon_s \rho$$

$Z_0$  – Not recommended? (confuses search with measuring  $\sigma_s$  once signal established)

C. Adam-Bourdarios et al., *The Higgs Machine Learning Challenge*, Proc. NIPS 2014 Workshop on High-Energy Physics and Machine Learning (HEPML2014), Montreal, Canada, PMLR 42 (2015) 19. <http://proceedings.mlr.press/v42/cowa14.html>

$Z_2$  – Most appropriate? (also used as “AMS2” in Higgs ML challenge)

$$Z_2 = \sqrt{2 \left( (S_{\text{sel}} + B_{\text{sel}}) \log\left(1 + \frac{S_{\text{sel}}}{B_{\text{sel}}}\right) - S_{\text{sel}} \right)} \implies (Z_2)^2 = 2S_{\text{tot}} \epsilon_s \left( \frac{1}{\rho} \log\left(\frac{1}{1-\rho}\right) - 1 \right) = S_{\text{tot}} \epsilon_s \rho \left( 1 + \frac{2}{3}\rho + \mathcal{O}(\rho^2) \right)$$

$$Z_3 = \frac{S_{\text{sel}}}{\sqrt{B_{\text{sel}}}} \implies (Z_3)^2 = S_{\text{tot}} \epsilon_s \frac{\rho}{1-\rho} = S_{\text{tot}} \epsilon_s \rho (1 + \rho + \mathcal{O}(\rho^2))$$

$Z_3$  (“AMS3” in Higgs ML) – Most widely used, but strictly valid only as an approximation of  $Z_2$  as an expansion in  $S_{\text{sel}}/B_{\text{sel}} \ll 1$ ?

$$\frac{S_{\text{sel}}}{B_{\text{sel}}} = \frac{\rho}{1-\rho} = \rho (1 + \rho + \mathcal{O}(\rho^2))$$

Expansion in  $\rho \ll 1$ ? – use the expression for  $Z_2$  if anything

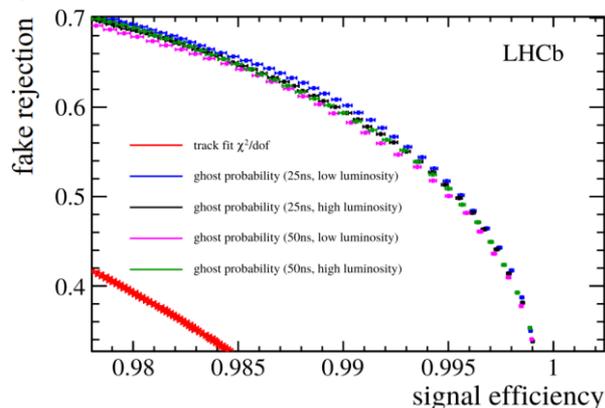
G. Punzi, *Sensitivity of searches for new signals and its optimization*, Proc. PhyStat2003, Stanford, USA (2003). [arXiv:physics/0308063v2](https://arxiv.org/abs/physics/0308063v2) [physics.data-an]  
 G. Cowan, E. Gross, *Discovery significance with statistical uncertainty in the background estimate*, ATLAS Statistics Forum (2008, unpublished). <http://www.pp.rhul.ac.uk/~cowan/stat/notes/SigCalcNote.pdf> (accessed 15 January 2018)

R. D. Cousins, J. T. Linnemann, J. Tucker, *Evaluation of three methods for calculating statistical significance when incorporating a systematic uncertainty into a test of the background-only hypothesis for a Poisson process*, Nucl. Instr. Meth. Phys. Res. A 595 (2008) 480. doi:10.1016/j.nima.2008.07.086  
 G. Cowan, K. Cranmer, E. Gross, O. Vitells, *Asymptotic formulae for likelihood-based tests of new physics*, Eur. Phys. J. C 71 (2011) 15. doi:10.1140/epjc/s10052-011-1554-0

- Several other interesting open questions → *beyond the scope of this talk*
  - optimization of systematics? → e.g. see AMS1 in Higgs ML challenge
  - predict significance in a binned fit? → integral over  $Z^2$  (=sum of log likelihoods)?

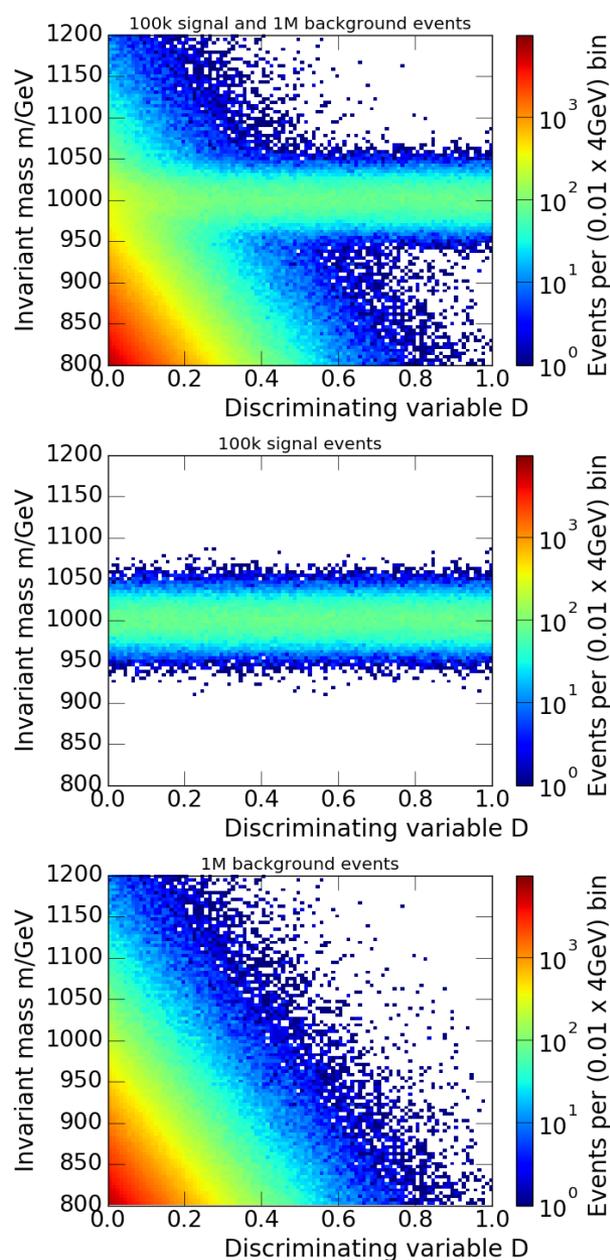
# Tracking and particle-ID

- ROCs irrelevant in event selection → but relevant in other HEP problems
- Event reconstruction and particle identification
  - Binary classifiers on a set of components of one event → not on a set of events
- Example: fake track rejection in LHCb
  - data set within one event: “track” objects created by the tracking software
    - True Positives: tracks that correspond to a charged particle trajectory in MC truth
    - True Negatives: tracks with no MC truth counterpart → relevant and well defined
- Binary classifier evaluation:  $\varepsilon_s$  and  $\varepsilon_b$  both relevant → ROC curve relevant
  - is AUC relevant? maximise physics performance? what if ROC curves cross?
  - these questions are *beyond the scope of this talk*



M. De Cian, S. Farry, P. Seyfert, S. Stahl, *Fast neural-net based fake track rejection in the LHCb reconstruction*, LHCb Public Note LHCb-PUB-2017-011 (2017).  
<https://cds.cern.ch/record/2255039>

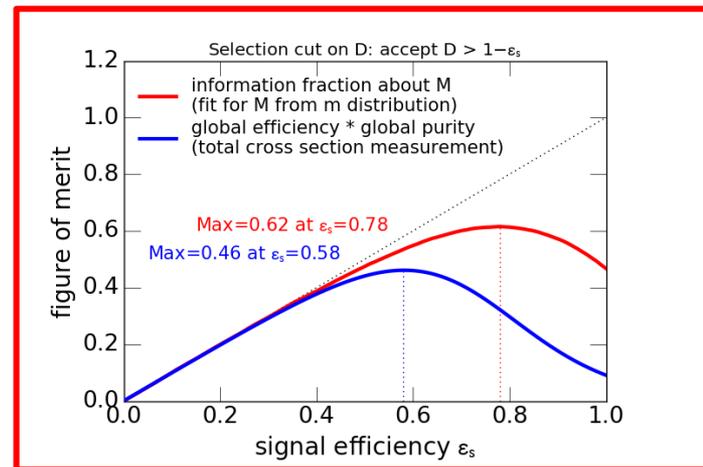
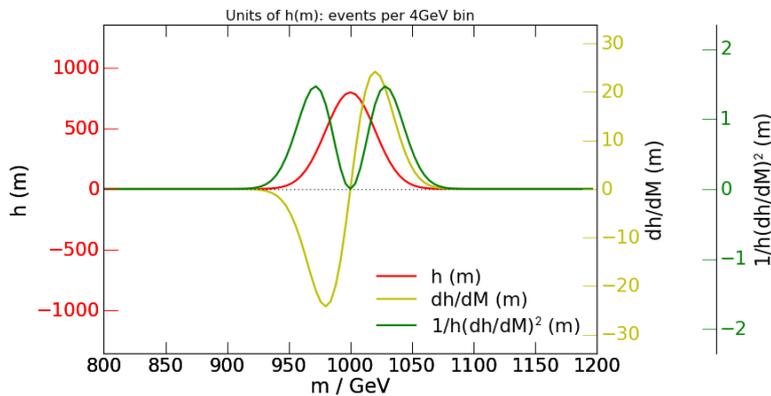
# Simple toy model



- Two independent observables  $\rightarrow f(m,D)=g(D)*h(m)$ 
  - discriminating variable  $D \rightarrow$  scoring classifier
  - invariant mass  $m \rightarrow$  used to fit signal mass  $M$
- Signal ( $\chi S=100$  fb): Gaussian peak in  $m$ , flat in  $D$ 
  - mass  $M=1000$  GeV, width  $W=20$  GeV
  - flat in  $D \rightarrow \epsilon_s=1-D_{\text{thr}}$  if accept events with  $D>D_{\text{thr}}$
- Background ( $\chi S=1000$  fb): exponential in both  $m$  and  $D$ 
  - cross-section 1000 fb  $\rightarrow B_{\text{tot}}=100k$
- Two measurements ( $\text{lumi}=100 \text{ fb}^{-1} \rightarrow S_{\text{tot}}=10k, B_{\text{tot}}=100k$ )
  - mass fit  $\rightarrow$  estimate  $\hat{M}$  (assuming  $\chi S, W$ )
  - cross section fit  $\rightarrow$  estimate  $\hat{\chi S}$  (assuming  $M, W$ )
  - counting, 1D and 2D fits, with/without cuts on  $D$
- Compare binary classifier to ideal case (no bkg):
  - ideal case  $\rightarrow \Delta \hat{M} = W/\sqrt{S_{\text{tot}}} = 0.200 \text{ GeV}$
  - ideal case  $\rightarrow \Delta \hat{\chi S} = \chi S/\sqrt{S_{\text{tot}}} = 1.00 \text{ fb}$

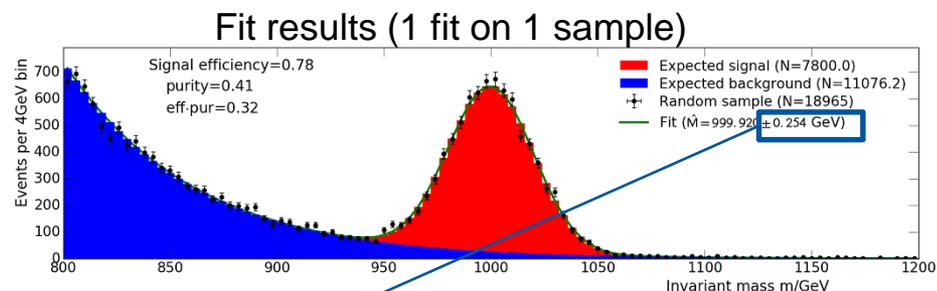
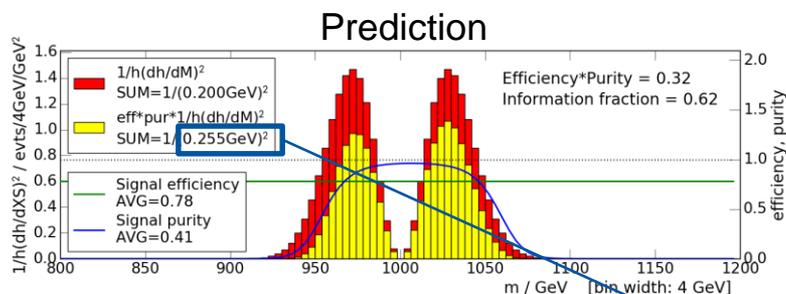
# M by 1D fit to m – optimizing the classifier

- Goal: fit true mass  $M$  from invariant mass  $m$  distribution after a cut on  $D$ 
  - Vary  $\epsilon_s = 1 - D_{\text{thr}}$  by varying cut  $D_{\text{thr}} \rightarrow$  compute information fraction on  $M$  for  $\epsilon_s \rightarrow$  maximum of information fraction:  $IF = 0.62$  ( $\Delta\hat{M} = 0.254 = \frac{0.200}{\sqrt{0.62}}$ ) at  $\epsilon_s = 0.78$
- Different measurements  $\rightarrow$  different metrics  $\rightarrow$  different optimizations
  - maximum of information for fit to  $M \rightarrow IF = 0.62$  ( $\Delta\hat{M} = 0.254 = \frac{0.200}{\sqrt{0.62}}$ ) at  $\epsilon_s = 0.78$
  - maximum of information for XS by counting  $\rightarrow \epsilon_s * \rho = 0.46$  at  $\epsilon_s = 0.58$
- To compute IF as sum over bins  $\rightarrow$  need average  $\frac{1}{h} \frac{\partial h}{\partial M}$  in each bin
  - proof-of-concept  $\rightarrow$  integrate by toy MC with event-by-event weight derivatives

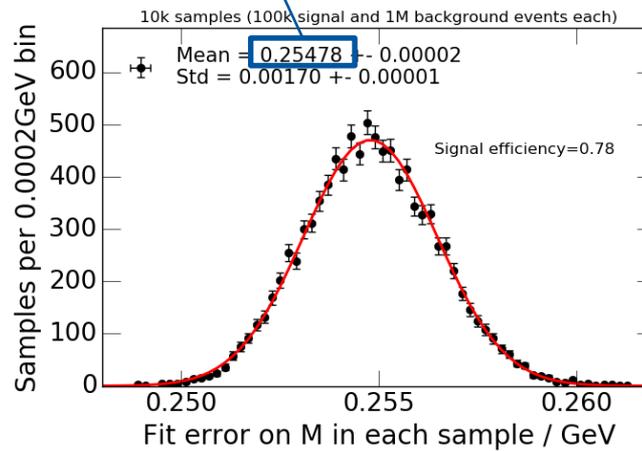
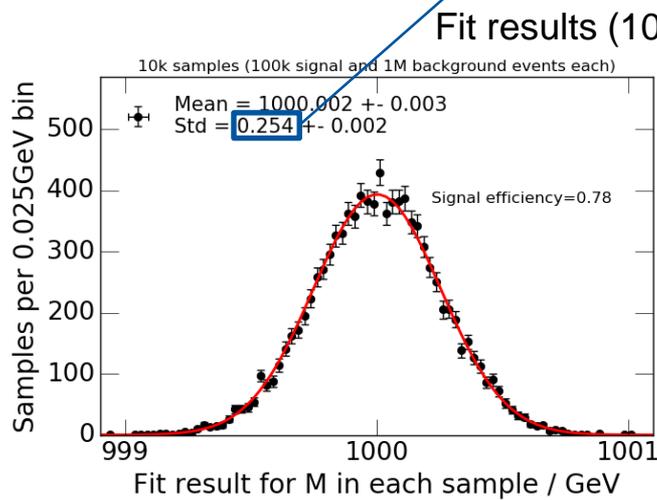


# M by 1D fit to m – cross-check

- Cross-check fit error returned by iminuit → repeat fit on 10k samples
  - check this only at the point of max information →  $\epsilon_s=0.78$  and  $\Delta\hat{M}=0.254$



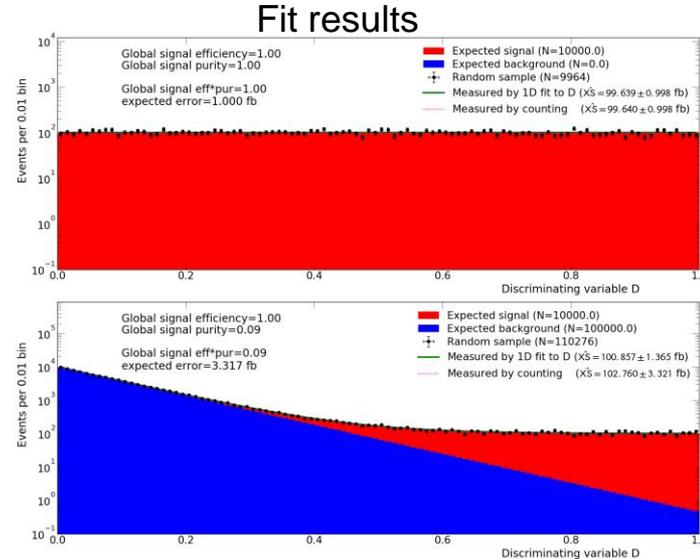
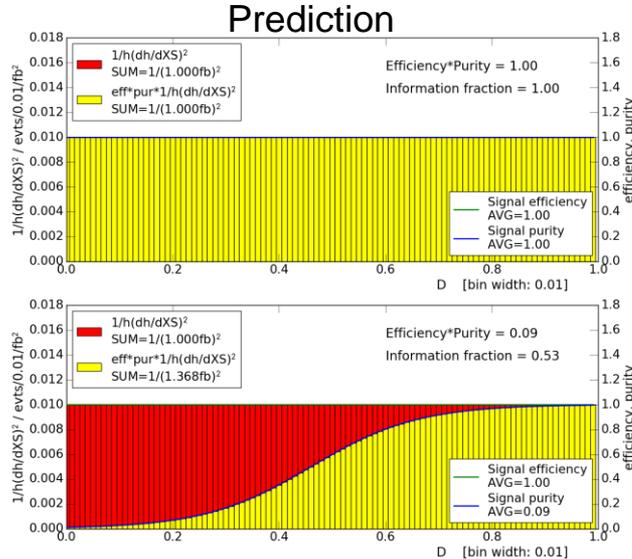
**OK!  $\Delta\hat{M}=0.254$  consistently**



# Cross-section by 1D fit to D

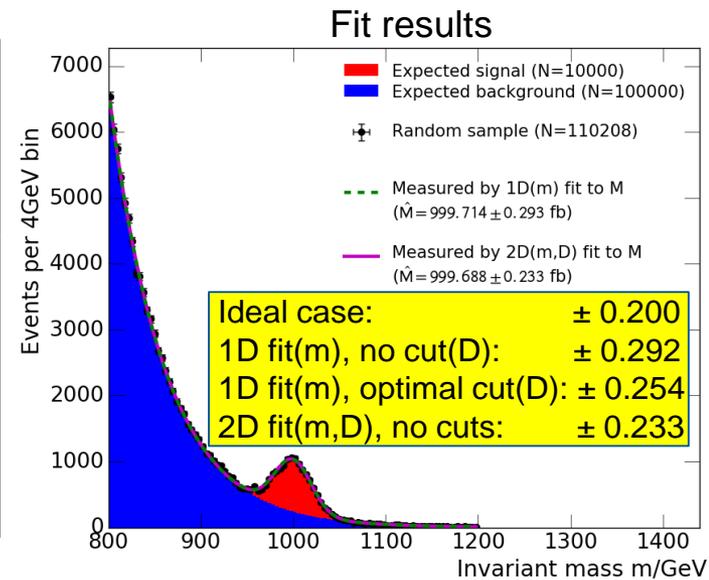
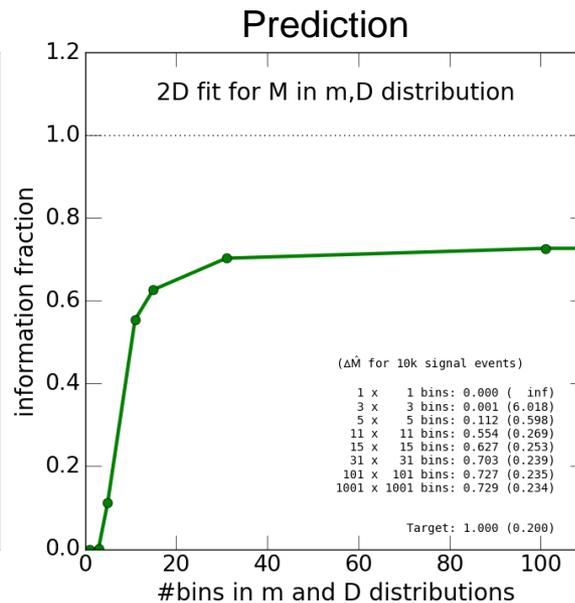
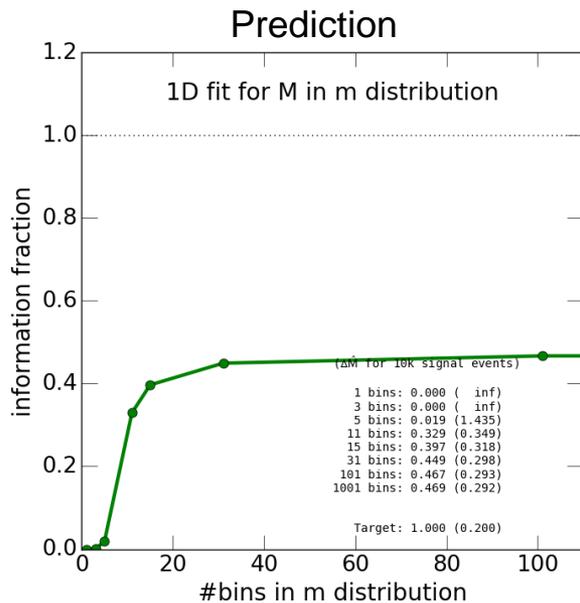
*i.e. the common practice of “BDT fits”*

- Cross-section fits analogous to mass fits but simpler
  - Differential cross-section proportional to total cross-section
  - $\frac{1}{s_i} \frac{\partial s_i}{\partial \sigma_s} = \frac{1}{\sigma_s}$  is constant  $\rightarrow \sum_i \frac{1}{s_i} \left( \frac{\partial s_i}{\partial \sigma_s} \right)^2 * \epsilon_i * \rho_i = \sum_i s_i * \epsilon_i * \rho_i$ 
    - special case : for a single bin (counting experiment)  $S_{\text{tot}} * \epsilon * \rho \rightarrow$  maximise global  $\epsilon * \rho$
- For simplicity show only fit in D (could fit m, or m and D) and no cuts
  - binning improves precision, also without cuts on D
  - use the scoring classifier D to partition data, not to reject events  $\rightarrow$  next slides



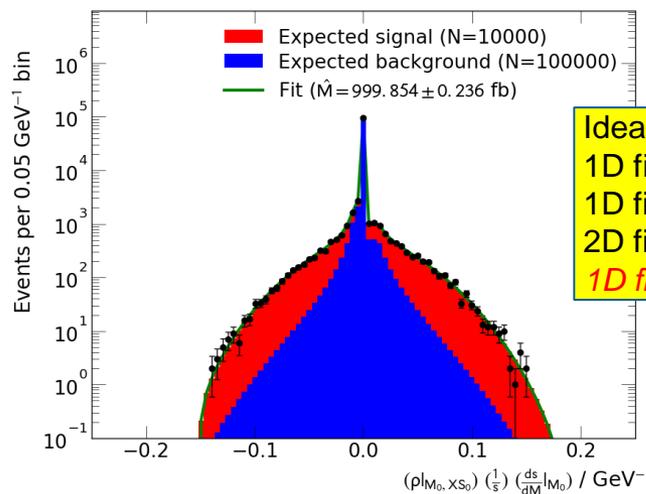
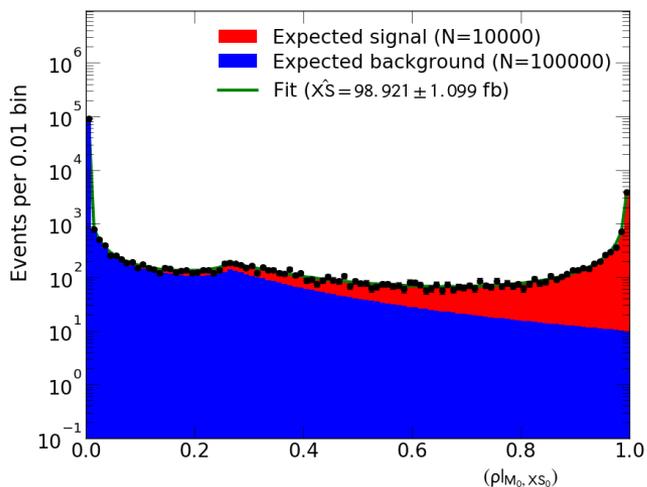
# M by 2D fit – use classifier to partition, not to cut

- Showed a fit for M on m, after a cut on D → can also fit in 2-D with no cuts
  - again, use the scoring classifier D to partition data, not to reject events
- Why is binning so important, especially using a discriminating variable?
  - next slide...



# Optimal partitioning – optimal variables

- How to partition the data into bins of equal  $\rho_i \frac{1}{s_i} \frac{\partial s_i}{\partial \sigma_s}$  ?
  - as a proof of concept → also made a 1D fit for  $\tilde{M}$  against this one variable “q”
  - not surprisingly, the precision is the same as that of the 2D fit on  $m, D$



Ideal case:	$\pm 0.200$
1D fit(m), no cut(D):	$\pm 0.292$
1D fit(m), optimal cut(D):	$\pm 0.254$
2D fit(m,D), no cuts:	$\pm 0.233$
1D fit(optimal q):	$\pm 0.236$

- In practice: train one ML variable to reproduce  $\frac{1}{s_i} \frac{\partial s_i}{\partial \sigma_s}$  ?

- Same general idea as the “optimal observables” technique

M. Davier, L. Duflot, F. LeDiberder, A. Rougé, *The optimal method for the measurement of tau polarization*, Phys. Lett. B 306 (1993) 411. doi:10.1016/0370-2693(93)90101-M  
M. Diel, O. Nachtmann, *Optimal observables for the measurement of three-gauge-boson couplings in  $e^+e^- \rightarrow W^+W^-$* , Z. Phys. C 62 (1994) 397. doi:10.1007/BF01555899  
O. Nachtmann, F. Nagel, *Optimal observables and phase-space ambiguities*, Eur. Phys. J. C40 (2005) 497. doi:10.1140/epjc/s2005-02153-9

# OLDER SLIDES

# HEP event selection properties

- Binary classifier optimisation goal: maximise physics reach at given budget
  - Trigger and computing → maximise signal event throughput within constraints
  - Physics analyses → maximise physics information from available data sets
- I will attempt a systematic analysis of properties:
  - 1. Qualitative class imbalance → signal relevant, background irrelevant
    - TN irrelevant and ill-defined (preselection, generator cuts) → only TP, FP, FN matter
  - 2. Extreme quantitative class imbalance → signal events swamped in background
  - 3. Prevalence largely constant in time → fixed by quantum physics cross section
    - Prevalence: known in advance for precision measurements; unknown for searches.
  - 4. Scale invariance (with two exceptions) → optimization based on 2 ratios  $\epsilon_s$ ,  $\rho$ 
    - Exception: trigger rate → constraint on throughput of FP(+TP) per unit time
    - Exception: total error (statistical + systematic) minimization also depends on scale L
  - 5. Fits to differential distributions → local  $\epsilon_s$ ,  $\rho$  relevant (global  $\epsilon_s$ ,  $\rho$  ~irrelevant)
- More details and examples in the following slides

M. Sokolova, G. Lapalme, *A Systematic Analysis of Performance Measures for Classification Tasks*, Information Processing and Management 45 (2009) 427.  
[doi:10.1016/j.ipm.2009.03.002](https://doi.org/10.1016/j.ipm.2009.03.002)

# Medical diagnostics (1) – accuracy

- **Medical Diagnostics (MED)**  
does Mr. A. have cancer?

- Binary classifier optimisation goal: maximise “diagnostic accuracy”
  - not obvious: many different specific goals → many different possible definitions
    - patient’s perspective → minimise diagnostic impact and impact of no/wrong treatment
    - society’s perspective: ethical and economic → allocate healthcare with limited budget
    - physician’s perspective → get knowledge of patient’s condition, manage patient

H. Sox, S. Stern, D. Owens, H. L. Abrams, *Assessment of Diagnostic Technology in Health Care: Rationale, Methods, Problems, and Directions*, The National Academies Press (1989). doi:10.17226/1432

- Most popular metric: “accuracy”, or “probability of correct test result”:

$$ACC = \frac{TP + TN}{TP + TN + FP + FN} = \pi_s \times TPR + (1 - \pi_s) \times TNR$$

<b>True Positives (TP)</b> (correctly diagnosed as ill)	<b>False Positives (FP)</b> (truly healthy, but diagnosed as ill)
<b>False Negatives (FN)</b> (truly ill, but diagnosed as healthy)	<b>True Negatives (TN)</b> (correctly diagnosed as healthy)

X. H. Zhou, D. K. McClish, N. A. Obuchowski, *Statistical Methods in Diagnostic Medicine* (Wiley, 2002). doi:10.1002/9780470317082

where “prevalence” is  $\pi_s = \frac{S_{tot}}{S_{tot} + B_{tot}}$

- Symmetric → all patients important, both truly ill (TP) and truly healthy (TN)

# Medical diagnostics (2) – from ACC to ROC

- ACC metric → widely used in medical diagnostics in the 1980-'90s (still now?)
  - Also “by far the most commonly used metric” in ML in the 1990s
- Limitation: ACC depends on relative prevalence
  - issue for imbalanced problems → diagnostic accuracy for rare diseases
  - issue if prevalence unknown or variable over time → disease epidemics
- Since the '90s → shift from ACC to ROC in MED and ML fields
  - TPR (sensitivity) and TNR (specificity) studied separately
    - reminder: all patients important, both truly ill (TP) and truly healthy (TN)
- Evaluation often based on the AUC → two advantages *for medical diagnostics*:
  - AUC interpretation: “probability that test result of randomly chosen sick subject indicates greater suspicion than that of randomly chosen healthy subject”
  - ROC comparison without prior  $D_{thr}$  choice (prevalence-dependent  $D_{thr}$  choice)

F. J. Provost, T. Fawcett, R. Kohavi, *The Case against Accuracy Estimation for Comparing Induction Algorithms*, Proc. 15th Int. Conf. on Machine Learning (ICML '98), Madison, USA (1998). <https://www.researchgate.net/publication/2373067>

J. A. Swets, *Measuring the accuracy of diagnostic systems*, Science 240 (1988) 1285. [doi:10.1126/science.3287615](https://doi.org/10.1126/science.3287615)

L. B. Lusted, *Signal Detectability and Medical Decision-Making*, Science 171 (1971) 1217 [doi:10.1126/science.171.3977.1217](https://doi.org/10.1126/science.171.3977.1217)

A. P. Bradley, *The use of the area under the ROC curve in the evaluation of machine learning algorithms*, Pattern Recognition 30 (1997) 1145. [doi:10.1016/S0031-3203\(96\)00142-2](https://doi.org/10.1016/S0031-3203(96)00142-2)

J. A. Hanley, B. J. McNeil, *The meaning and use of the area under a receiver operating characteristic (ROC) curve*, Radiology 143 (1982) 29. [doi:10.1148/radiology.143.1.7063747](https://doi.org/10.1148/radiology.143.1.7063747)

# Medical diagnostics (3) – from ROC to PRC?

- ROC and AUC metrics → currently widely used in medical diagnostics and ML
- Limitation: ROC-based evaluation questionable for *highly imbalanced data sets*
  - ROC may provide an overly optimistic view of performance with highly skewed data sets
- PRC may provide a more informative assessment of performance in this case
  - PRC-based reanalysis of some data sets in life sciences has been performed
- Very active area of research → other options proposed (CROC, cost models...)
  - Take-away message: ROC and AUC not always the appropriate solutions

J. Davis, M. Goadrich, *The relationship between Precision-Recall and ROC curves*, Proc. 23rd Int. Conf. on Machine Learning (ICML '06), Pittsburgh, USA (2006). [doi:10.1145/1143844.1143874](https://doi.org/10.1145/1143844.1143874)

C. Drummond, R. C. Holte, *Explicitly representing expected cost: an alternative to ROC representation*, Proc. 6th Int. Conf. on Knowledge Discovery and Data Mining (KDD-00), Boston, USA (2000). [doi:10.1145/347090.347126](https://doi.org/10.1145/347090.347126)

D. J. Hand, *Measuring classifier performance: a coherent alternative to the area under the ROC curve*, Mach Learn (2009) 77: 103. [doi:10.1007/s10994-009-5119-5](https://doi.org/10.1007/s10994-009-5119-5)

S. J. Swamidass, C.-A. Azencott, K. Daily, P. Baldi, *A CROC stronger than ROC: measuring, visualizing and optimizing early retrieval*, Bioinformatics 26 (2010) 1348. [doi:10.1093/bioinformatics/btq140](https://doi.org/10.1093/bioinformatics/btq140)

D. Berrar, P. Flach, *Caveats and pitfalls of ROC analysis in clinical microarray research (and how to avoid them)*, Briefings in Bioinformatics 13 (2012) 83. [doi:10.1093/bib/bbr008](https://doi.org/10.1093/bib/bbr008)

H. He, E. A. Garcia, *Learning from Imbalanced Data*, IEEE Trans. Knowl. Data Eng. 21 (2009) 1263. [doi:10.1109/TKDE.2008.239](https://doi.org/10.1109/TKDE.2008.239)

T. Saito, M. Rehmsmeier, *The Precision-Recall Plot Is More Informative than the ROC Plot When Evaluating Binary Classifiers on Imbalanced Datasets*, PLoS One 10 (2015) e0118432. [doi:10.1371/journal.pone.0118432](https://doi.org/10.1371/journal.pone.0118432)

# Simplest HEP example – total cross-section

- Total cross-section measurement in a counting experiment
- To minimize statistical errors: *maximise efficiency\*purity  $\epsilon_s * \rho$* 
  - well-known since decades
  - *global efficiency  $\epsilon_s = S_{sel}/S_{tot}$  and global purity  $\rho = S_{sel}/(S_{sel} + B_{sel})$  – “1 single bin”*

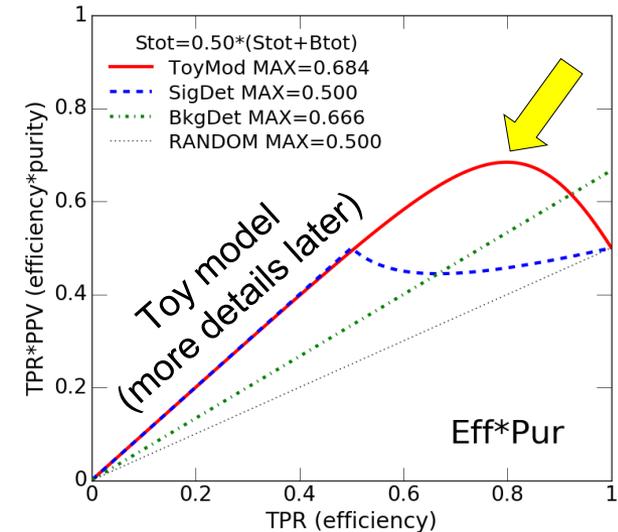
$$(\sigma_s)_{\text{meas}} = \frac{N_{\text{meas}} - \mathcal{L}\epsilon_b\sigma_b}{\mathcal{L}\epsilon_s}$$

$$\Delta\sigma_s = \frac{\Delta N_{\text{meas}}}{\mathcal{L}\epsilon_s} = \frac{1}{\mathcal{L}\epsilon_s} \sqrt{N_{\text{exp}}}$$

$$N_{\text{exp}} = S_{\text{sel}} + B_{\text{sel}} = \frac{S_{\text{sel}}}{\rho} = \frac{\mathcal{L}\sigma_s\epsilon_s}{\rho}$$

$$\frac{1}{(\Delta\sigma_s)^2} = \frac{1}{\sigma_s} \mathcal{L}\epsilon_s\rho = \frac{1}{\sigma_s^2} S_{\text{tot}}\epsilon_s\rho$$

- $\epsilon_s * \rho$ : metric between 0 and 1
  - qualitatively relevant (*only for this specific use case!*): the higher, the better
  - numerically: fraction of Fisher information ( $1/\text{error}^2$ ) available after selecting



# Predict and optimize statistical errors in binned fits

- Observed data: event counts  $n_i$  in  $m$  bins of a (multi-D) distribution  $f(x)$ 
  - expected counts  $y_i = \int f(x_i, \theta) dx \rightarrow$  depend on a parameter  $\theta$  that we want to fit
  - [NB here  $f$  is a differential cross section, it is not normalized to 1 like a pdf]

- Easy to show (backup slides) that minimum variance achievable is:

$$\boxed{(\Delta\hat{\theta})^2 = \text{var}(\hat{\theta}) \geq \frac{1}{\mathcal{I}_\theta}} \text{ (Cramer-Rao lower bound), where } \boxed{\mathcal{I}_\theta = \sum_{i=1}^m \frac{1}{y_i} \left( \frac{\partial y_i}{\partial \theta} \right)^2 = \int \frac{1}{f} \left( \frac{\partial f}{\partial \theta} \right)^2 dx} \text{ (Fisher information)}$$

- With an ideal classifier (or no background)  $\rightarrow y_i = S_i$  and  $\boxed{\mathcal{I}_\theta^{(\text{ideal classifier})} = \sum_{i=1}^m \frac{1}{S_i} \left( \frac{\partial S_i}{\partial \theta} \right)^2}$

- With a realistic classifier  $\rightarrow \boxed{y_i(\theta) = \epsilon_i S_i(\theta) + b_i}$  and  $\boxed{\mathcal{I}_\theta^{(\text{real classifier})} = \sum_{i=1}^m \epsilon_i \rho_i \times \frac{1}{S_i} \left( \frac{\partial S_i}{\partial \theta} \right)^2}$

–  $\epsilon_i$  and  $\rho_i \rightarrow$  local signal efficiency and purity in the  $i^{\text{th}}$  bin

- Binary classifier optimization  $\rightarrow$  maximise
  - higher is better

– interpretation:  $\boxed{(\Delta\hat{\theta}^{(\text{real classifier})})^2 \geq \frac{1}{\text{IF}} (\Delta\hat{\theta}^{(\text{ideal classifier})})^2}$

$$\boxed{\text{IF} = \frac{\mathcal{I}_\theta^{(\text{real classifier})}}{\mathcal{I}_\theta^{(\text{ideal classifier})}} = \frac{\sum_{i=1}^m \epsilon_i \rho_i \times \frac{1}{S_i} \left( \frac{\partial S_i}{\partial \theta} \right)^2}{\sum_{i=1}^m \frac{1}{S_i} \left( \frac{\partial S_i}{\partial \theta} \right)^2}}$$

# Optimal partitioning – information inflow

- Information about  $\theta$  in a binned fit  $\rightarrow \mathcal{I}_\theta = \sum_{i=1}^m \frac{1}{y_i} \left( \frac{\partial y_i}{\partial \theta} \right)^2$
- Do I gain anything by splitting bin  $y_i$  into two separate bins?  $y_i = w_i + z_i$ 
  - i.e. is the “information inflow” positive? A. van den Bos, *Parameter Estimation for Scientists and Engineers* (Wiley, 2007).

$$\frac{1}{w_i} \left( \frac{\partial w_i}{\partial \theta} \right)^2 + \frac{1}{z_i} \left( \frac{\partial z_i}{\partial \theta} \right)^2 - \frac{1}{w_i + z_i} \left( \frac{\partial (w_i + z_i)}{\partial \theta} \right)^2 = \frac{(w_i \frac{\partial z_i}{\partial \theta} - z_i \frac{\partial w_i}{\partial \theta})^2}{w_i z_i (w_i + z_i)} \geq 0$$
  - information increases (errors on parameters decrease) if  $\frac{1}{w_i} \frac{\partial w_i}{\partial \theta} \neq \frac{1}{z_i} \frac{\partial z_i}{\partial \theta}$
- Both  $w_i$  and  $z_i$  can be written as  $f = \epsilon s + b = \frac{\epsilon s}{\rho} \rightarrow \frac{\partial f}{\partial \theta} = \epsilon \frac{\partial s}{\partial \theta} \rightarrow \frac{1}{f} \frac{\partial f}{\partial \theta} = \rho \frac{1}{s} \frac{\partial s}{\partial \theta}$
- In summary: **try to partition the data into bins of equal  $\rho_i \frac{1}{s_i} \frac{\partial s_i}{\partial \sigma_s}$** 
  - for cross-section measurements (and searches?): split into bins of equal  $\rho_i$
  - **“use the scoring classifier  $D$  to partition the data, not to reject events”**
    - the BDT normally tries to represent a signal likelihood – i.e. ultimately the real  $\rho_i$