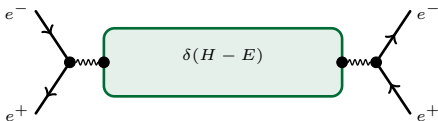
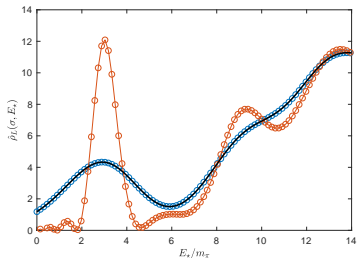


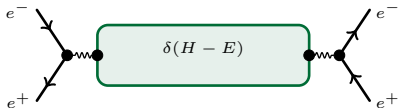
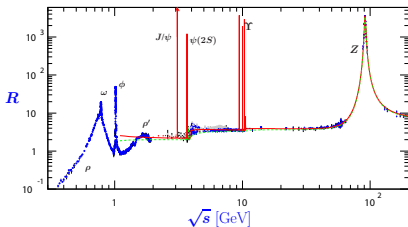
- why hadronic spectral densities
- the problem with lattice correlators: euclidean time, finite volume, statistical and systematic errors
- recovering the physical information from smeared spectral densities
- how to extract smeared spectral densities from noisy measurements
- examples in the case of a benchmark model
- examples in the case of a true lattice data
- conclusions and outlooks



- hadronic spectral densities are central objects in the calculations of physical observables associated with the continuum spectrum of the QCD Hamiltonian
- a notable classical example is the so-called R -ratio, i.e. the ratio of the differential cross-section for $e^+e^- \mapsto \text{hadrons}$ over the corresponding quantity for $e^+e^- \mapsto \mu^+\mu^-$

$$R \propto \underbrace{\langle 0 | J_{em}^k(0) \delta(H - E) \delta^3(\mathbf{P}) J_{em}^k(0) | 0 \rangle}_{\rho(E)}$$

- other important examples are hadronic τ decays, the flavour-changing non-leptonic decay-rates of kaons and heavy flavoured mesons, the deep inelastic scattering cross-section, and thermodynamic observables arising in the study of QCD at finite-temperature and of the quark-gluon plasma, etc.



- first-principles model-independent calculations of hadronic spectral densities can in principle be performed by recurring to non-perturbative lattice techniques
- the primary observables in a lattice calculations are **euclidean time-ordered correlators at discrete values of the coordinates and on a finite volume**

$$C(t) = \frac{1}{L^3} \sum_{\mathbf{x}} T \langle 0 | O(\mathbf{x}) \bar{O}(0) | 0 \rangle_L$$

- these can be rewritten in terms of the finite volume spectral densities

$$C(t) = \int_0^\infty dE \rho_L(E) e^{-tE} ,$$

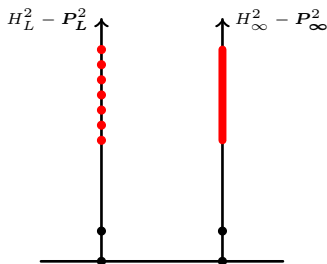
$$\rho_L(E) = \frac{1}{L^3} \sum_{\mathbf{x}} \langle 0 | O(0, \mathbf{x}) \delta(E - H_L) \bar{O}(0) | 0 \rangle_L$$

- now we see the problems:

$$C(t) = \int_0^\infty dE \rho_L(E) e^{-tE} + \delta C(t),$$

$$\begin{aligned} \rho_L(E) &= \frac{1}{L^3} \sum_{\mathbf{x}} \langle 0 | O(0, \mathbf{x}) \delta(E - H_L) \bar{O}(0) | 0 \rangle_L \\ &= \sum_n w_n(L) \delta(E - E_n(L)) \end{aligned}$$

- lattice correlators are unavoidably affected by **errors** and, in this case, the inverse Laplace-transform needed to extract the spectral densities becomes an **ill-posed numerical problem**
- even in the ideal case in which these can be computed exactly, **finite volume spectral densities cannot be associated with physical quantities**
- the finite volume hamiltonian has a discrete spectrum** and, consequently, the **finite volume spectral densities are distributions**, sums of isolated δ -function singularities



- in order to solve these problems one can conveniently consider **smeared spectral densities**

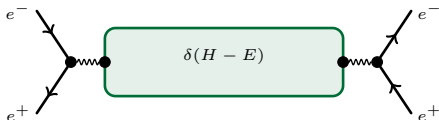
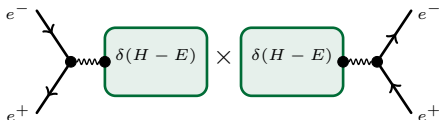
$$\hat{\rho}_L(\sigma, E_\star) = \int_0^\infty dE \Delta_\sigma(E_\star, E) \rho_L(E)$$

- the smearing function can be chosen to be peaked around E_\star and such that it becomes a Dirac δ -function when the smearing radius parameter σ is sent to zero
- smeared spectral densities are **smooth functions of the energy** and studying their infinite volume limit is a **well posed problem**; the physical information is recovered by taking the limits

$$\rho(E_\star) = \lim_{\sigma \rightarrow 0} \lim_{L \rightarrow \infty} \hat{\rho}_L(\sigma, E_\star)$$

in the specified order!

- notice that smeared spectral functions **must be introduced** in order to properly define cross-sections, this is the way we **avoid** the well-known issue of the **square of a δ -function** appearing at intermediate stages of the calculations



- in order to solve these problems one can conveniently consider **smeared spectral densities**

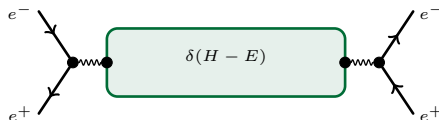
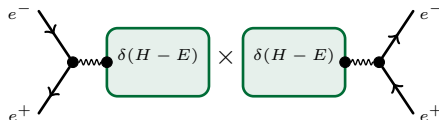
$$\hat{\rho}_L(\sigma, E_*) = \int_0^\infty dE \Delta_\sigma(E_*, E) \rho_L(E)$$

- the smearing function can be chosen to be peaked around E_* and such that it becomes a Dirac δ -function when the smearing radius parameter σ is sent to zero
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$$\rho(E_*) = \lim_{\sigma \rightarrow 0} \lim_{L \rightarrow \infty} \hat{\rho}_L(\sigma, E_*)$$

in the specified order!

- moreover, **experimental data** can be **smeared with the same function** used in the theoretical calculations



- m.t.hansen, h.b.meyer, d.robaina, PRD96 (2017) proposed to extract smeared spectral densities by using a classical method due to Backus and Gilbert (BG), g.backus, f.gilbert, Geophys.J.R.Astron.Soc.16 (1968)
- the central idea of BG is to search for a smearing function that lives in the space spanned by the **basis-functions** of the correlator

$$\Delta^{BG}(E_*, E) = \sum_{t=0}^{t_{max}} g_t(E_*) e^{-(t+1)E}$$

- once the coefficients $g_t(E_*)$ are known, the smeared spectral density is given by

$$C(t+1) = \int_0^\infty dE \rho_L(E) e^{-(t+1)E}$$

$$\begin{aligned} \hat{\rho}_L^{BG}(E_*) &= \sum_{t=0}^{t_{max}} g_t(E_*) C(t+1) \\ &= \int_0^\infty dE \rho_L(E) \Delta^{BG}(E_*, E) \end{aligned}$$

- in absence of errors on the correlator (an idealization), the coefficients $g_t(E_*)$ are obtained by minimizing the following functional

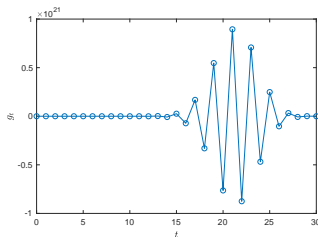
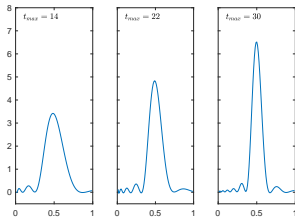
$$A_{BG}[g] = \int_0^\infty dE (E - E_*)^2 \left\{ \Delta^{BG}(E_*, E) \right\}^2$$

$$= \int_0^\infty dE (E - E_*)^2 \left\{ \sum_{t=0}^{t_{max}} g_t(E_*) e^{-(t+1)E} \right\}^2$$

under the unit-area constraint

$$\int_0^\infty dE \Delta^{BG}(E_*, E) = 1$$

- the width of the smearing function is **optimized** on the basis of the number of observations



- in the realistic case in which errors are present, the correlator has to be replaced with

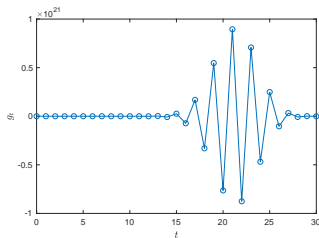
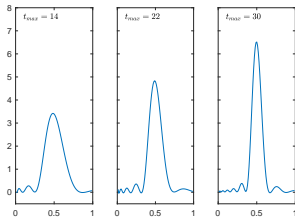
$$C_i(t) = \bar{C}(t) + \delta C_i(t), \quad i = 0, \dots, N - 1$$

- since **the coefficients are gigantic**, even a tiny deviation from the average is enormously amplified

$$\sum_{t=0}^{t_{max}} g_t(E_*) \delta C_i(t) \mapsto \infty$$

and **statistical errors also become gigantic**

- this is a manifestation of the fact that we are dealing here with a numerically ill-posed problem



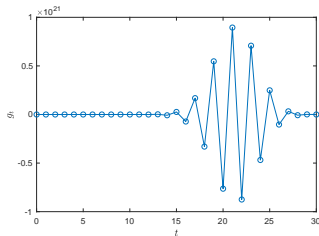
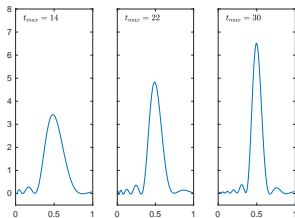
- the **very smart mechanism suggested by BG** to keep errors under control is to minimize the following functional

$$W[\lambda, g] = (1 - \lambda)A_{BG}[g] + \lambda B[g]$$

$$B[g] = \sum_{t,r=0}^{t_{max}} \text{Cov}_{tr} g_t(E_*) g_r(E_*)$$

$$\text{Cov}_{tr} = \frac{1}{N} \sum_{i=0}^{N-1} \delta C_i(t+1) \delta C_i(r+1)$$

- the presence of the error functional $B[g]$ **forbids solutions corresponding to gigantic values of the coefficients** and statistical errors are thus kept under control
- on the other hand, **the shape of the smearing function now depends**, in addition to the number of observations, also **on the associated errors**: this is a particularly unpleasant feature if the method has to be used in order to take the infinite volume limit
- moreover, there is **no natural way to set the trade-off parameter λ** , a part from trying to balance in a subjective way between resolution and errors



- we devised a method in which the target smearing function is an input of the procedure; in what follows

$$\Delta_{\sigma}(E_{\star}, E) = \frac{e^{-\frac{(E-E_{\star})^2}{2\sigma^2}}}{\int_0^{\infty} dE e^{-\frac{(E-E_{\star})^2}{2\sigma^2}}}$$

- the method searches for an optimal approximation of the target smearing function in the space of the basis functions

$$\bar{\Delta}_{\sigma}(E_{\star}, E) = \sum_{t=0}^{t_{max}} g_t(E_{\star}) e^{-(t+1)E}$$

- and again the coefficients are obtained by minimizing a convex combination of a deterministic and of the error functionals

$$W[\lambda, g] = (1 - \lambda)A[g] + \lambda \frac{B[g]}{C(0)^2}$$

under the unit area constraint

- but in our case the deterministic functional is a measure of the difference between the target and approximated smearing functions

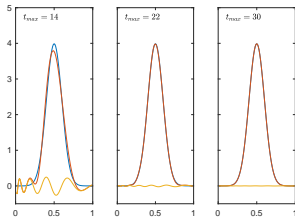
$$A[g] = \int_0^\infty dE |\bar{\Delta}_\sigma(E_\star, E) - \Delta_\sigma(E_\star, E)|^2$$

- but in our case the deterministic functional is a measure of the difference between the target and approximated smearing functions

$$A[g] = \int_0^\infty dE \left| \bar{\Delta}_\sigma(E_\star, E) - \Delta_\sigma(E_\star, E) \right|^2$$

- in absence of errors, our method is just a way to find an optimal polynomial approximation to a smooth function, $x = e^{-E}$

$$A[g] = \int_0^1 dx \left| \sum_{t=0}^{t_{max}} g_t x^t - \frac{\Delta_\sigma(E_\star, -\log(x))}{x} \right|^2$$



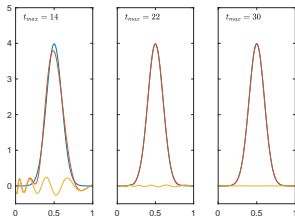
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- with our method, by increasing t_{max} the error in the approximation of the target smearing function can be made arbitrarily small



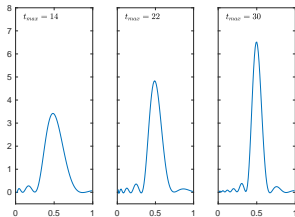
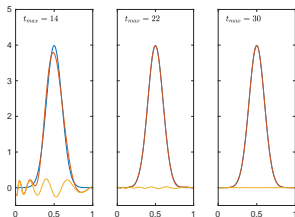
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- with our method, by increasing t_{max} the error in the approximation of the target smearing function can be made arbitrarily small
- this has to be compared with the BG method where by increasing t_{max} one gets a different (sharper) smearing function



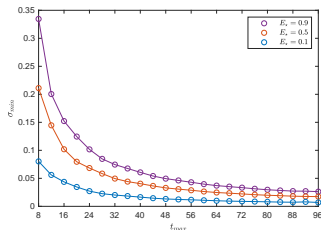
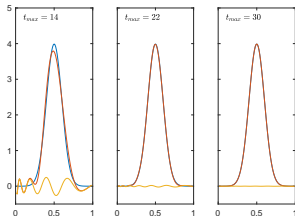
- furthermore, since at the end of the procedure the difference between the target and the approximated smearing function is known

$$\delta_\sigma(E_\star, E) = 1 - \frac{\bar{\Delta}_\sigma(E_\star, E)}{\Delta_\sigma(E_\star, E)}$$

- this information can be used in our method to estimate the systematic error on the estimated smeared spectral densities induced by this difference

$$\Delta^{bias} = \int_0^\infty dE \delta_\sigma(E_\star, E) \Delta_\sigma(E_\star, E) \rho_L(E)$$

$$\Delta^{syst} = |\delta_\sigma(E_\star, E_\star)| \hat{\rho}_L(\sigma, E_\star)$$



- furthermore, since at the end of the procedure the difference between the target and the approximated smearing function is known

$$\delta_\sigma(E_\star, E) = 1 - \frac{\tilde{\Delta}_\sigma(E_\star, E)}{\Delta_\sigma(E_\star, E)}$$

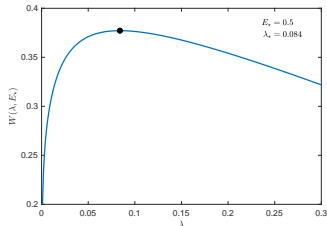
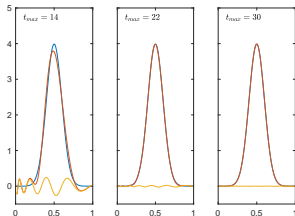
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$$\Delta^{syst} = |\delta_\sigma(E_\star, E_\star)| \hat{\rho}_L(\sigma, E_\star)$$

- finally, in our method there is a **natural way to set the trade-off parameter λ** by studying the functional $W[\lambda, E_\star]$ evaluated at the solution $g_\star(\lambda, E_\star)$ as a function of λ

$$\max_\lambda \left\{ (1 - \lambda)A[g_\star] + \lambda \frac{B[g_\star]}{C(0)^2} \right\} = W(\lambda_\star, E_\star)$$



- we have decided to test our method by using the same benchmark system previously proposed to test the BG method in the context of the extraction of hadronic spectral densities

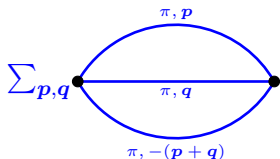
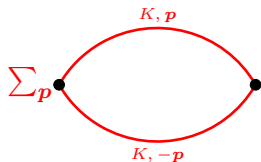
$$\mathcal{L}_{int}(x) = \frac{g_\pi}{6} \phi(x) \pi^3(x) + \frac{g_K m_\phi}{2} \phi(x) K^2(x),$$

$$3m_\pi < 2m_K < m_\phi$$

- we have considered a correlator having as finite volume spectral density

$$\rho_L(E) = \frac{g_K^2 m_\phi^2}{2(m_\pi L)^3} \sum_{\mathbf{p}} \frac{\delta(E - 2E_K(\mathbf{p}))}{4E_K^2(\mathbf{p})} + \frac{g_\pi^2}{48m_\pi^3 L^6} \sum_{\mathbf{p}, \mathbf{q}} \frac{\delta(E - E_\pi(\mathbf{p}) - E_\pi(\mathbf{q}) - E_\pi(\mathbf{p} + \mathbf{q}))}{E_\pi(\mathbf{p}) E_\pi(\mathbf{q}) E_\pi(\mathbf{p} + \mathbf{q})}$$

m.t.hansen, h.b.meyer, d.robaina, PRD96 (2017)
m.hansen, a.lupo, n.t. arXiv:1903.06476



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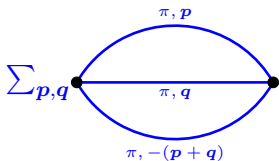
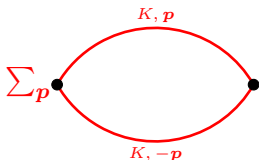
- that in the infinite volume limit becomes

$$\rho(E) = \frac{g_K^2 m_\phi^2}{32\pi^2 m_\pi^3} \sqrt{1 - \frac{4m_K^2}{m_\phi^2} \theta(E - 2m_K)}$$

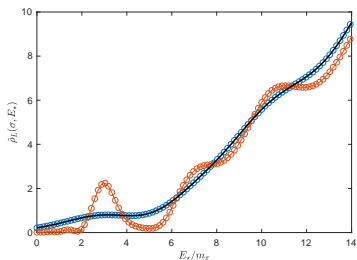
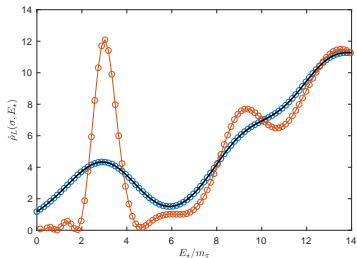
$$+ \frac{g_\pi^2}{3072\pi^4 m_\pi} \left(\frac{E}{m_\pi}\right)^2 \mathcal{F}\left(\frac{E}{m_\pi}\right) \theta(E - 3m_\pi)$$

$$\mathcal{F}(x) =$$

$$\frac{2}{x^4} \int_4^{(x-1)^2} dy \sqrt{(y-4) \left[\frac{(x^2-1)^2}{y} - 2(x^2+1) + y \right]}$$

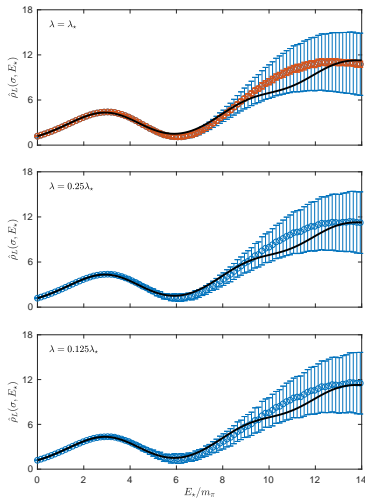


- the plots show the results obtained by using our method and the ones obtained by using the BG method
- both plots have been obtained by setting $\sigma = 0.1$ and $t_{max} = 30$; the one on the top corresponds to $L = 24$ while the one on the bottom to $L = 32$
- the blue points, obtained with our method, are in perfect agreement with the expected result that in this case is known exactly
- in the case of the BG (orange points) the smearing function is an output of the procedure, it can only be controlled by changing t_{max} and, moreover, it is different at different values of E_*



- the plots have been obtained by using our method on the volume $L = 24$ with $t_{max} = 30$ and $\sigma = 0.1$
- having a reliable estimate of the systematic errors, the results must be compatible at different values of λ within the total uncertainties

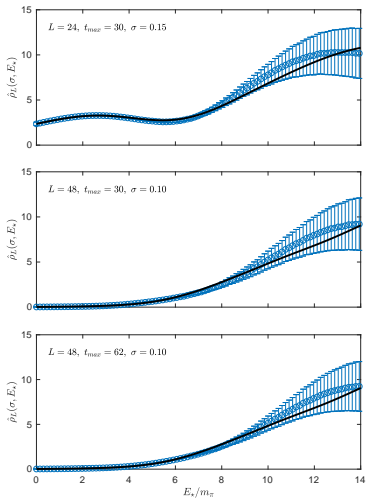
$$W[\lambda, g] = (1 - \lambda)A[g] + \lambda \frac{B[g]}{C(0)^2}$$



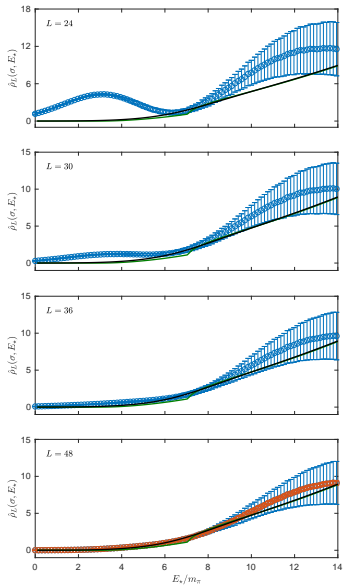
- when the smeared spectral density is smoother, either because the smearing radius is larger or because the volume is larger, the reconstruction works much better
- in these cases using

$$\Delta^{syst} = |\delta_{\sigma}(E_{\star}, E_{\star})| \hat{\rho}_L(\sigma, E_{\star})$$

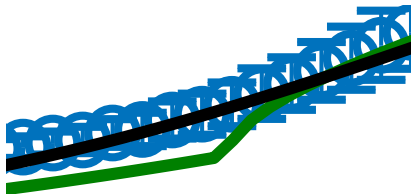
provides a **very conservative estimate of the systematic errors**



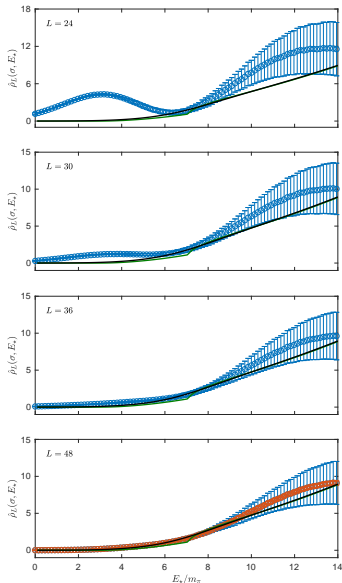
- the plots, obtained with $\sigma = 0.1$ and $t_{max} = 31$, show the approach to the infinite volume limit of the estimated smeared spectral functions

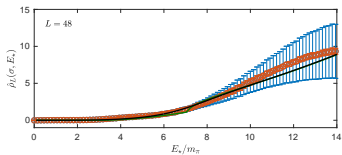
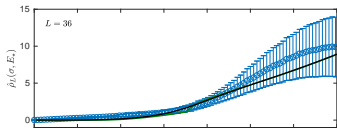
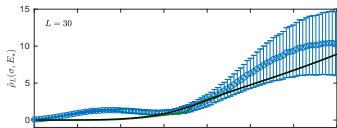
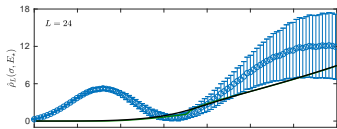
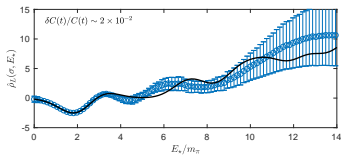
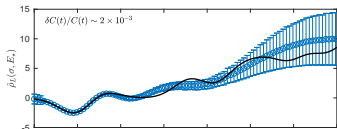
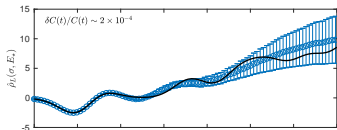
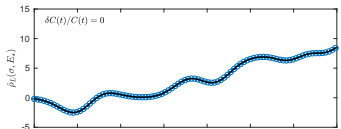


- the plots, obtained with $\sigma = 0.1$ and $t_{max} = 31$, show the approach to the infinite volume limit of the estimated smeared spectral functions
- the green curve is the exact infinite volume spectral density: this is a continuous function of the energy but has a cusp in correspondence of the two-kaons threshold
- in the infinite volume limit the data have to reproduce the black curve, the exact infinite volume smeared spectral density: this is a smooth curve



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- in the infinite volume limit the data have to reproduce the black curve, the exact infinite volume smeared spectral density: this is a smooth curve
- this already happens at $L = 36$ and the agreement is remarkably good (at the level of the statistical errors) at $L = 48$
- as already noticed, experimental data can be smeared with the same smearing function used in the theoretical calculations so that the results can directly be compared with measurements



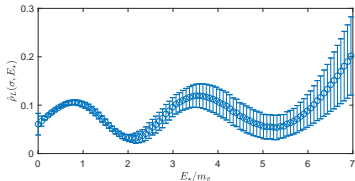
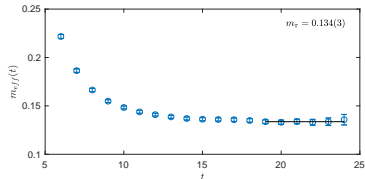


- we have applied our method to true lattice data in the case of a QCD pseudoscalar-pseudoscalar correlator

$$C_{\text{QCD}}(t) = \frac{1}{2L^3} \sum_{\mathbf{x}} T \langle 0 | P(0) P(\mathbf{x}) | 0 \rangle ,$$

$$P(x) = \{ \bar{d} \gamma_5 u + \bar{u} \gamma_5 d \} (x)$$

- the simulation has been performed on a lattice volume $L^3 \times T = 24^3 \times 48$ with equal (unphysical) masses for the dynamical up, down and strange quarks
- in this channel we expect a peak in correspondence of m_π and the next contribution to be at $E_* \simeq 3m_\pi$

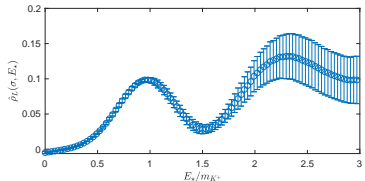
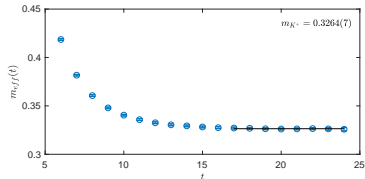


- we have applied our method to true lattice data also in the case of a QCD+QED pseudoscalar-pseudoscalar correlator

$$C_{\text{QCD+QED}}(t) = \frac{1}{2L^3} \sum_{\mathbf{x}} T\langle 0 | P(0) P(\mathbf{x}) | 0 \rangle ,$$

$$P(x) = \{ \bar{S} \gamma_5 U + \bar{U} \gamma_5 S \} (x)$$

- the simulation has been performed on a lattice volume $L^3 \times T = 24^3 \times 48$, at the unphysical value $\alpha_{em} = 0.05$ with dynamical up, down and strange quarks
- in this channel we expect a peak in correspondence of m_{K^+} and the next contribution to be at $E_{3K^+}/m_{K^+} \simeq 2.6$



- we have devised a new numerical method to cope with inverse problems
- the method inherits from the classical BG approach the very smart mechanism that allows to keep statistical errors under control
- in our method the smearing function is an input of the procedure and there is a natural way to chose the trade-off parameter λ
- by comparing results at sub-optimal values of λ one can asses the reliability of the estimated errors
- the method is general and can be applied to inverse problems arising in different research fields
- we look forward to many interesting applications: the R -ratio, hadronic τ decays, exotic spectroscopy, etc.

