Introduction to Bayesian methods

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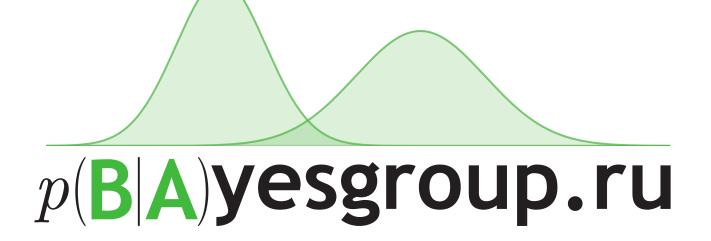
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Special thanks to Ekaterina Lobacheva for assistance with preparation of slides



Outline

- Bayesian framework
- Bayesian ML models
- Full Bayesian inference and conjugate distributions
- Approximate Bayesian inference

How to work with distributions?

Product rule

any joint distribution can be expressed as a product of one-dimensional conditional distributions

p(x, y, z) = p(x|y, z)p(y|z)p(z)

Conditional = $\frac{\text{Joint}}{\text{Marginal}}, \quad p(x|y) = \frac{p(x,y)}{p(y)}$

Sum rule

any marginal distribution can be obtained from the joint distribution by integrating out

$$p(y) = \int p(x, y) dx$$

Example

- We have a joint distribution over three groups of variables p(x,y,z)
- We observe \boldsymbol{x} and are interested in predicting \boldsymbol{y}
- Values of z are unknown and irrelevant to us
- How to estimate p(y|x) from p(x, y, z)?

The proups of variables p(x, y, z) is predicting y evant to us

Example

- We have a joint distribution over three groups of variables p(x, y, z)• We observe x and are interested in predicting y• Values of z are unknown and irrelevant to us • How to estimate p(y|x) from p(x, y, z)?

$$p(y|x) = \frac{p(x,y)}{p(x)} = \frac{\int p(x,y,z)dz}{\int p(x,y,z)dzdy}$$

Sum rule and product rule allow to obtain arbitrary conditional distributions from the joint one

Bayes theorem

Bayes theorem (follows from product and sum rules):

$$p(y|x) = \frac{p(x,y)}{p(x)} = \frac{p(x|y)p(y)}{p(x)} = \frac{p(x|y)p(y)}{\int p(x|y)p(y)dy}$$

Bayes theorem defines the rule for uncertainty conversion when new information arrives:

Posterior
$$=$$
 $-$

Likelihood \times Prior

Evidence

Statistical inference

Problem: given i.i.d. data $X = (x_1, ..., x_n)$ from distribution $p(x|\theta)$ one needs to estimate θ

Frequentist framework: use maxim

 $\theta_{ML} = \arg \max p(X|\theta) = \arg \max \theta_{ML}$

Bayesian framework: encode uncertainty about θ in a prior $p(\theta)$ and apply Bayesian inference $p(\theta|X) = \frac{\mathbf{II}}{\mathbf{\Gamma}}$

num likelihood estimation (MLE)

$$x \prod_{i=1}^{n} p(x_i|\theta) = \arg \max \sum_{i=1}^{n} \log p(x_i|\theta)$$

$$\frac{\prod_{i=1}^{n} p(x_i|\theta) p(\theta)}{\prod_{i=1}^{n} p(x_i|\theta) p(\theta) d\theta}$$

- We have a coin which may be fair or not
- The task is to estimate a probability $\boldsymbol{\theta}$ of landing heads up
- Data: 2 tosses with a result (H,H)

or not by θ of landing





Head (H) Tail (T)



- We have a coin which may be fair or not
- The task is to estimate a probability $\boldsymbol{\theta}$ of landing heads up
- Data: 2 tosses with a result (H,H)

Frequentist framework:

In all experiments the coin landed heads up $\theta_{ML} = 1$

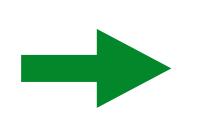
or not by θ of landing





Head (H) Tail (T)

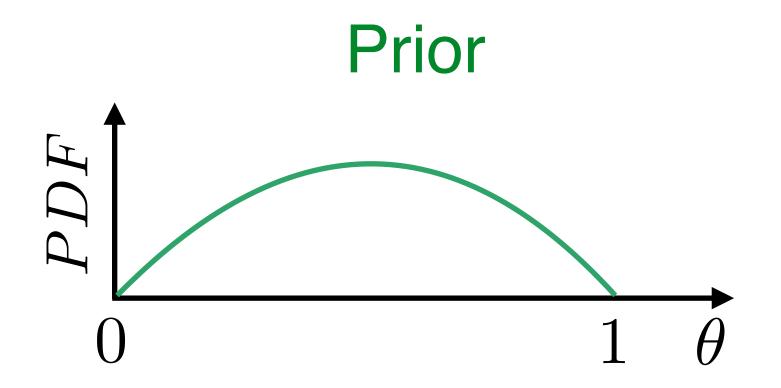
The coin is not fair and always lands heads up





- We have a coin which may be fair or not
- The task is to estimate a probability $\boldsymbol{\theta}$ of landing heads up
- Data: 2 tosses with a result (H,H)

Bayesian framework:

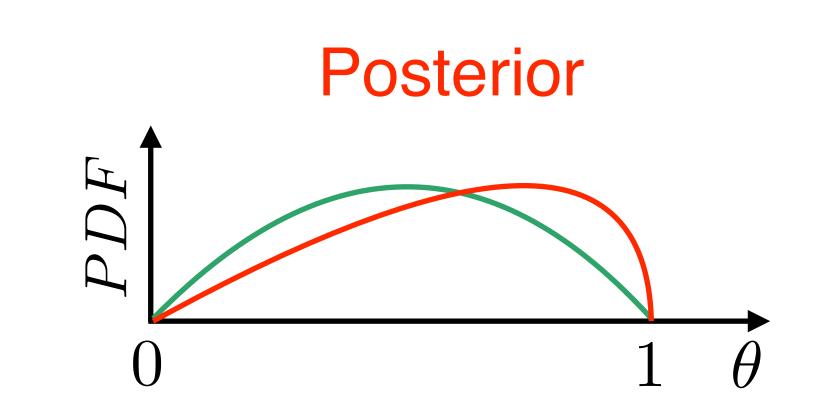


or not by θ of landing





Head (H) Tail (T)





- We have a coin which may be fair or not
- The task is to estimate a probability θ of landing heads up
- Data: 1000 tosses with a result (H,H,T,...) 489 tails and 511 heads





Head (H)

Tail (T)



- We have a coin which may be fair or not
- The task is to estimate a probability θ of landing heads up
- Data: 1000 tosses with a result (H,H,T,...) 489 tails and 511 heads

Both frameworks:

Sufficient amount of data matches our expectations





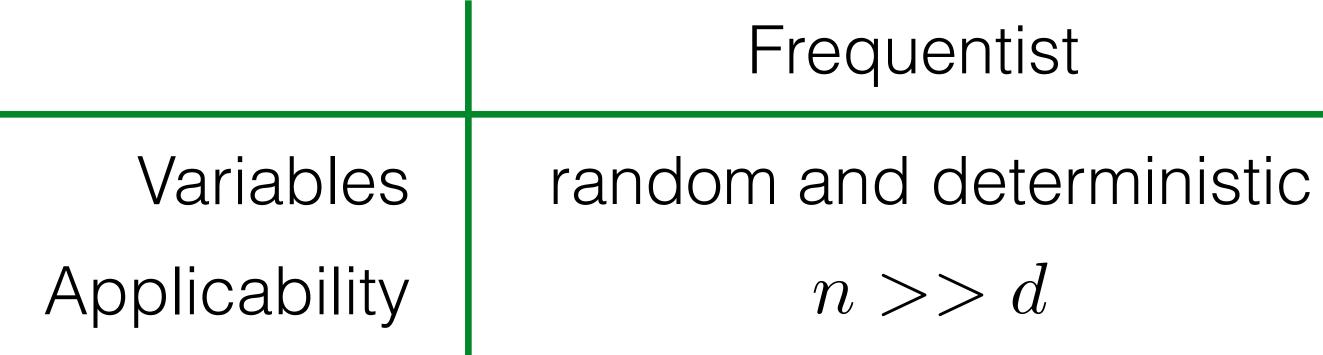


Tail (T) Head (H)

The coin is fair



Frequentist vs. Bayesian frameworks



- The number of tunable parameters in modern ML models is comparable with the sizes of training data
- Frequentist framework is a limit case of Bayesian one:

$$\lim_{n/d\to\infty} p\left(\theta|x_1,.\right.$$

ntist Bayesian leterministic everything is random > d $\forall n$

 $(\ldots, x_n) = \delta (\theta - \theta_{ML})$

Advantages of Bayesian framework

- We can encode our prior knowledge or desired properties of the final solution into a prior distribution
- Prior is a form of regularization
- the uncertainty of the estimate

Bayesian framework just provides an alternative point of view, it DOES NOT contradict or deny frequentist framework

• Additionally to the point estimate of θ posterior contains information about



Probabilistic ML model

For each object in the data:

- x set of observed variables (features)
- y set of hidden / latent variables (class label / hidden representation, etc.)

Model:

• θ — model parameters (e.g. weights of the linear model)

Discriminative probabilistic ML model

Models $p(y, \theta \mid x)$

Usually assumes that prior over θ does not depend on x:

Examples:

- the observed one)
- complexity)

Cannot generate new objects — needs x as an input

 $p(y, \theta \mid x) = p(y \mid x, \theta)p(\theta)$

Classification or regression task (hidden space is much easier than

Machine translation (hidden and observed spaces have the same

Generative probabilistic ML model

Models joint distribution $p(x, y, \theta) = p(x, y \mid \theta)p(\theta)$

Examples:

Generation of text, speech, images, etc.

Can generate new objects, i.e. pairs (x, y)

May be quite difficult to train since the observed space is usually much more complicated than the hidden one

Training Bayesian ML models

We are given training data (X_{tr}, Y_{tr}) and a discriminative model $p(y, \theta \mid x)$

Training stage — Bayesian inference over θ :

$$p(\theta \mid X_{tr}, Y_{tr}) = \frac{p(Y_{tr} \mid X_{tr}, \theta) p(\theta)}{\int p(Y_{tr} \mid X_{tr}, \theta) p(\theta) d\theta}$$

Result: ensemble of algorithms rather than a single one θ_{ML}

- model could extract and may be used as a new prior later
- Ensemble usually outperforms single best model Posterior capture all dependencies from the training data that the

Predictions of Bayesian ML models

Testing stage:

- From training we have a posterior distribution $p(\theta \mid X_{tr}, Y_{tr})$
- New data point x arrives

Ensembling w.r.t. posterior over the parameters θ : $p(y \mid x, X_{tr}, Y_{tr}) = \int p(y \mid x, \theta) p(\theta \mid X_{tr}, Y_{tr}) d\theta$

• We need to compute the predictive distribution on its hidden value y

Bayesian ML models

Training stage:

 $p\left(\theta \mid X_{tr}, Y_{tr}\right) = -$

Testing stage:

 $p(y \mid x, X_{tr}, Y_{tr}) = \int p(y \mid x, \theta) p(\theta \mid X_{tr}, Y_{tr}) d\theta$

$$\frac{p\left(Y_{tr} \mid X_{tr}, \theta\right) p(\theta)}{\int p\left(Y_{tr} \mid X_{tr}, \theta\right) p(\theta) d\theta}$$

Bayesian ML models

Training stage:

 $p\left(\theta \mid X_{tr}, Y_{tr}\right) =$

Testing stage:

 $p\left(y \mid x, X_{tr}, Y_{tr}\right) = \int_{-\infty}^{\infty} dx$

When are the integrals tractable? What can we do if they are intractable?

$$p(Y_{tr} \mid X_{tr}, \theta) p(\theta)$$
$$\int p(Y_{tr} \mid X_{tr}, \theta) p(\theta) d\theta$$

May be intractable

$$p(y \mid x, \theta) p\left(\theta \mid X_{tr}, Y_{tr}\right) d\theta$$

Distribution $p(\theta)$ and $p(x \mid \theta)$ are conjugate iff $p(\theta \mid x)$ belongs to the same parametric family as $p(\theta)$:

 $p(\theta) \in \mathcal{A}(\alpha), \quad p(x \mid \theta) \in \mathcal{B}(\theta) \longrightarrow p(\theta \mid x) \in \mathcal{A}(\alpha')$

parametric family as $p(\theta)$:

Intuition:

$$p(\theta \mid x) = \frac{p(x \mid x)}{\int p(x \mid x)}$$

Distribution $p(\theta)$ and $p(x \mid \theta)$ are conjugate iff $p(\theta \mid x)$ belongs to the same

 $p(\theta) \in \mathcal{A}(\alpha), \quad p(x \mid \theta) \in \mathcal{B}(\theta) \longrightarrow p(\theta \mid x) \in \mathcal{A}(\alpha')$

 $\frac{\theta}{\theta} p(\theta) \\ \frac{\theta}{\theta} p(\theta) d\theta$

parametric family as $p(\theta)$:

Intuition:

• Denominator is tractable since any distribution in \mathcal{A} is normalized

- Distribution $p(\theta)$ and $p(x \mid \theta)$ are conjugate iff $p(\theta \mid x)$ belongs to the same
 - $p(\theta) \in \mathcal{A}(\alpha), \quad p(x \mid \theta) \in \mathcal{B}(\theta) \longrightarrow p(\theta \mid x) \in \mathcal{A}(\alpha')$
 - $p(\theta \mid x) = \frac{\int p(x \mid \theta)p(\theta)}{\int p(x \mid \theta)p(\theta)d\theta} \leftarrow \text{conjugate}$

parametric family as $p(\theta)$:

Intuition: $p(\theta \mid x) = \frac{p(x \mid \theta)p(\theta)}{\int p(x \mid \theta)p(\theta)d\theta} \propto p(x \mid \theta)p(\theta)$

- Denominator is tractable since any distribution in \mathcal{A} is normalized
- All we need is to compute α'

Distribution $p(\theta)$ and $p(x \mid \theta)$ are conjugate iff $p(\theta \mid x)$ belongs to the same

 $p(\theta) \in \mathcal{A}(\alpha), \quad p(x \mid \theta) \in \mathcal{B}(\theta) \longrightarrow p(\theta \mid x) \in \mathcal{A}(\alpha')$

Full Bayesian inference

Training stage:

 $p(\theta \mid X_{tr}, Y_{tr}) = \frac{p(Y_{tr} \mid X_{tr}, \theta) p(\theta)}{\int p(Y_{tr} \mid X_{tr}, \theta) p(\theta) d\theta}$

Testing stage:

 $p(y \mid x, X_{tr}, Y_{tr}) = \left[\int p(y \mid x, \theta) p(\theta \mid X_{tr}, Y_{tr}) d\theta \right]$

Integrals are tractable if prior and likelihood are conjugate

Full Bayesian inference

- Easy to use analytical formulas for training and testing stages
- Strong assumptions on the model conjugacy of prior and likelihood
 - → Choose conjugate prior
 - → Only simple models (not flexible enough for most of the cases)

- We have a coin which may be fair or not
- The task is to estimate a probability θ of landing heads up
- Data: $X = (x_1, \ldots, x_n), \quad x \in \{0, 1\}$

Probabilistic model:

 $p(x,\theta) = p(x \mid \theta)p(\theta)$





Head (H)

Tail (T)



- We have a coin which may be fair or not
- The task is to estimate a probability θ of landing heads up
- Data: $X = (x_1, \ldots, x_n), \quad x \in \{0, 1\}$
- **Probabilistic model:**

 $p(x,\theta) =$

Likelihood: $Bern(x \mid \theta) =$





Head (H)

Tail (T)

$$= p(x \mid \theta)p(\theta)$$
$$\theta^{x}(1-\theta)^{1-x}$$



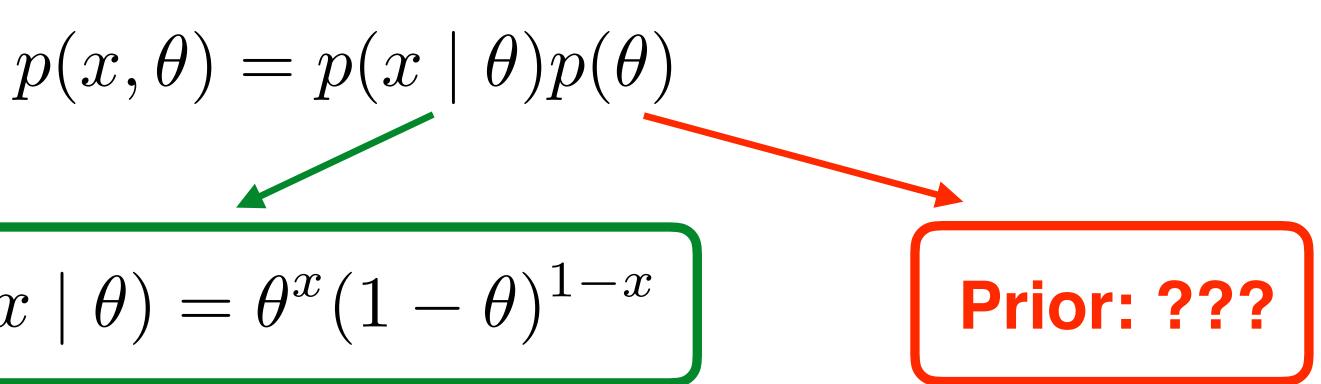
- We have a coin which may be fair or not
- The task is to estimate a probability θ of landing heads up
- Data: $X = (x_1, \ldots, x_n), \quad x \in \{0, 1\}$
- **Probabilistic model:**

Likelihood: $Bern(x \mid \theta) = \theta^x (1 - \theta)^{1-x}$





Tail (T) Head (H)





How to choose a prior?

- Correct domain: $\theta \in [0, 1]$
- Include prior knowledge: a coin is most likely fair
- Inference complexity: use conjugate prior

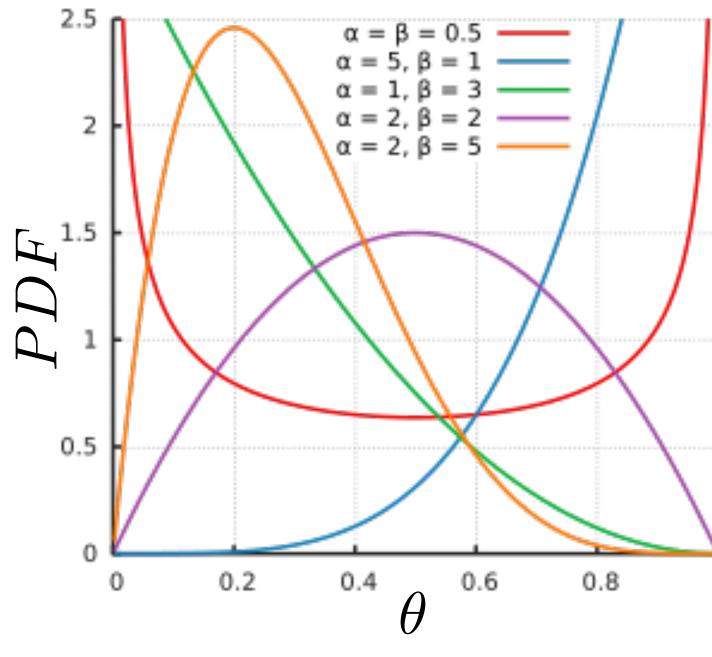
How to choose a prior?

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Beta distribution matches all requirements: $Beta(\theta \mid a, b) = \frac{1}{B(a, b)}$

$$\theta^{a-1}(1-\theta)^{b-1}$$

Beta distribution





Simplest way — approximate posterior with delta function in θ_{MP} :

 $\theta_{MP} = \arg \max p\left(\theta \mid X_{tr}, Y_{tr}\right) = \arg \max p\left(Y_{tr} \mid X_{tr}, \theta\right) p(\theta)$

Simplest way — approximate posterior with delta function in θ_{MP} :

 $\theta_{MP} = \arg \max p\left(\theta \mid X_{tr}, Y_{tr}\right) = \arg \max p\left(Y_{tr} \mid X_{tr}, \theta\right) p(\theta)$

On the testing stage: $p(y \mid x, X_{tr}, Y_{tr}) = \int p(y \mid x, \theta) p(\theta \mid X_{tr}, Y_{tr}) d\theta \approx p(y \mid x, \theta_{MP})$

Simplest way — approximate posterior with delta function in θ_{MP} : $\theta_{MP} = \arg\max p\left(\theta \mid X_{tr}, Y_{tr}\right)$

On the testing stage: $p(y \mid x, X_{tr}, Y_{tr}) = \int p(y \mid x, \theta) p(\theta \mid X_{tr}, Y_{tr}) d\theta \approx p(y \mid x, \theta_{MP})$

$$D = \arg \max p\left(Y_{tr} \mid X_{tr}, \theta\right) p(\theta)$$

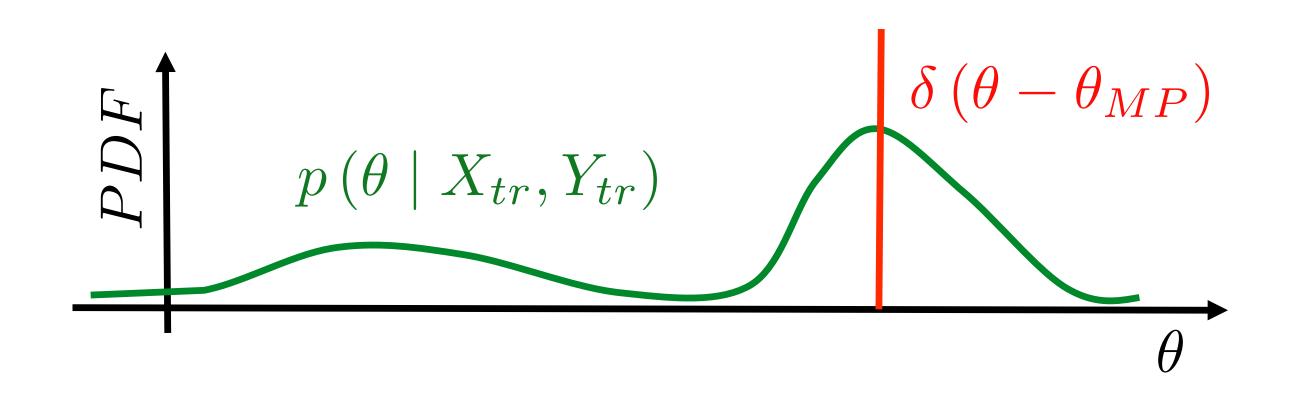
We do not need to calculate the normalisation constant



Simplest way — approximate posterior with delta function in θ_{MP} :

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What to do if there is no conjugacy?

Simplest way — approximate posterior with delta function in θ_{MP} :

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* Not the same as θ_{ML} — here we use prior

What to do if there is no conjugacy?

Simplest way — approximate posterior with delta function in θ_{MP} :

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On the testing stage: $p(y \mid x, X_{tr}, Y_{tr}) = \int p(y \mid x, \theta) p(\theta \mid X_{tr}, Y_{tr}) d\theta \approx p(y \mid x, \theta_{MP})$

More advanced techniques are needed!

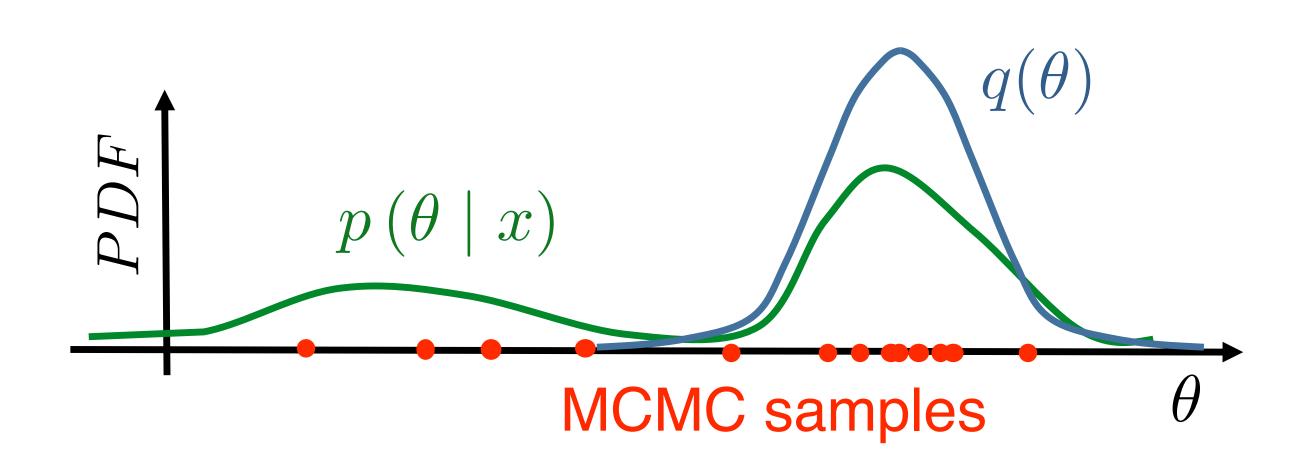
Approximate inference

Probabilistic model: $p(x, \theta) = p(x \mid \theta)p(\theta)$

Variational Inference

Approximate $p(\theta \mid x) \approx q(\theta) \in \mathcal{Q}$

- Biased
- Faster and more scalable



MCMC

Samples from unnormalized $p(\theta \mid$ $\mid \mathcal{X}$)

- Unbiased
- Need a lot of samples



Probabilistic model: $p(x, \theta) = p(x \mid \theta)p(\theta)$

Main idea: find posterior approximation $p(\theta \mid x) \approx q(\theta) \in Q$, using the following criterion function:

$$F(q) := KL(q(\theta) \| p(\theta \mid x)) \to \min_{q(\theta) \in \mathcal{Q}}$$

Kullback-Leibler divergence a good mismatch measure between two distributions over the same domain

eta) p(heta)tion $p(heta \mid x) pprox q(heta) \in \mathcal{Q}$, using

Kullback-Leibler divergence

A good mismatch measure between two distributions over the same domain

 $KL(q(\theta) \| p(\theta \mid x))$

Properties:

- $KL(q \parallel p) \ge 0$
- $KL(q \parallel p) = 0 \Leftrightarrow q = p$
- $KL(q \parallel p) \neq KL(p \parallel q)$

$$) = \int q(\theta) \log \frac{q(\theta)}{p(\theta \mid x)} d\theta$$

$$= \int q(\theta) \log \frac{q(\theta)}{p(\theta \mid x)} d\theta$$

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Probabilistic model: $p(x, \theta) = p(x \mid \theta)p(\theta)$

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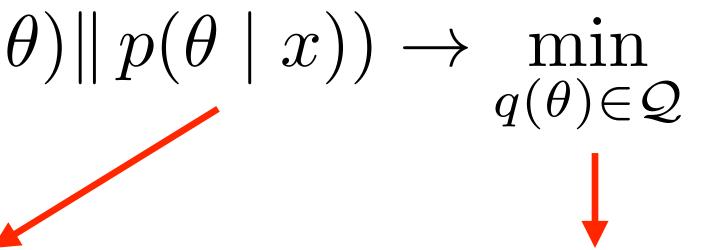
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We could not compute the posterior in the first place

eta) p(heta)tion $p(heta \mid x) pprox q(heta) \in \mathcal{Q}$, using



How to perform an optimization w.r.t. a distribution?

Mathematical magic

$$\log p(x) = \int q(\theta) \log p(x) d\theta =$$
$$= \int q(\theta) \log \frac{p(x,\theta)q(\theta)}{p(\theta \mid x)q}$$
$$= \int q(\theta) \log \frac{p(x,\theta)}{q(\theta)} d\theta$$

 $= \int q(\theta) \log \frac{p(x,\theta)}{p(\theta \mid x)} d\theta =$ $\frac{d\theta}{d(\theta)}d\theta = 0$ $d\theta + \int q(\theta) \log \frac{q(\theta)}{p(\theta \mid x)} d\theta = 0$

Mathematical magic

$$\log p(x) = \int q(\theta) \log p(x) d\theta =$$
$$= \int q(\theta) \log \frac{p(x,\theta)q(\theta)}{p(\theta \mid x)q}$$
$$= \int q(\theta) \log \frac{p(x,\theta)}{q(\theta)} d\theta$$
$$= \mathcal{L}(q(\theta)) + KL(q(\theta))$$

Evidence lower bound (ELBO)

 $= \int q(\theta) \log \frac{p(x,\theta)}{p(\theta \mid x)} d\theta =$ $\frac{q(\theta)}{q(\theta)}d\theta =$ $d\theta + \int q(\theta) \log \frac{q(\theta)}{p(\theta \mid x)} d\theta =$ $p(\theta \mid$ (x))

KL-divergence we need for VI

ELBO = Evidence Lower Bound $\log p(x) = \mathcal{L}(q(\theta)) + KL(q(\theta) || p(\theta | x))$

Evidence:

$$p(\theta \mid x) = \frac{p(x \mid \theta)p(\theta)}{p(x)} = \frac{p(x \mid \theta)p(\theta)}{\int p(x)}$$

Evidence of the probabilistic model shows the total probability of observing the data.

Lower Bound: KL is non-negative $\rightarrow \log p(x) \ge \mathcal{L}(q(\theta))$

$\frac{x \mid \theta) p(\theta)}{x \mid \theta) p(\theta) d\theta} = \frac{\text{Likelihood} \times \text{Prior}}{\text{Evidence}}$

Optimization problem with intractable posterior distribution:

$F(q) := KL(q(\theta) \| p(\theta \mid x)) \to \min_{q(\theta) \in \mathcal{Q}}$

Optimization problem with intractable posterior distribution:

Let's use our magic:

$F(q) := KL(q(\theta) \| p(\theta \mid x)) \to \min_{q(\theta) \in \mathcal{Q}}$

$\log p(x) = \mathcal{L}(q(\theta)) + KL(q(\theta) || p(\theta | x))$

Optimization problem with intractable posterior distribution:

Let's use our magic:

 $\log p(x) = \mathcal{L}(q(\theta))$ does not depend on

$F(q) := KL(q(\theta) \| p(\theta \mid x)) \to \min_{q(\theta) \in \mathcal{Q}}$

$$) + KL(q(\theta) \parallel p(\theta \mid x))$$

 q depend on q

Optimization problem with intractable posterior distribution:

Let's use our magic:

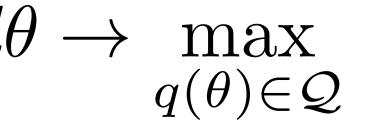
does not depend on q

$$KL(q(\theta) \| p(\theta \mid x)) \to \inf_{q(\theta)} q(\theta)$$

$F(q) := KL(q(\theta) \| p(\theta \mid x)) \to \min_{q(\theta) \in Q}$ $\log p(x) = \mathcal{L}(q(\theta)) + KL(q(\theta) || p(\theta | x))$ depend on q

 $\min_{\substack{\theta \in \mathcal{Q}}} \Leftrightarrow \mathcal{L}(q(\theta)) \to \max_{\substack{q(\theta) \in \mathcal{Q}}}$

$$\mathcal{L}(q(\theta)) = \int q(\theta) \log \frac{p(x,\theta)}{q(\theta)} d\theta$$



Variational inference: ELBO interpretation

$$\mathcal{L}(q(\theta)) = \int q(\theta) \log \frac{p(x,\theta)}{q(\theta)} d\theta = \int q(\theta) \log \frac{p(x \mid \theta)p(\theta)}{q(\theta)} d\theta =$$
$$= \int q(\theta) \log p(x \mid \theta) d\theta + \int q(\theta) \log \frac{p(\theta)}{q(\theta)} d\theta =$$

Variational inference: ELBO interpretation

$$\begin{aligned} \mathcal{L}(q(\theta)) &= \int q(\theta) \log \frac{p(x,\theta)}{q(\theta)} d\theta = \int q(\theta) \log \frac{p(x \mid \theta)p(\theta)}{q(\theta)} d\theta = \\ &= \int q(\theta) \log p(x \mid \theta) d\theta + \int q(\theta) \log \frac{p(\theta)}{q(\theta)} d\theta = \\ &= \mathbb{E}_{q(\theta)} \log p(x \mid \theta) - KL(q(\theta) \parallel p(\theta)) \end{aligned}$$
data term regularizer

$$\mathcal{L}(q(\theta)) = \int q(\theta) \log \frac{p(x,\theta)}{q(\theta)} d\theta$$



Final optimisation problem:

$$\mathcal{L}(q(\theta)) = \int q(\theta) \log \frac{p(x,\theta)}{q(\theta)} d\theta$$

- $q(\theta) = q(\theta)$
- Parametric family
- **Parametric approximation**

 $\theta \to \max_{q(\theta) \in \mathcal{Q}}$

How to perform an optimization w.r.t. a distribution?

Parametric approximation

Parametric family of variational distributions:

$$q(\theta) = q(\theta \mid \lambda),$$

Why is it a restriction? We choose a family of some fixed form:

- It may be too simple and insufficient to model the data • If it is complex enough then there is no guaranty we can train it
- well to fit the data

- - λ some parameters

Parametric approximation

Parametric family of variational distributions:

$$q(\theta) = q(\theta \mid \lambda),$$

Variational inference transforms to parametric optimization problem: $\mathcal{L}(q(\theta \mid \lambda)) = \int q(\theta \mid \lambda)$

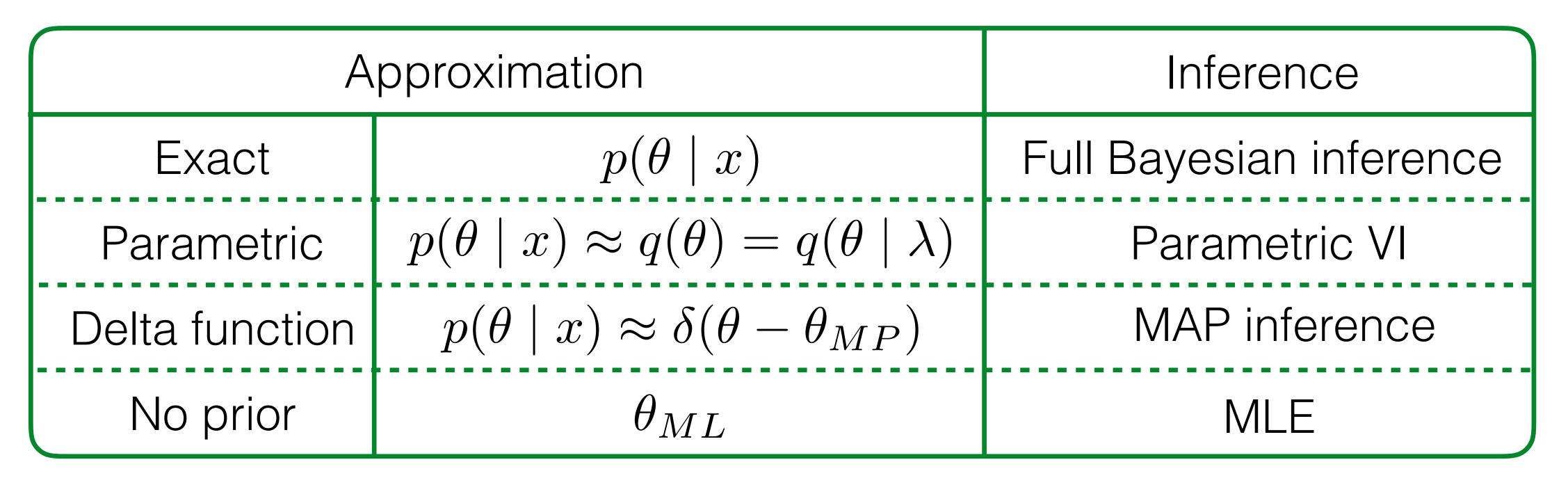
If we're able to calculate derivatives of ELBO w.r.t. λ then we can solve this problem using some numerical optimization solver.

- - λ some parameters

$$|\lambda) \log \frac{p(x,\theta)}{q(\theta \mid \lambda)} d\theta \to \max_{\lambda}$$

Inference methods: summary

Probabilistic model: $p(x, \theta)$



We want to compute: $p(\theta \mid x)$