

# Statistical methods and error analysis

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# Classical confidence intervals.

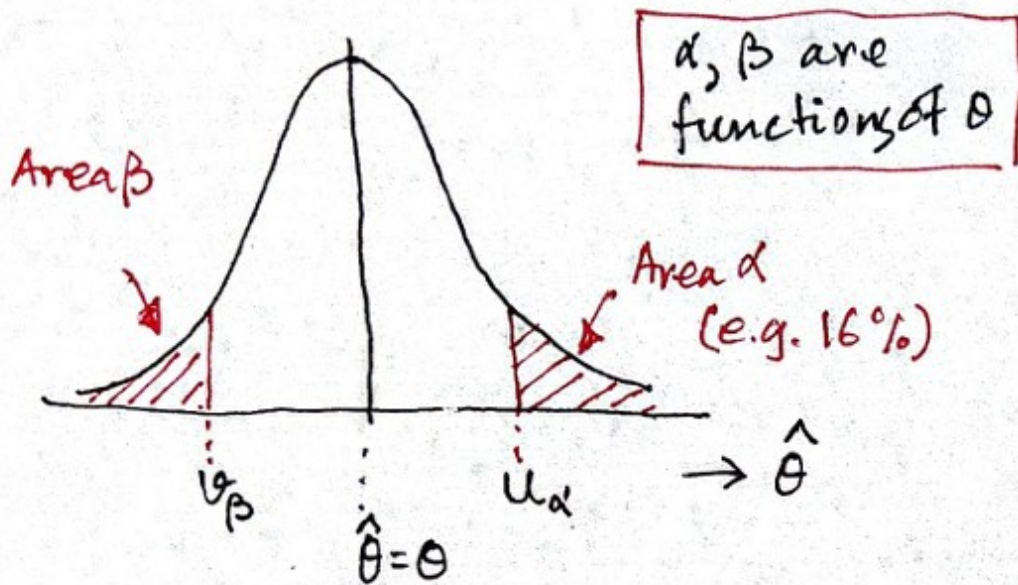
Source: G. Cowan


Due to Neyman, known as Neyman construction

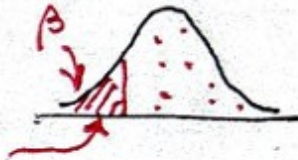
Consider a data set  $\{x_1, x_2, \dots, x_n\}$

From this set an estimator  $\hat{\theta}(x_1, x_2, \dots, x_n)$  is constructed for parameter  $\theta$

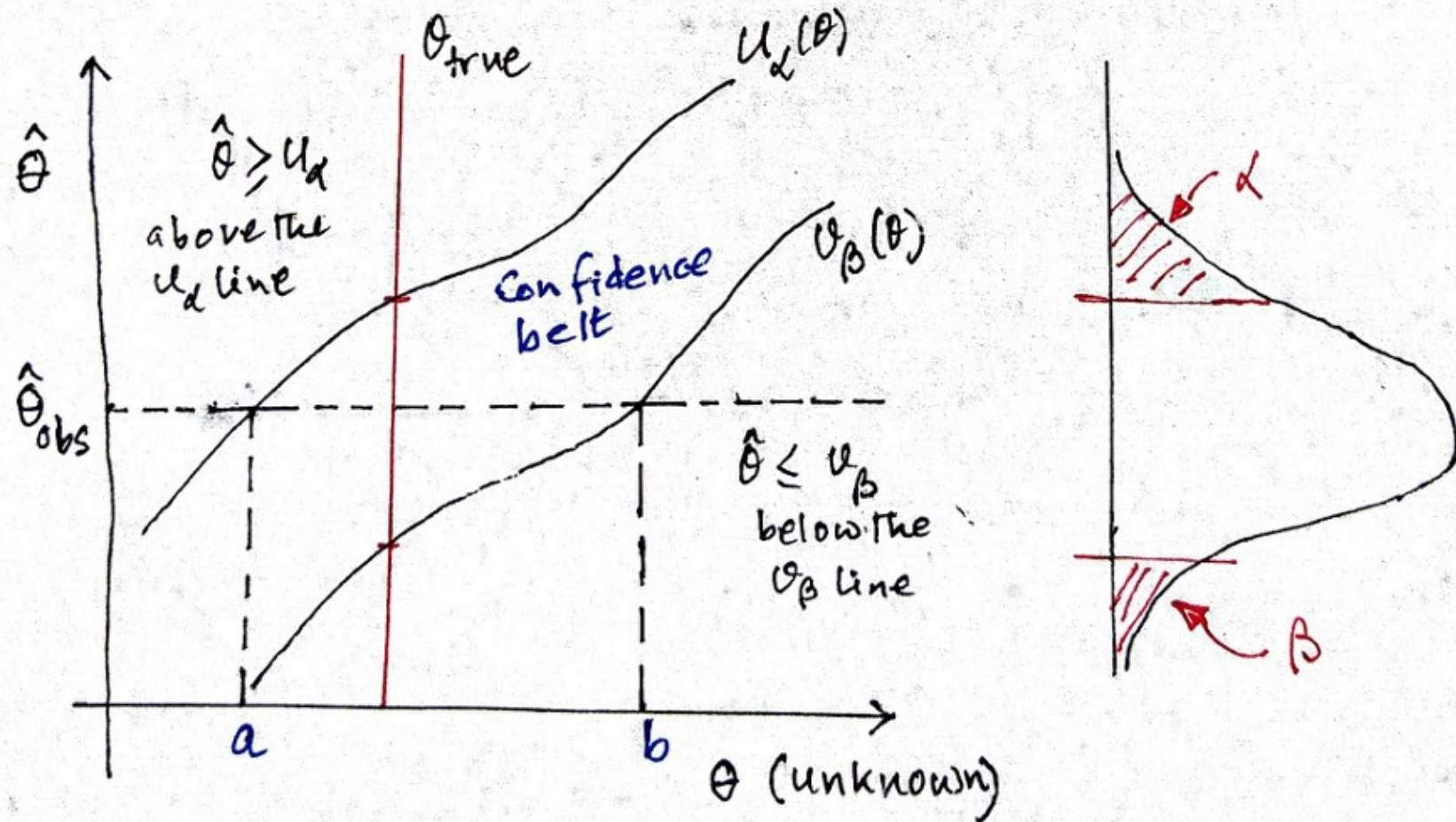
Assume:  $g(\hat{\theta}; \theta)$ , the p.d.f. of  $\theta$  is known



$$\alpha = P(\hat{\theta} \geq u_\alpha(\theta))$$
$$= \int_{u_\alpha}^{\infty} g(\hat{\theta}; \theta) d\hat{\theta} = 1 - G(u_\alpha; \theta)$$


$$\beta = P(\hat{\theta} \leq u_\beta(\theta))$$
$$= \int_{-\infty}^{u_\beta} g(\hat{\theta}; \theta) d\hat{\theta} = G(u_\beta; \theta)$$


# Neyman construction

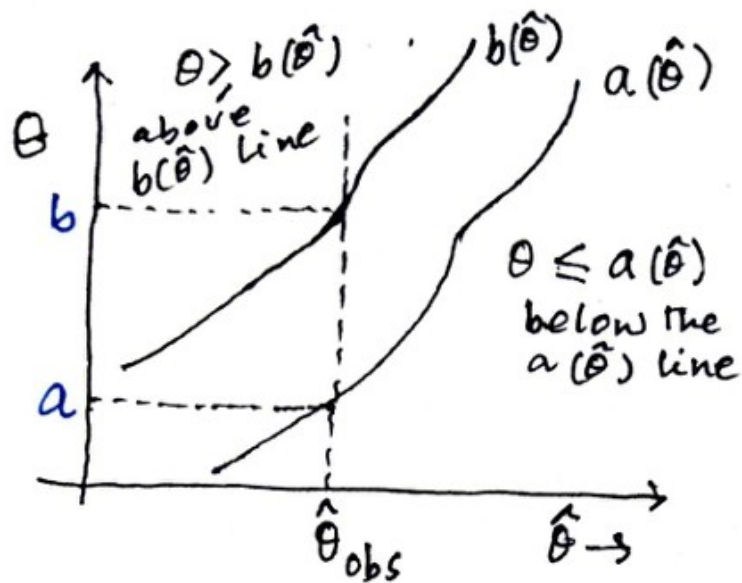


Probability of  $\hat{\theta}$  falling inside the interval

$$P(u_{\beta} \leq \hat{\theta} \leq u_{\alpha}) = 1 - \alpha - \beta$$

# Neyman construction - 2

If  $u_\alpha(\theta)$ ,  $u_\beta(\theta)$  are monotonically increasing functions of  $\theta$ ,  $u_\alpha$ ,  $u_\beta$  can be inverted



$$\underline{a(\hat{\theta})} = u_\alpha^{-1}(\hat{\theta})$$

$$\underline{b(\hat{\theta})} = u_\beta^{-1}(\hat{\theta})$$

$$\hat{\theta} \geq u_\alpha(\theta) \Rightarrow a(\hat{\theta}) \geq \theta$$

$$\hat{\theta} \leq u_\beta(\theta) \Rightarrow b(\hat{\theta}) \leq \theta$$

Therefore:  $P(a(\hat{\theta}) \leq \theta \leq b(\hat{\theta})) = 1 - \alpha - \beta$  ✓

$[a, b]$  is the classical confidence interval around the observed  $\hat{\theta} = \hat{\theta}_{obs}$ .

# Intervals and limits

$[a, b]$  is the classical confidence interval around the observed  $\hat{\theta} = \hat{\theta}_{obs}$ .

- ▶  $a, b$  are random numbers, they change with  $\hat{\theta}_{obs}$
- ▶ However we just showed that they have a coverage probability  $(1 - \alpha - \beta)$

i.e. if we construct a 68% interval in this way,  $(1 - \alpha - \beta) = 0.68$  then in 68% of repeat experiments  $[a, b]$  will contain the true but unknown value of  $\theta$ .

# Intervals and limits

If  $\beta = 0$ ,  $P(a \leq \theta) = 1 - \alpha$

lower limit on  $\theta$  at  
 $(1 - \alpha) \times 100\%$  confidence level  
for  $\alpha = 0.05 \rightarrow 95\%$  C.L.

If  $\alpha = 0$ ,  $P(\theta \leq b) = 1 - \beta$

Upper limit on  $\theta$  at  
 $(1 - \beta) \times 100\%$  confidence  
level

If  $\alpha = \beta = \frac{\gamma}{2}$ , central interval with C.L.  $(1 - \gamma) \times 100\%$

e.g.  $\gamma = 0.32 \Rightarrow \frac{(1 - 0.32) \times 100}{100} = 68\%$  C.L.

# Interval on Poisson mean

- Single observation  $n_{obs}$  of a Poisson process:  $\hat{\nu} = n_{obs}$

$$\alpha = P(\hat{\nu} \geq \hat{\nu}_{obs}; a),$$

$$\beta = P(\hat{\nu} \leq \hat{\nu}_{obs}; b),$$

$$\alpha = \sum_{n=n_{obs}}^{\infty} f(n; a) = 1 - \sum_{n=0}^{n_{obs}-1} f(n; a) = 1 - \sum_{n=0}^{n_{obs}-1} \frac{a^n}{n!} e^{-a},$$

$$\beta = \sum_{n=0}^{n_{obs}} f(n; b) = \sum_{n=0}^{n_{obs}} \frac{b^n}{n!} e^{-b}.$$

# Poisson interval - 2

$$\begin{aligned}\sum_{n=0}^{n_{\text{obs}}} \frac{\nu^n}{n!} e^{-\nu} &= \int_{2\nu}^{\infty} f_{\chi^2}(z; n_d = 2(n_{\text{obs}} + 1)) dz \\ &= 1 - F_{\chi^2}(2\nu; n_d = 2(n_{\text{obs}} + 1)),\end{aligned}$$

- Using the above relation between chi-squared and Poisson, we get

$$a = \frac{1}{2} F_{\chi^2}^{-1}(\alpha; n_d = 2n_{\text{obs}}),$$

$$b = \frac{1}{2} F_{\chi^2}^{-1}(1 - \beta; n_d = 2(n_{\text{obs}} + 1))$$



# Poisson Interval -3

- Poisson intervals are conservatively large

$$P(\nu \geq a) \geq 1 - \alpha,$$

$$P(\nu \leq b) \geq 1 - \beta,$$

$$P(a \leq \nu \leq b) \geq 1 - \alpha - \beta.$$

- Special case is zero observation --> upper limit  $\sim 3$

$$\beta = \sum_{n=0}^0 \frac{b^n e^{-b}}{n!} = e^{-b}$$

# Error on efficiency- CP interval

- Error on efficiency with binomial error does not work if efficiency is close to 1 or 0
- Clopper Pearson is a classical interval, known also as an exact interval

The Clopper-Pearson interval can be written as

$$S_{\leq} \cap S_{\geq}$$

or equivalently,

$$(\inf S_{\geq}, \sup S_{\leq})$$

with

$$S_{\leq} := \left\{ \theta \mid P[\text{Bin}(n; \theta) \leq x] > \frac{\alpha}{2} \right\} \text{ and } S_{\geq} := \left\{ \theta \mid P[\text{Bin}(n; \theta) \geq x] > \frac{\alpha}{2} \right\}$$

Source: Wikipedia

- There are other intervals, Agresti-Coull implemented in ROOT