Tuning the Parton Shower Parameters with the Marginal Likelihood

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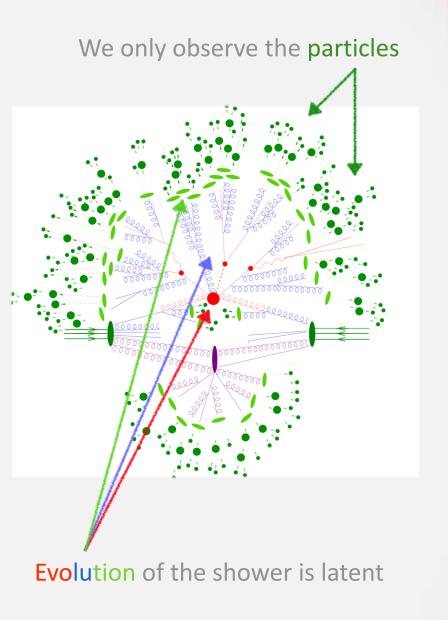
July 6th, 2021



Parton Shower Models

simulating jet physics

- We have generators which can simulate parton showers
 - e.g. Pythia, Sherpa, Herwig, etc.
- They depend on a set of physically motivated parameters
- These determine the distribution of final state particles
- The detector only observes the final state particles
- The evolution of the shower is of interest for reconstruction
- It also depends on the shower parameters





Monte Carlo Tuning

- Traditional approach:
 - Generate data using parton shower models
 - Make 1D projections of observables
 - Tune parameters by matching observables with data
 - Professor w/ Rivet
 - Buckley, et. al. <u>https://arxiv.org/abs/0907.2973</u>
- Inefficient and wastes information
- Traditional approaches don't have access to a likelihood *p(data | parameters)*
- Shower Deconstruction
 - Tractable likelihood, but brute force marginalization (scalability issues); focused on tagging not tuning
 - Phys. Rev. D 2011 <u>https://inspirehep.net/literature/889900</u>
 - Soper, Spannowsky

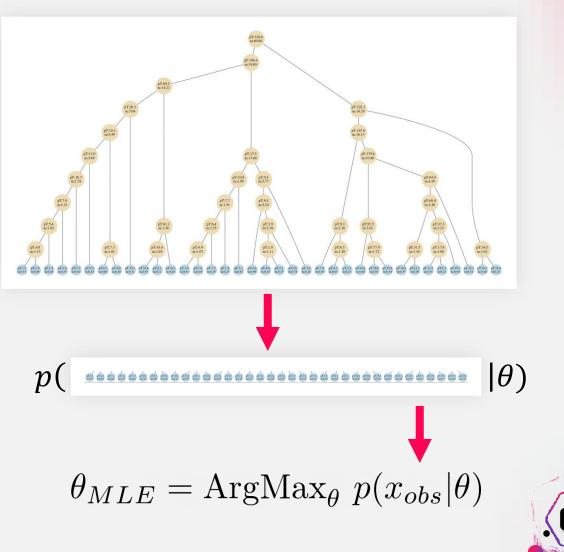
- Other ML related work based on surrogates for the intractable likelihood
 - Adversarial Variational Optimization
 - Non-differentiable simulator; GAN-like NN classifier
 - AISTATS 2019 <u>https://arxiv.org/abs/1707.07113</u>
 - Louppe, Hermans, Cranmer
 - DCTR
 - NN classifier as likelihood ratio surrogate; differentiable optimization
 - Phys. Rev. D 2020
 <u>https://inspirehep.net/literature/1744598</u>
 - Andreassen, Nachman



Statistical Framing of Tuning

maximum likelihood estimation

- Ideally, we would use the likelihood to fit the model to observed data
- Model parameters can be inferred through stastical estimation, e.g. MLE or Bayesian methods
- These use the full information of the model and are statistically efficient
- Ginkgo combined with the Trellis algorithm allows us to directly compute the likelihood for inference

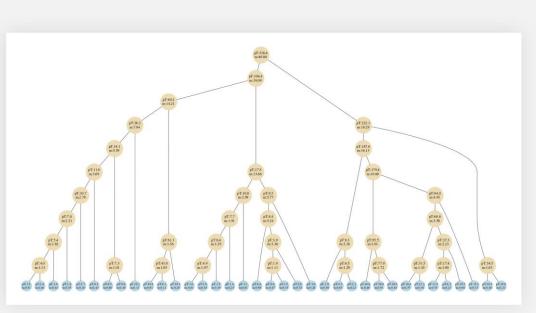


Ginkgo Refresher

a toy parton shower simulator

- Ginkgo uses a recursive algorithm to generate a binary tree where the leaves are the jet constituents
- Implements some key features needed for physics simulator:
 - Momentum conservation
 - A running splitting scale
- Model parameters that can be tuned
 - Threshold energy for stopping (like confinement)
 - Decay rate (jet-type dependent)

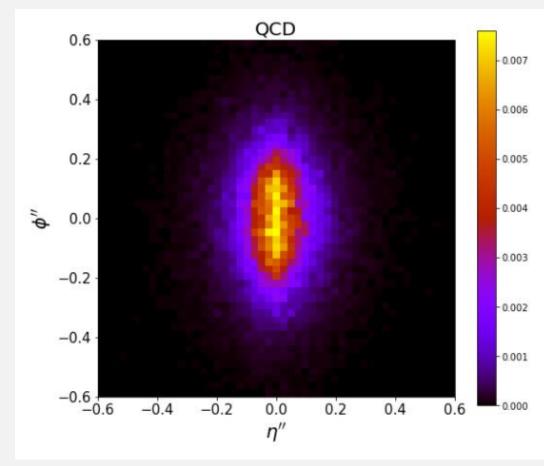
Kyle Cranmer, Sebastian Macaluso, Duccio Pappadopulo



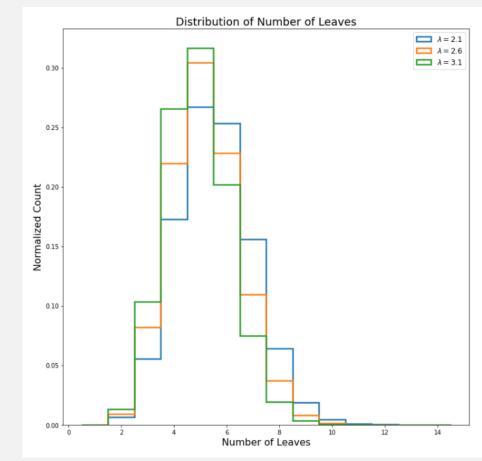


Ginkgo Jets

Phi vs Eta



Jet Multiplicity

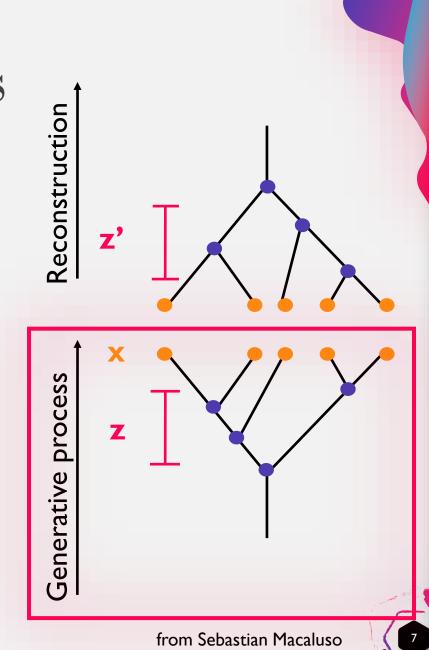




Ginkgo Generative Process

constructing a tractable likelihood

- The evolution at every step depends only on the parent 4-momentum $t_L \sim f(t|\lambda, t_P) = \frac{1}{1 - e^{-\lambda}} \frac{\lambda}{t_P} e^{-\frac{\lambda}{t_P}t} \quad t_R \sim f(t|\lambda, t_P, t_L) = \frac{1}{1 - e^{-\lambda}} \frac{\lambda}{(\sqrt{t_P} - \sqrt{t_L})^2} e^{-\frac{\lambda}{(\sqrt{t_P} - \sqrt{t_L})^2}t}$
- If $t < t_{cut}$ then stop, where t is the invariant mass squared
- Asymmetric under $L \leftrightarrow R$
- Particles are randomly permuted after splitting



Ginkgo Likelihood

constructing a tractable likelihood

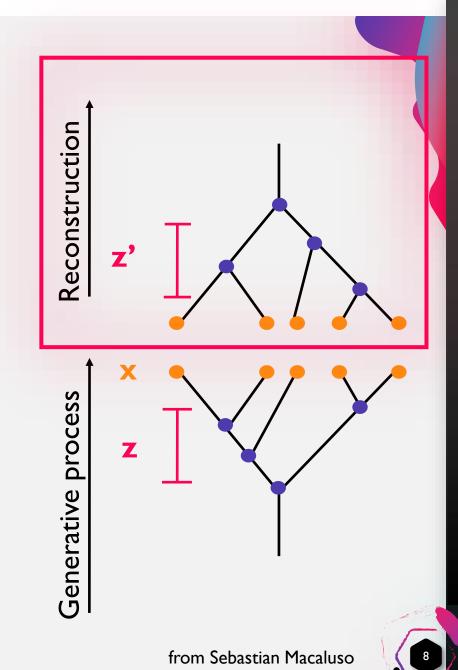
• The joint likelihood is factorized in an autoregressive form, depending on the particular tree, *z*

$$p(x, z|\theta) = \prod_{j} p(x_j|z_{\text{parent}(x_j)}, \theta) \prod_{i} p(z_i|z_{\text{parent}(z_i)}, \theta)$$

• In order to evaluate the probability of the splitting, we need to reconstruct the parent from the children

$$t_L \sim f(t|\lambda, t_{\rm P}) = \frac{1}{1 - e^{-\lambda}} \frac{\lambda}{t_{\rm P}} e^{-\frac{\lambda}{t_{\rm P}} t} \quad t_R \sim f(t|\lambda, t_{\rm P}, t_L) = \frac{1}{1 - e^{-\lambda}} \frac{\lambda}{(\sqrt{t_{\rm P}} - \sqrt{t_L})^2} e^{-\frac{\lambda}{(\sqrt{t_{\rm P}} - \sqrt{t_L})^2} t}$$

- Reconstruction probability is symmetrized
 - $\frac{1}{2}(f(t_L, t_R | \lambda, t_P) + f(t_R, t_L | \lambda, t_P))$



The Marginal Likelihood

how to integrate out the showering history

- Need to be able to evaluate this on observed data
- This also allows us to compute the exact likelihood ratio
 - See Lauren Greenspan's talk next!
- We must integrate out the showering history
- Typically, this is an intractable problem and where other approaches fail
 - Grows as (2N-3)!!
- Some machine learning approaches that approximate the marginal likelihood

$$p(x, z|\theta) = \prod_{j} p(x_j|z_{\text{parent}(x_j)}, \theta) \prod_{i} p(z_i|z_{\text{parent}(z_i)}, \theta)$$

$$\underline{p(x|\theta)} = \int dz \ \underline{p(x,z|\theta)}$$

$$\frac{1}{2} \quad \mathbf{known}$$

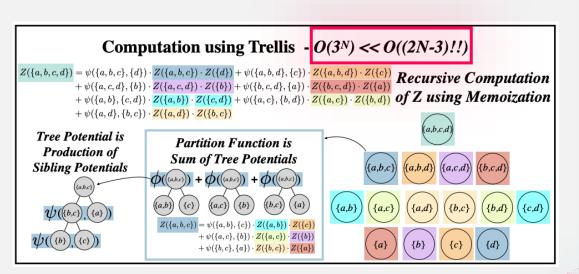
Hierarchical Cluster Trellis

how to integrate out the showering history

- The Hierarchical Cluster Trellis algorithm empowers this marginalization
- This is a model-independent algorithm for marginalizing over all binary tree histories
- Ginkgo is a perfect test case
- There are $(2N 3)!! \times 2^{N-1}$ possible binary trees including permuations
- Trellis reduces the computational complexity from brute force

https://github.com/SebastianMacaluso/ClusterTrellis

# Leaves	4	7	9	12
# Trees	120	~665k	~520M	~28T



from Sebastian Macaluso



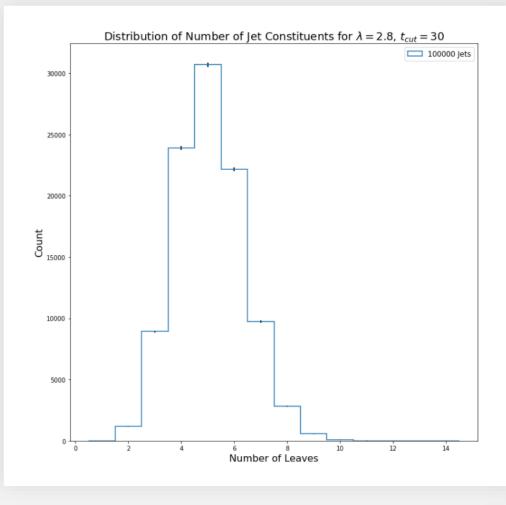
Results

Tuning using Maximum Likelihood Estimation



Observed Dataset

- Generated a dataset of 100k jets using Ginkgo
- Parameters fixed to $\lambda = 2.8$, $t_{cut} = 30 \ GeV^2$, with $|\vec{p}_{jet}| = 400 \ GeV$ and $m_{jet} = 30 \ GeV$
- Small number of jet constituents for proof of concept



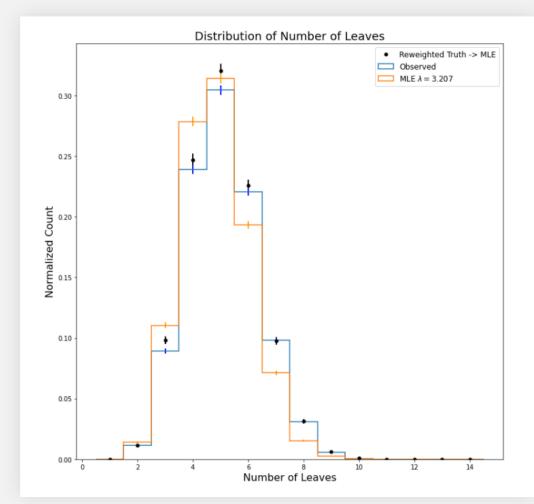
Subtle Normalization Issue

reweighting-based closure

- With the likelihood we can reweight the distribution from one value of λ to another
- This reveals a normalization problem with the likelihood, which we have not been able to track down yet
- This leads to a bias in the MLE
- But we can measure

 $f(\lambda,\lambda_t) = \mathbb{E}_{p(x|\lambda_t)}\left[p(x|\lambda)/p(x|\lambda_t)\right]$ which should be unity

• Correct the likelihood by 1/f

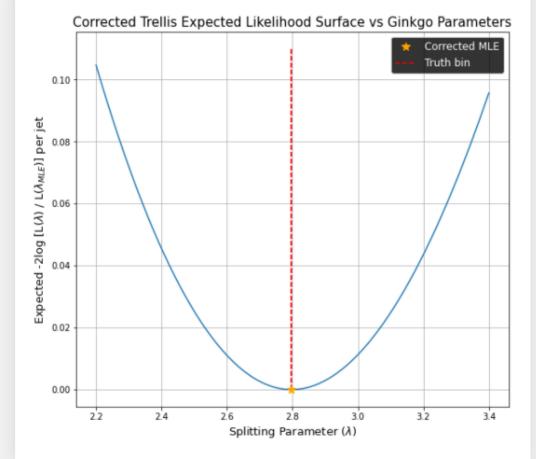


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Maximum Likelihood Estimate

tuning procedure

- A 1D grid search is performed for optimizing λ , the splitting rate
- Includes the normalization correction

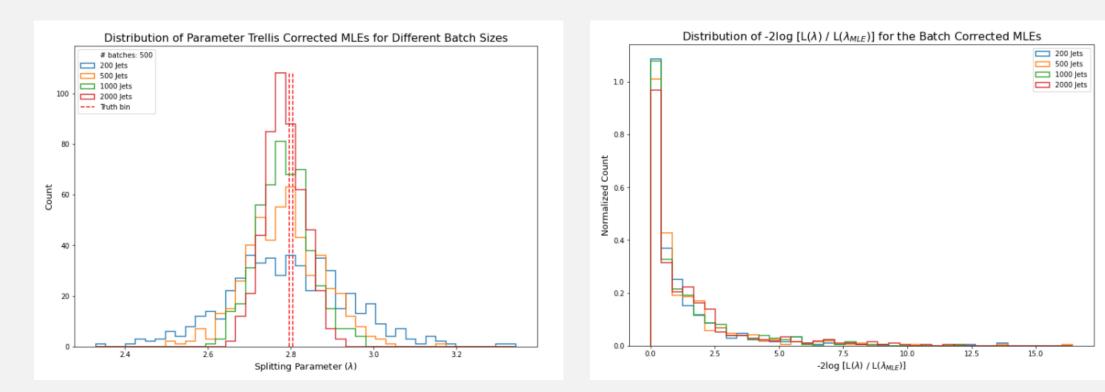




Maximum Likelihood Estimate

Distribution of the MLE

Distribution of the χ^2 test statistic



Summary

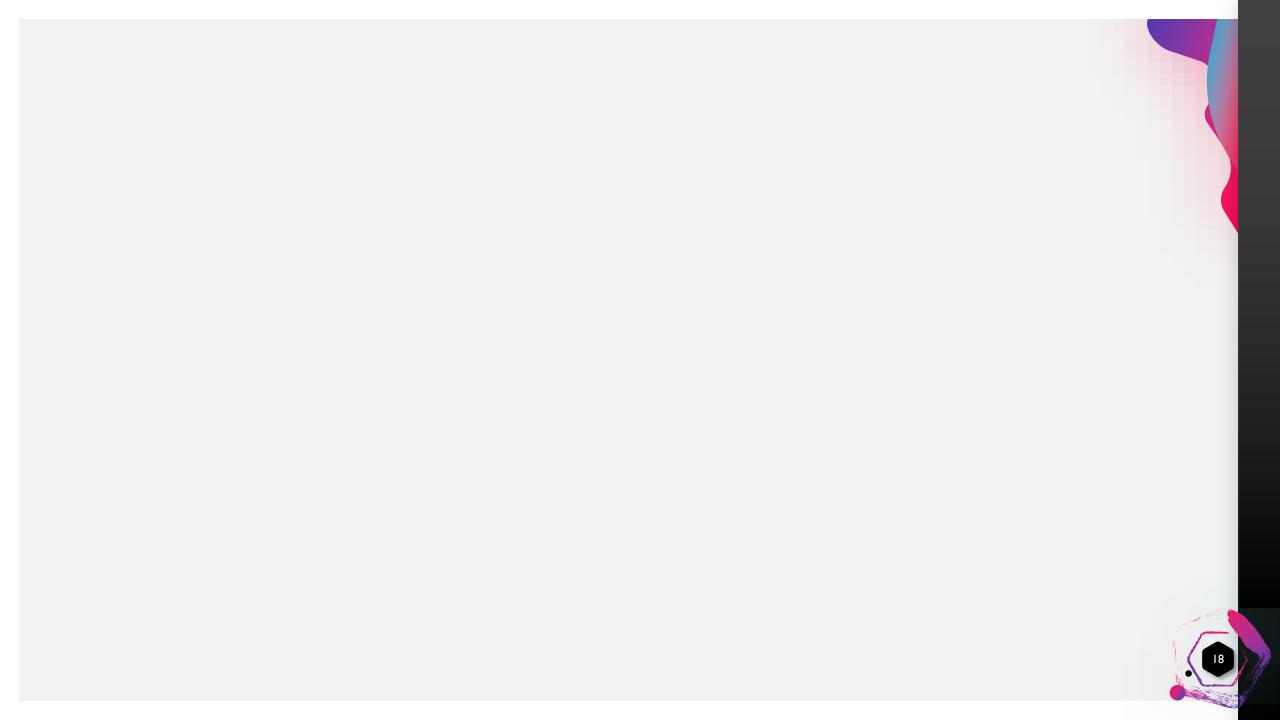
- Explored showering parameter tuning on Ginkgo, which has a tractable likelihood
- Implemented parameter tuning using the full information of the simulator model via maximum likelihood estimation
- The Trellis algorithm enabled marginalizing over possible jet clusterings
- Generalizes well to other probabilistic showering models
- To the best of our knowledge, this is the first demonstration of tuning the shower model using the full likelihood marginalized over all showering histories

Probabilistic **Parton Shower Cluster Trellis Tuning with** Marginal Likelihood

Thank You!

Some Github links

- Ginkgo: <u>https://github.com/SebastianMacalus</u> <u>o/ginkgo</u>
- Cluster Trellis: <u>https://github.com/SebastianMacalus</u> <u>o/ClusterTrellis</u>
- Ginkgo Inference: <u>https://github.com/mdkdrnevich/gink</u> <u>go-inference</u>



Backup

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Ginkgo Likelihood Math Written Out

$$p(x, z|\theta) = \prod_{j} p(x_j | z_{\text{parent}(x_j)}, \theta) \prod_{i} p(z_i | z_{\text{parent}(z_i)}, \theta)$$

$$t_L \sim f(t|\lambda, t_{\rm P}) = \frac{1}{1 - e^{-\lambda}} \frac{\lambda}{t_{\rm P}} e^{-\frac{\lambda}{t_{\rm P}}t} \quad t_R \sim f(t|\lambda, t_{\rm P}, t_L) = \frac{1}{1 - e^{-\lambda}} \frac{\lambda}{(\sqrt{t_{\rm P}} - \sqrt{t_L})^2} e^{-\frac{\lambda}{(\sqrt{t_{\rm P}} - \sqrt{t_L})^2}t}$$

$$p(z_i|z_{parent(z_i)}, \theta) = p(t_L, t_R|\lambda, t_P) = f(t_R|\lambda, t_P, t_L)f(t_L|\lambda, t_P)$$

$$\frac{1}{2}(f(t_L, t_R | \lambda, t_P) + f(t_R, t_L | \lambda, t_P))$$



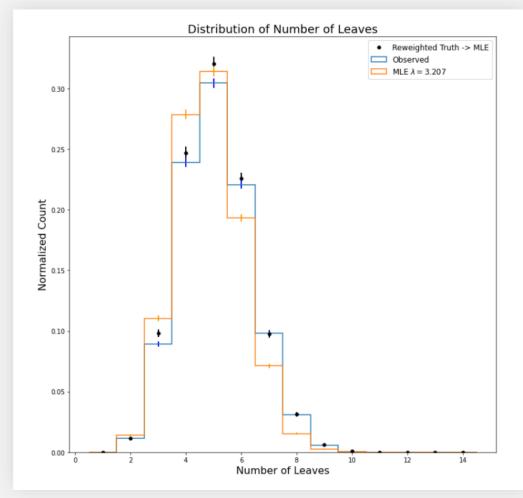
Iterative Reweighting Procedure

correcting the MLE

- The MLE is initially biased due to something missing in the Ginkgo likelihood
- We correct for this bias by introducing a reweighting procedure
- Build a simple approximation for $f(\lambda, \lambda_t) = \mathbb{E}_{p(x|\lambda_t)} \left[p(x|\lambda) / p(x|\lambda_t) \right]$
- Define an iterative correction

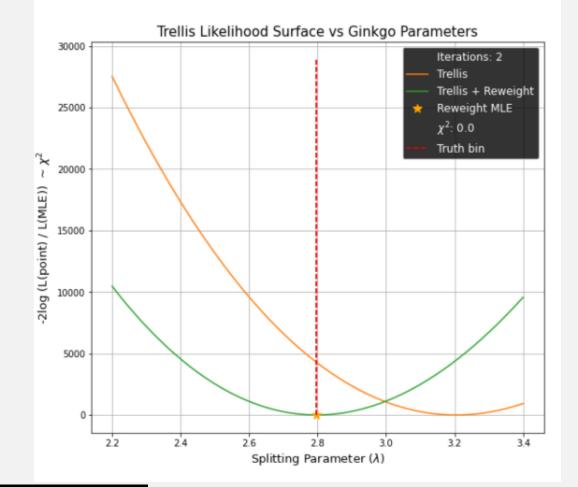
 $\lambda_{n+1}^* = \operatorname{ArgMax}_{\lambda} p(x_{obs}|\lambda) / f(\lambda, \lambda_n^*)$

• Fast and effective





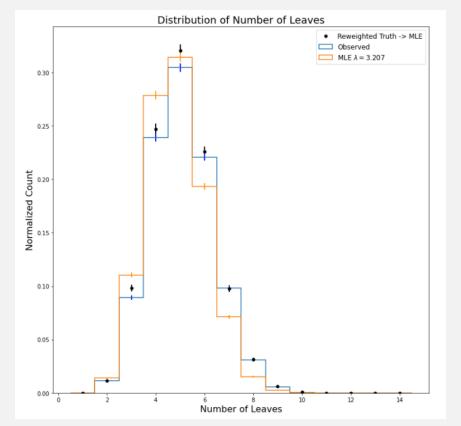
Iterative Reweighting Procedure



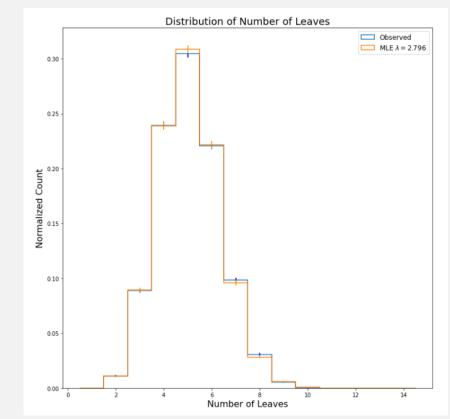
Likelihood Correction on Observed Data

Iterative Reweighting Procedure

Before



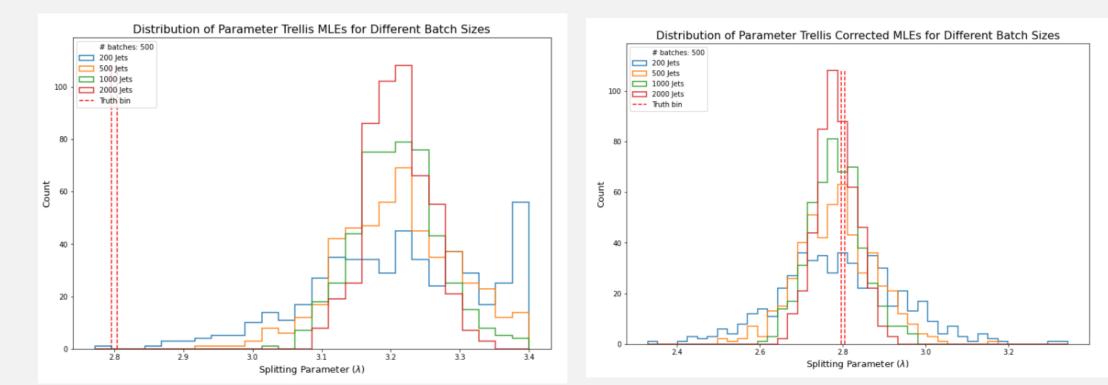
After 2 Iterations



MLE Distributions

Marginal MLE





χ^2 Test Statistic Distributions

Marginal χ^2 Test Statistic

Reweighted χ^2 **Test Statistic**

